



Peer review status:

This is a non-peer-reviewed preprint submitted to EarthArXiv.

1 **Influence of sea surface temperature patterns and mean warming on past**
2 **and future Atlantic tropical cyclone activity**

3 E. L. Levin^a G. A. Vecchi,^b W. Yang,^b

4 ^a *Program in Atmospheric and Oceanic Sciences, Princeton University*

5 ^b *Department of Geosciences, Princeton University*

6 *Corresponding author:* Emma Levin, emma.levin@princeton.edu

7 ABSTRACT: This study investigates the relative contributions of large-scale thermodynamic and
8 dynamic processes to decadal and multidecadal changes in Atlantic tropical cyclone (TC) activity,
9 spanning the historical record since the late 19th century, and extending to 2100 projections. We
10 employ a framework that decomposes TC counts into precursor disturbances that transition into
11 fully developed storms, applied to multi-ensemble simulations of two TC-permitting atmospheric
12 models. Using these models, we conduct controlled experiments with distinct SST forcings that
13 isolate the influence of SST spatial patterns from that of global-mean warming on Atlantic TC
14 activity. Our results show that decadal and multidecadal changes in TC frequency are primarily
15 governed by two thermodynamic variables: potential intensity and moist entropy deficit. In the
16 historical record, these variables reinforced one another, producing more robust trends in TC
17 activity. In contrast, future projections suggest opposing influences, with one variable (potential
18 intensity) becoming more favorable for TCs while the other (moist entropy deficit) becomes less
19 favorable, leading to increased uncertainty in TC projections. We trace this shift to differences in
20 relative warming between the tropical Atlantic and the broader tropics, underscoring that regional
21 SST patterns, rather than the global mean warming rate, control both past variability and projected
22 future changes in TC activity. Constraining future projected patterns of warming is therefore
23 essential for improving the reliability of TC projections.

24 **1. Introduction**

25 Responsible for over \$250 billion of inflation-adjusted losses in the U.S. between 2008 and 2017
26 (Klotzbach et al. (2018)), landfalling tropical cyclones (TCs) that form in the Atlantic Ocean are
27 a major hazard for both coastal and inland communities in the U.S. and Caribbean. Further, it is
28 crucial to understand the factors that have driven past fluctuations in Atlantic storm activity, and
29 the extent to which anthropogenic warming might influence TCs in the future.

30 While several studies attempting to reconstruct a reliable historical record of Atlantic TCs
31 generally disagree on the direction of the trend in Atlantic TC counts from the late-19th century
32 to the present (Emanuel (2021a), Vecchi and Knutson (2008), Vecchi and Knutson (2011), Vecchi
33 et al. (2021)), there is a growing understanding of a role for regional climate variability, warming
34 patterns, and aerosol concentrations, rather than global mean anthropogenic warming, in driving
35 past fluctuations in seasonal storm frequency. For instance, several studies indicate that natural
36 climate modes, such as the Atlantic Multi-decadal Oscillation (Goldenberg et al. (2001); Klotzbach
37 (2011b); Zhang and Delworth (2006)) and the El Niño Southern Oscillation (Klotzbach (2011a);
38 Klotzbach et al. (2022); Patricola et al. (2016); Pielke and Landsea (1999); Xie et al. (2005)),
39 along with their corresponding oscillating sea surface temperature patterns, can influence North
40 Atlantic TC activity. Additionally, another explanation links past multidecadal fluctuations in
41 Atlantic TC activity to the rise and subsequent decline of sulfate aerosol emissions, driven by
42 global environmental regulations (Booth et al. 2012; Dunstone et al. 2013; Mann and Emanuel
43 2006; Murakami and Wang 2022).

44 Although there is a growing consensus that global warming contributes to an increase in peak
45 TC windspeed and rainfall intensity (see reviews Knutson et al. (2020) and Walsh et al. (2019)
46 and references therein), it remains unclear how seasonal TC counts will change in a warmer world
47 (Knutson et al. (2020) and Sobel et al. (2021)). Most studies based on climate model simulations
48 project a decrease or minimal change in global TC count (see reviews Knutson et al. (2020) and
49 Walsh et al. (2019) and references therein), but there is much spread in projections for individual
50 basins (Kossin et al. 2020). However, some dynamical modeling studies (e.g., Bhatia et al. (2018),
51 Vecchi et al. (2019)) and statistical-dynamical downscaling efforts (e.g., Emanuel (2013), Emanuel
52 (2021b)) predict an increase in global TC frequency due to anthropogenic global warming.

53 Several factors may explain the discrepancies in TC frequency projections across studies. First,
54 these studies use various climate models with differing parametrization schemes and horizon-
55 tal resolutions, ranging from tens to hundreds of kilometers, that may or may not resolve TCs
56 and mesoscale processes (Camargo et al. (2020); Hsieh et al. (2023); Manganello et al. (2012)).
57 Inconsistencies in TC projections may also stem from uncertainties regarding the impact of thermo-
58 dynamic variables, particularly humidity, on TC genesis. For instance, saturation deficit represents
59 the difference between the specific humidity of the atmosphere and its specific humidity at satura-
60 tion (Olszewski (1986)). In a warmer world, the saturation deficit in the lower and mid-troposphere
61 increases (Held and Soden (2000); Emanuel (2010, 2021a)). A higher saturation deficit would
62 require an increased surface moisture flux to bring the atmospheric column to saturation, which
63 could inhibit deep convection and cyclogenesis (Sobel et al. (2021), Tang and Camargo (2014)).
64 This theory suggests that with a higher saturation deficit in a warmer world, TC activity would
65 decrease (Camargo et al. (2014); Emanuel (2021a); Emanuel (2010); Lee et al. (2020); Lee et al.
66 (2023); Vecchi et al. (2019)). However, several high-resolution models project an increase in TC
67 activity (Bhatia et al. (2018); Vecchi et al. (2019)), despite a rise in saturation deficit. These
68 studies suggest that the interplay between local thermodynamic and dynamic variables could play
69 an important role in driving changes in TC activity. In the models that show a TC frequency
70 increase, the reduced genesis favorability from changes in saturation is outweighed by an increase
71 in pre-TC disturbances (known as 'seeds') (Vecchi et al. 2019; Hsieh et al. 2020).

72 Finally, inherent model biases and variations in projections of regional climate and sea surface
73 temperature (SST) patterns, particularly in the Eastern Pacific and Atlantic basins (see review by
74 Shaw et al. (2024) and references therein), could lead to differences in future TC projections.
75 Atlantic TCs are sensitive to the relative average temperature of the tropical Atlantic compared to
76 the rest of the tropics (Vecchi and Soden (2007); Villarini et al. (2011b)). Villarini et al. (2011b)
77 show that the SST patterns explain much of the spread in projected North Atlantic TC frequency
78 projections.

79 To investigate the relationships between large-scale environmental factors and TC activity, several
80 indices have been developed (e.g., Bruyère et al. (2012); Emanuel and Nolan (2004); Emanuel
81 (2010); Tippett et al. (2011); Wang and Murakami (2020)). These indices, known as genesis
82 potential indices (GPIs), relate local environmental factors—such as potential intensity (PI), vertical

83 wind shear, and humidity—to TC activity. GPIs are typically calculated from monthly averages,
84 representing regional climatology rather than instantaneous weather conditions (Sobel et al. (2021)).
85 While GPIs have been used to predict the effects of future climate change on TC activity (e.g.,
86 Murakami and Wang (2022)), Camargo et al. (2014) found that most GPIs fail to capture future
87 TC projections from atmosphere-only climate models. Moreover, these TC proxies show low skill
88 in capturing interannual variability and multidecadal trends in TC activity compared to climate
89 models that explicitly resolve TCs (Cavicchia et al. 2023), likely because they do not account for
90 key physical processes involved in TC genesis, such as the formation of initial TC seeds.

91 As an alternative way to construct GPIs, Hsieh et al. (2020) present a framework for understanding
92 how large-scale local environmental fields influence TC formation, explicitly accounting for the
93 seed phase of formation. Unlike traditional GPIs (Emanuel (2022)), this approach divides tropical
94 cyclogenesis into distinct phases, identifying the specific large-scale local factors relevant to each
95 stage of TC development. This framework is motivated by a recognition that the feedbacks on
96 development of TCs are different at different stages (Hsieh et al. 2020; Zhang et al. 2021), and
97 thus the environmental controls on seeds and the likelihood of genesis to a mature TC may be
98 different. This technique of separating TCs into seeds and a nondimensional probability that a seed
99 transitions into a TC has proven useful for studying TCs across various timescales, including their
100 annual cycle (Yang et al. (2021)) and over longer idealized climatic periods (Hsieh et al. (2022);
101 Hsieh et al. (2023); Vecchi et al. (2019)). In this study, we extend the application of this framework
102 to the long observed historical record, where it has not yet been applied. Additionally, we focus
103 this approach on the Atlantic basin, where seed-like disturbances may be influenced by nonlocal
104 factors, such as African Easterly Waves (Patricola et al. (2018), Ritchie and Holland (1999)).

105 Through this work, we aim to address the following research questions:

- 106 1. To what extent does the theoretical framework proposed by Hsieh et al. (2020) capture long-
107 term observed and projected Atlantic TC variability?
- 108 2. To what extent do different environmental factors (both thermodynamic and dynamic) modu-
109 late Atlantic TCs in the observed and projected records, and how are these factors related to
110 regional warming patterns versus globally-uniform warming?

111 To address these questions, we employ the seed-probability framework within a suite of state-of-
112 the-art high-resolution atmosphere-only general circulation model experiments that have demon-

113 strated skill in simulating key characteristics of past TC activity (Chan et al. (2021), Chen and
114 Lin (2011), Zhao et al. (2009), and Zhao and Held (2010)). We seek to identify the dominant
115 large-scale environmental factors driving historical and future projected variations in Atlantic TC
116 frequency. By utilizing atmosphere-only models forced by observed SSTs, we effectively isolate
117 the influence SST patterns on TC activity.

118 **2. Methods and data**

119 *a. Models and experiments*

120 In this study, we employ two closely related global high-resolution atmospheric models from
121 the Geophysical Fluid Dynamics Laboratory (GFDL): AM2.5-C360 and HIRAM (Zhao et al.
122 (2009)), both of which have been used to study tropical cyclones (TCs). These models share
123 the same dynamical core but differ in a number of their parameterizations (e.g. convection and
124 cloud) and horizontal resolutions. AM2.5-C360 uses the relaxed Arakawa–Schubert convective
125 parametrization scheme, while HIRAM uses the shallow parametrization method described in
126 Bretherton et al. (2004) modified to allow for deep convection (Zhao et al. 2009). HIRAM, with
127 a horizontal resolution of 50 km, has demonstrated skill in simulating various aspects of past TC
128 variability (Chen and Lin (2011), Zhao et al. (2009), and Zhao and Held (2010)). AM2.5-C360,
129 on the other hand, has a higher resolution of 25 km, and has been employed in prior studies to
130 examine the influence of SSTs on tropical cyclone activity (Chan et al. 2021), the annual cycle
131 of TCs (Yang et al. 2021), and the evolution of storm tracks under warming conditions (Kortum
132 et al. 2024). The use of atmosphere-only models forced with prescribed SSTs provides a clean
133 framework for isolating the role of SST patterns in shaping TC behavior.

134 We conducted two sets of experiments (Table 1) to isolate the effects of spatial patterns of SST
135 change from global mean SST warming, one using the historical record and one using future
136 projections. For the past historical period, in the first experiment, hereafter named *obs sst*, we
137 generated ten ensemble members of AM2.5-C360 and five ensemble members of HIRAM. The
138 models are forced by bias-corrected observed monthly sea surface temperatures (Chan et al. 2021)
139 from the Hadley Centre Sea Ice and Sea Surface Temperature (HADISST) dataset for the period
140 1871–2019. Each ensemble member was initialized with different initial conditions but used the

141 same historical SST forcing. We decompose the SSTs in this experiment into several components:

$$SST_{\underline{obs\ sst}} = SST_{clim}(x, y, \text{mon}) + \langle SST' \rangle(t) + SST'(x, y, t), \quad (1)$$

142 where $SST_{clim}(x, y, \text{mon})$ represents the observed monthly climatology throughout the 1871-1890
143 period at each (longitude, latitude) point (x, y) for each month (January, February, ...), $\langle SST' \rangle(t)$
144 represents the observed spatially averaged mean anomaly for each month over the duration of the
145 record (January 1871, February 1871, ...) from the globally averaged 1871-1890 monthly clima-
146 tology, and $SST'(x, y, t)$ represents the observed spatially varying anomaly at each point (x, y) for
147 each month over the duration of the record (January 1871, February 1871, ...) from the globally
148 averaged 1871-1890 monthly climatology. Therefore, in the historical record that undergoes an-
149 thropogenic warming, $\langle SST' \rangle(t)$ represents the global mean warming while $SST'(x, y, t)$ represents
150 the spatially varying pattern of SST change.

151 The second experiment, hereafter named rm warming, is the same as the first, except the radiative
152 forcings are fixed at the year 1880 levels and the specified SSTs are modified in a way that the 15-year
153 lowpass of global mean monthly SST anomaly (relative to the 1871-1890 monthly climatology) is
154 subtracted at each ocean cell grid. Therefore, we remove the $\langle SST' \rangle(t)$ term in the obs sst record,
155 so the SSTs for this experiment can be written in their decomposed form as:

$$SST_{\underline{rm\ warming}} = SST_{clim}(x, y) + SST'(x, y, t), \quad (2)$$

156 where the only difference from the obs sst SST forcing is the removal of the globally averaged
157 warming signal $\langle SST' \rangle(t)$. Therefore, the differences between the obs sst and the rm warming
158 experiments can be attributed to spatially averaged mean warming and the direct atmospheric
159 effect of changing radiative forcing. Thus, if we find minimal differences in the TC activity and
160 their environmental controls between these simulations, we can argue that mean warming played
161 a minimal role in driving TC variability during the historical record. We run this experiment in
162 AM2.5-C360 model with 5 ensemble members.

163 The second set of experiments isolates the effect of SST spatial patterns versus globally averaged
164 warming in the future projected period. The first experiment, hereafter referred to as rcp4.5 flor,
165 represents a time-varying projected late 21st-century warming scenario, which is run in both

166 AM2.5-C360 (6 ensembles) and HIRAM (3 ensembles) during the period 2021-2100. AM2.5-
 167 C360 and HIRAM use transient RCP4.5 radiative forcings from CMIP5 (Taylor et al. 2012) in
 168 these experiments. The SST forcing in this experiment is constructed using observed 2001-2020
 169 climatologies combined with modeled SST anomalies from the FLOR model (Vecchi et al. 2014).
 170 FLOR is a fully coupled atmosphere–ocean general circulation model developed at GFDL. The
 171 model features an ocean component with an approximately $1^\circ \times 1^\circ$ spatial resolution, and an
 172 atmospheric component with the same physics and dynamical core as AM2.5-C360, but run at a
 173 lower atmospheric spatial resolution of approximately 50 km. Using the same SST decomposition
 174 convention as above, we write the SST forcing of the rcp4.5 flor experiment:

$$SST_{rcp4.5\ flor} = \underbrace{SST_{clim}(x, y, \text{mon})}_{Observed} + \underbrace{\langle SST' \rangle(t) + SST'(x, y, t)}_{Modeled}, \quad (3)$$

175 where $SST_{clim}(x, y, \text{mon})$ is the observed 2001–2020 monthly climatology from the bias-corrected
 176 SST dataset of (Chan et al. 2021), which is the same observed SST dataset used in the historical
 177 obs_sst and rm_warming experiments. The use of an observed climatology, rather than the model’s
 178 simulated climatology, minimizes the impact of model mean-state biases. Both the globally
 179 averaged warming anomaly $\langle SST' \rangle(t)$ and the spatially varying anomaly $SST'(x, y, t)$ are derived
 180 from the FLOR RCP4.5 projection experiments. As a result, this experiment includes both global
 181 mean SST warming and spatially patterned SST changes.

182 To further isolate the effects of spatially patterned versus globally averaged SST warming on TCs
 183 during the future period, we also perform a set of idealized experiments consisting of a control
 184 simulation and a perturbed simulation forced by prescribed SSTs. These experiments are described
 185 in Hsieh et al. (2020) and Hsieh et al. (2022). In each case, repeating climatological monthly SST
 186 patterns are prescribed and held fixed throughout the simulation. All simulations are run for 50
 187 years, such that each experiment experiences 50 repetitions of the same annual SST cycle. First,
 188 the control simulation, hereafter named cntl, is forced by the observed 1986–2005 average annual
 189 SST cycle, where the decomposed SSTs only have an observed climatological component,

$$SST_{cntl} = \underbrace{SST_{clim}(x, y, \text{mon})}_{Observed}. \quad (4)$$

TABLE 1. Controlled experiments used in this study.

Experiment	Period	Climatological Component	Mean Warming Component	Patterned Warming Component
<i>obs_sst</i>	1871-2019	Yes, observed	Yes, observed	Yes, observed
<i>rm_warming</i>	1871-2019	Yes, observed	No	Yes, observed
<i>rcp4.5_flor</i>	2020-2100	Yes, observed	Yes, modeled	Yes, modeled
<i>cntl</i>	1986-2005	Yes, observed	No	No
<i>2K_2CO2</i>	1986-2005	Yes, observed	Yes, +2K	No

190 The second idealized experiment, hereafter referred to as *2K 2CO2*, includes the same observed
 191 1986–2005 climatological SST component plus a uniform global mean SST warming of 2 K. The
 192 SST forcing for this experiment is therefore

$$SST_{\underline{2K\ 2CO_2}} = \underbrace{SST_{clim}(x, y, \text{mon})}_{\text{Observed}} + \langle SST' \rangle, \quad (5)$$

193 where the global mean warming anomaly $\langle SST' \rangle$ is fixed at +2 K. In addition, atmospheric CO_2
 194 concentrations are doubled relative to the *cntl* experiment in order to represent enhanced greenhouse
 195 gas radiative forcing. Differences between the *cntl* and *2K 2CO2* experiments therefore represent
 196 the response to a purely global mean SST warming without any change in SST spatial patterns and
 197 direct atmospheric effects of increasing carbon dioxide (e.g. Zhao et al. (2009)). By comparing
 198 these responses with those obtained from the *rcp4.5_flor* experiment, which includes both global
 199 mean warming and spatially patterned SST changes, we can isolate the relative roles of mean versus
 200 patterned SST warming in modulating TC activity and its environmental controls.

201 *b. Data*

202 For the historical record of Atlantic tropical cyclone (TC) frequencies from 1871 to 2020, we
 203 utilize the adjusted dataset developed by Vecchi and Knutson (2008). To compute observed
 204 seasonal large-scale environmental factors, we use the ERA5 reanalysis dataset (Hersbach et al.
 205 2020) during the period 1979–2020 and the MERRA2 reanalysis dataset (Gelaro et al. 2017) during
 206 the period 1980-2020.

207 *c. TC and seed tracking*

208 To track tropical cyclones (TCs), we employ the algorithm developed by Harris et al. (2016),
209 setting the specific thresholds for wind speed, minimum sea level pressure, lifetime, and warm-
210 core characteristics as in Chan et al. (2021). The algorithm utilizes the following 6-hourly inputs:
211 sea level pressure, 850 hPa vorticity, 10-m wind speed, and mid-tropospheric (300–500 hPa) air
212 temperature. The process begins by searching for local minima in sea level pressure and then
213 applying an 850 hPa vorticity threshold of $1.5 \times 10^{-4} \text{s}^{-1}$, which filters out disorganized or weak
214 systems. For the remaining storms, additional filters ensure longevity and robustness, including a
215 minimum total lifetime of 72 hours (Villarini et al. (2011a)), at least 48 hours with a warm core
216 (defined as a maximum 300–500 hPa temperature encircled by a 2° contour within 500 km of the
217 storm’s minimum sea level pressure), and at least 36 consecutive hours with both a warm core and
218 maximum 10-m winds exceeding 15 m s^{-1} . We require that at least one time step along the storm’s
219 trajectory that the maximum wind speed exceeds 17 m s^{-1} .

220 For the HIRAM model, we implemented minor modifications to the algorithm to ensure that
221 the tracked TC frequency in the reference experiment closely matches the globally averaged TC
222 frequency. These are the same adjustments for HIRAM are used in Yang et al. (2021). The
223 warm-core temperature contour is increased from 2 to 2.5° and the required proximity of the warm
224 core to the storm center is decreased from 500 km to 110 km. For the AM2.5-C360 model, we
225 retain the original thresholds without modifications.

226 Seed detection is performed using the aforementioned TC tracking algorithm, with the approach
227 described in Yang et al. (2021). We examine candidate points as local pressure minimal that exceed
228 the a $1.5 \times 10^{-4} \text{s}^{-1}$ vorticity threshold. A seed must span a radius of at least 50 km and exhibit a
229 maximum 850 hPa relative vorticity of at least $4 \times 10^{-4} \text{s}^{-1}$ during its lifetime. The seed tracking
230 algorithm remains unchanged for both atmospheric models.

