

The debt burden of tropical cyclones and climate change

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Abstract

Addressing climate change, through both mitigation and adaptation, is anticipated to require global investments of more than \$6 trillion annually by 2035. However, many countries face significant barriers to accessing the finance needed for these investments, due to low or absent credit ratings, large debt burdens, and high borrowing costs. There is concern that climate change, through its economic impacts, may amplify these barriers, potentially locking countries into a “vicious cycle” in which mounting economic losses further constrain countries’ capacity to invest in adaptation and mitigation. We provide evidence that the cost and availability of capital for many countries have already been shaped by their historical exposure to tropical cyclones (TCs) and warming temperatures. Our empirically derived estimates suggest that, across all TC-exposed countries, debt-to-GDP ratios are on average 30% higher due to the cumulative effects of TCs since 1990. GDP levels are on average 10% lower due to the combined impacts of TCs and warming temperatures across all countries. We estimate that because of these impacts, hotter countries are more likely to receive credit ratings below investment grade (< BBB–), and borrowing costs are at least 1 basis point (0.01%) higher in 28 countries and 5 basis points higher in highly-exposed countries. Future increases in temperature and TC activity will likely worsen countries’ credit, potentially undermining both countries’ abilities to address climate change and their long-run development prospects.

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1 Introduction

2 Efforts to reduce greenhouse gas emissions and adapt to the impacts of climate change are esti-
3 mated to require global investments of more than US\$6 trillion annually by 2035 (Bhattacharya
4 et al. (2024)). Among emerging markets and developing countries (EMDCs), excluding China,
5 estimated investment needs are US\$2.4 trillion by 2030—more than ten times the climate fi-
6 nance that these countries currently receive (Climate Policy Initiative (2025)). While increasing
7 financial support to developing countries has been a key priority in global climate negotiations
8 since 2009 (UNFCCC. Conference of the Parties serving as the meeting of the Parties to the Paris
9 Agreement (CMA) (2024)), the bulk of climate finance has remained concentrated in advanced
10 economies while the rest of the world has received only a small share, and largely in the form
11 of debt (Extended Data Fig. 1).

12 The large regional disparity in finance flows is not new (Lucas (1990)), but likely persists due
13 to differences in the strength and quality of institutions, the maturity of financial markets,
14 and other indicators of economic development which are ultimately reflected in sovereign
15 credit ratings (United Nations, Inter-agency Task Force on Financing for Development (2022),
16 Kowalewski et al. (2025)). Ratings issued by agencies such as Standard & Poor's, Moody's, and
17 Fitch continue to play an influential role in shaping the lending decisions of global investors,
18 despite concerns regarding the role of rating agencies in contributing to the global financial cri-
19 sis in 2009. Because sovereign ratings serve as ceilings for domestic economic actors, a down-
20 grade at the sovereign level can significantly raise financing costs across the entire economy
21 (Almeida et al. (2017)). Low-rated sovereigns face the difficult choice of borrowing from in-
22 ternational capital markets and being exposed to currency risks, or borrowing from domestic
23 markets at potentially higher rates at short-term maturities. On the other hand, countries with-
24 out a credit rating are effectively locked out of international capital markets (fig. 1, United Na-
25 tions, Inter-agency Task Force on Financing for Development (2022), UNCTAD (2024)). These
26 constraints are compounded by high sovereign debt levels which are projected to reach 100%
27 of global GDP by 2030 (International Monetary Fund (2024)). Among low-income countries,
28 interest payments currently exceed their combined health and education spending and are es-
29 timated to contribute to net negative transfers to Multilateral Development Banks (MDBs) (G20
30 Independent Expert Group (2023)). Most of these countries have been in sovereign debt default
31 over the past two decades (fig. S1, Beers et al. (2021)).

32 There is growing evidence that many of these countries also face substantial macroeconomic
33 risk from climate change itself, especially lower-income and hotter countries that are differ-
34 entially exposed and vulnerable to a range of climate threats and their impacts (Burke et al.
35 (2015), Burke et al. (2018), Nath et al. (2023), Bilal and Käning (2024), Acevedo et al. (2020),
36 Mohan and Strobl (2021)). As global warming approaches 1.5°C above pre-industrial levels
37 (Masson-Delmotte et al. (2021), Diffenbaugh and Barnes (2023), Bevacqua et al. (2025)), there
38 is growing concern that credit ratings and the overall ability of countries to access financing
39 will be negatively impacted by these economic impacts of a warming climate, contributing to a

“vicious cycle” by which impacted countries are increasingly challenged to secure the finance necessary for addressing climate change and other development objectives (Kling et al. (2018), Volz et al. (2020)). Consistent with these growing concerns, data show a strong correlation between countries’ credit ratings, mean temperatures, and tropical cyclone (TC) exposure, with hotter and more TC-exposed countries more likely to have speculative grade ratings (< BBB-) (fig. 1). It is projected that climate change will increase the likelihood of major category TCs and TC-induced precipitation, although significant uncertainties remain regarding TC formation and overall frequency (Masson-Delmotte et al. (2021), Knutson et al. (2020), Knutson et al. (2015), van Oldenborgh et al. (2017), Sobel et al. (2016), Bhatia et al. (2022)). Concern about the vicious cycle is especially acute for small island states such as Jamaica, Barbados, and Grenada, where TCs are the primary source of negative financial shock (fig. 1, fig. S7). Because of their small land area, high population density, coastal infrastructure and exposure to sea level rise, a small island state’s entire economy can be exposed to the impacts of a single TC event (Brownbridge and Canagarajah (2024), Hsiang and Jina (2014)). Small island states also face higher reconstruction costs and significant constraints in mobilizing post-disaster financing due to their remote locations (Slany (2020)).

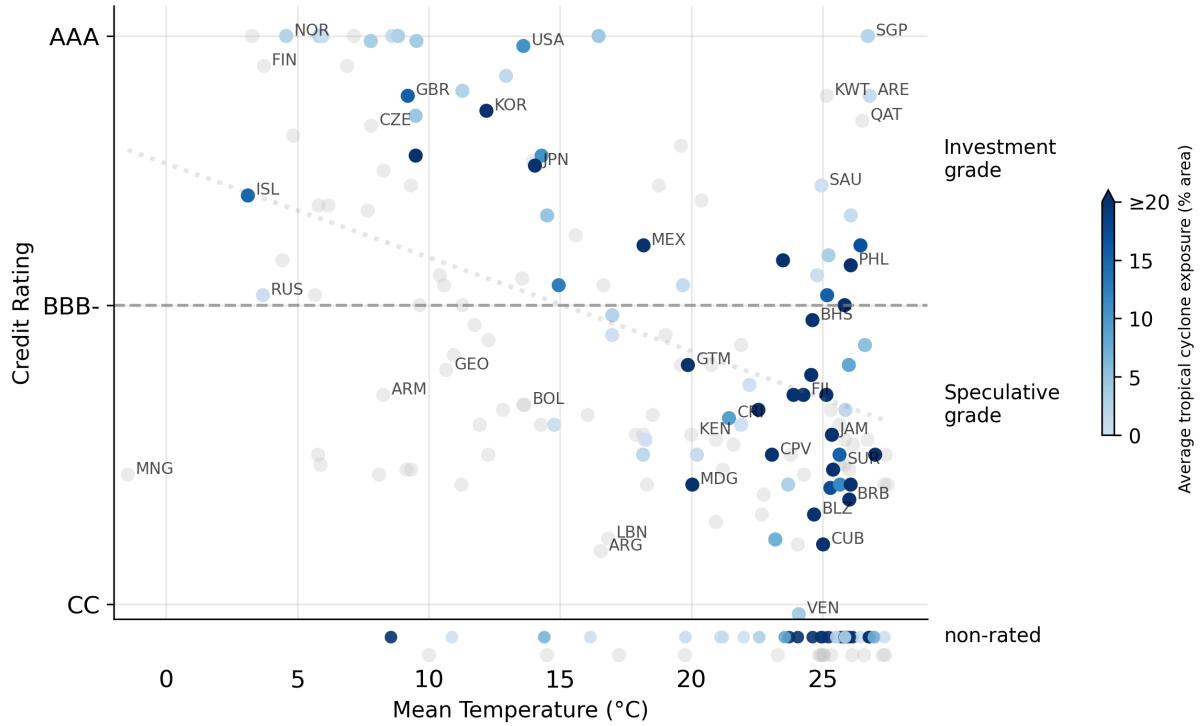


Figure 1. Scatter plot of country mean temperatures, average tropical cyclone exposure, and sovereign credit ratings in 2019. Mean temperatures are on the x-axis, and credit ratings on the y-axis are the average of ratings from the “Big Three” agencies (S&P, Moody’s, Fitch). Colors indicate the average tropical cyclone exposure in terms of land area affected. Scatter plot along the x-axis shows the temperature distribution and tropical cyclone exposure of 46 countries that have never received an official rating from the three major credit rating agencies.

56 A growing body of research investigates how the actual and perceived risks of climate change
57 may impact sovereign ratings and bond yields. These studies largely focus on the temperature-
58 GDP impact channel, the association between ratings and aggregate indicators of climate risk
59 vulnerability (Cevik and Jalles (2020), Bolton et al. (2023), Klusak et al. (2021), Beirne et al.
60 (2021)), or the impact of disasters in general (Fisera et al. (2023), Deryugina (2022)). Some studies
61 use theoretically modeled estimates to understand how the combination of high sovereign
62 debt and TCs may slow post-disaster recovery, impact governments' ability to issue debt and
63 increase borrowing costs (Phan and Schwartzman (2024), Bakkensen and Barrage (2025), Mal-
64 lucci (2022)), without explicitly accounting for the temperature-GDP impact channel. Mean-
65 while, empirical evidence on the impact of TCs has been mixed, with some finding that severe
66 TCs do not increase debt nor cause long-run impacts on GDP growth (International Monetary
67 Fund (2014), Cavallo et al. (2013)), and others finding long-run growth impacts (Hsiang and
68 Jina (2014), Cabezon et al. (2015), Brownbridge and Canagarajah (2024), Slany (2020)).

69 This study provides the first global assessment of how historical exposure to TCs and long-
70 term country-level warming has shaped the cost and availability of capital. We do this by
71 empirically estimating how combined impacts of TCs and rising temperatures affect debt and
72 GDP. We further provide evidence of how these impacts have shaped countries' credit ratings
73 and borrowing costs, which may compound countries' debt burdens and limit their capacity to
74 respond to climate change.

75 We first estimate the impact of TCs and warming temperatures on two macroeconomic vari-
76 ables that are key determinants of sovereign credit ratings: the debt-to-GDP ratio (hereafter
77 "debt ratio") and GDP (Cantor and Packer (1996), Afonso et al. (2011)). The debt ratio is a
78 widely used indicator for understanding a country's debt burden that facilitates cross-country
79 comparisons by scaling debt relative to economic output. It also serves as an important indica-
80 tor for understanding a government's capacity to implement fiscal policy measures in response
81 to financial crises or economic downturns (Romer and Romer (2019), Jordà et al. (2016)). To esti-
82 mate the impacts of TCs and temperature on debt ratios and GDP, we employ a local projection
83 model commonly used in applied macroeconomic settings to investigate the long-run impact
84 of exogenous shocks (Methods).

85 Next, to understand how the macroeconomic impacts from TCs and temperature change could
86 affect countries' credit rating and borrowing costs, we estimate a range of regression and ma-
87 chine learning-based prediction models that relate credit ratings to observed macroeconomic
88 factors. We then combine these models with our estimates of the impact of TCs and tempera-
89 ture on the macro-economy to estimate counterfactual credit ratings and borrowing costs had
90 these shocks not occurred.

2 Results

We find clear evidence that exposure to TCs affects subsequent debt ratios, with the effect size depending on the amount of land area exposed to high wind speeds (fig. 2). For instance, exposure of ten percent of a country's land area to tropical storm-level winds (>18 m/s) increases the debt ratio by 2.5% after ten years, while equivalent exposure to category 1–3 winds (33–50 m/s) raises it by 3.5%, and exposure to major TCs (>50 m/s) raises it by 5%. Impacts peak 5–10 years after exposure and fade after 15 years. The impact of each wind speed intensity scales with the share of land area exposed (fig. 2, panel b & d), which varies greatly depending on a country's size and average TC exposure. For example, up to 100% of Barbados' land area has historically been exposed to Category 3 or higher winds, compared with less than 5% of the United States (see fig. 3, panel b; fig. S7). Our baseline model includes temperature impacts as controls based on literature documenting robust evidence of temperature impacts on GDP (Methods). We do not detect a clear impact of temperature on the debt ratio separate from TCs, as coefficients are highly sensitive to the inclusion of different samples and time trends (fig. S3).