231 *d. Seed and probability indices*

232 We use Hsieh et al. (2020)’s probabilistic framework, which decomposes annual TC counts into
233 two stages: precursor seed disturbances and fully developed TCs. Based on these two stages,
234 Hsieh et al. (2020) infer that the annual frequency of North Atlantic TCs, N_{TC} follows a binomial

235 distribution:

$$N_{TC} \sim \text{binom}(N_s, P), \quad (6)$$

236 where N_s is the total number of first stage rotating seeds present in the basin for a given year,
237 which has units of storm count, and P is the dimensionless basin-aggregated probability that a
238 first-stage seed transitions into a second-stage TC. Consequently, the expected value of N_{TC} , which
239 has dimensions of storm count, is given by

$$N_{TC} = N_s \times P. \quad (7)$$

240 For instance, if a given season produces 75 seeds ($N_s = 75$), of which 15 transition into TCs
241 ($N_{TC=15}$), the basin-aggregated empirical transition probability from the first stage to the second
242 stage for that season would be $P = 15/75 = 0.2$.

243 Hsieh et al. (2020)'s ansatz parametrizes N_s and P as functions of large-scale environmental
244 conditions. First, they develop a proxy for N_s , known as the seed propensity index (SPI):

$$N_s \approx SPI = (\kappa) \cdot (-\omega) \cdot \frac{1}{1 + Z^{-1/\sigma}}, \quad (8)$$

245 where κ is a constant proportionality fitted parameter in units of storms per Pa s^{-1} , ω is the
246 monthly mean 500 hPa vertical velocity in pressure coordinates (in units of Pa s^{-1} where $\omega > 0$
247 is for downward motion), $\sigma = 0.69$ is a constant nondimensional fitting parameter, and Z is
248 nondimensional parameter that represents the ability of the low-level vorticity to spinup and is a
249 function of low level vorticity (Ikehata and Satoh (2021)),

$$Z = \frac{f + \zeta}{\sqrt{|\beta + \partial_y \zeta| U}}, \quad (9)$$

250 where f represents the Coriolis parameter and β represents its meridional gradient, ζ is the monthly
251 mean relative vorticity at 850 hPa, and U is assumed to be a constant wind speed of 20 m/s which
252 is empirically fit using aquaplanet model simulations.

253 Next, the probability that a weakly rotating seed develops into a strongly rotating TC (P ; a
 254 dimensionless value between 0 and 1) is parameterized as a probability index ($P(\Lambda)$):

$$P \approx P(\Lambda) = \frac{1}{1 + (\Lambda_0/\Lambda)^{1/\gamma}}, \quad (10)$$

255 where $\Lambda_0 = 0.014$ and $\gamma = -0.9$ are constant dimensionless fitting parameters, and Λ is the venti-
 256 lation index defined by Tang and Emanuel (2010) and Tang and Emanuel (2012), which measures
 257 the degree to which the influx and circulation of cold dry air into the storm's convective plume can
 258 inhibit the storm's strength. The ventilation index is a non-dimensional metric:

$$\Lambda = \frac{v_s \cdot \chi}{PI}, \quad (11)$$

259 where v_s is the vertical wind shear between 850 hPa and 250 hPa in units of m s^{-1} , PI is potential
 260 intensity (the theoretical upper limit of a TC's wind speed based on temperature contrasts between
 261 the sea surface and upper troposphere in units of m s^{-1}), and χ is a dimensionless parameter
 262 representing moist entropy deficit. This parameter is defined as

$$\chi = \frac{s_m^* - s_m}{s_0^* - s_b}, \quad (12)$$

263 where s_m^* is the saturation moist entropies at 600 hPa in the inner core of a TC, s_m is the envi-
 264 ronmental entropy at 600 hPa, s_0^* are the saturation moist entropy at the sea surface, and s_b is the
 265 entropy of the boundary layer. The numerator of the ventilation index equation (11) represents
 266 the difference in midlevel entropy between the TC and its environment, while the denominator
 267 represents the air-sea disequilibrium.

268 Thus, equations (8) and (10) serve as proxies for N_s and P expressed as functions of large-scale
 269 environmental factors that can be derived from climate model simulations. This framework offers
 270 a valuable approach for studying TCs using climate models, particularly those that lack the spatial
 271 resolution required to directly resolve TCs and seeds. We compute the mean Atlantic basin SPI
 272 over 5–30°N and $P(\Lambda)$ over 10–30°N. The slightly more equatorward bound for the SPI index
 273 reflects the sequential nature of TC development: storms first emerge as precursor disturbances
 274 before transitioning into fully developed TCs. Because the TC main development region, where

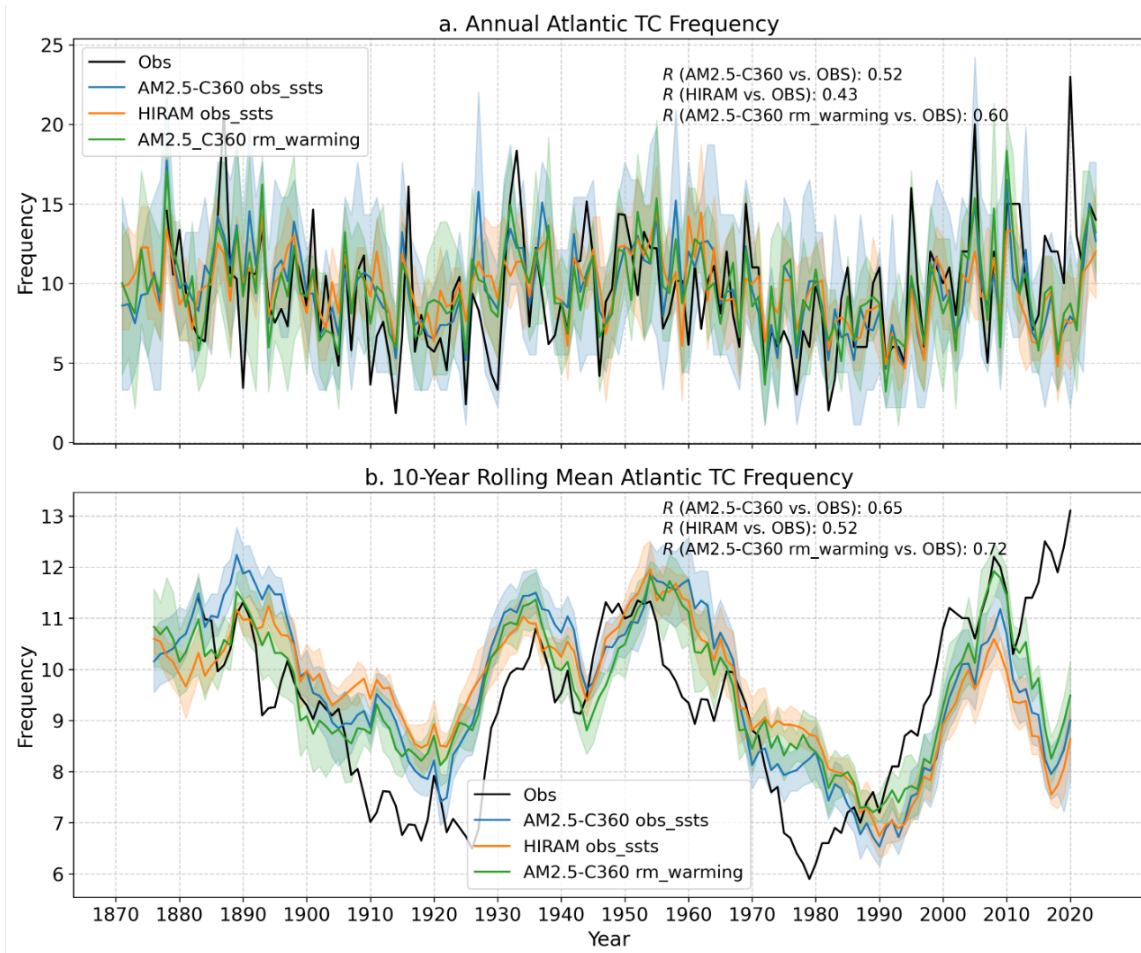
275 most Atlantic TCs reach full maturity, is typically defined north of 10°N (Vecchi et al. 2011)),
276 extending the SPI domain southward to 5°N accounts for precursor seed formation that occurs
277 prior to TC genesis. Our results are not qualitatively sensitive to this choice, as using a 10–30°N
278 domain for both SPI and $P(\Lambda)$ yields similar conclusions. Following the methodology of Vecchi
279 et al. (2019), we compute the TC framework and evaluate the monthly large-scale environmental
280 controls on Atlantic TC activity during the climatologically most active portion of the season,
281 June–November.

282 3. Results

283 a. Evaluating the models for historical Atlantic TCs

284 First, we assess how well the historical simulations (*obs sst* and *rm warming*) reproduce past
285 observed TC variability. Figure 1a compares observed annual Atlantic TC counts to those simulated
286 by both models. Overall, the *obs sst* simulation in both models and AM2.5-C360 *rm warming*
287 simulation effectively capture interannual fluctuations in TC activity. For instance, all simulations
288 show a local peak in 2005 and a trough in 1982, corresponding to historically active and inactive
289 years, respectively. The annual ensemble-mean correlation between simulated *obs sst* record and
290 observed TC frequency, without detrending the record, is slightly higher in AM2.5-C360 compared
291 to HIRAM, though both exhibit similar values (~ 0.5). Interestingly, the annual ensemble-mean
292 correlation between simulated *rm warming* record and observed TC frequency is slightly higher
293 ($r = 0.6$) than that for the *obs sst* simulations.

302 For most years, observed TC fluctuations (black line) remain within the 95% inter-ensemble
303 spread of all model simulations (blue, orange, and green shading), suggesting that the model
304 simulations reproduce realistic levels of TC variability. However, the exceptionally active 2020
305 season, and several seasons immediately preceding it, are not well captured, with observed TC
306 counts lying outside the estimated 95% confidence interval of both models. The observed TC
307 count also lies outside the estimated 95% confidence interval in other years, such as 1890, 1916,
308 1933, and 1982. Work is better underway to better understand the extent to which this discrepancy,
309 which represents a deficiency in the models or the forcing, is due to ambiguity in the observed
310 target or is likely to be best understood as one of the unlikely outcomes that even a well-calibrated
311 system will exhibit.



294 FIG. 1. (a) Annual TC frequency from 1871 to 2023, including the adjusted observational record from Vecchi and Knutson (2011) (black), AM2.5-C360 obs_sst simulation (blue), AM2.5-C360 rm_warming simulation
 295 (green), and HIRAM obs_sst simulation (orange). The shading represents the ensemble minimum and maximum
 296 values for each simulation. (b) Ten-year rolling mean of the TC frequency shown in (a). Solid lines indicate
 297 ensemble means, while shaded regions represent the bootstrapped 95% confidence intervals derived from re-
 298 sampling ensemble members. Confidence intervals are obtained by computing the 10-year rolling mean for
 299 each ensemble member, resampling the ensemble with replacement 1,000 times, and taking the 2.5th and 97.5th
 300 percentiles of the resulting distribution, which are represented by the bounds of the shading.
 301

312 We propose several hypotheses for this model–observation discrepancy. First, aerosol changes
 313 during this period may not have been fully represented in the SST-forced simulations. The onset
 314 of the COVID-19 pandemic in early 2020 sharply reduced global travel and industrial activity,

315 potentially lowering aerosol emissions or changing their geographic distribution. At the same
316 time, new International Maritime Organization regulations, effective January 1, 2020, reduced
317 the sulfur content of shipping fuels from 3.5% to 0.5%. While intended to improve air quality
318 and public health, this reduction diminished marine aerosol loading, brightening fewer clouds and
319 decreasing reflected solar radiation. The resulting radiative imbalance likely accelerated short-
320 term warming, creating atmospheric conditions more favorable for TC development (Diamond
321 2023; Jordan and Henry 2024; Zhang et al. 2025), yet this signal should be constrained in the
322 observed SST forcing. Still, such effects cannot account for the discrepancies observed in the years
323 immediately prior to 2020.

324 Additional possible causes include model and storm-tracking limitations within AM2.5-C360
325 and HIRAM, as well as evolving observational practices since the advent of the satellite era. For
326 example, modern satellites may have become increasingly capable of detecting weaker disturbances,
327 inflating the observed storm record relative to earlier periods. Finally, internal climate variability
328 and random weather fluctuations may also have played a role (e.g. Kortum et al. (2024)). A
329 fuller exploration of these hypotheses lies beyond the scope of this report but remains an important
330 direction for future research.

331 Figure 1b evaluates the ability of the model simulations to reproduce historical low-frequency
332 variability in Atlantic TC activity using a 10-year rolling mean. This filtering suppresses interannual
333 variability while retaining a combination of decadal (10–20 year) and multidecadal (30+ year)
334 variability. The *obs_sst* simulations in both models, as well as the AM2.5-C360 *rm_warming*
335 simulation, broadly reproduce the observed low-frequency fluctuations in Atlantic TC activity
336 since the late nineteenth century. In particular, these simulations capture decadal-scale peaks in
337 the 1930s and 1950s and a relative minimum during the early twentieth century (1910–1930).
338 They also reproduce multidecadal variations, including elevated activity during 1871–1900 and
339 1930–1970 and reduced activity during 1900–1930 and 1970–2000.

340 A key exception is the models' inability to reproduce the recent decadal-scale increase in TC
341 activity from 2010–2020. During 2010–2018, the simulated and observed 10-year rolling-mean
342 TC counts exhibit a pronounced anticorrelation: the observations show increasing decadal-mean
343 activity, whereas all three model simulations display a decline over the same period. Although
344 discrepancies between simulated and observed interannual TC counts are also evident during

345 2015–2019 (Fig. 1a), we attribute the decadal-scale anticorrelation primarily to the models’ failure
346 to capture the exceptionally active 2020 hurricane season, as discussed above. The 2020 season
347 was anomalous, producing 23 named storms, whereas the ensemble-mean simulations generate
348 approximately eight storms. This extreme observed outlier likely inflates the observed 10-year
349 rolling mean trend over 2010–2020, while the comparatively weak simulated activity in 2020
350 contributes to a declining trend in the modeled decadal means.

351 Consistent with the annual analysis, the ensemble spread for the 10-year rolling mean TC record
352 is larger in the AM2.5-C360 *obs sst* simulation than in the HIRAM *obs sst* simulation, reflecting
353 the larger ensemble size in AM2.5-C360 (ten members) compared to HIRAM (five members).
354 Notably, correlation coefficients for the 10-year rolling mean TC counts are modestly higher than
355 those obtained for annual TC frequencies, indicating improved model performance at capturing
356 decadal-scale variability. The correlations between observed and simulated 10-year rolling mean
357 TC counts are higher for AM2.5-C360 ($r = 0.65$ and $r = 0.72$) than for HIRAM *obs sst* ($r = 0.52$),
358 suggesting greater skill in AM2.5-C360 for reproducing decadal Atlantic TC variability.

359 In both Fig. 1a and Fig. 1b, the strongest agreement between observed TC counts and the simulated
360 ensemble mean occurs in the AM2.5-C360 *rm warming* simulation. Importantly, this experiment
361 closely reproduces the decadal and multidecadal variability evident in the observational record,
362 and its variability is nearly indistinguishable from that of the *obs sst* simulations. These results
363 suggest that historical TC variability is primarily driven by the spatial pattern of ocean warming
364 rather than the global-mean warming trend, as the *rm warming* simulation closely matches the
365 *obs sst* simulation despite the removal of the mean warming signal.

366 *b. Mean warming versus patterned warming*

367 In this section we discuss the effect of regional warming patterns versus global mean warming
368 on Atlantic TCs and their environmental controls during both the historical and projected record,
369 leveraging all these experiments described in Section 2.

370 *(i) Historical record*

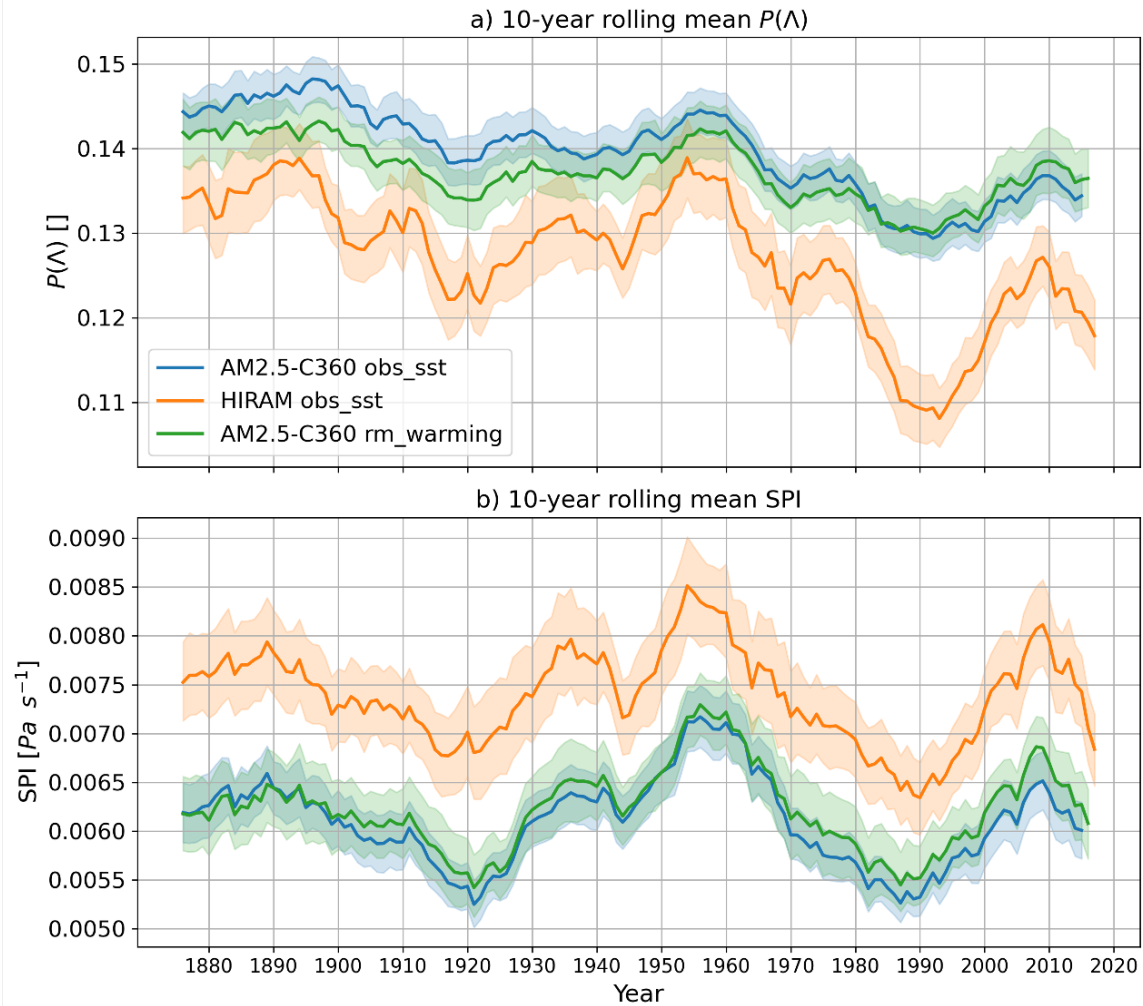
371
372 First, Fig. 1 indicates that regional sea surface temperature (SST) patterns, rather than the global-
373 mean warming trend, primarily drove interannual, decadal, and multidecadal variability in Atlantic

374 tropical cyclone (TC) activity during the historical period. This inference follows from the strong
375 similarity between the AM2.5-C360 obs sst and rm warming simulations, both of which exhibit
376 comparable TC variability and moderate correlations with the observed TC record. Because the
377 sole difference between these simulations is the removal of the global-mean SST warming trend in
378 the rm warming experiment, any systematic divergence between the two would be attributable to
379 mean warming. The absence of such divergence therefore suggests that historical TC variability
380 cannot be explained by the global-mean warming trend alone and instead reflects the influence of
381 regional SST patterns.

382 To further assess whether regional SST patterns drive historical variability in the large-scale
383 environmental controls on TCs, we examine the two components of the TC proxy proposed by
384 Hsieh et al. (2020), SPI and $P(\Lambda)$, across three simulations: the obs sst simulations in both models
385 and the AM2.5-C360 rm warming simulation. Figure 2 shows 10-year rolling mean time series
386 of basin-averaged SPI and $P(\Lambda)$ for these simulations over the historical period. Substantial
387 differences between the rm warming and obs sst simulations would indicate a role for global-mean
388 warming in modulating the environmental controls on TC activity.