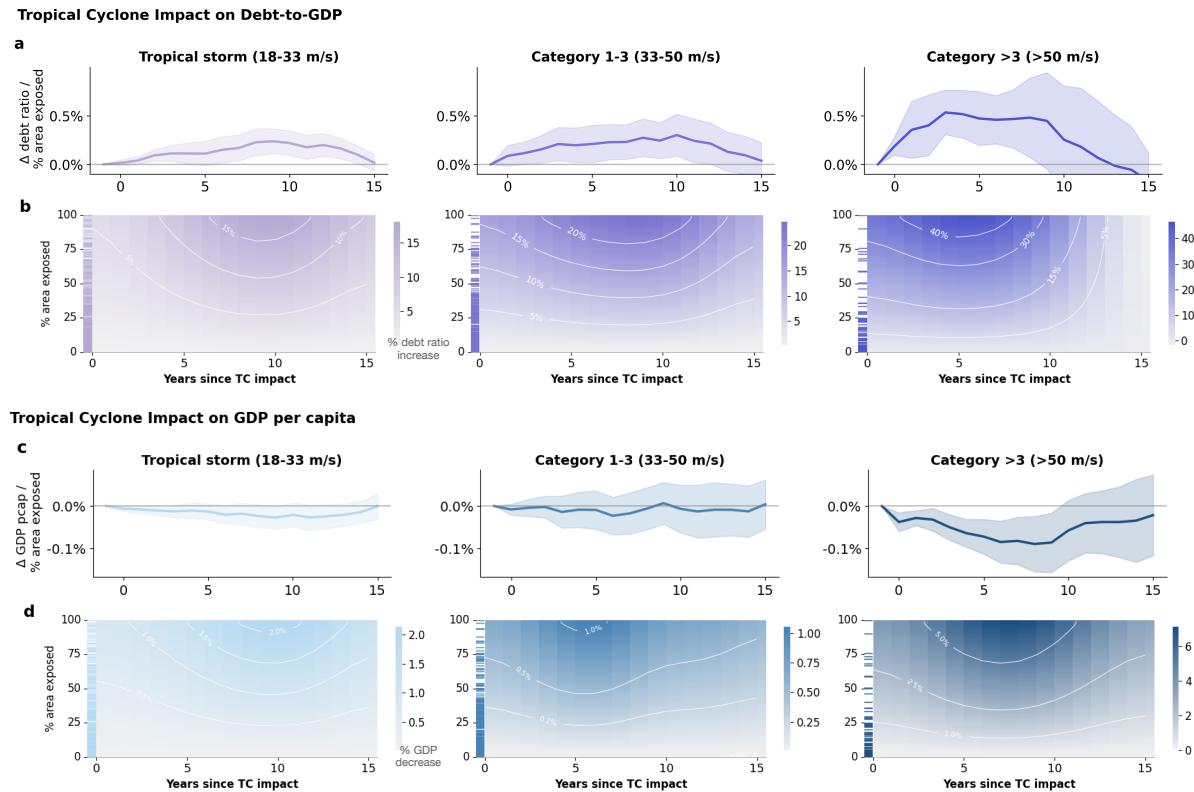


Figure 2. Impact of tropical cyclones on debt-to-GDP ratios and GDP. Panel **a** plots the response of the debt ratio to 1% of land area exposed to wind speeds of increasing intensity. Panel **b** scales each response estimate by the land area exposed to each wind speed category, from 0 to 100%. The distribution of observed exposure values from the data are shown as rug plots along the y-axis. Panel **c** plots the response of GDP per capita to 1% of land area exposed to wind speeds of varying intensity. Panel **d** scales the response by land area exposed, as in Panel b.

When considering GDP as the outcome, we again find that the impact of TCs depends on the land area impacted by wind speeds of different intensities (fig. 2). Exposure of 10% of a country's land area to tropical storm-level winds (>18 m/s) reduces GDP by 0.2% after ten years, while similar exposure to major tropical cyclones (>50 m/s) reduces GDP by 1% at its peak before returning to trend. We also confirm that, consistent with literature (Burke et al. (2015), Nath et al. (2023)), country-level warming impacts GDP growth and that this impact is independent from TC impacts. A 1°C hotter year relative to the country's mean temperature reduces GDP by 1% after four years, with impacts weakly persisting even after a decade (fig. S4). These impacts are highly dependent on the country's mean temperature, consistent with Burke et al. (2015) and Nath et al. (2023). We test three different methods for isolating temperature shocks, each showing statistically significant and persistent effects of a hotter year lasting up to six years for countries with mean historical temperature of 25°C (fig. S4).

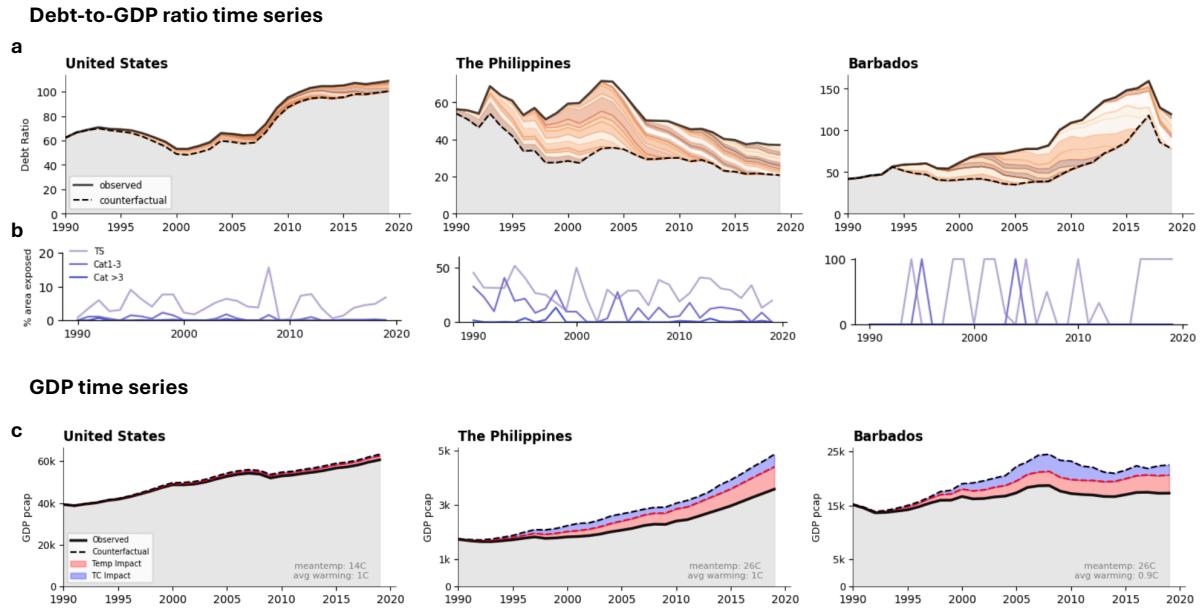


Figure 3. Panel a (observed vs counterfactual debt ratio time series) plots the time series of the observed debt ratio (black line) and counterfactual debt ratios (colored bars) for three countries with varying exposures. Each shaded bar represents the cumulative impact of all tropical cyclones in a given year. The black dotted line represents the counterfactual time series after removing the impact of all storms from 1990 onward. **Panel b (exposures)** plots the land area exposure values for each wind speed category. The color bars and axes ranges are different in each subplot. **Panel c (observed vs counterfactual GDP time series)** plots the time series of the observed GDP (black line) and counterfactual GDP that removes the impact of all tropical cyclone occurrences (blue) and temperature impacts (red) from 1990 onwards. The black dotted line plots the counterfactual scenario with both tropical cyclone and temperature impacts removed.

We then use these estimates to calculate counterfactual GDPs and debt ratios had observed warming and TC landfalls since 1990 did not occur. We find that, on average across TC-exposed countries, debt ratios in 2019 are 30% higher due to observed TCs, and GDP levels are approximately 10% lower due to the combined impacts of TCs and country-level warming (fig. 3,

121 fig. 4). Among small island sovereign states, on average 50% of the public debt burden in 2019
 122 is attributable to TCs, and GDP levels are 20% lower due to the combined impact of TCs and
 123 country-level warming. A map showing the global distribution of debt ratio impacts is shown
 124 in fig. 4 (GDP impacts are shown in fig. S6). In terms of GDP impacts, country-level warming
 125 accounts for a larger share of impacts than TCs (fig. 3, panel c).

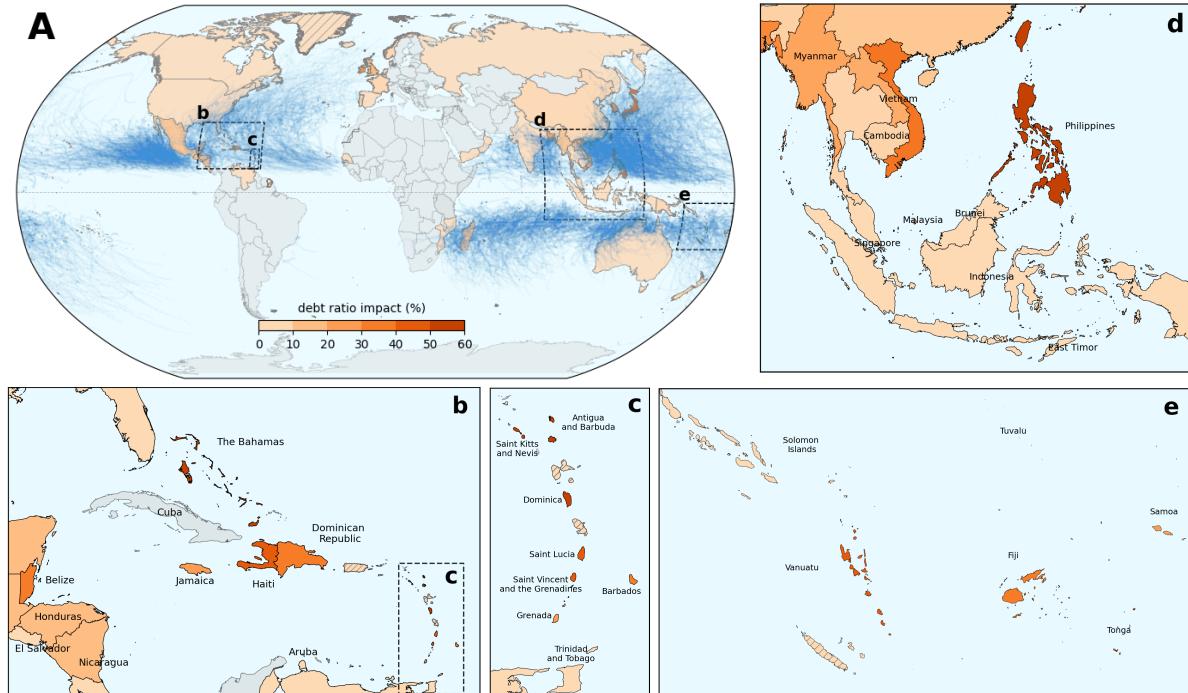


Figure 4. Share of countries' debt ratio in 2019 attributable to tropical cyclone exposure from 1990 onwards. Panel A shows the global map, where blue lines show tropical cyclone tracks from 1990-2019. Sub-panels b-e provide a zoomed-in view of regions that are highly exposed to tropical cyclones. Hatch marks indicate sovereign territories and dependencies without a debt record, and grey indicates countries either without tropical cyclone exposure or debt records.

126 Next, we investigate how the long-run impact on debt ratios and GDP may be affecting sovereign
 127 credit ratings and borrowing costs today. To do this, we predict the probability distribution of
 128 counterfactual ratings that countries would have been assigned based on their counterfactual
 129 debt ratios and GDP in 2019 (Methods). We estimate that 85 countries are more likely to be
 130 assigned below-investment-grade ratings (<BBB-) and 93 countries are less likely to receive a
 131 rating upgrade as a consequence of these macroeconomic impacts (fig. S16).

132 These rating changes are associated with a change in the borrowing costs of countries, here
 133 measured as the basis point difference in the coupon rates of sovereign bonds (Methods). Many
 134 countries with below investment-grade ratings are likely incurring additional borrowing costs
 135 of 1 to 35 basis points due to the impact of TCs and warming temperatures (fig. 5). The ad-
 136 dditional borrowing costs are greater for countries that predominantly issue bonds in local cur-
 137 rency than for countries issuing in a dominant currency (USD, EUR, GBP, CHF, CAD)(Extended

Data Fig. 2).

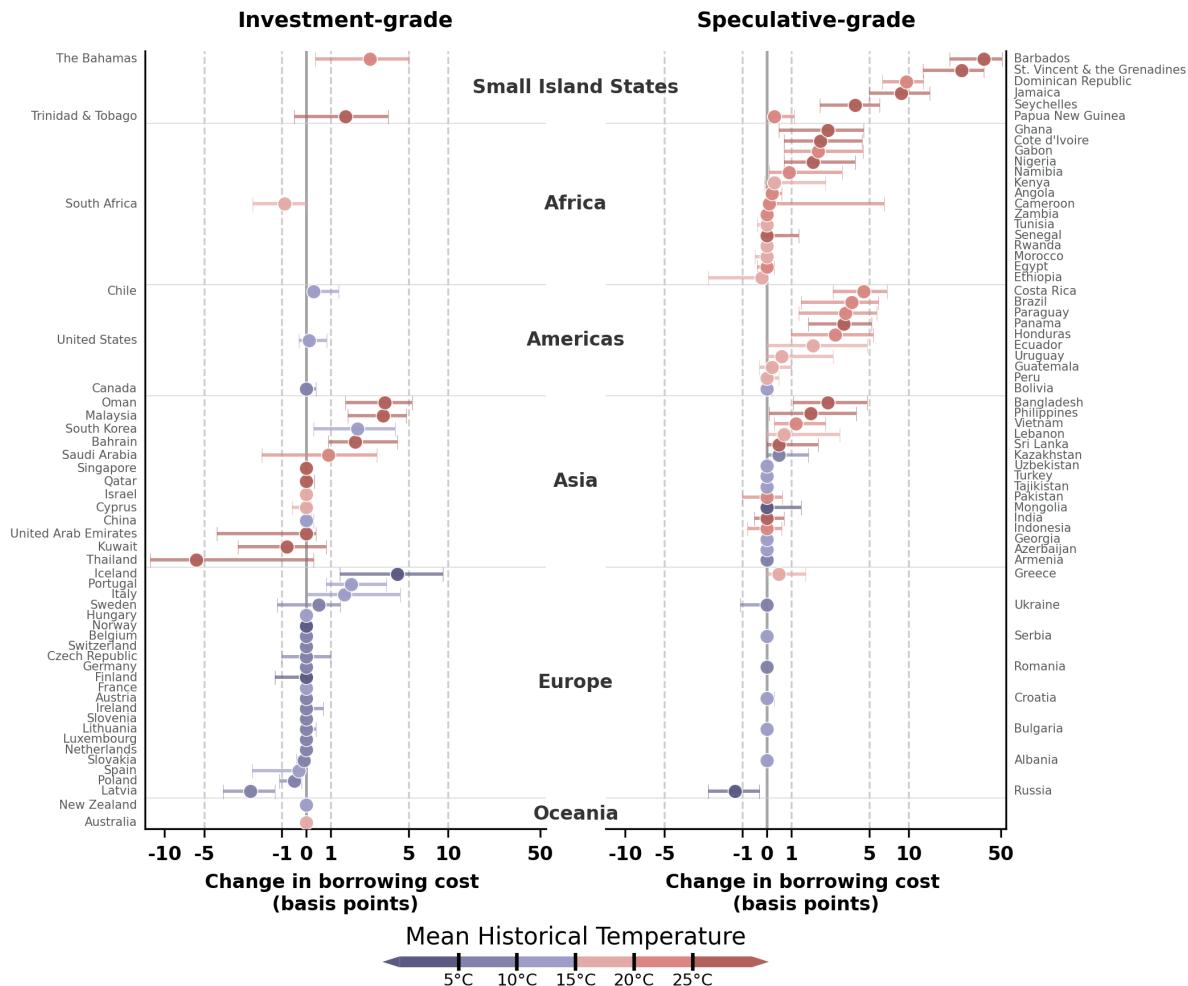


Figure 5. The change in borrowing cost attributed to historical tropical cyclones and climate change. Values are plotted separately for countries with investment-grade (left) and speculative-grade ratings (right), and non-rated countries. Colors are based on each country's mean historical temperature. Error bars show the inter-quartile range of predictions derived from 500 randomized training samples in the rating prediction model, with each sample holding out 20% of countries.

3 Discussion

The scale of investment required to reduce emissions and adapt to the impacts of climate change is estimated at trillions of dollars annually in the coming decade (Bhattacharya et al. (2024)). Global policy efforts have emphasized the critical role of the private sector in mobilizing finance at this scale, yet many countries face persistent challenges in accessing capital markets due to weak or non-existent credit ratings, large debt burdens, and high borrowing costs (see fig. 1, Extended Data Fig. 1, Bhattacharya et al. (2024)). Understanding the dynamic impact

146 of TCs and warming temperatures on countries' overall debt burdens and the additional costs
147 that this imposes is important to ensure the design of policies that can deliver climate finance
148 equitably and in a manner that supports each country's needs.

149 We provide a global-scale estimate of how past exposure to TCs and warming temperatures
150 may have impacted countries' borrowing costs today, including predictions for countries that
151 have never received a credit rating. While research suggests that climate risks may already be
152 reflected in credit ratings, or may lead to downgrades in the future (Cevik and Jalles (2020),
153 Bolton et al. (2023), Klusak et al. (2021)), they do not clarify the underlying macroeconomic
154 channels through which credit ratings may be affected. Cappiello et al. (2025) suggest that
155 countries with high exposure to physical risks (e.g. temperature anomalies and disasters) or
156 transition risks (e.g. dependence on fossil fuel revenues) are associated with lower credit rat-
157 ings. Other studies directly estimate climate change impacts on bond yields, albeit for a limited
158 sample of countries for which these data are available (Kling et al. (2018), Beirne et al. (2021)).
159 We find that hotter countries with low credit ratings today have an increased likelihood of ex-
160 periencing a rating downgrade and higher borrowing costs, while some colder countries with
161 higher credit ratings today have an increased likelihood of a rating upgrade as warming tem-
162 peratures positively impact GDP growth. Meanwhile, countries that do not have a credit rating
163 today may find it increasingly difficult to attain one.

164 Our analysis also contributes to the literature on the macroeconomic impacts of TCs. Existing
165 estimates of TC damages measure exposure primarily as a function of wind speed (e.g. Hsiang
166 and Jina (2014), Bakkensen and Mendelsohn (2016), Noy (2009)). A frequently cited estimate
167 of TC-induced GDP damages provided by Hsiang and Jina (2014) estimates that an additional
168 m/s of wind speed per unit area causes output loss of 0.09% five years after a storm. The
169 intuition for using wind speed as a metric of TC exposure draws from literature demonstrating
170 that direct losses are a power function of the maximum wind speed affecting a given unit area
171 or property (Pielke (2007), Nordhaus (2010), Emanuel (2005), Emanuel (2011), Southern (1979)).
172 However, the maximum wind speed metric reflects only the peak intensity of a storm and fails
173 to capture the broader spatial extent of exposure that may be more relevant for understanding
174 economic losses.

175 In contrast, we measure TC exposure as the percent of land area of the country impacted by
176 different wind speeds, to capture the multiple hazards associated with the spatial structure of
177 the storm while also accounting for differences in country size. Our metric thus captures both
178 the scale and distribution of potential damages, and is more suitable for capturing non-linear
179 impacts. It also better aligns with the mechanism by which TCs disrupt economic activity and
180 trigger reconstruction costs, following widespread damage to infrastructure, agriculture, and
181 settlements. Our results indeed confirm that there are meaningful debt and GDP impacts in
182 areas exposed to low wind speeds.

183 A major caveat of our analysis is that we do not directly measure the impacts of TC-induced
184 precipitation or storm surge. While significant uncertainties remain regarding how climate

185 change will impact the frequency or landfall location of TCs, evidence points to the increasing
186 intensity and heavy precipitation associated with TCs globally (Sobel et al. (2023), Masson-
187 Delmotte et al. (2021), Knutson et al. (2015), Khouakhi et al. (2017)). Furthermore, recent events
188 have led to significant damages even in areas exposed to low wind speeds, due to heavy precip-
189 itation and inland flooding (Schleypen et al. (2024), Bakkensen et al. (2018)). Likewise, increases
190 in sea level have already led to increasing risk of storm-surge flooding from landfalling TCs (Xi
191 et al. (2023); Lin et al. (2016); Glavovic et al. (2022)), and intensifying precipitation and sea level
192 rise can create non-linear increases in compound flood hazard (Wahl et al. (2015), Moftakhari
193 et al. (2017), Bevacqua et al. (2019)). In our current analysis, we do not directly test the impact
194 of TC-induced precipitation or storm surge separate from the the winds due to the large uncer-
195 tainties in their measurement and the persistent challenges in both satellite and ground-based
196 precipitation measurements at the global scale. Our exclusive focus on TC winds was neces-
197 sary in order to maintain global coverage with a consistent exposure metric, with the necessary
198 trade-off being that we do not isolate the impacts of TC-induced precipitation or storm surge
199 separate from TC winds. Because stronger winds are highly correlated with greater precip-
200 itation and storm surge, damages attributed to wind intensity partly reflect these non-wind
201 hazards. Even so, disentangling and quantifying damages attributable to wind, precipitation,
202 and storm surge remains an important avenue for future work.

203 We also provide a novel empirical estimate of TC impacts on a country's debt ratio. Empirical
204 estimates of TC impacts have focused largely on measuring direct losses, GDP, mortality or
205 well-being (Cavallo et al. (2013), Hsiang and Jina (2014), Bakkensen and Mendelsohn (2016),
206 Young and Hsiang (2024), Rappaport (2014), Hallegatte et al. (2016)). Among the few empirical
207 studies that have focused on the debt impact of disasters (Noy and Nualsri (2011); Melecky and
208 Raddatz (2011); Zhang and Chang (2020), see overview in Deryugina (2022)), TCs are not the
209 main focus and these studies rely on a database of damage estimates (e.g. EM-DAT) to identify
210 disaster occurrences, which introduces endogeneity concerns. For instance, only disasters that
211 cross a specific threshold for damage are included in the database (e.g. fatalities greater than
212 10, damages greater than 0.5% of GDP), which means that estimated impacts may potentially
213 be driven by other factors that led the storm to be included in the database in the first place,
214 rather than characteristics of the storm itself (Botzen et al. (2019)). One study that examines
215 TC impacts using a wind field measure considers only countries in the Caribbean region and
216 uses debt service costs as a proxy for measuring debt burdens (Ouattara and Strobl (2013)). Yet
217 other studies have utilized theoretical models to conclude that recovery and access to capital
218 will be negatively impacted among countries following TCs (Phan and Schwartzman (2024),
219 (Bakkensen and Barrage (2025), Mallucci (2022)). Our study focuses explicitly on the impact of
220 TCs using a physical measure (land area affected by varying wind speed intensities) to directly
221 estimate impacts on the debt burden associated with different storm intensities, and utilize
222 comprehensive debt data available for 190 countries (Mbaye et al. (2018)). Thus, we recover
223 a direct empirical understanding of how countries' debt ratios evolve in the aftermath of TCs
224 across countries with varying exposure profiles.

225 The differential responses to TCs compared to warming temperatures suggest that these fac-

226 tors impact the economy through different channels. TCs have been associated with capital
227 destruction and the need for increased capital expenditures during the recovery and recon-
228 struction phase (Melecky and Raddatz (2011)). In contrast, warming temperatures have been
229 associated with GDP impacts mainly through the productivity channel, for instance in terms
230 of agricultural yields or labor productivity Burke et al. (2015). This potentially explains why
231 we recover a clear temperature impact on GDP but not for debt ratios, while TC impacts are
232 recovered for both GDP and debt ratios.

233 There have been increasing calls for sovereign debt relief in recent years, whether through
234 debt cancellations, restructuring, or deferred payment options for countries facing increasing
235 debt burdens (UNCTAD (2023), Government of Barbados (2024), Jubilee Commission (2025)).
236 Several banks have started offering products and new policies for communities impacted by
237 catastrophes (UNEP Finance Initiative, Munich Re (2024)). At the sovereign level, however,
238 no overarching institution exists to coordinate debt relief across official (government) and pri-
239 vate creditors. For example, the IMF-led Debt Service Suspension Initiative (DSSI) during the
240 COVID-19 pandemic only included official creditors, and several eligible countries declined to
241 participate due to concerns that doing so could raise their borrowing costs from private cred-
242 itors (Lang et al. (2023)). Similar concerns have limited the effectiveness of the G20 Common
243 Framework for Debt Treatments (Jubilee Commission (2025), Paris Club and G20 (2020)). Fur-
244 ther complicating debt restructuring efforts is that China, now the largest bilateral creditor to
245 many developing countries, is not a member of the Paris Club, which has been the principal
246 forum for sovereign debt negotiations since 1956 (Horn et al. (2021), United Nations Develop-
247 ment Programme (2025)).

248 To address sovereign debt challenges, The Bridgetown Initiative launched in 2022 calls for in-
249 ternational financial architecture reform, including debt relief measures, transparency in credit
250 rating agencies, and inclusion of 'debt-pause' clauses, or debt repayment suspension for coun-
251 tries following disasters (Government of Barbados (2024)). Grenada's debt-pause clause was
252 the first to be activated in 2024 following Hurricane Beryl, after it was inserted as part of its
253 debt structuring negotiations in 2015 (Asonuma et al. (2018)). Our results suggest these mea-
254 sures could lower sovereign debt burdens by providing immediate liquidity after major TCs,
255 easing the need for new borrowing. For countries with weak or absent credit ratings, a tem-
256 porary suspension of debt-service payments can free scarce capital for disaster response and
257 reduce reliance on high-cost borrowing. Other policy options may involve increasing the share
258 of concessional lending or grants in climate finance alongside innovative mechanisms such as
259 debt-for-nature swaps.