389 We find that AM2.5-C360 and HIRAM exhibit systematic differences in the mean magnitudes
390 of SPI and $P(\Lambda)$. In particular, $P(\Lambda)$ is consistently higher in both AM2.5-C360 simulations
391 than in HIRAM by approximately 0.1, whereas SPI is lower in AM2.5-C360 by about 1.5×10^{-3} .
392 Despite these mean offsets, SPI exhibits no persistent long-term trend in any simulation. Instead,
393 SPI is dominated by pronounced multidecadal (30–60 yr) variability, with elevated values during
394 1880–1910 and 1930–1970 and reduced values during 1910–1930 and 1970–2000. In contrast,
395 $P(\Lambda)$ displays a markedly different low-frequency evolution. While superimposed with decadal
396 (10–20 yr) fluctuations, including relative maxima in the 1890s, 1950s, 1970s, and 2000–2020
397 and a minimum during 1910–1930, $P(\Lambda)$ is characterized by a gradual decline over much of the
398 historical period. Most notably, we observe a pronounced multidecadal decrease in $P(\Lambda)$ between
399 approximately 1960 and 1990.

400 Crucially, the time series of both SPI and $P(\Lambda)$ are nearly identical between the AM2.5-C360
401 obs sst and rm warming simulations, with their 95% confidence intervals overlapping for almost the
402 entire record. This close agreement demonstrates that the historical variability of the large-scale
403 environmental controls on Atlantic TC activity is governed primarily by regional SST patterns



408 FIG. 2. 10-year rolling mean of a) $P(\Lambda)$ and b) SPI, from the AM2.5-C360 *obs_sst* simulation (blue), the
 409 AM2.5-C360 *rm_warming* simulation (green) and the HIRAM *obs_sst* simulation (orange) during the 1871-2020
 410 historical period. Shaded regions indicate the bootstrapped 95% confidence intervals as in Fig. 1, derived from
 411 resampled ensemble members, while solid lines represent the ensemble means.

404 rather than by the global-mean warming trend. The similarity of these proxy components in
 405 simulations with and without mean warming reinforces the conclusion that spatial SST structure,
 406 rather than uniform warming, dominated TC-relevant environmental variability over the historical
 407 period.

412 (ii) *Projected record*

413

414 To evaluate the relative roles of regional SST patterns and global-mean warming in shaping
415 future Atlantic tropical cyclone (TC) activity and its environmental controls, we analyze the
416 sensitivity of both annual TC counts and the TC proxy ($SPI \times P(\Lambda)$) to mean SST warming
417 across two experimental frameworks in both models. Specifically, we examine sensitivities in
418 the rcp4.5 flor simulations, which include both patterned and mean SST warming, and contrast
419 them with sensitivities derived from the idealized comparison between the cntl and 2K 2CO2
420 experiments, which isolate the response to a spatially uniform +2 K global-mean warming. Because
421 the rcp4.5 flor simulations incorporate both components of warming, whereas the 2K 2CO2-cntl
422 comparison reflects only the effect of global-mean warming, systematic differences in TC and
423 TC-proxy sensitivities between these two cases can be attributed to the influence of regional SST
424 patterns. Accordingly, large discrepancies between the sensitivities diagnosed from the rcp4.5 flor
425 simulations and those from the idealized uniform-warming experiments indicate that, during the
426 future projected period, variability in Atlantic TC activity is governed primarily by regional SST
427 patterns rather than by global-mean SST warming alone.

428 Figure 3 shows the sensitivity of Atlantic tropical cyclone (TC) counts to global-mean SST
429 warming relative to the 1986–2005 climatology. We first note that the sign and magnitude of the
430 diagnosed sensitivities differ across models, even for the same experimental configurations. In the
431 AM2.5-C360 uniform-warming experiments, based on the comparison between the 2K 2CO2 and
432 cntl simulations, no statistically robust increase or decrease in Atlantic TC counts is evident in
433 response to uniform global-mean warming. In contrast, the corresponding HIRAM experiments
434 yield a negative sensitivity of approximately -1.5 to -2 TCs per K, indicating a projected decline in
435 TC activity under uniform warming in that model. Model-dependent differences are also apparent
436 in the rcp4.5 flor simulations, which include both mean and patterned SST warming. In AM2.5-
437 C360, these simulations indicate a modest positive sensitivity of roughly +0.5 to +1.5 TCs per K,
438 suggesting a slight increase in TC activity with warming. By contrast, the HIRAM rcp4.5 flor
439 simulation exhibits a wide range of sensitivities, spanning approximately -1.0 to +2.3 TCs per K,
440 with no consistent sign. The substantial discrepancies between models highlight the considerable
441 uncertainty in projections of future Atlantic TC activity in response to SST warming.

442 For each model, comparisons of TC sensitivity to mean warming between experiments that
443 include regional patterned warming (rcp4.5 flor) and those that exclude it (cntl and 2K 2CO2)

444 reveal pronounced differences in both the sign and magnitude of the projected response. In the
445 AM2.5-C360 controlled experiments without patterned warming, TC sensitivity to mean warming
446 is not statistically distinguishable from zero. In contrast, the AM2.5-C360 experiment that includes
447 patterned warming exhibits a clear positive sensitivity of TC activity to mean warming. A similar
448 contrast emerges in HIRAM: the controlled experiments without patterned warming show a robust
449 negative sensitivity of TC activity to mean warming, whereas the experiment that includes patterned
450 warming displays no consistent trend. These sign reversals between experiments with and without
451 regional SST patterns indicate that patterned warming, rather than global-mean warming alone,
452 plays a dominant role in shaping projected Atlantic TC variability.

453 We next examine the sensitivity of the TC proxy ($SPI \times P(\Lambda)$) to mean warming in simulations
454 that do and do not include regional patterned warming in both AM2.5-C360 and HIRAM (Fig. 3b).
455 Consistent with the TC count results, we find substantial differences in proxy sensitivity between
456 experiments with and without patterned warming. In AM2.5-C360, both experimental configura-
457 tions show a decrease in the TC proxy with increasing mean warming; however, the magnitude
458 of this sensitivity is nearly twice as large in the uniform-warming experiment compared to the
459 experiment that includes patterned warming. In HIRAM, the sign of the proxy sensitivity differs
460 between experiments. The uniform-warming experiment exhibits a decrease in the TC proxy with
461 mean warming, whereas the experiment that includes patterned warming shows no clear trend. To-
462 gether, these results indicate that regional SST patterns exert a critical influence on the large-scale
463 environmental controls of future Atlantic TC activity, beyond the effect of global-mean warming
464 alone.

465 Finally, we find that the simulated sensitivities of TC counts and the TC proxy are consistent
466 in HIRAM but diverge in AM2.5-C360. In HIRAM, the uniform-warming experiment shows a
467 decrease in TC counts with increasing mean warming, accompanied by a corresponding decrease in
468 the TC proxy. Likewise, in the experiment that includes both mean and patterned warming, where
469 TC counts exhibit no clear sensitivity to mean warming, the TC proxy also shows no discernible
470 trend. This internal consistency suggests that, in HIRAM, the proxy reliably reflects the model's
471 simulated TC response to warming. In contrast, AM2.5-C360 exhibits a mismatch between TC
472 count and proxy sensitivities. In the uniform-warming experiment, TC counts show no statistically
473 robust response to mean warming, whereas the TC proxy indicates a declining sensitivity. Con-

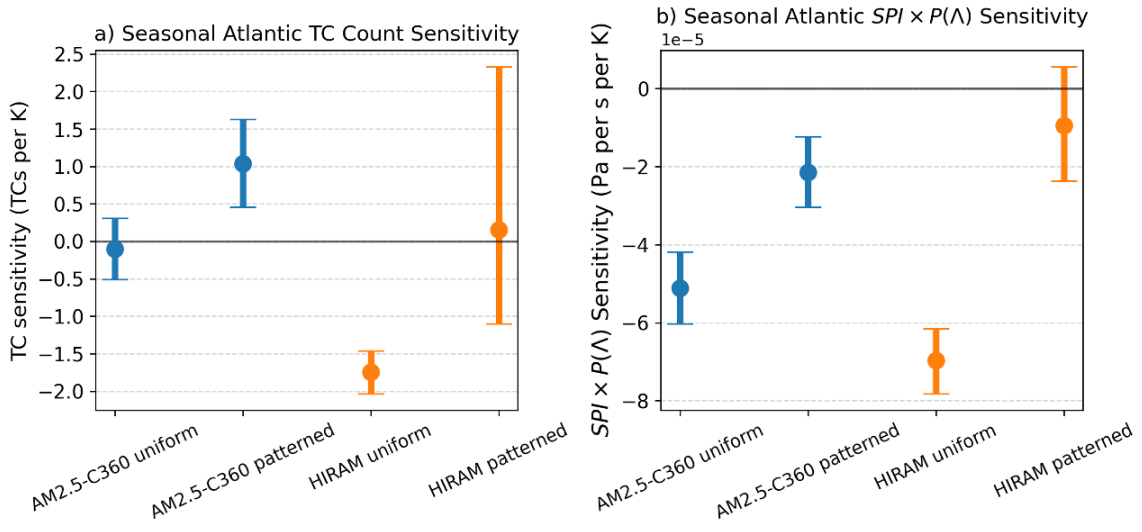
474 versely, in the experiment that includes both patterned and mean warming, TC counts exhibit a
475 positive sensitivity to mean warming, while the TC proxy indicates a negative response. Although
476 the TC proxy in the AM2.5-C360 rcp4.5 flor simulation remains positively correlated with the
477 projected TC count time series (Fig. 4), this discrepancy suggests a systematic bias in the proxy’s
478 representation of the mean response during the future projection period. Nonetheless, the proxy
479 continues to capture the decadal variability of TC activity, indicating that the bias primarily affects
480 the mean trend rather than the temporal variability. We further explore the ability of the TC proxy
481 to capture TC variability throughout the past and projected record in the following section.

498 *c. TC seed and probability proxies and environmental controls*

499 In this section, we evaluate the extent to which the seed-probability framework ($SPI \times P(\Lambda)$)
500 captures TC variability in our temporally varying simulations during both the historical and
501 projected periods using the obs sst and rcp4.5 flor experiments. Due to the similarity in results
502 between the obs sst and rm warming experiments, we only include the results from the obs sst
503 experiment. First, in Section a, we demonstrate that the SPI framework faithfully reproduces both
504 the spatial climatology and temporal variability of tracked TC seeds. This agreement provides
505 confidence in applying the full theoretical framework, $SPI \times P(\Lambda)$, to approximate TC counts.

506 Next, to evaluate the full theoretical framework, we examine the correlation between the
507 parametrized TC estimate ($SPI \times P(\Lambda)$) and the simulated TCs from the models. It is important to
508 note that the proxies SPI and $P(\Lambda)$ are not independent variables, but exhibit a spatial-temporal
509 covariation (Fig. A3). Visual inspection of Figure 4 indicates that the framework captures key fea-
510 tures of low-frequency TC variability, encompassing decadal (10–20 yr) and multidecadal (30–60
511 yr) timescales, in both models during the historical and projected periods. In particular, the frame-
512 work reproduces decadal-scale maxima in TC activity during the 1930s, 1950s, 2000–2010, 2020s,
513 2040–2055, and 2060–2075, as well as decadal-scale minima around 1920, 1980–2000, 2040, and
514 2060. In addition, the framework captures a pronounced multidecadal decline in TC activity from
515 approximately 1955 to 1990.

516 In both observed and projected periods, the parametrized TCs remain within the 95% confidence
517 interval of the simulated TCs, indicating that the framework provides a realistic representation of
518 TC variability. However, when comparing model performance, the framework exhibits stronger



482 FIG. 3. Sensitivity of (a) Atlantic tropical cyclone (TC) counts and (b) the Atlantic TC proxy ($SPI \times P(\Lambda)$)
 483 to mean warming across experiments with and without patterned SST warming during the future projection
 484 period. AM2.5-C360 experiments are shown in blue and HIRAM experiments in orange. Experiments labeled
 485 “uniform” on the x-axis represent TC sensitivity to uniform mean warming, computed by comparing the *cntl*
 486 and *2K 2CO2* experiments. Sensitivity is calculated as the difference in TC counts or proxy values between
 487 the *2K 2CO2* and *cntl* simulations, divided by 2 to account for the imposed +2 K warming. The bars denote
 488 95% confidence intervals estimated via bootstrapping. For the uniform experiments, confidence intervals are
 489 obtained by randomly pairing one year from the *cntl* experiment and one year from the *2K 2CO2* experiment
 490 (each consisting of 50 years of identical forcing), repeating this procedure 1,000 times without replacement, and
 491 computing the 2.5th, mean, and 97.5th percentiles of the resulting sensitivity distribution. Experiments labeled
 492 “patterned” represent TC sensitivity to mean warming in the *rcp4.5 flor* simulations. Because this experiment
 493 is transient rather than idealized, sensitivity is estimated by regressing seasonal Atlantic TC counts (or proxy
 494 values) against seasonal mean warming relative to the 1986–2005 climatology. Confidence intervals are obtained
 495 by bootstrapping the ensemble-member regression slopes 1,000 times with replacement and extracting the 2.5th
 496 and 97.5th percentiles. In both panels, the lower bound of each bar corresponds to the 2.5th percentile, the central
 497 marker denotes the bootstrap mean, and the upper bound indicates the 97.5th percentile.

519 agreement with HIRAM than AM2.5-C360, as reflected in the higher correlation coefficient be-
 520 tween parametrized and simulated ensemble mean TC count (0.94 vs. 0.78). In AM2.5-C360, the
 521 framework systematically underestimates projected TC activity, a behavior consistent across both

522 the uniform and patterned warming experiments shown in Figure 3. Overall, these results indicate
523 a smaller bias between tracked TCs and the TC proxy in HIRAM than in AM2.5-C360 during the
524 projected period.

537 Next, we assess the variability of SPI and $P(\Lambda)$ in both the historical and projected records.
538 Figure 5 illustrates the temporal evolution of SPI and $P(\Lambda)$ across the entire combined record in
539 both models' obs_sst and rcp45_flor experiments. As we observe in Figure 2, the models exhibit
540 distinct magnitudes of SPI and $P(\Lambda)$, with $P(\Lambda)$ higher in AM2.5-C360 than in HIRAM, while SPI
541 remains lower in AM2.5-C360. The magnitudes of these variables remain distinct throughout the
542 historical and projected record. Examining SPI in Figure 5b, we find no discernible long-term trend
543 but rather multidecadal and decadal variability. As shown in Figure 2b, during the historical period
544 (1871–2019), SPI displays prominent multidecadal fluctuations, with peaks from 1880-1910 and
545 1930-1970 and reduced values from 1910-1930 and 1970-2000. In contrast, the projected record
546 reveals predominantly decadal variability, characterized by peaks recurring approximately every
547 20 years. Meanwhile, $P(\Lambda)$ follows a distinct pattern: during the historical period, it exhibits a
548 steady decline, most pronounced 1960 to 1990 (Figure 2), before stabilizing in the projected record,
549 where variability remains relatively stagnant, except for one period of elevated values around 1970.
550 The most striking trend in the entire record for both SPI and $P(\Lambda)$ is a sharp decline from 1960
551 to 1990, followed by a notable rebound and increase from 1990 to 2010.

561 (i) *Thermodynamic variables*

562 To understand why $P(\Lambda)$ exhibits a decreasing trend throughout the historical period but remains
563 relatively stable in the projected period, we examine its thermodynamic components, χ and PI, both
564 of which respond directly to anthropogenic warming. Figure 6 illustrates the relationship between
565 PI and χ^{-1} in both the obs_sst and rcp4.5_flor experiments, as well as in several reanalysis datasets.
566 During the historical period (blue) and across the duration of the reanalysis data (orange and olive),
567 we observe a direct relationship between PI and χ^{-1} , indicating that their effects compounded each
568 other. When PI was higher and more conducive to TCs, χ was lower, further enhancing storm
569 activity. In contrast, during the projected period (green), this relationship reverses: as one variable
570 becomes more favorable for TC activity, the other becomes less favorable, suggesting that their
571 effects offset each other. Additionally, the models exhibit distinct magnitudes of PI and χ^{-1} . PI
572 is notably higher in HIRAM by approximately 7 m s^{-1} . While PI values in HIRAM are closer to

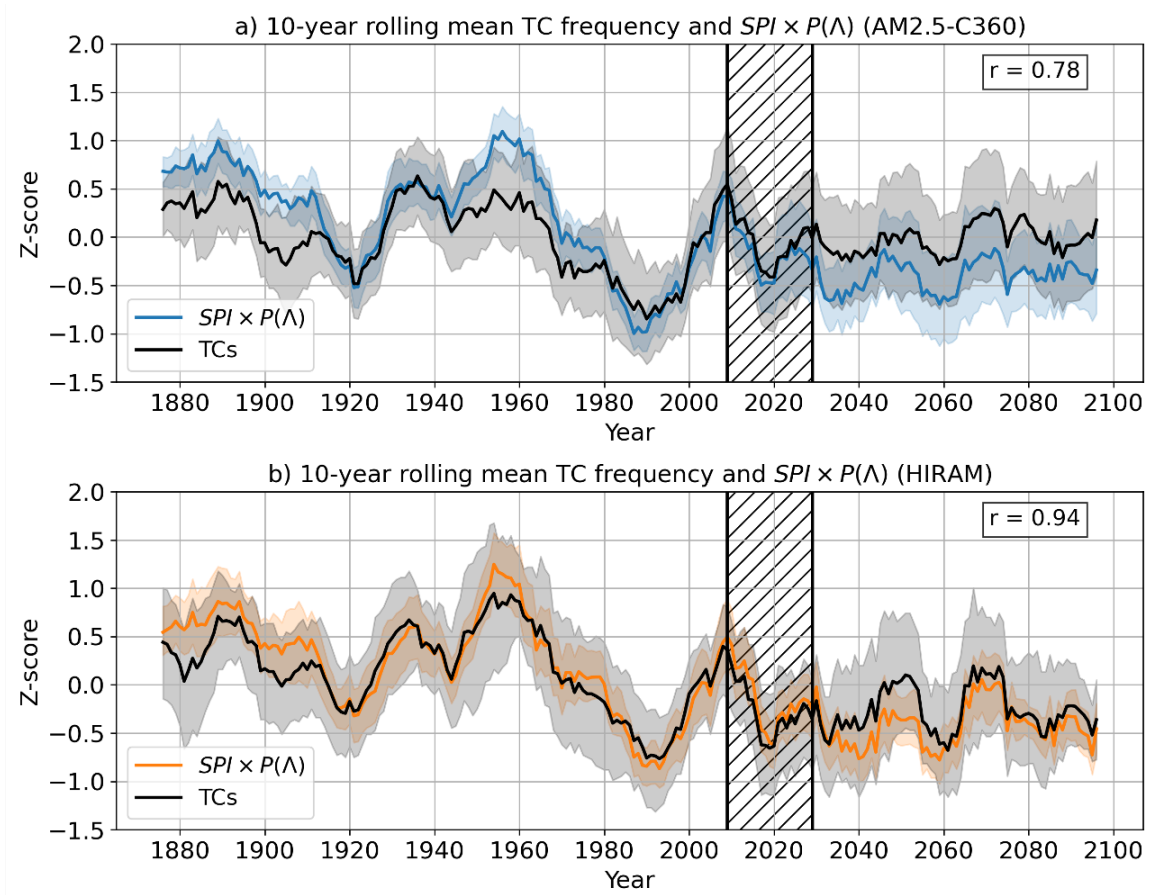
573 those found in reanalysis products, they still fall short by about 5 m s^{-1} . Meanwhile, χ^{-1} is slightly
574 higher in AM2.5-C360 than in HIRAM, though the reanalysis datasets display significant variation
575 in their projections of χ^{-1} .

583 To further investigate the reversal of the relationship between χ and PI in the historical and
584 projected periods, we plot their time series in Figure 7. The results show that χ increases
585 consistently throughout the entire record, in addition to some decadal variability, suggesting
586 that anthropogenic warming enhances the moisture deficit between the storm's inner core and its
587 environment. Meanwhile, the trend lines in Figure 7 reveal that on top of some decadal fluctuations,
588 PI decreases throughout most of the historical period, from 1871 to roughly 1980, but increases
589 after 1980 throughout the rest of the projected period. This shift in the sign of the PI trend around
590 1980 is the primary driver of the observed reversal in the PI and χ relationship.