260 4 Conclusion

261 Our findings show that exposure to TCs and warming temperatures over the past three decades
262 has already shaped the cost and availability of capital for many countries. As financing from

263 global capital markets becomes more costly, countries are likely to become further trapped in a
264 “vicious cycle” of debt, unable to access the upfront finance needed to reduce climate impacts
265 in the first place. Ultimately, these results underscore the urgent need to address the growing
266 financial costs borne by countries that have contributed least to historical emissions yet face the
267 greatest impacts of climate change.

Methods

269 Our approach can be broadly summarized as containing three parts. In the empirical part, we
 270 estimate the impact of climate shocks on two key macroeconomic determinants of a country's
 271 sovereign credit ratings: the debt-to-GDP ratio and GDP (Cantor and Packer (1996); Mellios
 272 and Paget-Blanc (2006); Afonso et al. (2011); see also fig. S14). Building on the empirical re-
 273 sults, we then derive counterfactual scenarios for the two variables in the absence of tropical
 274 cyclones and rising temperatures. Finally, we train a prediction model to estimate how the
 275 macroeconomic impacts may have affected sovereign credit ratings and borrowing costs for
 276 countries.

277 **4.1 Empirical model**

278 We employ a local projections (LP) with long differences model (Jordà (2005), Jorda and Taylor
 279 (2025)), an approach that is increasingly common in the macroeconomic literature to directly
 280 estimate the dynamic and cumulative long-run response of outcomes to a shock (e.g. Bilal
 281 and Känzig (2024), Nath et al. (2023), Romer and Romer (2019)). Unlike Vector Autoregres-
 282 sive (VAR) models, LPs estimate impacts sequentially at each step of the forecast horizon with
 283 separate regressions, making them more robust to potential biases arising from model mis-
 284 specification (Montiel Olea and Plagborg-Møller (2021), Montiel Olea et al. (2025), Jorda and
 285 Taylor (2025)).

286 Our main model is specified as below:

$$y_{i,t+h} - y_{i,t-1} = \beta_t^h \cdot \mathbf{tc}_{it} + \delta_t^h \cdot \tau_{it} + \mathbf{x}_{it} + \epsilon_{it} \quad \text{for } h = 0, 1, \dots, H \quad (1)$$

287 The model allows us to estimate the impact of tropical cyclones (β_t) and temperature (δ_t) oc-
 288 ccurring in year t on outcomes (y) for country (i) at increasing time horizons ($t + h$), controlling
 289 for \mathbf{x}_{it} where:

$$\mathbf{x}_{it} = \sum_{l=1}^n \beta_l^h \cdot \mathbf{tc}_{i,t-l} + \sum_{l=1}^n \delta_l^h \cdot \tau_{i,t-l} + \alpha_i + \gamma_t + \phi_i(t)$$

$$\beta_l^h = \begin{bmatrix} \beta_{1,l}^h & \beta_{2,l}^h & \beta_{3,l}^h \end{bmatrix}, \quad \mathbf{tc}_{i,t-l} = \begin{bmatrix} tc_{1,i,t-l} \\ tc_{2,i,t-l} \\ tc_{3,i,t-l} \end{bmatrix}$$

$$\delta_l^h = \begin{bmatrix} \delta_{0,l}^h & \delta_{1,l}^h \end{bmatrix}, \quad \tau_{i,t-l} = \begin{bmatrix} \tau_{i,t-l} \\ \tau_{i,t-l} \cdot \bar{T}_i \end{bmatrix}$$

$$\phi_i(t) \in \{0, \phi_i t, \phi_{1i} t + \phi_{2i} t^2\}$$

290 The vector of coefficients β_t^h and δ_t^h form the impulse response functions (IRF) that estimate
 291 the dynamic cumulative effect of the shock occurring at time (t). x_t represents a vector of controls,
 292 which includes lags of both tropical cyclone and temperature shocks, α_i and γ_t represent
 293 country and year fixed effects, respectively. Country-specific time trends are denoted by $\phi_i(t)$.

294 To credibly isolate the impact of a shock from a given year, we must account for serial corre-
 295 lations in both the outcome variable and the shock variables (Jorda and Taylor (2025)). Even
 296 after isolating the shock variable following the methods outlined below (4.2.1, 4.2.2), weak se-
 297 rial correlation can persist in the data. To account for this, we include up to two lags of the
 298 lagged difference in the outcome variable ($y_{t-1} - y_{t-2}, y_{t_2} - y_{t_3}$), up to 10 lags of the TC shock
 299 variables, and 2 lags of the temperature shock variables. Finally, we also test the model us-
 300 ing Driscoll-Kraay standard errors to address any spatial and temporal serial correlations in
 301 the panel data, which can additionally account for cross-sectional dependencies across units.
 302 We additionally test empirical bootstrap and wild cluster bootstrap as alternative methods for
 303 generating standard errors. A list of robustness checks conducted are shown in fig. S2 and ??.

304 4.1.1 Tropical cyclone exposure

305 To characterize the tropical cyclone shock, we use tropical cyclone wind fields generated from
 306 a parametric wind model (Chavas et al. (2015) that captures the full wind extent of a storm
 307 and explicitly accounts for the asymmetrical structure of tropical cyclones (Chen et al. (2023))
 308 as it evolves over land (Jing et al. (2024)). Wind fields for each tropical cyclone are generated at
 309 30-arcsec spatial resolution (approximately $10 \times 10 \text{ km}^2$). We combine the wind fields from all
 310 tropical cyclone occurrences within a year to compute the maximum wind speed experienced
 311 over land in each grid cell. Wind speeds are then classified into three intensity categories: 18–33
 312 m/s (tropical storms), 33–50 m/s (Category 1–2 storms), and >50 m/s (major storms, Category
 313 3 and above) according to the Saffir-Simpson Hurricane Wind Scale. For each country and
 314 year, we estimate the share of land area exposed to each category of wind speed to ensure
 315 comparability across countries of varying sizes.

316 Let ws_{gy} denote the maximum wind speed in grid cell g during year y , across all storms. A_G
 317 denotes the total land area of a country computed as the total number of grid cells falling within
 318 country borders. The share of land area exposed to wind speed category $c \in \{1, 2, 3\}$ in year y
 319 is computed as:

$$tc_{c,y} = \frac{1}{A_G} \sum_{g=1}^G \begin{cases} 1 & \text{if } ws_{gy} \in T_c \\ 0 & \text{otherwise} \end{cases}$$

320 where the wind speed categories are defined as:

$$T_1 = (18, 33) \text{ m/s}, \quad T_2 = [33, 50) \text{ m/s}, \quad T_3 = [50, \text{max}) \text{ m/s}.$$

321 This yields $tc_{c,y} \in [0, 1]$, representing the proportion of a country's land area exposed to wind
322 speeds in each category during year y . This measure of tropical cyclone exposure does not
323 exhibit serial correlation, unlike temperature, which we take an additional step to isolate the
324 shock as described below (see fig. S5).

325 4.1.2 Temperature exposure

326 We use the ERA5 gridded 2-meter surface temperature dataset ($0.25^\circ \times 0.25^\circ$, 31 km) and con-
327 struct population-weighted, country-level annual temperature exposures for all years. The
328 population weights are from Rossi-Hansberg and Zhang (2025).

329 We build on existing literature (Burke et al. (2015), Nath et al. (2023)) showing that the impact
330 of a hotter or cooler year depends nonlinearly on a country's average temperature, and that
331 serial correlations in temperature should be accounted for. We follow Nath et al. (2023) in
332 constructing temperature shocks as the residuals from a nonlinear autoregressive model that
333 includes lagged temperature terms:

$$T_{it} = \sum_{j=1}^p \gamma_j T_{i,t-j} + \sum_{j=1}^p \theta_j T_{i,t-j} \cdot \bar{T}_i + \mu_i + \mu_t + \tau_{it} \quad (2)$$

334 The residual τ_{it} is the estimated temperature shock.

335 In addition, we consider two alternative ways of isolating the temperature shock: accounting
336 for the persistence of shocks by applying a Hamilton filter as in Bilal and Känzig (2024), and
337 simply removing a country-specific time quadratic trend from the temperature time series. In
338 models that do not specify a country-specific time trend or include a linear time trend, isolating
339 the temperature shock through this latter method yields qualitatively similar results as the
340 other two methods (fig. S4).

341 4.1.3 Economic data

342 For debt-to-GDP ratio we use the Global Debt Database from the IMF, which provides com-
343 prehensive and harmonized data on public and private sector debt for 190 countries with time
344 series extending to the 1950s for advanced economies (Mbaye et al. (2018)). Public sector debt
345 is defined as all debt held by the public sector, including the total gross debt of central, state,
346 and local governments, and social security funds. Public sector debt data is available for more

347 than 40 continuous years for 119 countries and more than 30 years for 142 countries. For coun-
 348 tries that do not have aggregate public sector data reported, we use the general government
 349 data or central government data. General government debt data is recorded for 88 countries
 350 and central government debt data is recorded for a wider sample of 174 countries. We also
 351 test the model using the sample of countries with only the central government data, as well as
 352 government debt data compiled by the World Bank (Kose et al. (2022)) and the Global Macro
 353 Database (Müller et al. (2025)), and find qualitatively similar results (fig. S2). For GDP data we
 354 use GDP per capita in constant 2015 USD from the World Development Indicators as in Burke
 355 et al. (2015) (World Bank (2025)).

356 4.2 Constructing counterfactual scenarios

357 We construct counterfactual scenarios by sequentially removing the influence of tropical cy-
 358 clones and long-term temperature changes. For the debt ratio, counterfactual scenarios are
 359 generated by setting the occurrence of tropical cyclones to zero in each year. For instance, when
 360 constructing counterfactual scenarios beginning in 1990, the impact of all tropical cyclones oc-
 361 ccurring from 1990 onward are removed sequentially, as shown in the main text (fig. 3). This
 362 approach also allows us to selectively remove individual hurricane seasons to assess their spe-
 363 cific impact.

364 Let \mathcal{C} denote the set of countries i and $\mathcal{Y} = \{1990, 1991, \dots, n\}$ the study period. In the main
 365 text we restrict the sample to $n = 2019$. Although year fixed effects absorb economic impacts of
 366 global shocks such as the COVID-19 pandemic that are common across all countries, they may
 367 not capture any idiosyncratic country-specific impacts. Nonetheless, we find that extending
 368 the study period to 2022 does not change our main empirical result.

369 For each country i , we construct a matrix $\mathbf{V}_i \in \mathbb{R}^{(n+1) \times T}$, where each column $\mathbf{v}_{i,y}$ represents a
 370 year-specific counterfactual vector capturing the dynamic effect of a shock occurring in year y .
 371 Each row corresponds to a calendar year $k \in \{1990, \dots, n\}$. Each vector $\mathbf{v}_{i,y}$ is constructed as:

$$v_{i,y}[k] = \begin{cases} \sum_{c=1}^3 \beta_{c,l}^h \cdot ws_{c,i,y}, & \text{if } k = y + h \\ 0, & \text{if } k < y \end{cases}$$

372 where $\beta_{c,l}^h$ denotes the estimated impulse response coefficient at horizon h for wind category c ,
 373 and $ws_{c,i,y}$ denotes the share of land area in country i exposed to wind category c in year y .

374 To compute the total dynamic impact of shocks over the entire study period, we sum the
 375 columns of \mathbf{V}_i to obtain a single cumulative counterfactual vector:

$$\mathbf{v}_i^{\text{total}} = \mathbf{V}_i \cdot \mathbf{1}_T$$

376 where $\mathbf{1}_T \in \mathbb{R}^{T \times 1}$ is a column vector of ones and $T = |\mathcal{Y}|$. For each $i \in \mathcal{C}$, the resulting vector
 377 $\mathbf{v}_i^{\text{total}} \in \mathbb{R}^{n+1}$ represents the cumulative impact of all shocks experienced by country i over the
 378 study period. In the final step, to construct the counterfactual outcomes for each country, we
 379 remove the cumulative impact of shocks from the observed values:

$$y_{it}^{\text{cf}} = \frac{y_{it}^{\text{obs}}}{1 + \mathbf{v}_i^{\text{total}}[t]}$$

380 where $\mathbf{v}_i^{\text{total}}[t]$ denotes the cumulative percentage change in the outcome due to all tropical
 381 cyclone shocks affecting country i up to year t .

382 4.3 Prediction model

383 We estimate the implications of these macroeconomic impacts in terms of countries' credit rat-
 384 ings today and their cost of borrowing. In short, we test the hypothesis that due to the long-run
 385 economic impacts from tropical cyclones and warming temperatures, the ratings agencies may
 386 be assigning ratings to countries that are lower than they would have been absent those shocks.

387 4.3.1 Credit rating data

388 Credit rating data from the "Big Three" rating agencies (Moody's, S&P, Fitch) are obtained
 389 from Bloomberg Finance L.P. and complemented by additional observations from Trading Eco-
 390 nomics (Bloomberg L.P. (2023), Trading Economics (2024)). The data include rating changes as
 391 well as outlook announcements, which are both known to influence market perceptions and
 392 sovereign bond yields (Cantor and Packer (1996), Kenourgios et al. (2020)). Following existing
 393 literature, the ratings are converted into a descending linear scale, with the highest rated bonds
 394 (AAA) ranked 21 and the lowest rated bonds (C) ranked 1 (e.g. Afonso et al. (2012) Kenourgios
 395 et al. (2020)). We add further variation by including the outlook change announcements as 0.5
 396 changes in the rating, as in Eichengreen et al. (2007). For example, an AA rating (rank 19) that
 397 receives a negative outlook is ranked as 18.5. The full conversion table is available in table 2.
 398 Many countries received their first rating starting in the late 1990s (fig. S10) and 46 countries
 399 that have never received a credit rating from the "Big Three" agencies are by definition not
 400 included in the data.

401 4.3.2 Sovereign bond data

402 Sovereign bond data are from Bloomberg and Refinitiv (now part of London Stock Exchange
 403 Group (LSEG)). Bloomberg provides yield data for 69 advanced and emerging market economies.

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Historic yield data are available for this limited set of countries because their market size, trad-
ing volume and liquid currencies makes it possible to construct yield data from the secondary
market. For most countries, however, bond issuance is infrequent and these bonds are not ac-
tively traded in the secondary market. Thus, to include a larger set of countries in our analysis,
we focus on predicting the annual coupon rates of bonds issued in the primary market, or at
the date of issuance.

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Sovereign bond data from Bloomberg and Refinitiv are available at the individual issuance
level. Combined, we recover 48,000 sovereign bond issuances from 132 countries (27,861 ex-
cluding bonds issued by China and Japan), which we use to derive the credit rating and coupon
rate relationship for 10-year maturity bonds during the relatively low interest rate period from
2011-2019 (see section 4.3.4). In our study, data have been filtered to include only bonds with
fixed coupon rates, bullet maturities, and non-zero-coupon bonds. Private placement bonds
are excluded. The dataset includes information on issuance currency, maturity, and issuance
size.

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4.3.3 Types of prediction models

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We compare three different models for prediction: ordered response models (probit), ordered
random forest, and two types of gradient boosted decision-tree models (XGBoost) (table 3).
Each model has strengths and weaknesses in terms of performing the task at hand.

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The ordered probit model is used to predict the probability of an observation being assigned to
a category that is ordinal, conditional on a set of predictor variables. Because sovereign credit
credit ratings follow an ordinal logic, and the spacing between the categories are not equal (i.e.
category thresholds are unequally spaced on the latent scale), ordered probit models have been
commonly employed to understand the determinants of ratings (Blanchard (2022), Ardagna
(2018)). However, the ordered probit model requires the parallel regression assumption, which
means that the marginal effect of predictor variables are consistent across all categories of the
outcome variable. This assumption is often violated in practice: for instance, an increase in
GDP will not have the same effect on the likelihood of a rating upgrade for a country with a BB
rating versus an AA rating. Therefore, we also test Ordered Forest models, a modified form of
random forest models, which allows for nonlinear combinations of variables.

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Extreme gradient boosting (XGBoost) offers a more powerful way to train a prediction model,
where decision trees are used to sequentially minimize a loss function. Even though using
XGBoost does not explicitly recognize the ordinal nature of the dependent variables at the
outset, we find that it is nonetheless able to learn that the rating categories are ordinal during
the training process. On average, our XGBoost model achieves above 70% accuracy and >90%
accuracy with tolerance (± 2) in held out test samples (fig. S13).

439
One might assume that a practical way to implement XGBoost for ordinal outcomes is to treat

440 the ordered categories as continuous variables in a regression task. Intuitively, this mimics the
441 behavior of an ordered probit model that estimates a latent variable to be classified into differ-
442 ent categories based on different thresholds. In ordered probit models, the distance between
443 category thresholds can vary flexibly to account for unequal distances between categories on a
444 latent scale (e.g. moving from B to BB is easier compared to moving from BBB to A). However,
445 because XGBoost regression assumes equal spacing between categories based on the provided
446 inputs, it cannot replicate the uneven spacing across categories.

447 Therefore, we implement XGBoost as a classification task, which has the added benefit of pro-
448 viding a probability distribution across all possible ratings. While we cannot impose the ordinal
449 nature of rating variables in this implementation, we nonetheless find that the model is able to
450 recover the ordinal nature of the ratings (fig. S13).

451 We additionally test XGBOrdinal, a new package that transforms the ordinal classification task
452 into a series of binary classification tasks using XGBoost (Kahl et al. (2025), Frank and Hall
453 2001). While inefficient (the time to train the model increases exponentially with the number
454 of boosting iterations added), XGBOrdinal is better able to capture the ordered nature of the
455 outcomes compared to the basic XGBoost Classification. However, the embedded assumptions
456 in XGBOrdinal, on rare occasions, lead to a violation of the Kolmogorov probability axioms
457 (e.g. avoiding negative probabilities and ensuring all probabilities sum to one) and spurious
458 predictions. We therefore use XGBoost Classification as our main model.

459 After training on forty years of macroeconomic indicators and other variables referenced by
460 rating agencies (e.g. S&P's rating methodology considers variables across five 'pillars': institu-
461 tional, economic, external, fiscal, and monetary). We predict what each country's rating would
462 have been in 2019 using the counterfactual debt-to-GDP ratios and GDP values estimated if
463 countries had not been exposed to climate shocks. The list of variables used in our prediction
464 model are shown in table 4.

465 4.3.4 Translating to borrowing costs

466 Rating changes are translated into borrowing costs, measured as the basis point change in the
467 coupon rate for a fixed 10-year bond. The change in borrowing costs are calculated using
468 the difference in the probability distribution of predicted credit ratings using observed versus
469 counterfactual data. The observed values are from a range of monetary, fiscal, and institu-
470 tional variables in 2019 (table 4). For the counterfactual data, we replace the debt-to-GDP ratio
471 and GDP per capita data with the counterfactual estimates derived from our empirical model,
472 which removes the effect of all land-falling TCs from 1990 onwards and country-specific warm-
473 ing.

474 We use the observed and counterfactual probability distributions of credit ratings to calcu-
475 late the change in expected value of borrowing costs, using a generalized relationship be-

476 tween credit rating and coupon rates for sovereign bonds. This relationship is derived us-
477 ing Bloomberg and Refinitiv data of 9,253 sovereign bonds issued between 2011-2019 by 124
478 countries (17,261 including bonds issued by China and Japan) (fig. S11, fig. S12). This rela-
479 tionship holds across different time periods, maturities, and interest rate environments. JPY-
480 denominated bonds are excluded from the sample as rates have been kept artificially low by
481 the central bank. Utilizing this relationship allows us to generate predictions for both coun-
482 tries with sparse sovereign bond issuances as well as countries that do not have a credit rating
483 (Extended Data Fig. 2). We run the XGB model with 500 random seeds, each time holding out
484 a random 20% of countries, to derive 500 probability distributions of credit ratings for both
485 observed and counterfactual scenarios.

486 We additionally consider the change in borrowing costs for bonds denominated in dominant
487 currencies versus local currencies. Among sovereign bonds, countries issuing in their local
488 home currency have faced coupon rates higher than those issued in a dominant currency (USD,
489 EUR, GBP, CHF, CAD). Between 2011-2019, bonds denominated in local currency have paid
490 coupon rates that are on average 170 basis points, or 1.7% higher than bonds denominated in a
491 dominant or liquid currency (fig. S12).

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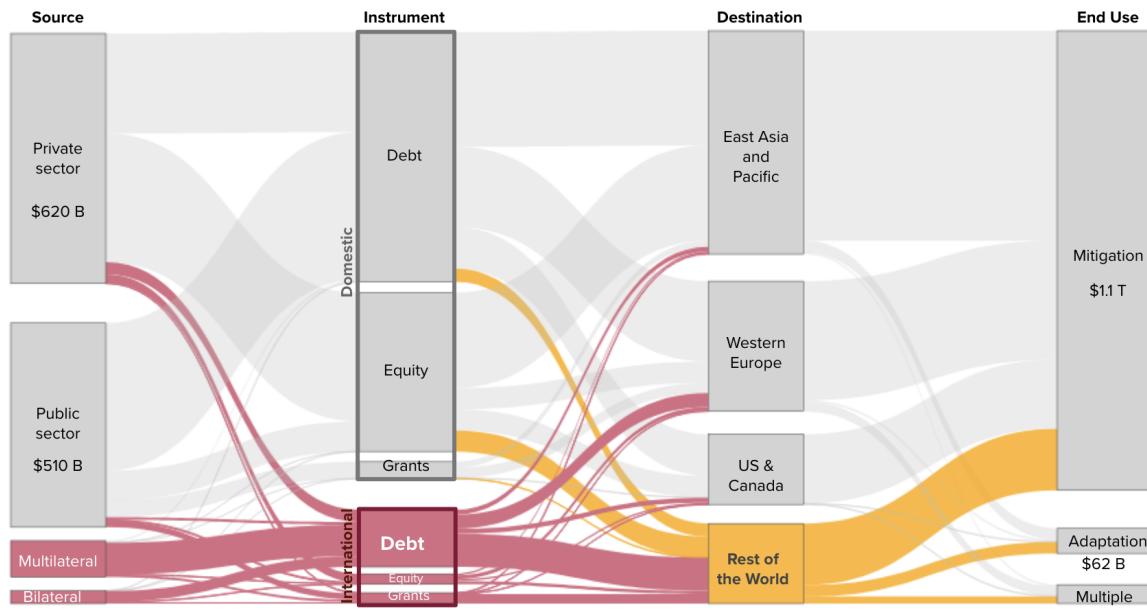
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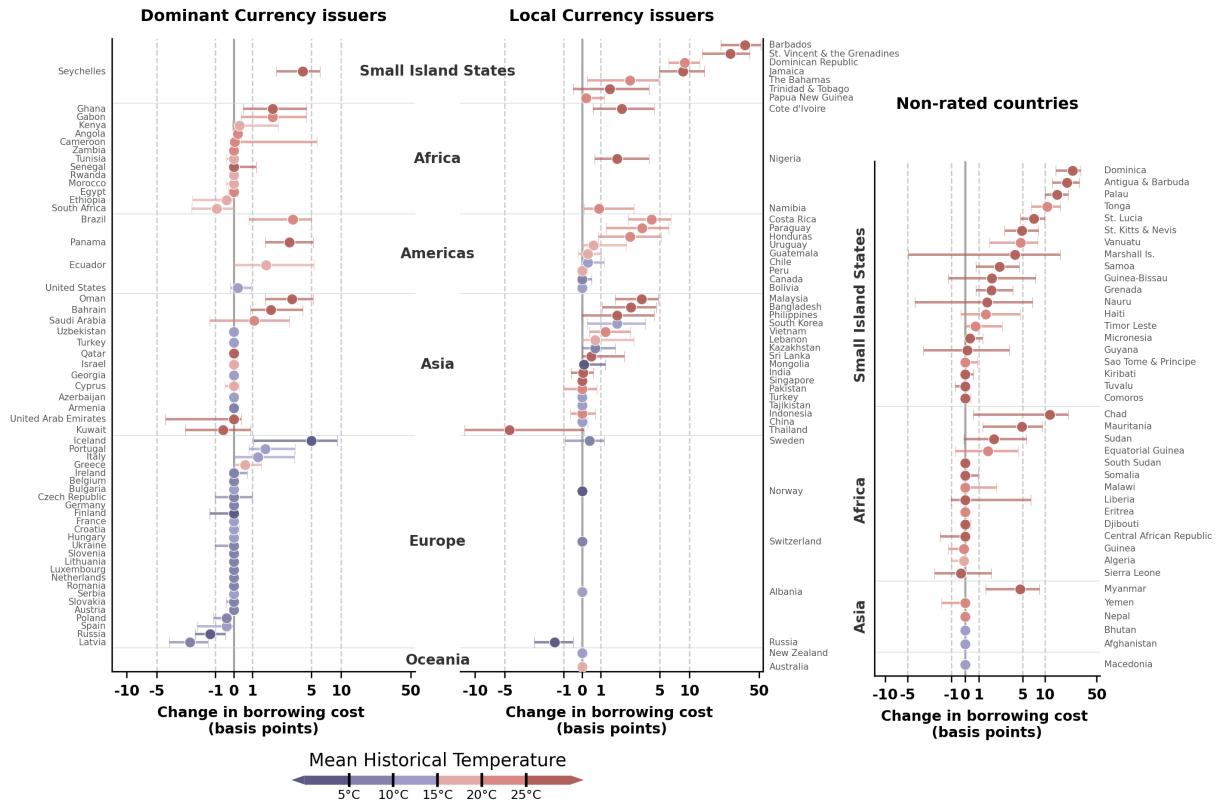
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Supplementary Figures



Extended Data Fig. 1. Global climate finance flows in 2022/2023. Data from Climate Policy Initiative.