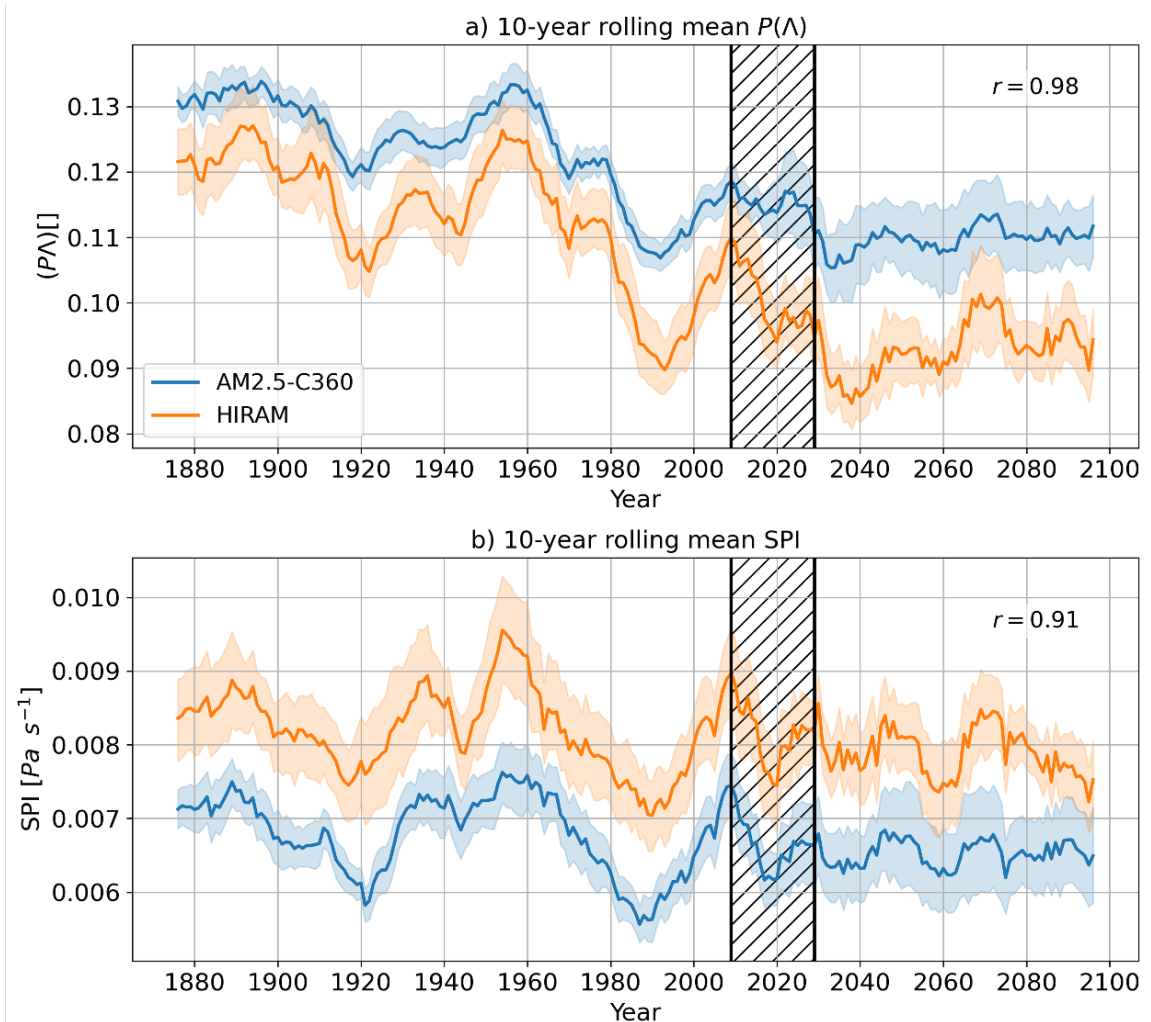
594 To better understand the discrepancy between observed and projected trends in PI, we analyze
595 trends in tropical sea surface temperatures (SSTs) during both periods. Tropical SSTs exert a
596 strong influence on upper-tropospheric temperatures, which in turn regulate PI (Emanuel et al.
597 (2013); Eusebi et al. (2025); Ramsay and Sobel (2011); Vecchi and Soden (2007); Vecchi et al.
598 (2013)), since PI is fundamentally determined by the temperature contrast between the surface and
599 the upper troposphere. The primary mechanism linking SSTs to free-tropospheric temperatures is
600 deep convection (Flannaghan et al. (2014); Fueglistaler et al. (2015); Sobel et al. (2002)), which
601 is most vigorous in regions of the warmest tropical SSTs (Zhang (1993)). Consequently, the
602 warmest tropical SSTs, typically associated with the western Pacific warm pool, effectively set
603 upper-tropospheric temperatures across the tropics through convective adjustment.

604 In the present climate, Sobel et al. (2002) showed that the tropical-mean SST is closely tied to
605 the warmest tropical SSTs. Thus, the rate at which local SSTs warm relative to the tropical-mean
606 warming rate provides a key indicator of PI. Motivated by this framework, we compare Atlantic
607 TC-season SST warming rates with those of the broader tropics during both the observed (Fig. 8)
608 and projected (Fig. 9) periods. Although the tropical Atlantic warmed during the historical period
609 (Fig. 8a), it did so more slowly than the tropical mean (Fig. 8b). In contrast, in the projected period,
610 tropical Atlantic SSTs are expected to warm not only in absolute terms but also more rapidly
611 than the tropical mean. These results imply that the Atlantic surface-to-troposphere temperature
612 contrast weakened in the historical record but is projected to strengthen in the future. We therefore

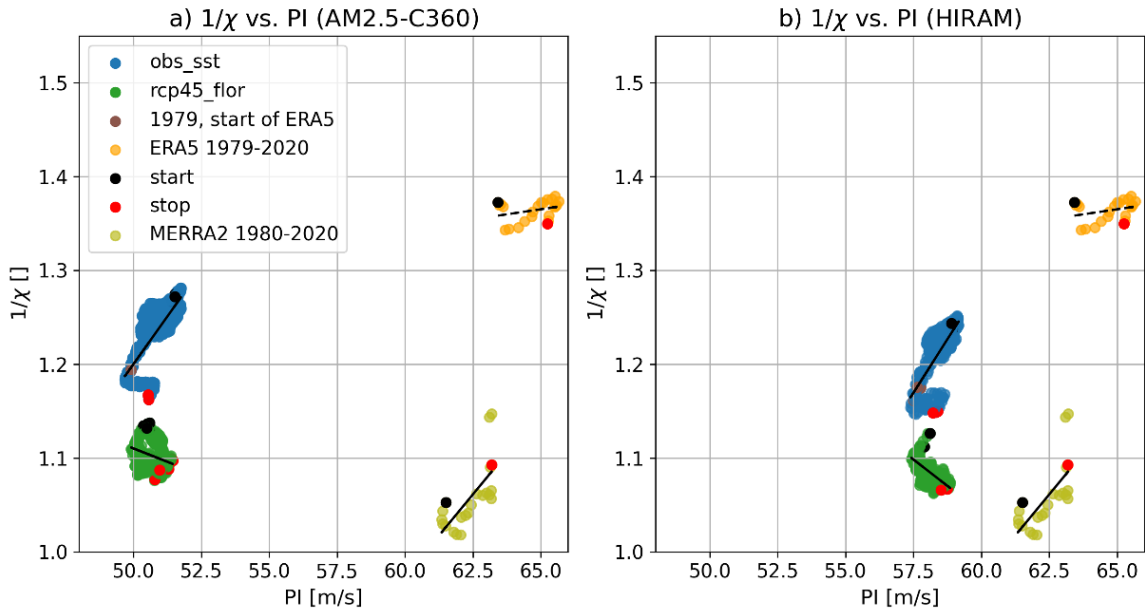
613 propose that the observed decline in PI, versus its projected increase, arises from differences in
614 relative SST changes. In particular, the spatial pattern of Atlantic SST warming relative to the
615 tropics plays a central role in shaping PI trends and their linkage to χ , both of which are critical
616 drivers of changes in $P(\Lambda)$.



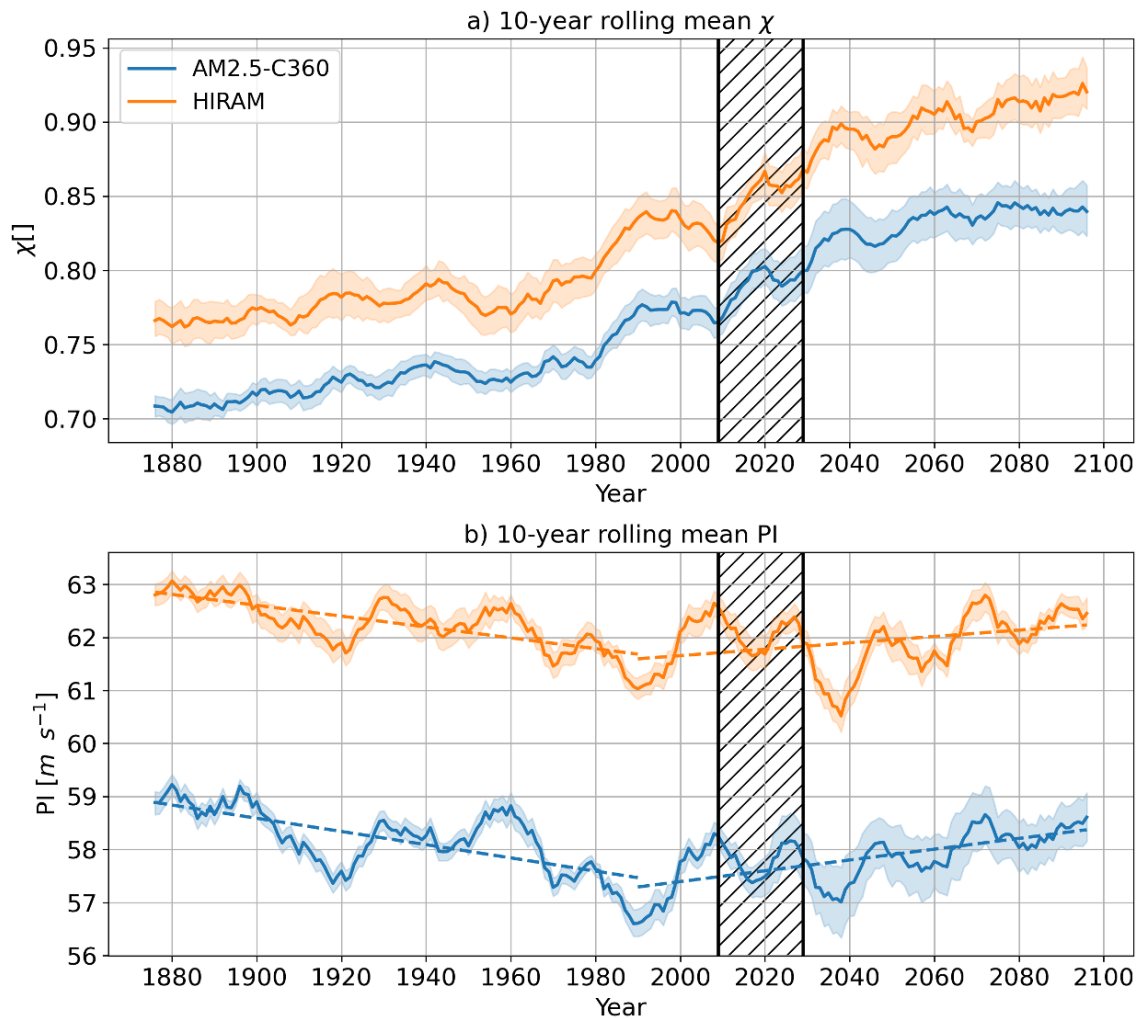
525 FIG. 4. Normalized 10-year rolling mean of simulated Atlantic tropical cyclone (TC) frequency (N_{TC} ; black)
 526 and parameterized TC frequency ($SPI \times P(\Lambda)$; colored) for (a) AM2.5-C350 (blue) and (b) HIRAM (orange).
 527 Shaded regions indicate the bootstrapped 95% confidence intervals as in Fig. 1, derived from resampled ensemble
 528 members, while solid lines represent the ensemble means. The historical period (1871–2019) is based on the
 529 *obs SST* simulation, whereas the future period (2020–2100) is derived from the *rcp4.5 SST* simulation. These
 530 two periods are combined into a continuous time series, with both simulated and parameterized TC frequencies
 531 standardized using a Z-score normalization: $Z(x) = (x - \mu_x) / \sigma_x$, where x represents the data point, μ_x is the
 532 mean, and σ_x is the standard deviation over the full record. Correlation coefficients between the ensemble mean
 533 simulated and parameterized TC frequencies are displayed in the top right corner of each panel. Values to the
 534 left of the hatched region correspond to simulations forced with observed SSTs (*obs_sst*), while values to the
 535 right correspond to simulations forced with observed climatology plus projected SST patterns (*rcp45_flor*). The
 536 hatched region marks a smooth transition that blends both forcing regimes.



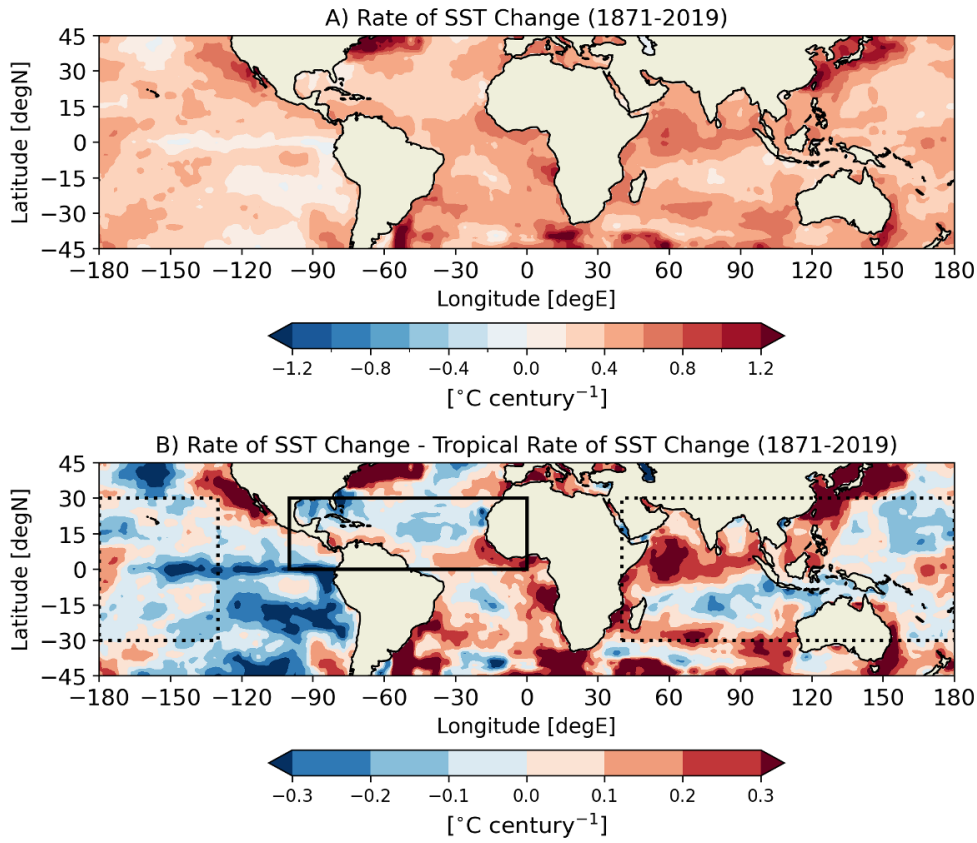
552 FIG. 5. 10-year rolling mean of a) $P(\Delta)$ and b) SPI, from AM2.5-C360 (blue) and HIRAM (orange). Shaded
 553 regions indicate the bootstrapped 95% confidence intervals as in Fig. 1, derived from resampled ensemble
 554 members, while solid lines represent the ensemble means. The historical period (1871–2019) is based on the
 555 *obs SST* simulation, whereas the future period (2020–2100) is derived from the *rcp4.5 SST* simulation. These two
 556 periods are combined into a continuous time series. The correlation coefficients between the ensemble mean of
 557 each model is shown in the upper right corner of each panel. Values to the left of the hatched region correspond to
 558 simulations forced with observed SSTs (*obs sst*), while values to the right correspond to simulations forced with
 559 observed climatology plus projected SST patterns (*rcp45 flor*). The hatched region marks a smooth transition
 560 that blends both forcing regimes.



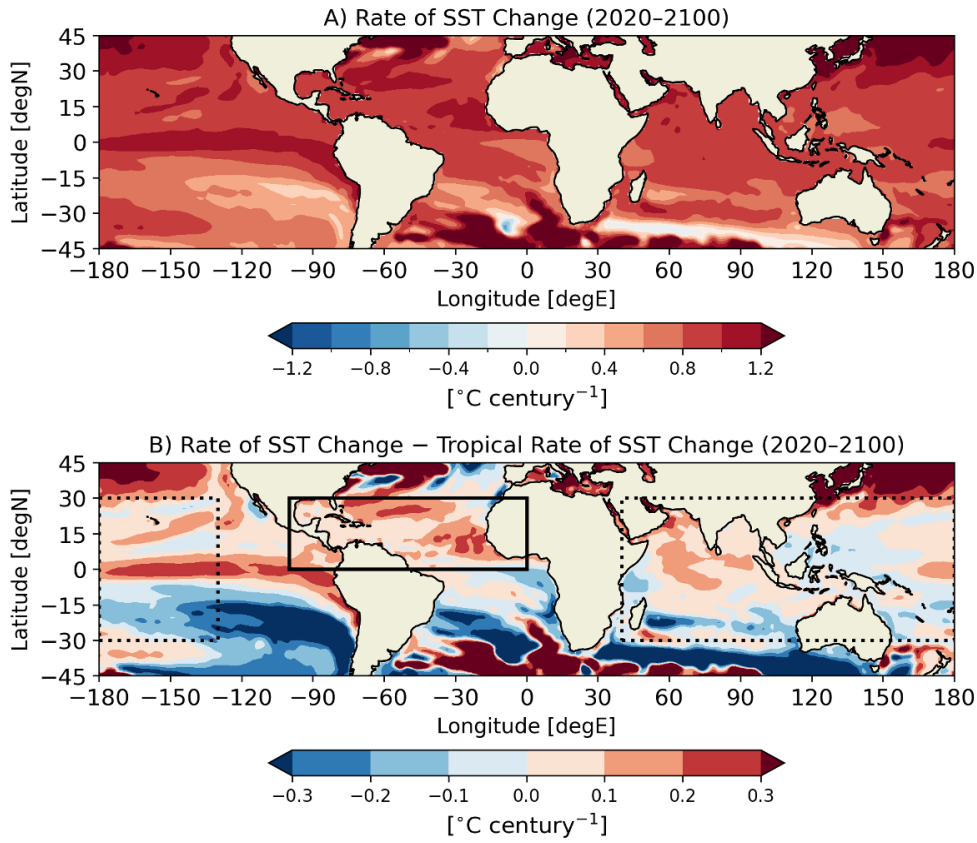
576 FIG. 6. 10-year rolling mean of TC-season basin-averaged χ^{-1} and PI for (a) AM2.5-C360 and (b) HIRAM. To
 577 clarify their relationship, both axes are oriented so that increasing values indicate more favorable conditions for
 578 TC activity. Blue and green dots represent historical (1871–2019) and projected (2020–2100) periods, and the
 579 *obs_sst* and *rcp4.5_flor* experiments, respectively, while orange and olive dots correspond to ERA5 (1979–2020)
 580 and MERRA5 (1980–2020) reanalysis data. Black and red dots denote the start and end of each period. Trend
 581 lines are plotted for each period, with solid lines indicating statistically significant trends ($p < 0.05$) and dashed
 582 lines otherwise. Positive (negative) trends indicate periods where PI and χ effects reinforce (offset) each other.



591 FIG. 7. Similar to Figure 5, but for the thermodynamic variables: (a) χ and (b) PI. Dotted trend lines are fitted
 592 for PI over the distinct periods with opposite trends (1871-1990 with a decreasing trend and 1990-2100 with an
 593 increasing trend) in both models.



617 FIG. 8. (a) The rate of historical observed TC season SST changes from 1871 to 2019, with SSTs sourced from
 618 Chan et al. (2021). (b) The difference between (a) and the mean tropical TC season SST change rate over the
 619 same period. In (b), the solid black box highlights the tropical Atlantic basin, while the dotted black box marks
 620 the Indo-Pacific Warm Pool region, as defined by Weller et al. (2016), which typically encompasses the warmest
 621 ocean surface waters on Earth.



622 FIG. 9. Same as Fig. 8, but for the future projected period 2020–2100. Both panels show projected SSTs used
 623 to force the *rcp45 flor* experiment, detailed in Section 2.