Extended Data Fig. 2. The change in borrowing cost attributed to historical tropical cyclones and climate change. Values are plotted separately for countries with that issue predominantly in a dominant currency, local currency, and estimates for non-rated countries. Colors are based on each country's mean historical temperature. Error bars show the inter-quartile range of predictions derived from 500 randomized training samples in the rating prediction model, with each sample holding out 20% of countries.

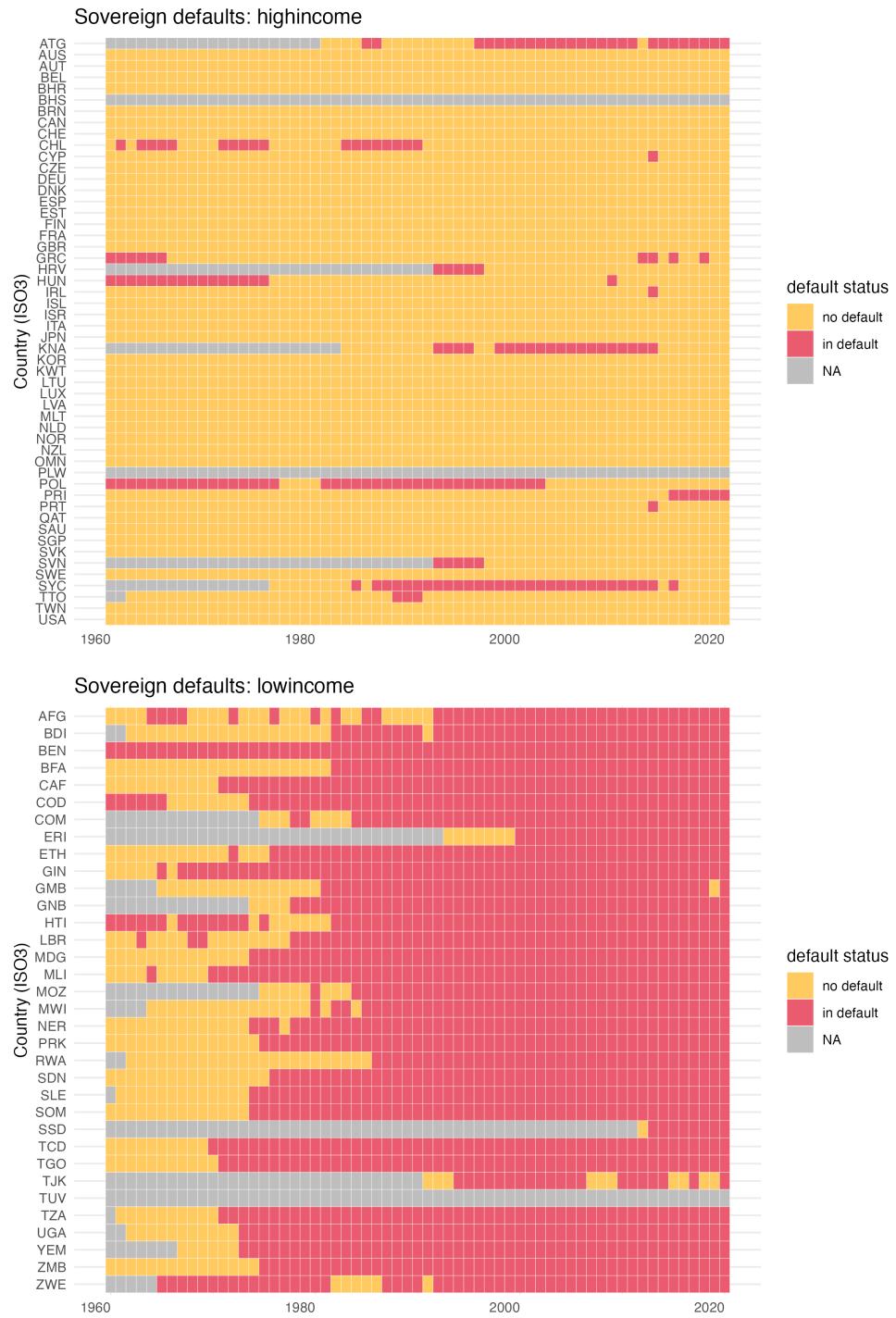


Figure S1. Sovereign default status of high-income (Panel A) and low-income (Panel B) countries over time. Default episodes include any missed payments on loans or bonds falling outside the grace period, owed to both official and private creditors. Data from Bank of Canada and Bank of England (Beers et al. (2021)).

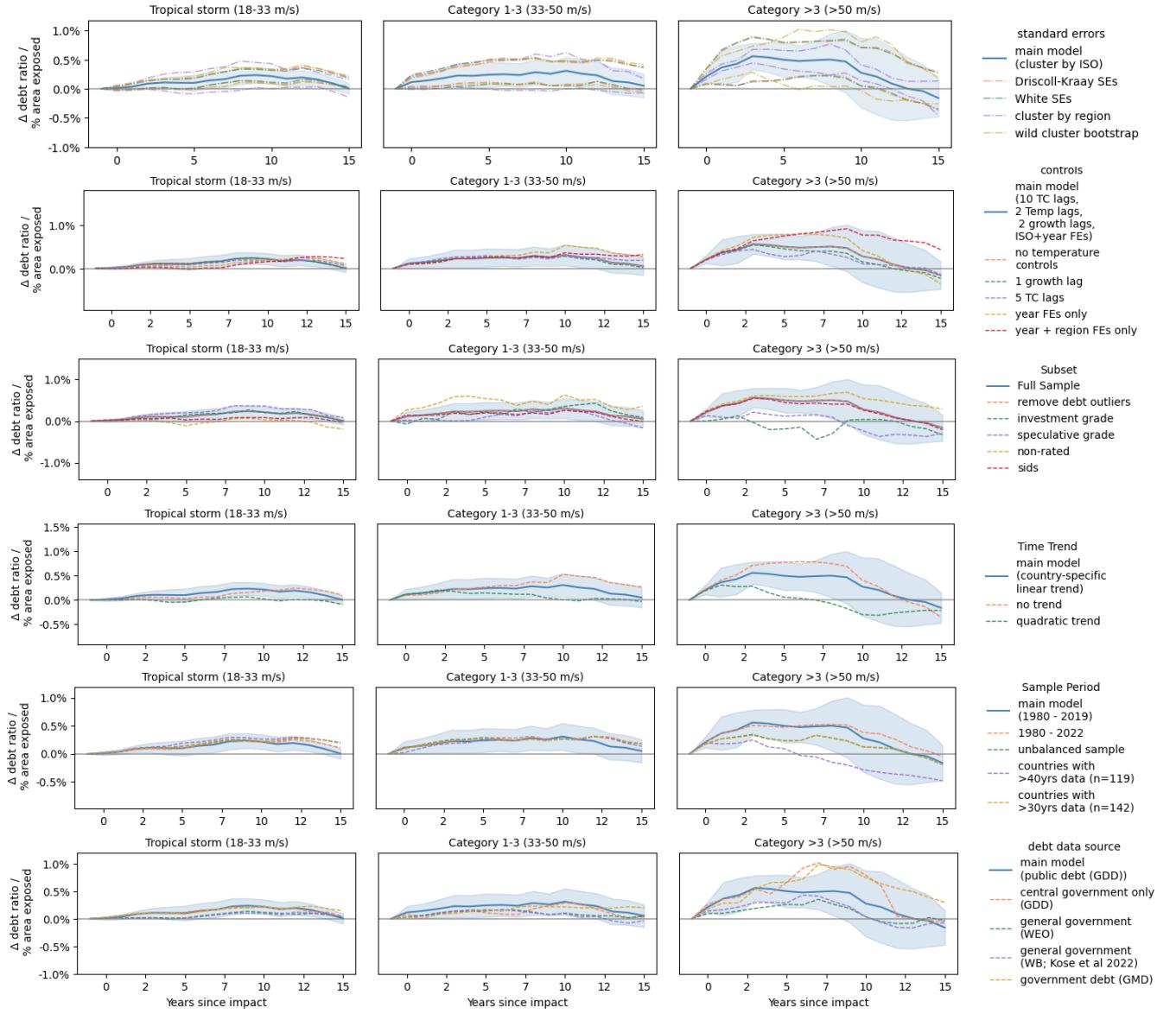


Figure S2. Tropical cyclone impacts on debt ratio. Columns plot the impulse response functions following 1% land area affected by increasing wind speed intensities. Rows are based on robustness checks, by standard error, inclusion of different controls, subsets, time trend, sample period, and the debt data source.

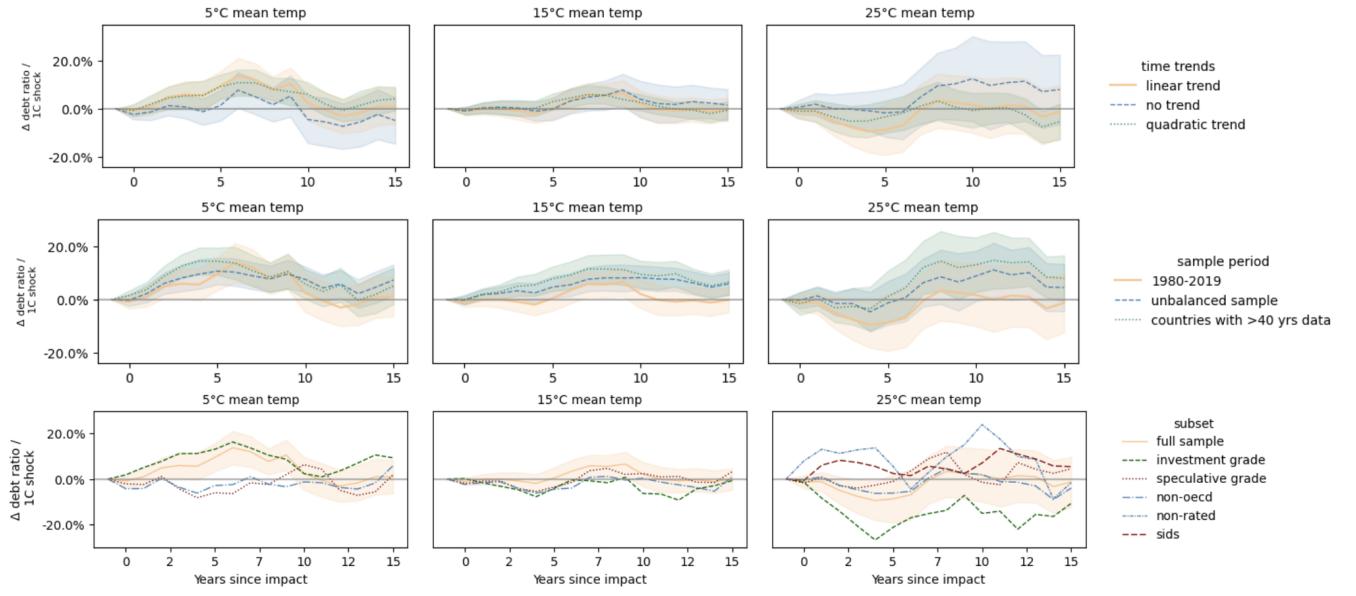


Figure S3. Temperature impacts on debt ratio. Columns plot the impulse response functions following 1°C hotter year, conditional on the countries' historical average temperature (selectively constructed for 5°C, 15°C, 25°C). Rows show sensitivity to time trends, sample period, and subset of countries. The distribution of countries' annual mean temperature is shown in fig. S8.