624 (ii) *Dynamic variables*

625 While the relationship between thermodynamic variables is a primary driver of the potential trend
 626 (or lack thereof) in $P(\Lambda)$, we also examine the role of the ventilation index (Eq. 11), specifically
 627 vertical wind shear, as a potential contributor to changes in $P(\Lambda)$. Figure 10b presents the time
 628 series of vertical wind shear across both models throughout the entire record. Both models simulate
 629 similar magnitudes of wind shear in the historical and projected periods. During the historical
 630 period, especially from 1950 to 2000, wind shear increases, making conditions progressively less
 631 favorable for storm activity. In contrast, the projected period shows no clear trend in wind shear: in
 632 AM2.5-C360, wind shear remains relatively stable, whereas in HIRAM, it exhibits multidecadal and
 633 decadal variability. In HIRAM’s projected period, wind shear is elevated between approximately

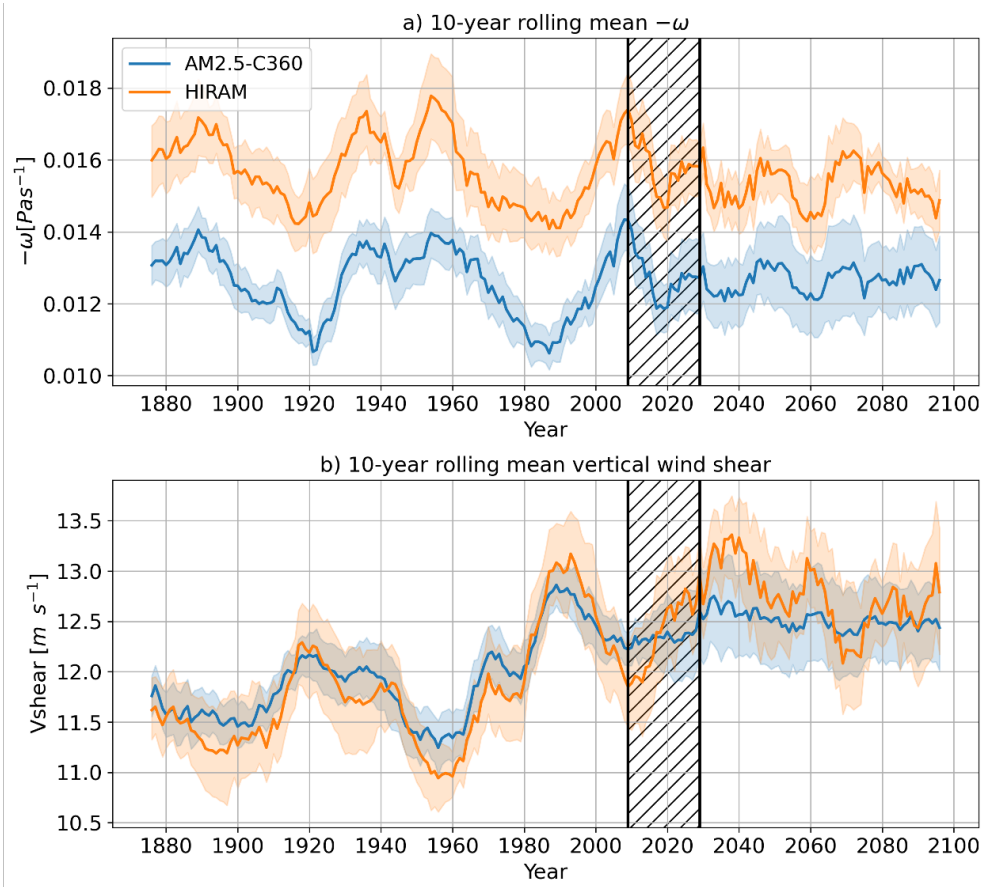


FIG. 10. Similar to Fig. 7, but for the dynamic variables: (a) $-\omega$ and (b) vertical wind shear.

634 2020 and 2070 and is characterized by decadal-scale fluctuations, with relative maxima around
 635 2050, 2060, and 2080 and minima near 2050 and 2070.

636 The increasing wind shear during the historical period, especially during the 1960–1990 period,
 637 likely amplified the already declining TC favorability, as both thermodynamic and dynamic con-
 638 ditions became less conducive to storm development. In the projected period, however, where
 639 thermodynamic variables counterbalance each other, wind shear has the potential to influence the
 640 overall trend in $P(\Lambda)$. Yet, since wind shear remains largely steady, $P(\Lambda)$ also remains unchanged
 641 in the projected period.

642 To investigate the physical mechanisms underlying changes in vertical wind shear in the tropical
 643 Atlantic TC basin, particularly the increase during 1960–1990 and the relative stability in the future
 644 projections, we examine ensemble-mean, zonally averaged atmospheric temperature changes in
 645 both models (Figure 11). Because our analysis focuses on seasonal and ensemble means over

646 the broad tropical Atlantic region (10°N–30°N), we assume the large-scale flow is in approximate
 647 geostrophic balance. Under this assumption, the thermal-wind relationship provides a useful
 648 approximation for the zonal component of vertical wind shear in terms of the meridional temperature
 649 gradient. We approximate the zonal component of the vertical wind shear between 850 hPa and
 650 250 hPa using the thermal-wind balance:

$$u(250) - u(850) = -\frac{R}{f} \ln\left(\frac{850}{250}\right) \frac{\partial \bar{T}}{\partial y}, \quad (13)$$

651 where $u(250) - u(850)$ represents the approximated zonal wind shear between 250 hPa and 850 hPa,
 652 R is the gas constant for dry air, f is the Coriolis parameter, and $\partial \bar{T} / \partial y$ is the meridional gradient
 653 of the layer-mean temperature.

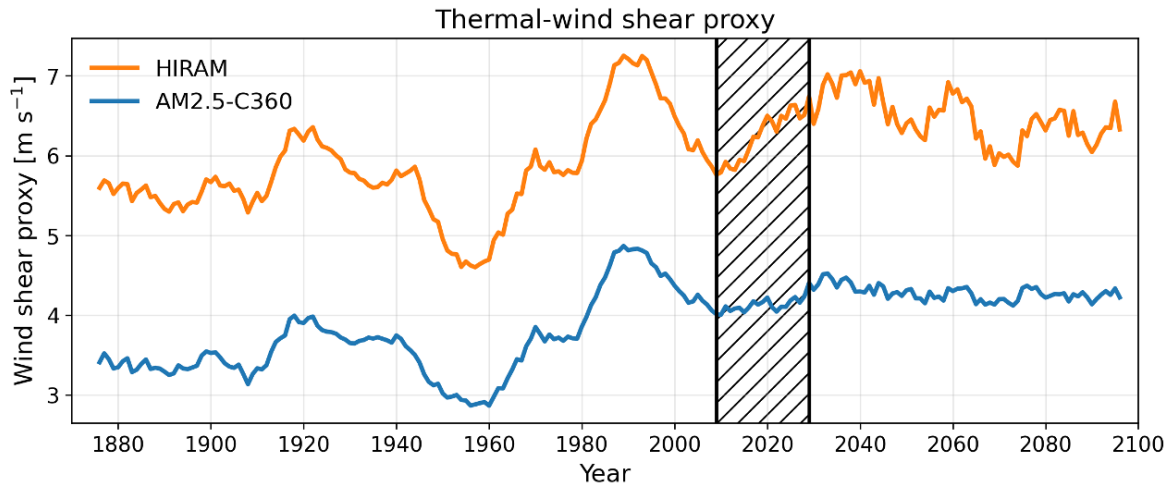
654 The layer-mean temperature \bar{T} is defined as the log-pressure–weighted mean temperature between
 655 850 hPa and 250 hPa:

$$\bar{T} = \frac{1}{\ln\left(\frac{850}{250}\right)} \int_{250}^{850} T(p) d \ln p, \quad (14)$$

656 where $T(p)$ denotes temperature at pressure level p . This definition is equivalent to the temperature
 657 that governs the thickness of the 850–250 hPa layer and therefore directly controls the vertical wind
 658 shear through the thermal-wind relationship. We compute this shear proxy using ensemble-
 659 mean, seasonally averaged, and zonally averaged temperature profiles from each model during
 660 both the historical and future projection periods. The meridional temperature gradient is then
 661 evaluated and averaged over 10°N–30°N using cosine-latitude weighting. Both models resolve
 662 11 atmospheric pressure levels between 850 hPa and 250 hPa, allowing for a robust calculation
 663 of the log-pressure–weighted mean temperature. The resulting time series of the ensemble-mean
 664 thermal-wind shear proxy for both models is shown in Figure 11.

665 First, we note that the magnitude of the thermal-wind shear proxy differs by approximately an
 666 order of magnitude from the explicitly simulated vertical wind shear shown in Figure 10, and that
 667 the proxy magnitudes also differ between the two models. In AM2.5-C360, the proxy shear ranges
 668 from approximately 3–4 m s⁻¹, while in HIRAM it ranges from roughly 5–7 m s⁻¹. In contrast,
 669 the vertically resolved wind shear simulated directly by both models ranges from approximately
 670 11–13 m s⁻¹.

671 Despite these differences in magnitude, the temporal evolution of the shear proxy closely tracks
672 that of the explicitly simulated vertical wind shear. Both the proxy and the simulated shear exhibit
673 predominantly multidecadal variability during the historical period, including a pronounced peak
674 between 1910 and 1950 and a marked increase during 1960–1990, features that are evident in both
675 models. The projected evolution of the shear proxy similarly mirrors that of the explicitly simulated
676 shear. In HIRAM, both records display a clear multidecadal maximum from 2020–2050, followed
677 by decadal-scale peaks during 2050–2070 and 2070–2090. In contrast, the AM2.5-C360 future
678 projections show relatively stable wind shear in both the proxy and the explicitly simulated fields,
679 with little evidence of pronounced decadal or multidecadal variability. Together, these results
680 suggest that changes in the spatial structure of meridional temperature gradients largely govern
681 the decadal and multidecadal evolution of vertical wind shear in both models. More broadly, they
682 reinforce the conclusion that spatial patterns of temperature change, rather than basin-wide mean
683 warming alone, play a central role in modulating environmental conditions, such as vertical wind
684 shear, that influence TC development.



685 FIG. 11. Ten-year rolling mean time series of the thermal-wind shear proxy defined in Eq. 13 for AM2.5-C360
 686 (blue) and HIRAM (orange). Lines show the ensemble-mean zonal shear between 250 and 850 hPa averaged over
 687 10°–30°N. The historical segment (1871–2019) is derived from the *obs_sst* simulations, while the future segment
 688 (2020–2100) is derived from the *rcp4.5_sst* simulations. The two periods are concatenated to form a continuous
 689 time series. Values to the left of the hatched region correspond to simulations forced with observed SSTs
 690 (*obs_sst*), while values to the right correspond to simulations forced with observed climatology plus projected
 691 SST patterns (*rcp45_flor*). The hatched region marks a smooth transition that blends both forcing regimes.

692 Next, because changes in SPI are primarily driven by variations in $-\omega$, we analyze the temporal
693 evolution of this variable in Figure 10a. Just as SPI exhibits no clear trend across the full record
694 in either model (Fig. 5b), $-\omega$ similarly shows no discernible long-term trend. However, both
695 variables display pronounced multidecadal variability in the historical period, with a maxima from
696 1930-1970 and minima from 1900-1930 and 1960-2010. In the future projections, we observe
697 decadal-scale peaks and troughs. Additionally, while the typical magnitude of $-\omega$ is greater in
698 HIRAM than in AM2.5-C360, both models exhibit comparable temporal variability. As detailed
699 in Appendix Section c, since vorticity does not play an important role in driving TC changes, we
700 do not include a discussion of this dynamic variable here.

701 4. Discussion and conclusions

702 As shown in Figure 4, the TC framework described in Hsieh et al. (2020) ($SPI \times P(\Lambda)$) suc-
703 cessfully reproduces Atlantic TC variability in simulations forced with observed SSTs, as well
704 as in projections under an RCP4.5 SST increase. Its performance is slightly stronger in HIRAM
705 than in AM2.5-C360, particularly during projected period, during which the theoretical framework
706 in AM2.5-C360 tends to underestimate the projected TCs (Figures 3 and 4). Importantly, the
707 framework achieves this skill without calibration or tuning, underscoring its ability to capture TC
708 variability directly from local large-scale environmental conditions. This is particularly noteworthy
709 in the Atlantic basin, where TCs frequently originate from nonlocal disturbances such as African
710 easterly waves (AEWs; Emanuel (2022)). In contrast to other widely used TC proxies such as
711 genesis potential indices (GPIs), the framework presented by Hsieh et al. (2020) explicitly incorpo-
712 rates the initial seed disturbances associated with AEWs that initiate TC formation (Cavicchia et al.
713 2023). Most GPIs require calibration prior to application in future climate scenarios (Camargo
714 et al. 2014) and exhibit limited skill in reproducing interannual and multidecadal TC variabil-
715 ity (Cavicchia et al. 2023). Together, these findings highlight the robustness of the $SPI \times P(\Lambda)$
716 framework for assessing Atlantic TC variability across both historical and projected climates.

717 Using the theoretical framework and explicitly tracked TCs from a suite of idealized simulations
718 in both AM2.5-C360 and HIRAM, we assess the relative roles of regional SST pattern changes and
719 global mean warming in modulating Atlantic TC activity during both the historical and projected
720 periods. During the historical period, we compare an experiment that includes both global mean

721 and patterned warming (*obs_sst*) with an experiment that isolates patterned warming and removes
722 the global mean signal (*rm_warming*). We find that both the tracked TC counts and the framework-
723 derived TC variability are indistinguishable between these experiments, indicating that global mean
724 warming played a minimal role in historical Atlantic TC variability. Instead, regional SST patterns
725 dominated past variations in storm activity.

726 In the future projected period, we evaluate TC sensitivity to global mean warming by contrasting
727 an experiment that includes both patterned and mean warming (*rcp45_flor*) with experiments that
728 include only uniform global mean warming (*cntl* and *2K_2CO2*). Across both models and TC rep-
729 resentations (explicit tracking and the theoretical framework), we identify statistically significant
730 differences in TC sensitivity to mean warming between experiments that do and do not include
731 patterned warming. In HIRAM, in particular, the sign of the TC trend differs between experi-
732 ments: when both patterned and mean warming are present, TC counts exhibit no statistically
733 significant trend, whereas under uniform mean warming alone, TC counts decline. Together, these
734 results underscore the critical role of regional warming patterns in shaping Atlantic TC activity.
735 Assessments of future TC changes cannot rely solely on global mean warming trends; accounting
736 for spatial patterns of warming is essential for drawing robust statistical conclusions about TC
737 sensitivity to climate change.

738 Our results emphasizing the importance of regional SST warming patterns in driving changes
739 in Atlantic TC activity are consistent with the findings of Vecchi et al. (2019), Villarini et al.
740 (2011b), Knutson et al. (2015), Vecchi and Soden (2007). These studies further suggest that SST
741 spatial patterns do not exert a first-order influence on TC responses to warming in other ocean
742 basins. They argue that this distinction arises because relative SST anomalies in the North Atlantic
743 under transient $2\times\text{CO}_2$ forcing in AM2.5-C360 and HIRAM are substantially larger than those in
744 other TC basins. We find similar behavior in the future SST projections used in the *rcp45_flor*
745 experiment, which are derived from the FLOR RCP4.5 model. As shown in Figure 13, the projected
746 rate of relative SST change in the Atlantic main development region is markedly larger ($+0.25\text{ }^\circ\text{C}$
747 per century) than in other basins, including the western Pacific ($-0.05\text{ }^\circ\text{C}$ per century) and the
748 Indian Ocean ($+0.10\text{ }^\circ\text{C}$ per century). Together, these results support the conclusion that regional
749 warming patterns play a particularly critical role in shaping TC variability in the Atlantic, with
750 a weaker influence in other ocean basins where relative SST contrasts are smaller. This pattern

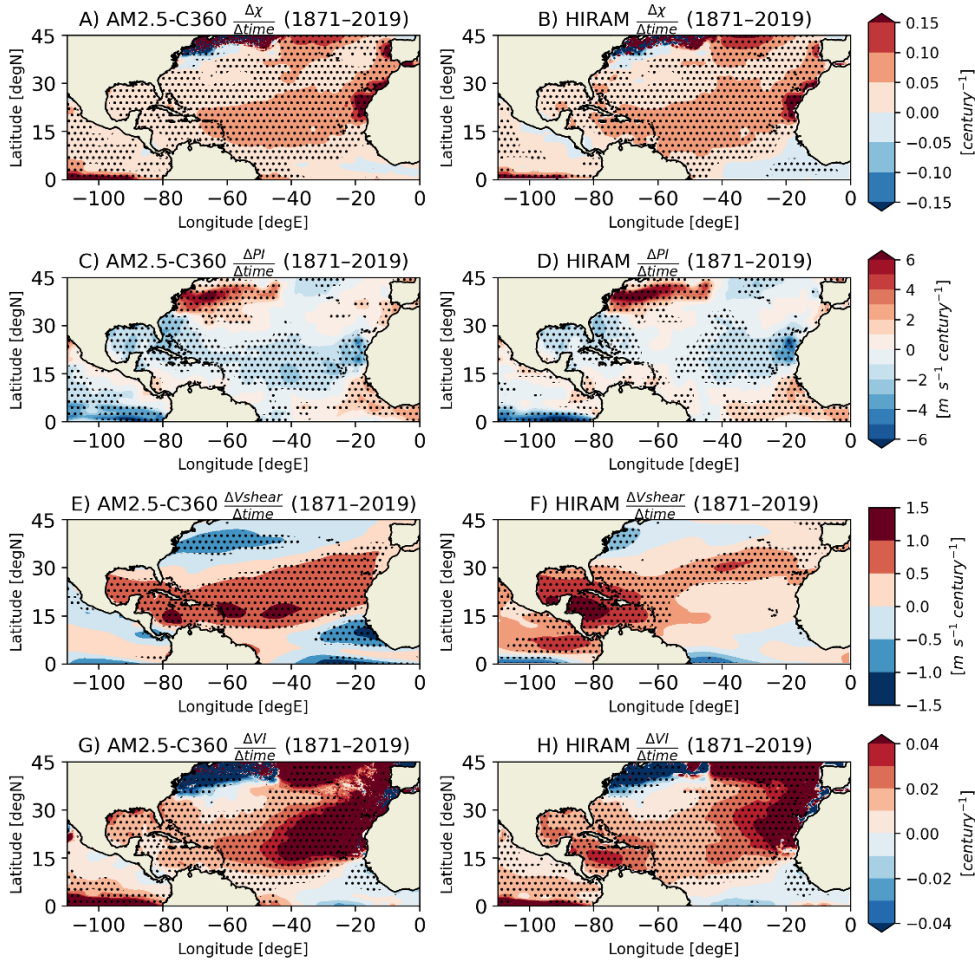
751 likely includes a radiatively-forced component, but there are currently large disagreements about
752 the sign and magnitude of such forcing (e.g. Zhuo et al. (2025)).

753 Next, we analyze how the large-scale dynamic and thermodynamic variables contribute to changes
754 in parameterized TCs during both the historical and projected periods. Examining the temporal
755 evolution of SPI (Fig. 5b), we find no clear long-term trend. Instead, SPI is governed by
756 multidecadal and decadal fluctuations, largely modulated by variability in $-\omega$. In contrast, the
757 temporal evolution of $P(\Lambda)$ is characterized by a distinct decreasing trend during the historical
758 period, followed by a relatively stable plateau in the projected period (Fig. 5a). Notably, from
759 2020-2100 $P(\Lambda)$ exhibits little decadal variability, particularly in AM2.5-C360. The difference
760 in $P(\Lambda)$'s variability between both record periods is primarily driven by the relationship between
761 its thermodynamic components: moist entropy deficit (χ) and potential intensity (PI). During the
762 historical period, particularly from 1871-1990, PI and χ had a compounding effect, both becoming
763 increasingly unfavorable for TC activity and reinforcing the decline in $P(\Lambda)$. However, in the
764 future projection with FLOR SSTs, their opposing influences cancel out, leading to a stable $P(\Lambda)$
765 with no clear trend.

766 This shift in the thermodynamic relationship is evident in spatial analyses, where we observe a
767 statistically significant increasing trend in Λ during the historical period (Fig. 12), thus becoming
768 less favorable for TCs, whereas no such trend is apparent in the projected period (Fig. 13). This
769 highlights how the changing interaction between PI and χ fundamentally alters the behavior of Λ ,
770 $P(\Lambda)$, and parametrized TCs, over time.

774 The reversal in the relationship between χ and PI across the historical and projected periods
775 arises from a shift in PI trends (Figure 7). While χ increases steadily throughout both periods, PI
776 declines during the historical interval from 1871 to 1990 and increases thereafter. In the projected
777 period, both PI and moisture deficit increase, consistent with prior studies that project rising global
778 PI and moisture deficit through the late 21st century (e.g., Emanuel et al. (2008), Emanuel (2021b),
779 Lee et al. (2020)).