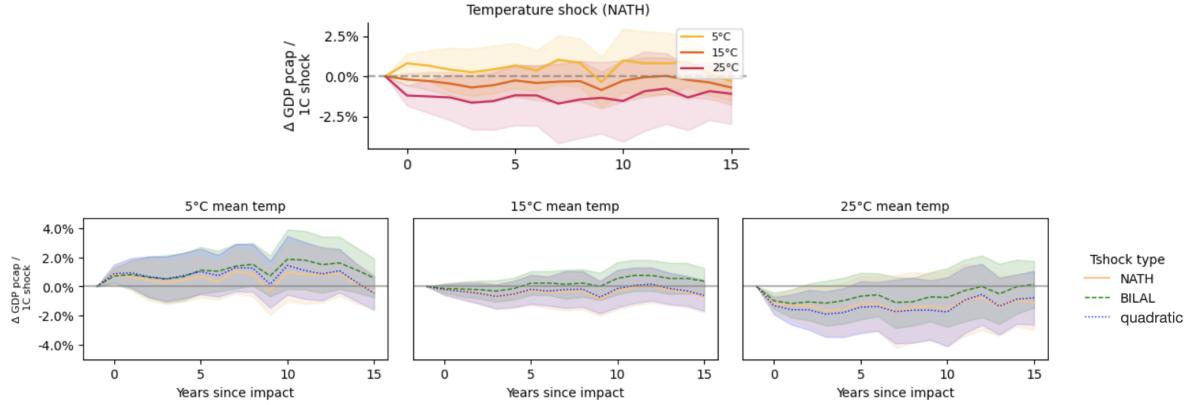


Figure S4. Temperature impacts on GDP. Top panel shows the GDP IRF following a 1°C hotter year, with temperature shocks isolated using method from Nath et al 2024. Bottom panel plots the IRFs (selectively constructed for 5°C, 15°C, 25°C), comparing three different methods of isolating the temperature shock.

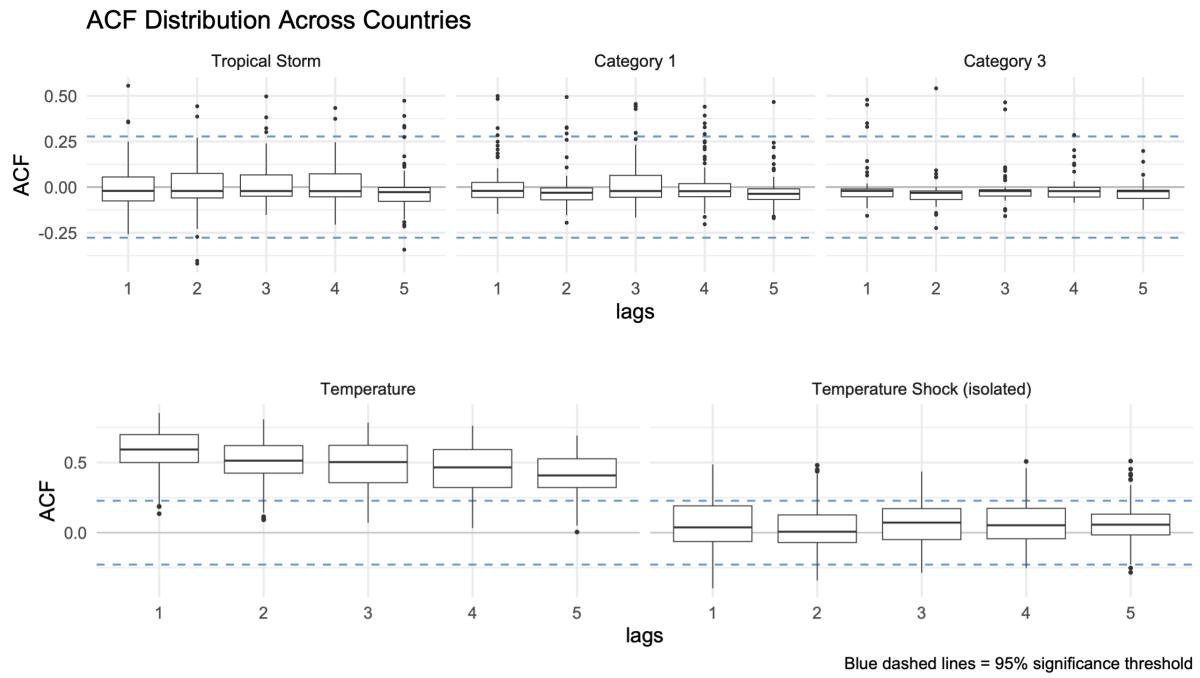


Figure S5. Distribution of autocorrelation function (ACF) plots across countries, up to 5 lags. Top panel shows the ACF plots for each of the TC exposure metric (% land area exposed to different wind intensities). Bottom panel shows the ACF plots for annual temperature, where the temperature shock is not isolated (left) and isolated as in Nath et al 2024 (right).

Change in GDP per capita in 2019 attributable to tropical cyclone and temperature exposure

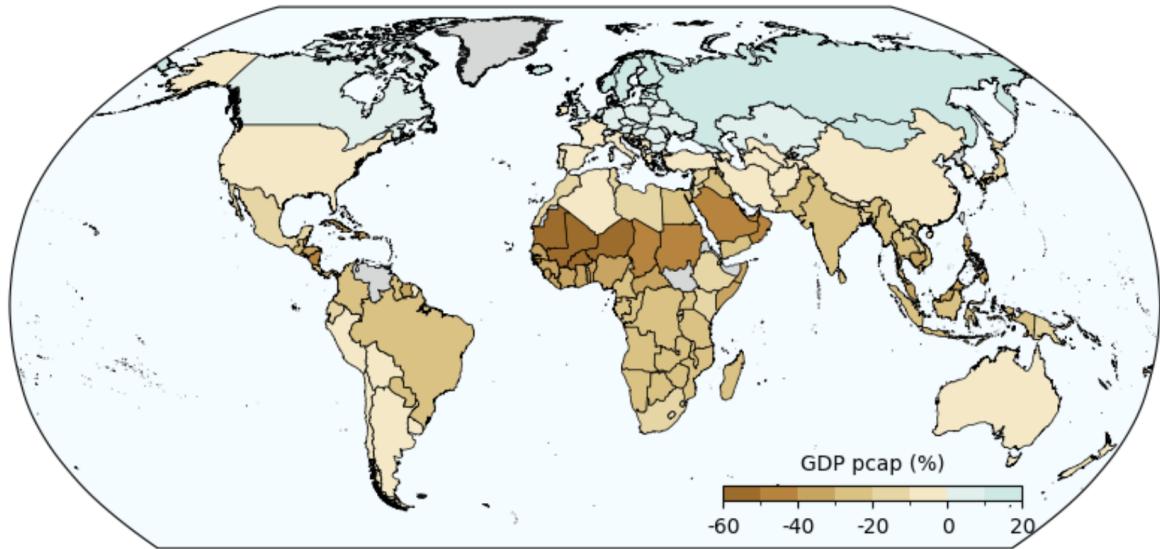


Figure S6. Global map of the share of countries' GDP in 2019 attributable to tropical cyclones and temperature shocks from 1990 onward.

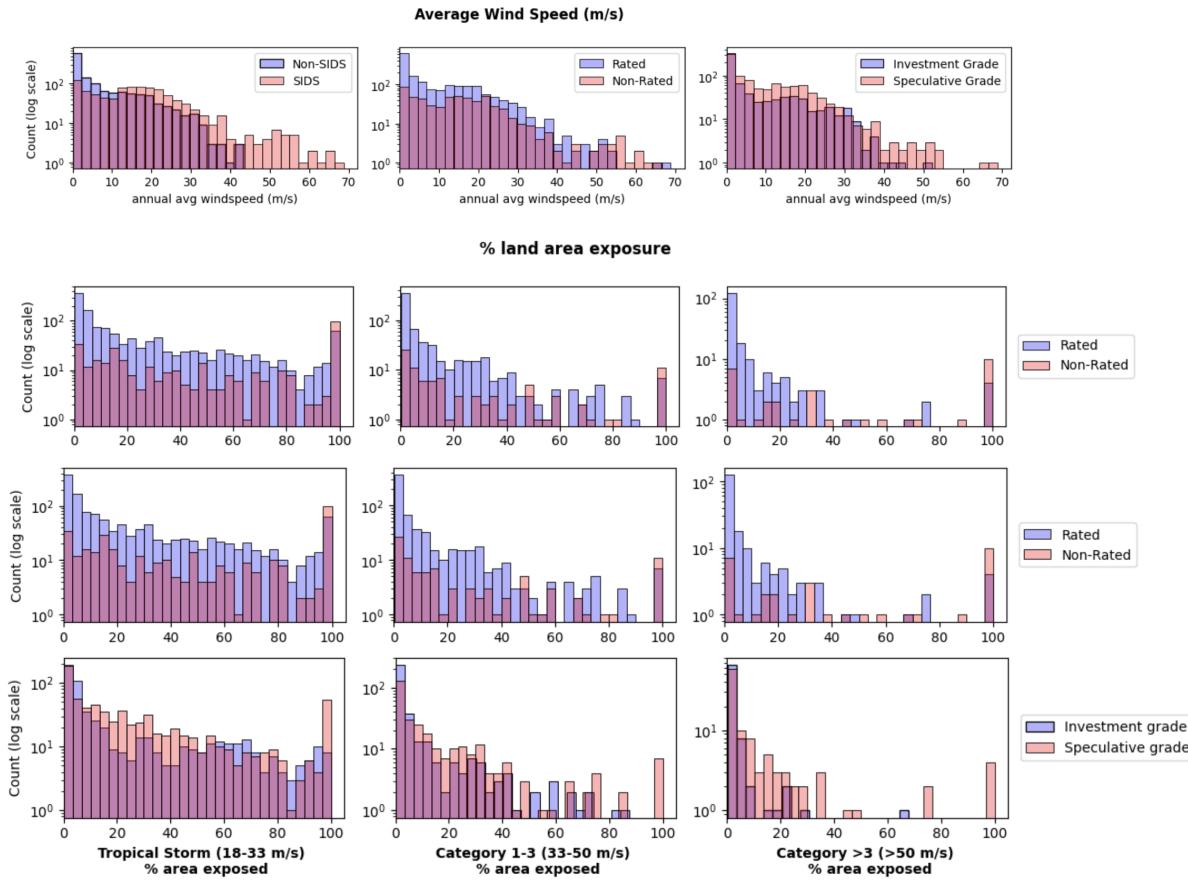


Figure S7. Distribution of annual wind speed exposure by country subsets (Small Island Developing States (SIDS) vs non-SIDS, rating vs non-rated countries, investment grade vs speculative grade countries).

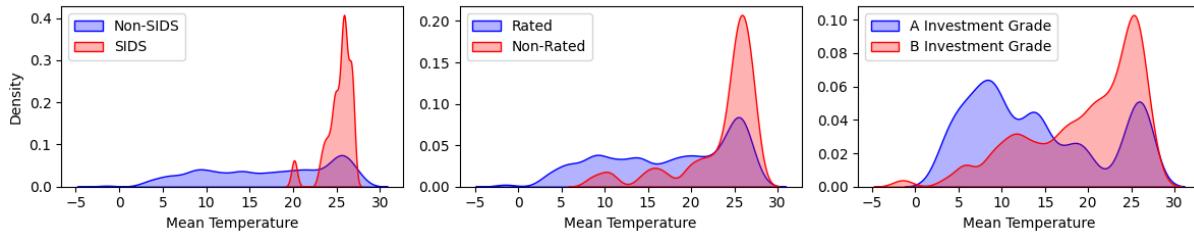


Figure S8. Distribution of annual average temperatures by country subsets (Small Island Developing States (SIDS) vs non-SIDS, rating vs non-rated countries, investment grade vs speculative grade countries).

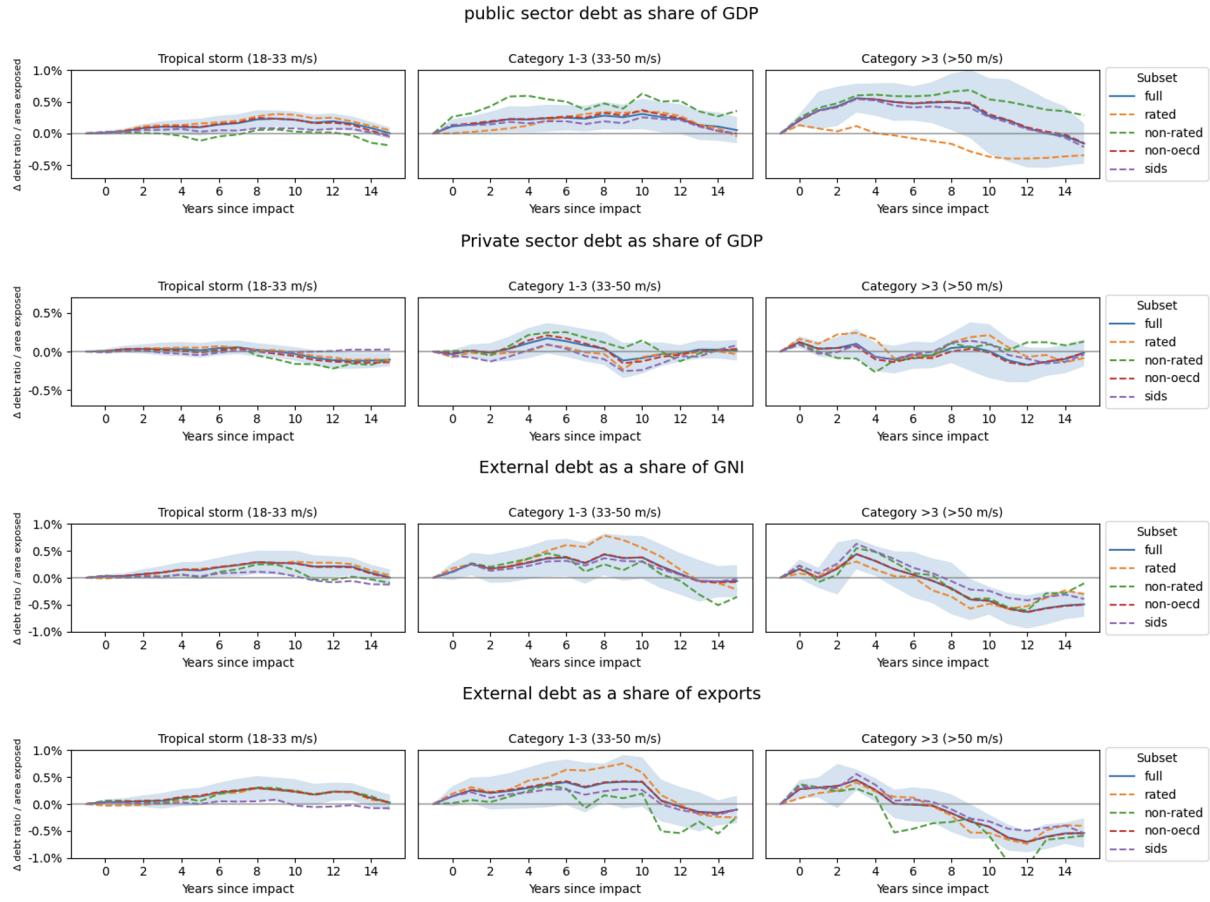


Figure S9. Impulse response functions for alternative debt metrics.