780 Our historical results also show a pronounced increase in PI from 1990 to 2020 (Figure 7), in
781 agreement with Kossin and Camargo (2009) and Emanuel (2021a), who analyze PI evolution using
782 reanalysis products and likewise identify a sharp post-1980/1990 increase. The NCEP/NCAR
783 reanalysis used in Kossin and Camargo (2009) spans 1948–2008 and therefore captures only the



771 FIG. 12. Linear regression coefficients for the century-scale TC-season rate of change in (a, b) χ , (c, d) PI, (e,
 772 f) vertical wind shear, and (g, h) ventilation index (Λ) during the historical period (1871–2019) for AM2.5-C360
 773 and HIRAM. Dotted regions indicate statistically significant trends ($p < 0.05$).

784 latter portion of our multi-century record; nevertheless, both studies exhibit decreasing PI from
 785 roughly 1950 to 1990 followed by an increase thereafter. Emanuel (2021a), using the ECMWF
 786 CERA-20C reanalysis (Laloyaux et al. 2018) covering 1901–2010, find multidecadal PI variability
 787 from 1901 to 1980 without a clear long-term trend, with a peak near 1950 and troughs around 1920
 788 and 1970. Our results similarly show decadal PI variability over this period, with troughs near 1920
 789 and 1970, but with a peak closer to 1960 and an overall declining trend from 1901 to 1980. While
 790 both studies agree on the timing of PI minima and the rapid increase after 1980-1990, differences

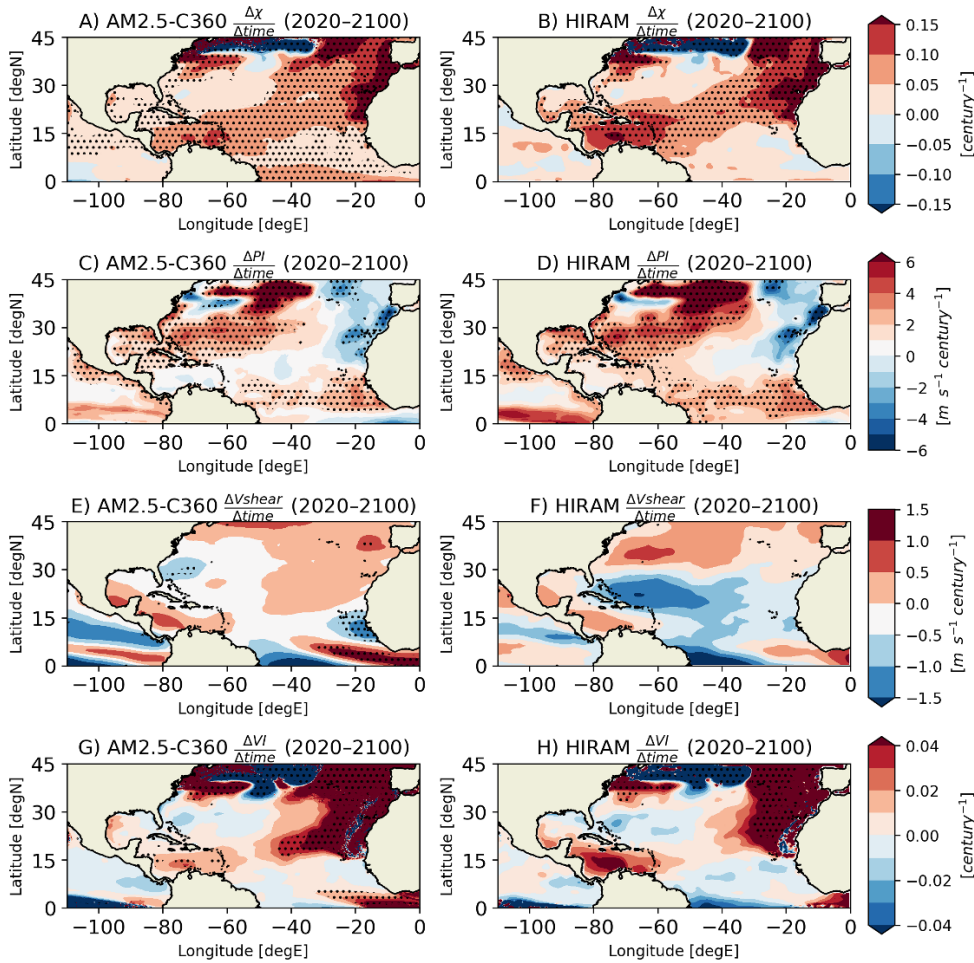


FIG. 13. Same as Fig. 12 but for the projected period (2020-2100).

791 in peak timing and long-term trend prior to 1980 likely reflect biases in reanalysis products during
 792 the pre-satellite era.

793 We attribute the reversal in PI trends between the historical and projected periods to differences
 794 in the relative warming rates of the tropical Atlantic and the broader tropics, which are known
 795 to be important in modulating TC activity (Vecchi and Soden (2007); Villarini et al. (2010)).
 796 During the historical period, the tropical Atlantic warmed more slowly than the rest of the tropics,
 797 likely weakening the temperature contrast between the sea surface and the upper troposphere.
 798 Conversely, in the projected period, Atlantic SSTs are expected warm faster than the rest of the
 799 tropics, enhancing the surface-to-upper troposphere temperature contrast and fostering conditions

800 more favorable for increases in PI and TC activity. This underscores the critical role of regional
801 SST patterns, rather than just mean warming trends, in shaping Atlantic TC activity.

802 These findings highlight the deep uncertainty in future TC projections (Knutson et al. 2020), a
803 challenge that dates back to the earliest model-based study of TC frequency (Broccoli and Manabe
804 1990). The strong compensation between thermodynamic variables, consistent with results from
805 downscaled CMIP6 simulations (Emanuel 2021b), suggests that these variables, directly influenced
806 by warming, are unlikely to be the primary drivers of TC activity changes. Instead, dynamic
807 variables such as vertical wind shear, which are indirectly linked to warming through circulation
808 changes, emerge as key determinants of future TC trends. However, because these dynamic
809 variables are not directly constrained by warming, their future behavior remains highly uncertain
810 (Pfahl et al. 2017).

811 Regional temperature patterns play a fundamental role in shaping both thermodynamic and
812 dynamic variables, including wind shear and $-\omega$, through thermal wind balance and by setting the
813 most active convective regions. Stronger meridional temperature gradients enhance vertical wind
814 shear, while high relative SST patterns intensify deep tropical convection. Ultimately, regional
815 SST patterns and the processes that control the vertical structure of warming may greatly influence
816 changes in the dynamic variables, reinforcing the importance of spatial temperature distributions
817 in determining future TC activity.

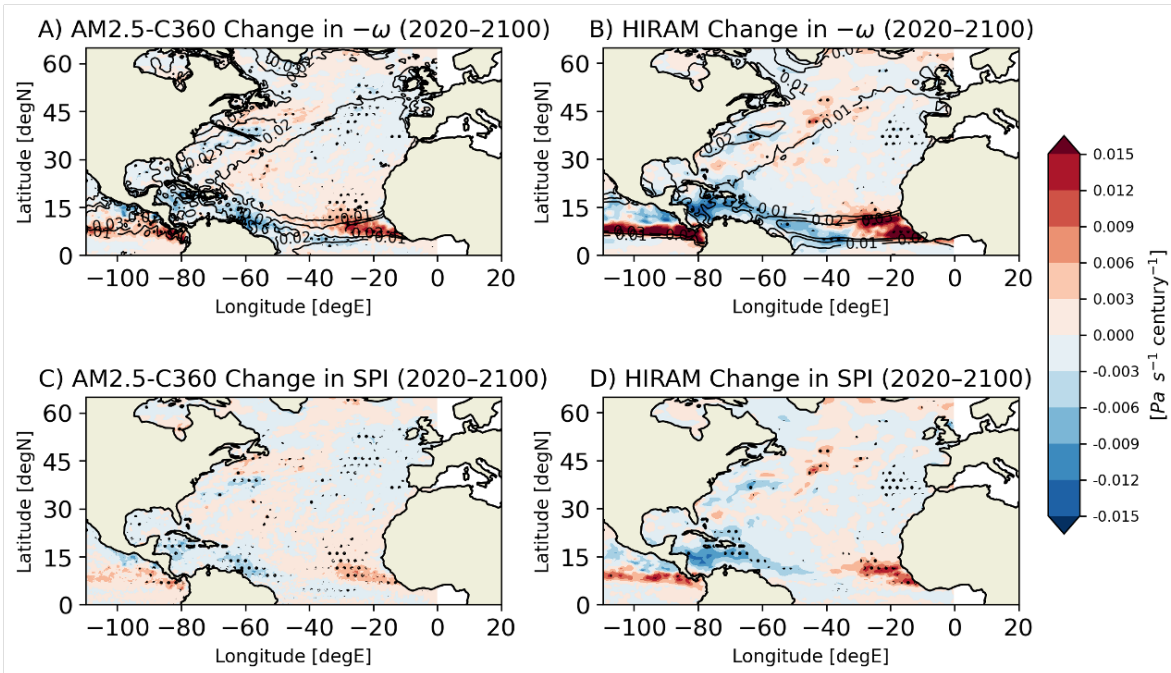
818 Building on our argument that the compensation of thermodynamic variables enables dynamic
819 factors to influence TC activity, we observe that the trend in $P(\Lambda)$ closely mirrors that of vertical
820 wind shear in model simulations, which is primarily determined by temperature patterns. As shown
821 in Figure 11, vertical wind shear increases during the historical period, driven by a strengthened
822 atmospheric meridional temperature gradient. However, these trends do not persist in future
823 projections (Figure 13e), and substantial inter-model differences arise. For example, AM2.5-C360
824 forecasts a reduction in shear along the U.S. East Coast, whereas HIRAM predicts an increase in the
825 same region. Since neither model shows a consistent basin-wide trend in shear, the projected trend
826 in $P(\Lambda)$ remains similarly uncertain. This ambiguity in shear trends is likely due to differences in
827 projected atmospheric temperature.

828 The dominant role of thermodynamic variable compensation in regulating $P(\Lambda)$ reinforces the
829 idea that, given the uncertainty in $P(\Lambda)$ projections, future TC activity will be primarily modulated

830 by changes in SPI, which represents the number of basin-wide seeds. This aligns with previous
831 studies such as Hsieh et al. (2020), Vecchi et al. (2019), Sugi et al. (2020), and Yamada et al. (2021),
832 which emphasize the critical role of TC seeds in shaping future TC activity. Our findings support
833 this conclusion, as the projected stability of $P(\Lambda)$ (Fig. 5a) suggests that the decadal variability of
834 parametrized TCs ($SPI \times P(\Lambda)$; Fig. 4) largely follows the variability of SPI (Fig. 5) and $-\omega$ (Fig.
835 10a).

836 While some previous studies suggest that the global number of TC seeds may decrease in a
837 warmer climate (Hsieh et al. (2020), Vecchi et al. (2019), Sugi et al. (2020), Yamada et al. (2021)),
838 our findings indicate that SPI, our seed proxy, does not exhibit a clear Atlantic basin-wide trend
839 in either the historical or projected periods. Instead, Atlantic SPI is characterized by decadal
840 and multidecadal variability driven by fluctuations in $-\omega$. Although no basin-wide trends in $-\omega$
841 emerge, we identify localized trends within the basin, particularly along the ITCZ and within the
842 TC main development region (Fig. 14). Specifically, an eastward shift in $-\omega$ is evident in the
843 deep tropics, with a decreasing trend in the western basin and an increasing trend in the eastern
844 basin, which could be as result higher rate of warming in the eastern tropical Atlantic than the
845 western tropical Atlantic (Fig. 9b). This pattern suggests a potential eastward shift in TC genesis
846 under warming conditions (Kortum et al. (2024)), and mirrors some observed eastward shifts in
847 Atlantic TC tracks (Colbert and Soden 2012; Vecchi and Knutson 2011). However, significant
848 uncertainties remain in $-\omega$, as vertical velocity and seed development could be influenced by
849 radiative feedbacks linked to cloud processes or unresolved mechanisms, as described in the gross
850 moist stability framework (Hsieh et al. (2023), Neelin and Held (1987)).

856 It is important to recognize the inter-model differences present in our results, such as variations in
857 the framework's ability to capture TC activity in idealized simulations and discrepancies in future
858 projections of vertical wind shear, which we suggest have to do with differences in the pattern
859 of warming and TC seed distribution. This study examines only two atmosphere-only models,
860 both of which have been shown to predict the highest number of future TCs compared to other
861 model simulations (Bhatia et al. (2018), Hsieh et al. (2022)). However, as noted by Knutson et al.
862 (2020), most dynamical TC-resolving models project a decrease in storm activity in a warming
863 climate. Furthermore, because the models used here are atmosphere-only and forced exclusively
864 by SSTs, they may highlight the limitations of relying on SSTs alone for TC predictions, an issue



851 FIG. 14. Linear regression coefficients for the century-scale TC-season rate of change in (a, b) $-\omega$ and (c, d)
 852 SPI during the projected period (2020–2100) for AM2.5-C360 and HIRAM. Dotted regions indicate statistically
 853 significant trends ($p < 0.05$). In (a) and (b), black contours show the mean TC-season $-\omega$ during the historical
 854 period (1871–2019). Red (blue) shading represents regions with increasing ascent (descent) in the projection
 855 period.

856 that became particularly evident during the 2020 TC season (Fig. 1). As suggested by Kortum
 857 et al. (2024), factors such as atmospheric chaos, aerosols, and other environmental influences could
 858 play a critical role in modulating TC activity, potentially introducing biases in SST-forced models.

859 Beyond the aforementioned caveats, our results highlight that regional SST and atmospheric
 860 temperature patterns, rather than the global-mean warming trend alone, play a central role in shaping
 861 both historical and future Atlantic tropical cyclone (TC) activity. In particular, the relationship
 862 between PI and χ is strongly modulated by the relative rate of warming in the tropical Atlantic
 863 compared to the broader tropics. Likewise, the behavior of key dynamical variables is governed
 864 by spatial temperature gradients rather than uniform warming.

865 Substantial uncertainties remain in projections of future SST patterns, especially in the Atlantic
 866 and other tropical basins such as the eastern Pacific (Seager et al. (2022); Shaw et al. (2024)). In

876 the Atlantic, for example, the magnitude of future Atlantic Meridional Overturning Circulation
877 (AMOC) weakening is poorly constrained, with potentially important implications for regional
878 SST evolution (Cheng et al. (2013); Srokosz and Bryden (2015)). More broadly, most climate
879 models project a shift toward an El Niño–like mean state, characterized by enhanced warming
880 in the eastern Pacific, whereas observations suggest a tendency toward a La Niña–like state with
881 comparatively weaker eastern Pacific warming (Zhuo et al. 2025). Consistent with the multi-model
882 mean, our simulations exhibit an El Niño–like shift (Figure 13), with eastern Pacific SSTs warming
883 more rapidly than the rest of the basin.

884 Changes in the mean state of the tropical Pacific may influence Atlantic TC seed formation;
885 for example, Hsieh et al. (2022) show that La Niña conditions are more favorable for Atlantic
886 SPI. However, the impact of Pacific SST changes on PI is likely more muted. This is because
887 the warmest tropical SSTs, typically associated with the western Pacific warm pool, largely set
888 upper-tropospheric temperatures across the tropics through convective adjustment, thereby exerting
889 a dominant control on PI. Consistent with this mechanism, Eusebi et al. (2025) demonstrate that
890 the relative difference between tropical Atlantic SSTs and the warmest tropical SSTs, rather than
891 the tropical-mean SST, provides a more robust predictor of Atlantic TC PI. Together, these results
892 underscore the importance of accurately representing regional temperature patterns across the
893 global tropics, as their influence on Atlantic TCs can arise through distinct physical pathways.
894 Improving the reliability of future TC projections therefore requires tighter constraints on projected
895 SST pattern changes across a diverse ensemble of climate models.

896 *Acknowledgments.* We thank Gabriele Villarini, Stephan Fueglistaler, Andrew Williams, Nadir
897 Jeevanjee, and Hiroyuki Murakami for discussion and comments. This work has been supported
898 in part by a grant from the Heising-Simons foundation and the Carbon Mitigation Initiative at
899 Princeton University, funded by BP. E.L. was supported by the National Science Foundation
900 Graduate Research Fellowship. The simulations were performed on computational resources
901 managed and supported by Princeton Research Computing, a consortium of groups including
902 the Princeton Institute for Computational Science and Engineering, the Office of Information
903 Technology’s High Performance Computing Center, and the Visualization Laboratory at Princeton
904 University.

905 *Data availability statement.* Source code of the HiRAM model is available from <https://www.gfdl.noaa.gov/hiram-quickstart>. ERA5 data are available from <https://cds.climate.copernicus.eu/datasets/reanalysis-era5-pressure-levels-monthly-means?tab=overview>, and MERRA2 data are available from https://gmao.gsfc.nasa.gov/gmao-products/merra-2/data-access_merra-2/.

910 APPENDIX

911 **Further evaluation of the seed-probability framework and its components**

912 *a. Verifying SPI Theory*

913 To assess whether the SPI framework can faithfully reproduce the climatology and variability
914 of tracked TC seeds in both models, we compare the spatial distribution of tracked seeds and SPI
915 during the historical period (Figure A1) and examine time series of tracked seeds and SPI for
916 both the historical and projected periods (Figure A2). We first examine the spatial climatology
917 of tracked seeds and SPI during the historical period. HIRAM exhibits a stronger sensitivity to
918 seed formation than AM2.5-C360: the climatological mean seed density in AM2.5-C360 ranges
919 from approximately 20–60 seed days per year per $10^\circ \times 10^\circ$ grid box, whereas in HIRAM it ranges
920 from roughly 5–15 seed days per year per $10^\circ \times 10^\circ$ grid box. As described in Section 2, the same
921 seed-tracking algorithm is applied to both models. Consequently, these differences arise from
922 distinct realizations of the atmospheric vorticity field, which is the key dynamical component of
923 the seed-tracking framework. Consistent with the seed climatology, the climatological mean SPI

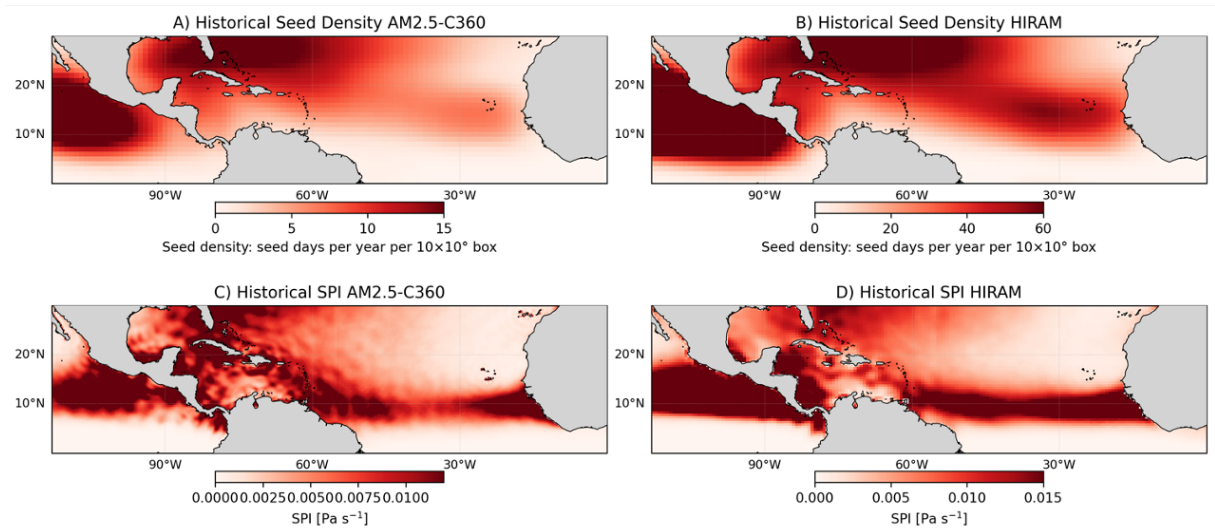
924 also exhibits stronger sensitivity in HIRAM than in AM2.5-C360, with typical SPI values in active
925 regions of 0.0075–0.0125 Pa s⁻¹ in HIRAM compared to 0.010–0.015 Pa s⁻¹ in AM2.5-C360.

926 Importantly, the SPI framework successfully reproduces the primary spatial features of seed
927 formation in the tropical Atlantic. In both models, SPI captures two prominent maxima: one in
928 the main development region near the West African coast between approximately 5°–15°N, and a
929 second in the central Atlantic between 25°–30°N. The framework also correctly identifies regions
930 of suppressed seed formation, including north of Colombia and Venezuela, south of Cuba, and in
931 the western Gulf of Mexico. In addition, SPI reproduces a pronounced minimum in seed activity
932 off the west coast of subtropical Africa between 15°–30°N. Together, these results demonstrate that
933 the SPI framework faithfully captures the observed spatial distribution of TC seeds in both models.

934 We next examine normalized time series of historical and future projected SPI and tracked
935 seed counts. In both models, SPI closely reproduces the decadal and multidecadal variability in
936 seed formation. Specifically, the framework captures the multidecadal peak in seed activity from
937 approximately 1920–1980, including an enhanced decadal maximum during 1950–1960. SPI also
938 reproduces several decadal peaks throughout both the historical and projected periods, including a
939 pronounced increase from 2000–2020 in both models. The strong correlation between ensemble-
940 mean SPI and tracked seed counts ($r = 0.86$ in AM2.5-C360 and $r = 0.73$ in HIRAM) further
941 demonstrates the ability of the SPI framework to capture variability in seed formation across the
942 extended record. Overall, the results shown in Figures A1 and A2 indicate that the SPI theory
943 provides a robust representation of TC seed formation in both models. This gives us confidence in
944 applying the full theoretical framework, $SPI \times P(\Lambda)$, to approximate TC counts.

960 *b. Covariance between seed and probability proxies*

961 We assess the correlation coefficient between the annual TC-season basin-averaged *SPI* and
962 $P(\Lambda)$ (Fig. A3) for both the observed historical period (1871-2019) and the projected future period
963 (2020-2100). Our analysis reveals that, across both models and time periods, the ensemble mean
964 of *SPI* and $P(\Lambda)$ exhibit a moderate correlation. Specifically, for each model, the ensemble mean
965 correlation between the proxies is consistent across the observed and projected periods ($r \sim 0.5$ for
966 AM2.5-C360 and $r \sim 0.7$ for HIRAM), with HIRAM generally showing stronger correlations. In
967 AM2.5-C360, the correlation varies more widely across ensemble members, ranging from $r = 0.25$

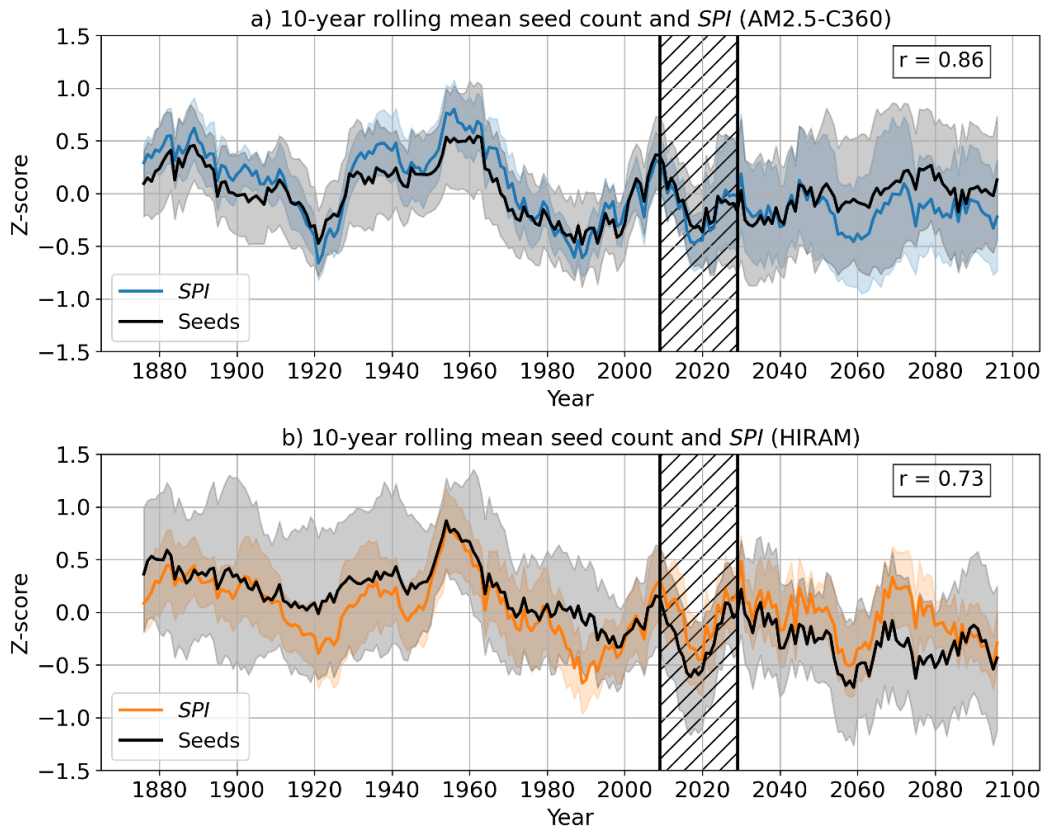


945 FIG. A1. Mean spatial density of Atlantic tropical cyclone seed locations during the historical period
 946 (1871–2019) in (a) AM2.5-C360 and (b) HIRAM. Density is defined as the annual mean number of seed
 947 days within each model grid cell, aggregated onto $10^\circ \times 10^\circ$ latitude–longitude bins. Panels (c) and (d) show
 948 the mean seasonally averaged (June–November) seed probability index (SPI) during the historical period for
 949 AM2.5-C360 and HIRAM, respectively.

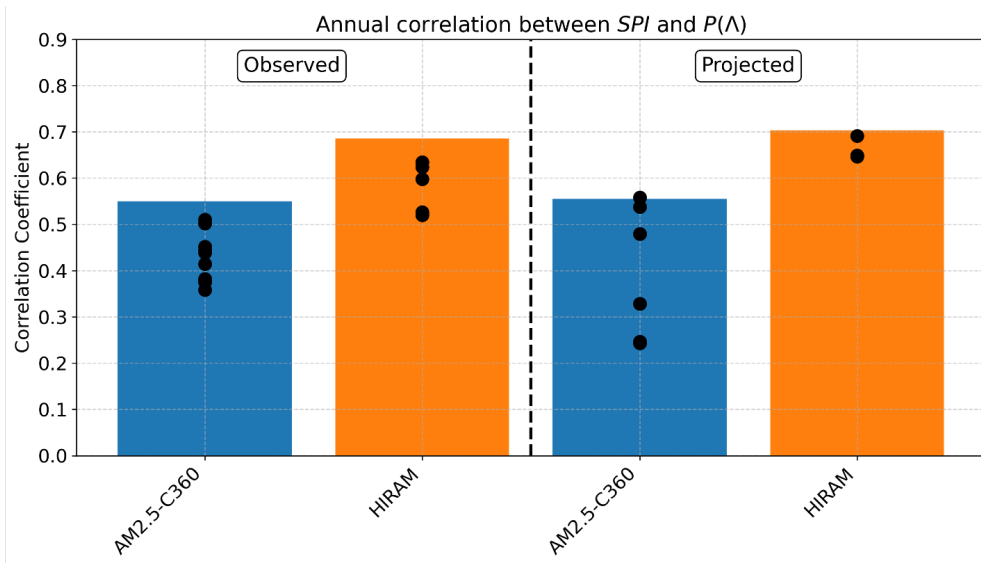
968 to $r = 0.55$, while in HIRAM, the spread is narrower, with correlations ranging from $r = 0.5$ to
 969 $r = 0.7$. Despite the variation in ensemble spread and inter-model differences, it is evident that
 970 SPI and $P(\Lambda)$ are not independent, but rather covary in a manner that reflects a spatio-temporal
 971 relationship, with their combined behavior serving as a proxy for TC activity.

975 *c. Vorticity*

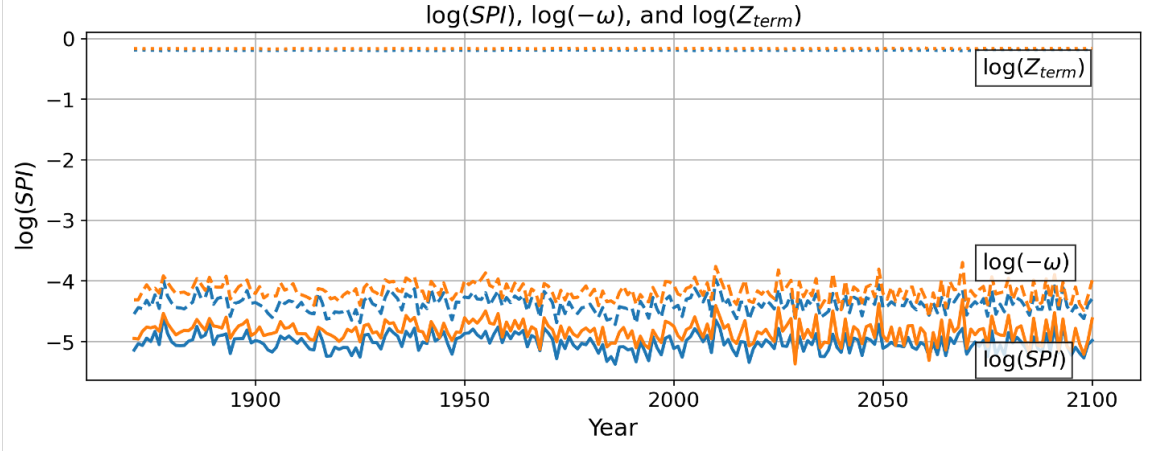
976 To show that changes in the vorticity term of the SPI have a negligible impact on variations in TC
 977 activity, we perform a logarithmic decomposition of the SPI , allowing the terms to be expressed
 978 additively:



950 FIG. A2. Normalized 10-year rolling mean of simulated Atlantic TC seed frequency (black) and parameterized
 951 seed frequency (SPI; colored) for (a) AM2.5-C350 (blue) and (b) HIRAM (orange). Shaded regions indicate
 952 the bootstrapped 95% confidence intervals as in Fig. 4. The historical period (1871–2019) is based on the
 953 obs SST simulation, whereas the future period (2020–2100) is derived from the rcp4.5 SST simulation. These
 954 two periods are combined into a continuous time series, with both simulated and parameterized TC frequencies
 955 standardized using a Z-score normalization, as in Fig. 4. Correlation coefficients between the ensemble mean
 956 simulated and parameterized TC frequencies are displayed in the top right corner of each panel. Values to the
 957 left of the hatched region correspond to simulations forced with observed SSTs (obs_sst), while values to the
 958 right correspond to simulations forced with observed climatology plus projected SST patterns (rcp45_flor). The
 959 hatched region marks a smooth transition that blends both forcing regimes.



972 FIG. A3. Annual TC-season basin mean correlation coefficient between SPI and $P(\Lambda)$ in AM2.5-C360 (blue)
 973 and HIRAM (orange) in the *obs SST* simulation (left) and *rcp4.5 SST* simulation (right). Bars indicate the
 974 ensemble mean correlation, while black dots represent individual ensemble members.



987 FIG. A4. Annual ensemble mean seasonally and basin-averaged $\log(SPI)$ (solid lines), $\log(-\omega)$ (dashed lines),
 988 and $\log(Z_{term})$ (dotted lines) in AM2.5-C350 (blue) and HIRAM (orange). The historical period (1871–2019)
 989 is based on the obs SST simulation, whereas the future period (2020–2100) is derived from the rcp4.5 SST
 990 simulation. These two periods are combined into a continuous time series.

$$\begin{aligned}
 SPI &= \underbrace{-\omega}_{\text{vertical velocity}} \times \underbrace{\frac{1}{1 + Z^{-1/0.69}}}_{\text{vorticity term } (Z_{term})} \\
 &= -\omega \times Z_{term}
 \end{aligned} \tag{A1}$$

$$\log(SPI) = \log(-\omega) + \log(Z_{term}).$$

979 This decomposition demonstrates that changes in $\log(SPI)$ are driven equally by additive contribu-
 980 tions from changes in $\log(-\omega)$ and $\log(Z_{term})$. To assess these contributions over time, we analyze
 981 fluctuations in all three terms across the historical and projected record (Fig. A4). We find that
 982 $\log(Z_{term})$ exhibits minimal annual and decadal variability. Additionally, interannual and decadal
 983 variations in $\log(SPI)$ closely track those in $\log(-\omega)$, indicating that changes in SPI —and con-
 984 sequently in parameterized TC activity—are primarily driven by fluctuations in vertical velocity
 985 rather than vorticity. Given this negligible role of vorticity, we largely disregard its influence on
 986 approximated TC activity in this report.

References

- Bhatia, K., G. A. Vecchi, H. Murakami, S. Underwood, and J. P. Kossin, 2018: Projected response of tropical cyclone intensity and intensification in a global climate model. *Journal of Climate*, **31** (20), 8281–8298, <https://doi.org/10.1175/JCLI-D-17-0898.1>.
- Booth, B. B. B., N. J. Dunstone, P. R. Halloran, T. Andrews, and N. Bellouin, 2012: Aerosols implicated as a prime driver of twentieth-century north atlantic climate variability. *Nature*, **484** (7393), 228–232, <https://doi.org/10.1038/nature10946>.
- Bretherton, C. S., J. R. McCaa, and H. Grenier, 2004: A new parameterization for shallow cumulus convection and its application to marine subtropical cloud-topped boundary layers. part i: Description and 1d results. *Monthly Weather Review*, **132** (4), 864–882, [https://doi.org/10.1175/1520-0493\(2004\)132<0864:ANPFSC>2.0.CO;2](https://doi.org/10.1175/1520-0493(2004)132<0864:ANPFSC>2.0.CO;2).
- Broccoli, A. J., and S. Manabe, 1990: Can existing climate models be used to study anthropogenic changes in tropical cyclone climate? *Geophysical Research Letters*, **17** (11), 1917–1920, <https://doi.org/10.1029/GL017i011p01917>.
- Bruyère, C. L., G. J. Holland, and E. Towler, 2012: Investigating the use of a genesis potential index for tropical cyclones in the north atlantic basin. *Journal of Climate*, **25** (24), 8611–8626, <https://doi.org/10.1175/JCLI-D-11-00619.1>.
- Camargo, S. J., M. K. Tippett, A. H. Sobel, G. A. Vecchi, and M. Zhao, 2014: Testing the performance of tropical cyclone genesis indices in future climates using the hiram model. *Journal of Climate*, **27** (24), 9171–9196, <https://doi.org/10.1175/JCLI-D-13-00505.1>.
- Camargo, S. J., and Coauthors, 2020: Characteristics of model tropical cyclone climatology and the large-scale environment. *Journal of Climate*, **33** (11), 4463–4483, <https://doi.org/10.1175/JCLI-D-19-0500.1>.
- Cavicchia, L., E. Scoccimarro, G. Ascenso, A. Castelletti, M. Giuliani, and S. Gualdi, 2023: Tropical cyclone genesis potential indices in a new high-resolution climate models ensemble: Limitations and way forward. *Geophysical Research Letters*, **50** (11), e2023GL103001, <https://doi.org/10.1029/2023GL103001>.

- 1018 Chan, D., G. A. Vecchi, W. Yang, and P. Huybers, 2021: Improved simulation of 19th- and 20th-
1019 century north atlantic hurricane frequency after correcting historical sea surface temperatures.
1020 *Science Advances*, **7** (26), eabg6931, <https://doi.org/10.1126/sciadv.abg6931>.
- 1021 Chen, J.-H., and S.-J. Lin, 2011: The remarkable predictability of interannual variability of
1022 atlantic hurricanes during the past decade. *Geophysical Research Letters*, **38** (11), L11 804,
1023 <https://doi.org/10.1029/2011GL047629>.
- 1024 Cheng, W., J. C. H. Chiang, and D. Zhang, 2013: Atlantic meridional overturning circulation
1025 (AMOC) in CMIP5 models: Rcp and historical simulations. *Journal of Climate*, **26** (18), 7187–
1026 7198, <https://doi.org/10.1175/JCLI-D-12-00496.1>.
- 1027 Colbert, A. J., and B. J. Soden, 2012: Climatological variations in north atlantic tropical cyclone
1028 tracks. *Journal of Climate*, **25** (2), 657–673, <https://doi.org/10.1175/JCLI-D-11-00034.1>.
- 1029 Diamond, M. S., 2023: Detection of large-scale cloud microphysical changes within a major
1030 shipping corridor after implementation of the international maritime organization 2020 fuel
1031 sulfur regulations. *Atmospheric Chemistry and Physics*, **23** (14), 8259–8269, <https://doi.org/10.5194/acp-23-8259-2023>.
- 1033 Dunstone, N. J., D. M. Smith, B. B. B. Booth, L. Hermanson, and R. Eade, 2013: Anthropogenic
1034 aerosol forcing of atlantic tropical storms. *Nature Geoscience*, **6** (7), 534–539, <https://doi.org/10.1038/ngeo1854>.
- 1036 Emanuel, K., 2010: Tropical cyclone activity downscaled from noaa–cires reanalysis, 1908–1958.
1037 *Journal of Advances in Modeling Earth Systems*, **2** (1), 1–12, <https://doi.org/10.3894/JAMES.2010.2.1>.
- 1039 Emanuel, K., 2021a: Atlantic tropical cyclones downscaled from climate reanalyses show in-
1040 creasing activity over past 150 years. *Nature Communications*, **12** (1), 7027, <https://doi.org/10.1038/s41467-021-27364-8>.
- 1042 Emanuel, K., 2021b: Response of global tropical cyclone activity to increasing co₂: Results
1043 from downscaling cmip6 models. *Journal of Climate*, **34** (1), 57–70, <https://doi.org/10.1175/JCLI-D-20-0367.1>.
- 1044

- 1045 Emanuel, K., 2022: Tropical cyclone seeds, transition probabilities, and genesis. *Journal of*
1046 *Climate*, **35 (11)**, 3629–3643, <https://doi.org/10.1175/JCLI-D-21-0922.1>.
- 1047 Emanuel, K., and D. Nolan, 2004: Tropical cyclone activity and global change. *Proceedings, 26th*
1048 *Conference on Hurricanes and Tropical Meteorology*, Miami, FL, American Meteorological
1049 Society, P2.2, URL https://ams.confex.com/ams/26HURR/techprogram/paper_75463.htm.
- 1050 Emanuel, K., S. Solomon, D. Folini, S. Davis, and C. Cagnazzo, 2013: Influence of tropical
1051 tropopause layer cooling on atlantic hurricane activity. *Journal of Climate*, **26 (7)**, 2288–2303,
1052 <https://doi.org/10.1175/JCLI-D-12-00242.1>.
- 1053 Emanuel, K., R. Sundararajan, and J. Williams, 2008: Hurricanes and global warming: Results
1054 from downscaling IPCC AR4 simulations. *Bulletin of the American Meteorological Society*,
1055 **89 (3)**, 347–367, <https://doi.org/10.1175/BAMS-89-3-347>.
- 1056 Emanuel, K. A., 2013: Downscaling CMIP5 climate models shows increased tropical cyclone
1057 activity over the 21st century. *Proceedings of the National Academy of Sciences*, **110 (30)**,
1058 12 219–12 224, <https://doi.org/10.1073/pnas.1301293110>.
- 1059 Eusebi, R., W. Yang, G. A. Vecchi, and S. Fueglistaler, 2025: Statistical modeling of north atlantic
1060 hurricane frequency and the impact and role of patterned warming. *Journal of Climate*, **38 (19)**,
1061 5391–5410, <https://doi.org/10.1175/JCLI-D-24-0647.1>.
- 1062 Flannaghan, T. J., S. Fueglistaler, I. M. Held, S. Po-Chedley, B. Wyman, and M. Zhao, 2014:
1063 Tropical temperature trends in atmospheric general circulation model simulations and the impact
1064 of uncertainties in observed SSTs. *Journal of Geophysical Research: Atmospheres*, **119 (23)**,
1065 13 327–13 337, <https://doi.org/10.1002/2014JD022365>.
- 1066 Fueglistaler, S., C. Radley, and I. M. Held, 2015: The distribution of precipitation and the spread
1067 in tropical upper tropospheric temperature trends in CMIP5/AMIP simulations. *Geophysical*
1068 *Research Letters*, **42 (14)**, 6000–6007, <https://doi.org/10.1002/2015GL064966>.
- 1069 Gelaro, R., and Coauthors, 2017: The modern-era retrospective analysis for research and ap-
1070 plications, version 2 (MERRA-2). *Journal of Climate*, **30 (14)**, 5419–5454, <https://doi.org/10.1175/JCLI-D-16-0758.1>.
- 1071

- 1072 Goldenberg, S. B., C. W. Landsea, A. M. Mestas-Nuñez, and W. M. Gray, 2001: The recent
1073 increase in atlantic hurricane activity: Causes and implications. *Science*, **293** (5529), 474–479,
1074 <https://doi.org/10.1126/science.1060040>.
- 1075 Harris, L. M., S.-J. Lin, and C. Tu, 2016: High-resolution climate simulations using GFDL
1076 HiRAM with a stretched global grid. *Journal of Climate*, **29** (11), 4293–4314, <https://doi.org/10.1175/JCLI-D-15-0389.1>.
- 1078 Held, I. M., and B. J. Soden, 2000: Water vapor feedback and global warming. *Annual Review of*
1079 *Environment and Resources*, **25**, 441–475, <https://doi.org/10.1146/annurev.energy.25.1.441>.
- 1080 Hersbach, H., and Coauthors, 2020: The ERA5 global reanalysis. *Quarterly Journal of the Royal*
1081 *Meteorological Society*, **146** (730), 1999–2049, <https://doi.org/10.1002/qj.3803>.
- 1082 Hsieh, T.-L., G. A. Vecchi, W. Yang, I. M. Held, and S. T. Garner, 2020: Large-scale control
1083 on the frequency of tropical cyclones and seeds: a consistent relationship across a hierarchy
1084 of global atmospheric models. *Climate Dynamics*, **55** (11), 3177–3196, <https://doi.org/10.1007/s00382-020-05446-5>.
- 1086 Hsieh, T.-L., W. Yang, G. A. Vecchi, and M. Zhao, 2022: Model spread in the tropical cyclone
1087 frequency and seed propensity index across global warming and enso-like perturbations. *Geo-*
1088 *physical Research Letters*, **49** (7), e2021GL097157, <https://doi.org/10.1029/2021GL097157>.
- 1089 Hsieh, T.-L., B. Zhang, W. Yang, G. A. Vecchi, M. Zhao, B. J. Soden, and C. Wang, 2023: The
1090 influence of large-scale radiation anomalies on tropical cyclone frequency. *Journal of Climate*,
1091 **36** (16), 5431–5441, <https://doi.org/10.1175/JCLI-D-22-0449.1>.
- 1092 Ikehata, K., and M. Satoh, 2021: Climatology of tropical cyclone seed frequency and survival rate
1093 in tropical cyclones. *Geophysical Research Letters*, **48** (18), e2021GL093626, <https://doi.org/10.1029/2021GL093626>.
- 1095 Jordan, G., and M. Henry, 2024: IMO2020 regulations accelerate global warming by up to 3 years
1096 in UKESM1. *Earth's Future*, **12** (8), e2024EF005011, <https://doi.org/10.1029/2024EF005011>.
- 1097 Klotzbach, P. J., 2011a: El niño–southern oscillation’s impact on atlantic basin hurricanes and u.s.
1098 landfalls. *Journal of Climate*, **24** (4), 1252–1263, <https://doi.org/10.1175/2010JCLI3799.1>.

- 1099 Klotzbach, P. J., 2011b: The influence of el niño–southern oscillation and the atlantic multi-
1100 decadal oscillation on caribbean tropical cyclone activity. *Journal of Climate*, **24** (3), 721–731,
1101 <https://doi.org/10.1175/2010JCLI3705.1>.
- 1102 Klotzbach, P. J., S. G. Bowen, R. Pielke, and M. Bell, 2018: Continental u.s. hurricane landfall
1103 frequency and associated damage: Observations and future risks. *Bulletin of the American*
1104 *Meteorological Society*, **99** (7), 1359–1376, <https://doi.org/10.1175/BAMS-D-17-0184.1>.
- 1105 Klotzbach, P. J., K. M. Wood, C. J. Schreck III, S. G. Bowen, C. M. Patricola, and M. M. Bell,
1106 2022: Trends in global tropical cyclone activity: 1990–2021. *Geophysical Research Letters*,
1107 **49** (6), e2021GL095774, <https://doi.org/10.1029/2021GL095774>.
- 1108 Knutson, T., and Coauthors, 2020: Tropical cyclones and climate change assessment: Part II:
1109 Projected response to anthropogenic warming. *Bulletin of the American Meteorological Society*,
1110 **101** (3), E303–E322, <https://doi.org/10.1175/BAMS-D-18-0194.1>.
- 1111 Knutson, T. R., J. J. Sirutis, Ming Zhao, R. E. Tuleya, M. Bender, G. A. Vecchi, G. Villarini, and
1112 D. Chavas, 2015: Global projections of intense tropical cyclone activity for the late twenty-
1113 first century from dynamical downscaling of CMIP5/RCP4.5 scenarios. **28** (18), 7203–7224,
1114 <https://doi.org/10.1175/JCLI-D-15-0129.1>.
- 1115 Kortum, G., G. A. Vecchi, T.-L. Hsieh, and W. Yang, 2024: Influence of weather and climate
1116 on multidecadal trends in atlantic hurricane genesis and tracks. *Journal of Climate*, **37** (5),
1117 1501–1522, <https://doi.org/10.1175/JCLI-D-23-0088.1>.
- 1118 Kossin, J. P., and S. J. Camargo, 2009: Hurricane track variability and secular potential intensity
1119 trends. *Climatic Change*, **97** (1), 329–337, <https://doi.org/10.1007/s10584-009-9748-2>.
- 1120 Kossin, J. P., K. R. Knapp, T. L. Olander, and C. S. Velden, 2020: Global increase in major
1121 tropical cyclone exceedance probability over the past four decades. **117** (22), 11975–11980,
1122 <https://doi.org/10.1073/pnas.1920849117>.
- 1123 Laloyaux, P., and Coauthors, 2018: CERA-20C: a coupled reanalysis of the twentieth century. *J.*
1124 *Adv. Model. Earth Syst.*, **10**, 1172–1195, <https://doi.org/10.1029/2018MS001273>.

- 1125 Lee, C.-Y., S. J. Camargo, A. H. Sobel, and M. K. Tippett, 2020: Statistical–dynamical downscaling
1126 projections of tropical cyclone activity in a warming climate: Two diverging genesis scenarios.
1127 *Journal of Climate*, **33** (11), 4201–4221, <https://doi.org/10.1175/JCLI-D-19-0452.1>.
- 1128 Lee, C.-Y., A. H. Sobel, M. K. Tippett, S. J. Camargo, M. Wüest, M. Wehner, and H. Murakami,
1129 2023: Climate change signal in atlantic tropical cyclones today and near future. *Earth's Future*,
1130 **11** (11), e2023EF003 539, <https://doi.org/10.1029/2023EF003539>.
- 1131 Manganello, J. V., and Coauthors, 2012: Tropical cyclone climatology in a 10-km global at-
1132 mospheric gcm: Toward weather-resolving climate modeling. *Journal of Climate*, **25** (11),
1133 3867–3893, <https://doi.org/10.1175/JCLI-D-11-00346.1>.
- 1134 Mann, M. E., and K. A. Emanuel, 2006: Atlantic hurricane trends linked to climate change.
1135 *Eos, Transactions American Geophysical Union*, **87** (24), 233–241, [https://doi.org/10.1029/](https://doi.org/10.1029/2006EO240001)
1136 [2006EO240001](https://doi.org/10.1029/2006EO240001).
- 1137 Murakami, H., and B. Wang, 2022: Patterns and frequency of projected future tropical cyclone
1138 genesis are governed by dynamic effects. *Communications Earth & Environment*, **3** (1), 1–10,
1139 <https://doi.org/10.1038/s43247-022-00410-z>.
- 1140 Neelin, J. D., and I. M. Held, 1987: Modeling tropical convergence based on the moist static en-
1141 ergy budget. *Monthly Weather Review*, **115** (1), 3–12, [https://doi.org/10.1175/1520-0493\(1987\)](https://doi.org/10.1175/1520-0493(1987)115(0003:MTCBOT)2.0.CO;2)
1142 [115\(0003:MTCBOT\)2.0.CO;2](https://doi.org/10.1175/1520-0493(1987)115(0003:MTCBOT)2.0.CO;2).
- 1143 Olszewski, K., 1986: Saturation deficit in various air masses. *Miscellanea Geographica*, **2** (1),
1144 79–84, <https://doi.org/10.2478/mgrsd-1986-020111>.
- 1145 Patricola, C. M., P. Chang, and R. Saravanan, 2016: Degree of simulated suppression of at-
1146 lantic tropical cyclones modulated by flavour of el niño. *Nature Geoscience*, **9** (2), 155–160,
1147 <https://doi.org/10.1038/ngeo2624>.
- 1148 Patricola, C. M., R. Saravanan, and P. Chang, 2018: The response of atlantic tropical cy-
1149 clones to suppression of african easterly waves. *Geophysical Research Letters*, **45** (1), 471–479,
1150 <https://doi.org/10.1002/2017GL076081>.

- 1151 Pfahl, S., P. A. O’Gorman, and E. M. Fischer, 2017: Understanding the regional pattern of
1152 projected future changes in extreme precipitation. *Nature Climate Change*, **7 (6)**, 423–427,
1153 <https://doi.org/10.1038/nclimate3287>.
- 1154 Pielke, R. A., and C. N. Landsea, 1999: La niña, el niño, and atlantic hurricane damages in the
1155 united states. *Bulletin of the American Meteorological Society*, **80 (10)**, 2027–2033.
- 1156 Ramsay, H. A., and A. H. Sobel, 2011: Effects of relative and absolute sea surface temperature
1157 on tropical cyclone potential intensity using a single-column model. *Journal of Climate*, **24 (1)**,
1158 183–193, <https://doi.org/10.1175/2010JCLI3690.1>.
- 1159 Ritchie, E. A., and G. J. Holland, 1999: Large-scale patterns associated with tropical cyclogenesis
1160 in the western pacific. *Monthly Weather Review*, **127 (9)**, 2027–2043, [https://doi.org/10.1175/
1161 1520-0493\(1999\)127<2027:LSPAWT>2.0.CO;2](https://doi.org/10.1175/1520-0493(1999)127<2027:LSPAWT>2.0.CO;2).
- 1162 Seager, R., N. Henderson, and M. Cane, 2022: Persistent discrepancies between observed and mod-
1163 eled trends in the tropical pacific ocean. *Journal of Climate*, **35 (14)**, 4571–4584, [https://doi.org/
1164 10.1175/JCLID210648.1](https://doi.org/10.1175/JCLID210648.1).
- 1165 Shaw, T. A., and Coauthors, 2024: Regional climate change: consensus, discrepancies, and ways
1166 forward. *Frontiers in Climate*, **6**, 1–14, <https://doi.org/10.3389/fclim.2024.1391634>.
- 1167 Sobel, A. H., I. M. Held, and C. S. Bretherton, 2002: The enso signal in tropical tropospheric
1168 temperature. *Journal of Climate*, **15 (18)**, 2702–2708, [https://doi.org/10.1175/1520-0442\(2002\)
1169 015<2702:TESITT>2.0.CO;2](https://doi.org/10.1175/1520-0442(2002)015<2702:TESITT>2.0.CO;2).
- 1170 Sobel, A. H., A. A. Wing, S. J. Camargo, C. M. Patricola, G. A. Vecchi, C. Lee, and M. K. Tippett,
1171 2021: Tropical cyclone frequency. *Earth’s Future*, **9 (12)**, e2021EF002 275, [https://doi.org/
1172 10.1029/2021EF002275](https://doi.org/10.1029/2021EF002275).
- 1173 Srokosz, M., and H. Bryden, 2015: Observing the atlantic meridional overturning circulation
1174 yields a decade of inevitable surprises. *Science*, **348 (6241)**, 1255–1257, [https://doi.org/10.1126/
1175 science.1255575](https://doi.org/10.1126/science.1255575).
- 1176 Sugi, M., Y. Yamada, K. Yoshida, R. Mizuta, M. Nakano, C. Kodama, and M. Satoh, 2020: Future
1177 changes in the global frequency of tropical cyclone seeds. *SOLA*, **16**, 70–74, [https://doi.org/
1178 10.2151/sola.2020012](https://doi.org/10.2151/sola.2020012).

- 1179 Tang, B., and S. J. Camargo, 2014: Environmental control of tropical cyclones in cmip5: A ventila-
1180 tion perspective. *Journal of Advances in Modeling Earth Systems*, **6** (1), 115-128, [https://doi.org/](https://doi.org/10.1002/2013MS000294)
1181 10.1002/2013MS000294.
- 1182 Tang, B., and K. Emanuel, 2010: Midlevel ventilation's constraint on tropical cyclone intensity.
1183 *Journal of the Atmospheric Sciences*, **67** (6), 1817-1830, <https://doi.org/10.1175/2010JAS3318>.
1184 1.
- 1185 Tang, B., and K. Emanuel, 2012: A ventilation index for tropical cyclones. *Bulletin of the American*
1186 *Meteorological Society*, **93** (12), 1901-1915, <https://doi.org/10.1175/BAMSD1100165.1>.
- 1187 Taylor, K. E., R. J. Stouffer, and G. A. Meehl, 2012: An overview of CMIP5 and the experiment
1188 design. <https://doi.org/10.1175/BAMS-D-11-00094.1>.
- 1189 Tippett, M., S. Camargo, and A. Sobel, 2011: A poisson regression index for tropical cyclone
1190 genesis and the role of large-scale vorticity in genesis. *Journal of Climate*, **24** (9), 2335-2357,
1191 <https://doi.org/10.1175/2010JCLI3811.1>.
- 1192 Vecchi, G. A., S. Fueglistaler, I. M. Held, T. R. Knutson, and M. Zhao, 2013: Impacts of
1193 atmospheric temperature trends on tropical cyclone activity. *Journal of Climate*, **26** (11), 3877–
1194 3898, <https://doi.org/10.1175/JCLI-D-12-00503.1>.
- 1195 Vecchi, G. A., and T. R. Knutson, 2008: On estimates of historical north atlantic tropical cyclone
1196 activity. *Journal of Climate*, **21** (14), 3580–3600, <https://doi.org/10.1175/2008JCLI2178.1>.
- 1197 Vecchi, G. A., and T. R. Knutson, 2011: Estimating annual numbers of atlantic hurricanes missing
1198 from the hurdat database (1878–1965) using ship track density. *Journal of Climate*, **24** (6),
1199 1736–1746, <https://doi.org/10.1175/2010JCLI3810.1>.
- 1200 Vecchi, G. A., C. Landsea, W. Zhang, G. Villarini, and T. Knutson, 2021: Changes in atlantic
1201 major hurricane frequency since the late-19th century. *Nature Communications*, **12** (1), 4054,
1202 <https://doi.org/10.1038/s41467-021-24268-5>.
- 1203 Vecchi, G. A., and B. J. Soden, 2007: Effect of remote sea surface temperature change on
1204 tropical cyclone potential intensity. *Nature*, **450** (7172), 1066–1070, [https://doi.org/10.1038/](https://doi.org/10.1038/nature06423)
1205 nature06423.

- 1206 Vecchi, G. A., M. Zhao, H. Wang, G. Villarini, A. Rosati, A. Kumar, I. M. Held, and R. Gudgel,
1207 2011: Statistical–dynamical predictions of seasonal north atlantic hurricane activity. *Monthly*
1208 *Weather Review*, **139** (4), 1070–1082, <https://doi.org/10.1175/2010MWR3499.1>.
- 1209 Vecchi, G. A., and Coauthors, 2019: Tropical cyclone sensitivities to co2 doubling: roles of
1210 atmospheric resolution, synoptic variability and background climate changes. *Climate Dynamics*,
1211 **53** (9), 5999–6033, <https://doi.org/10.1007/s00382-019-04913-y>.
- 1212 Villarini, G., G. A. Vecchi, T. R. Knutson, and J. A. Smith, 2011a: Is the recorded increase
1213 in short-duration north atlantic tropical storms spurious? *Journal of Geophysical Research:*
1214 *Atmospheres*, **116**, D10 112, <https://doi.org/10.1029/2010JD015493>.
- 1215 Villarini, G., G. A. Vecchi, T. R. Knutson, M. Zhao, and J. A. Smith, 2011b: North atlantic tropical
1216 storm frequency response to anthropogenic forcing: Projections and sources of uncertainty.
1217 *Journal of Climate*, **24** (13), 3224–3248, <https://doi.org/10.1175/2011JCLI3853.1>.
- 1218 Villarini, G., G. A. Vecchi, and J. A. Smith, 2010: Modeling the dependence of tropical storm
1219 counts in the north atlantic basin on climate indices. *Monthly Weather Review*, **138** (7), 2681–
1220 2705, <https://doi.org/10.1175/2010MWR3315.1>.
- 1221 Walsh, K. J. E., S. J. Camargo, T. R. Knutson, J. Kossin, T.-C. Lee, H. Murakami, and C. Patricola,
1222 2019: Tropical cyclones and climate change. *Tropical Cyclone Research and Review*, **8** (4),
1223 240–250, <https://doi.org/10.1016/j.tcrr.2020.01.004>.
- 1224 Wang, B., and H. Murakami, 2020: Dynamic genesis potential index for diagnosing present-day
1225 and future global tropical cyclone genesis. *Environmental Research Letters*, **15** (11), 114 008,
1226 <https://doi.org/10.1088/1748-9326/abbb01>.
- 1227 Weller, E., S.-K. Min, W. Cai, F. W. Zwiers, Y.-H. Kim, and D. Lee, 2016: Human-caused
1228 indo-pacific warm pool expansion. *Science Advances*, **2** (7), e1501 719, <https://doi.org/10.1126/sciadv.1501719>.
- 1230 Xie, L., T. Yan, L. J. Pietrafesa, J. M. Morrison, and T. Karl, 2005: Climatology and interan-
1231 nual variability of north atlantic hurricane tracks. *Journal of Climate*, **18** (24), 5370–5381,
1232 <https://doi.org/10.1175/JCLI3560.1>.

- 1233 Yamada, Y., and Coauthors, 2021: Evaluation of the contribution of tropical cyclone seeds to
1234 changes in tropical cyclone frequency due to global warming in high-resolution multi-model
1235 ensemble simulations. *Progress in Earth and Planetary Science*, **8 (1)**, 11, [https://doi.org/](https://doi.org/10.1186/s40645-020-00397-1)
1236 [10.1186/s40645-020-00397-1](https://doi.org/10.1186/s40645-020-00397-1).
- 1237 Yang, W., T.-L. Hsieh, and G. A. Vecchi, 2021: Hurricane annual cycle controlled by both seeds and
1238 genesis probability. *Proceedings of the National Academy of Sciences*, **118 (41)**, e2108397 118,
1239 <https://doi.org/10.1073/pnas.2108397118>.
- 1240 Zhang, B., B. J. Soden, G. A. Vecchi, and W. Yang, 2021: The role of radiative interactions in
1241 tropical cyclone development under realistic boundary conditions. *Journal of Climate*, **34 (6)**,
1242 2079–2091, <https://doi.org/10.1175/JCLI-D-20-0574.1>.
- 1243 Zhang, C., 1993: Large-scale variability of atmospheric deep convection in relation to sea surface
1244 temperature in the tropics. *Journal of Climate*, **6 (10)**, 1898–1913, [https://doi.org/10.1175/](https://doi.org/10.1175/1520-0442(1993)006(1898:LSVOAD)2.0.CO;2)
1245 [1520-0442\(1993\)006\(1898:LSVOAD\)2.0.CO;2](https://doi.org/10.1175/1520-0442(1993)006(1898:LSVOAD)2.0.CO;2).
- 1246 Zhang, J., Y.-S. Chen, E. Gryspeerdt, T. Yamaguchi, and G. Feingold, 2025: Radiative forcing
1247 from the 2020 shipping fuel regulation is large but hard to detect. *Communications Earth &*
1248 *Environment*, **6 (1)**, 1–11, <https://doi.org/10.1038/s43247-024-01911-9>.
- 1249 Zhang, R., and T. L. Delworth, 2006: Impact of atlantic multidecadal oscillations on india/sahel
1250 rainfall and atlantic hurricanes. *Geophysical Research Letters*, **33 (17)**, L17 712, [https://doi.org/](https://doi.org/10.1029/2006GL026267)
1251 [10.1029/2006GL026267](https://doi.org/10.1029/2006GL026267).
- 1252 Zhao, M., and I. M. Held, 2010: An analysis of the effect of global warming on the intensity
1253 of atlantic hurricanes using a gcm with statistical refinement. *Journal of Climate*, **23 (23)**,
1254 6382–6395, <https://doi.org/10.1175/2010JCLI3837.1>.
- 1255 Zhao, M., I. M. Held, S.-J. Lin, and G. A. Vecchi, 2009: Simulations of global hurricane clima-
1256 tology, interannual variability, and response to global warming using a 50-km resolution gcm.
1257 *Journal of Climate*, **22 (24)**, 6653–6678, <https://doi.org/10.1175/2009JCLI3049.1>.
- 1258 Zhuo, J.-Y., C.-Y. Lee, A. Sobel, R. Seager, S. J. Camargo, Y.-H. Lin, B. Fosu, and K. A. Reed,
1259 2025: A more La Niña–like response to radiative forcing after flux adjustment in CESM2.
1260 *Journal of Climate*, **38 (4)**, 1321–1346, <https://doi.org/10.1175/JCLI-D-24-0331.1>.