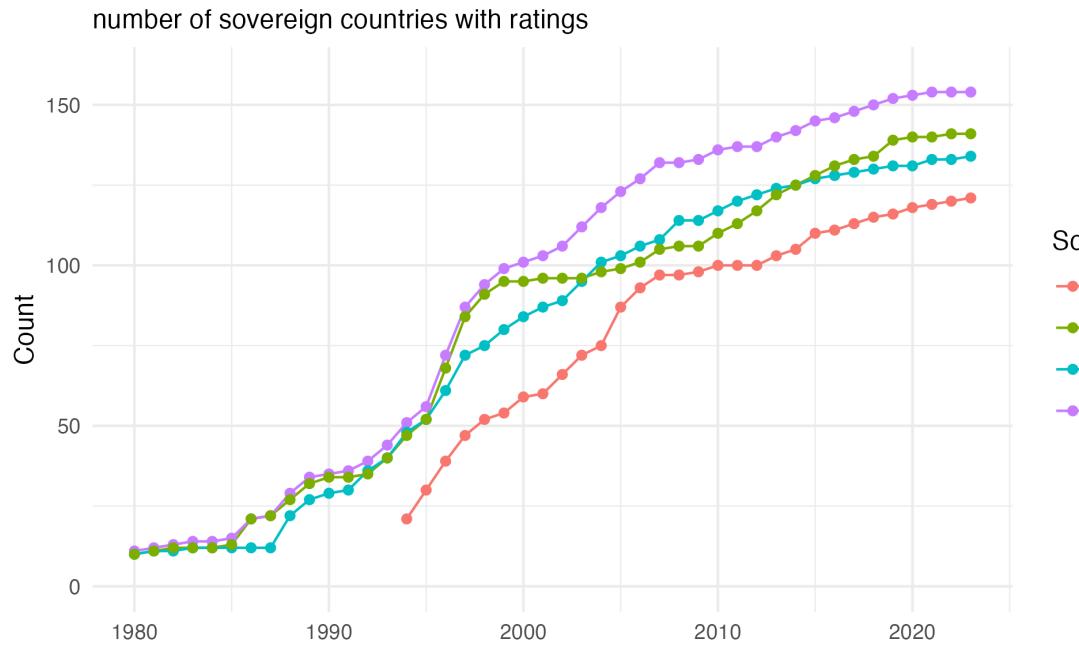


Figure S10. The number of unique sovereign countries with a rating from the three major credit rating agencies. The total number of countries with an assigned rating is plotted separately, as some countries do not have ratings from all agencies.

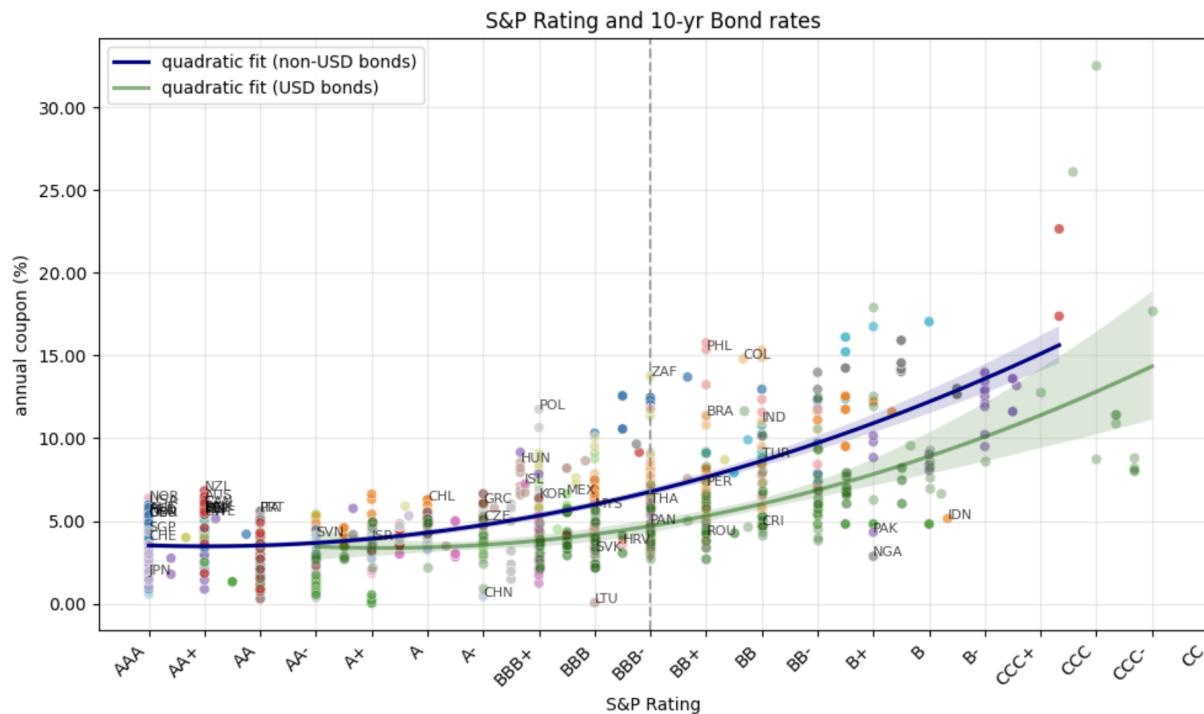


Figure S11. Scatter plot of sovereign credit ratings and annual coupon rates associated with 10-yr maturity bond issuances. Data from Bloomberg available for 40 developed and emerging market countries.

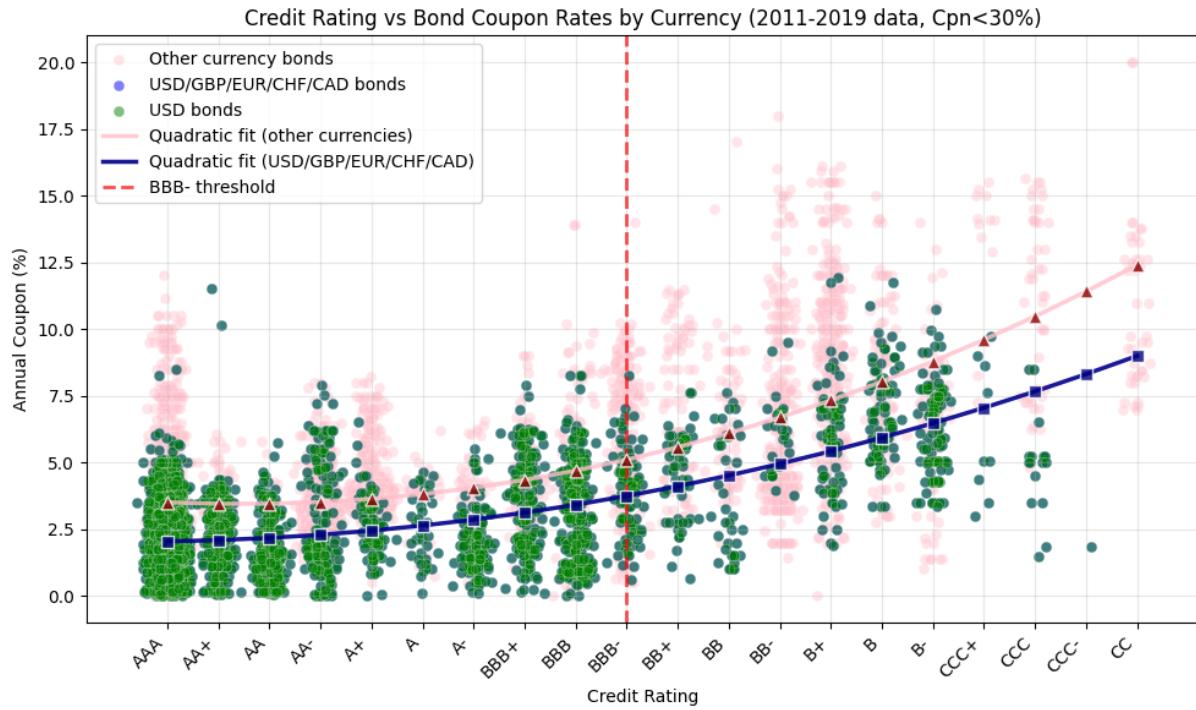


Figure S12. Fitted relationship between sovereign credit ratings and coupon rates, by currency of denomination. Dominant currency are bonds denominated in USD, EUR, GBP, CHF or CAD and local currency are bonds denominated in any other currency. Data from London Stock Exchange Group available for 124 countries with bond issuances between 2011-2019.

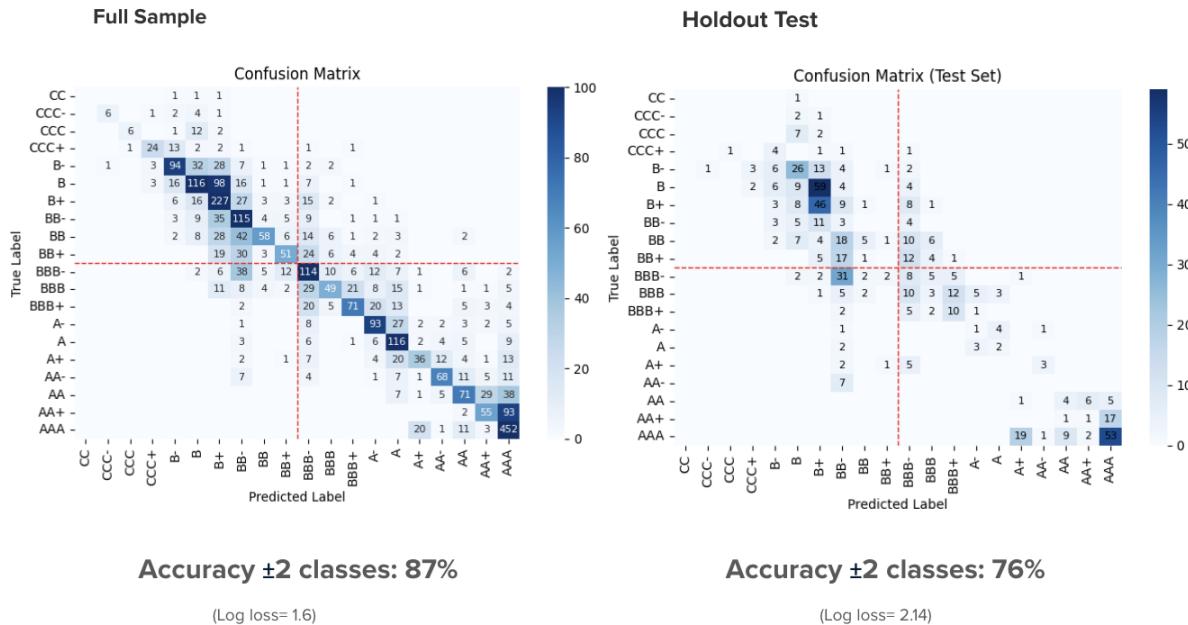


Figure S13. Confusion matrix for the sovereign credit rating prediction model using XGBoost for one random seed. Left panel is for the full sample of countries, right panel is for 20% of held out countries.

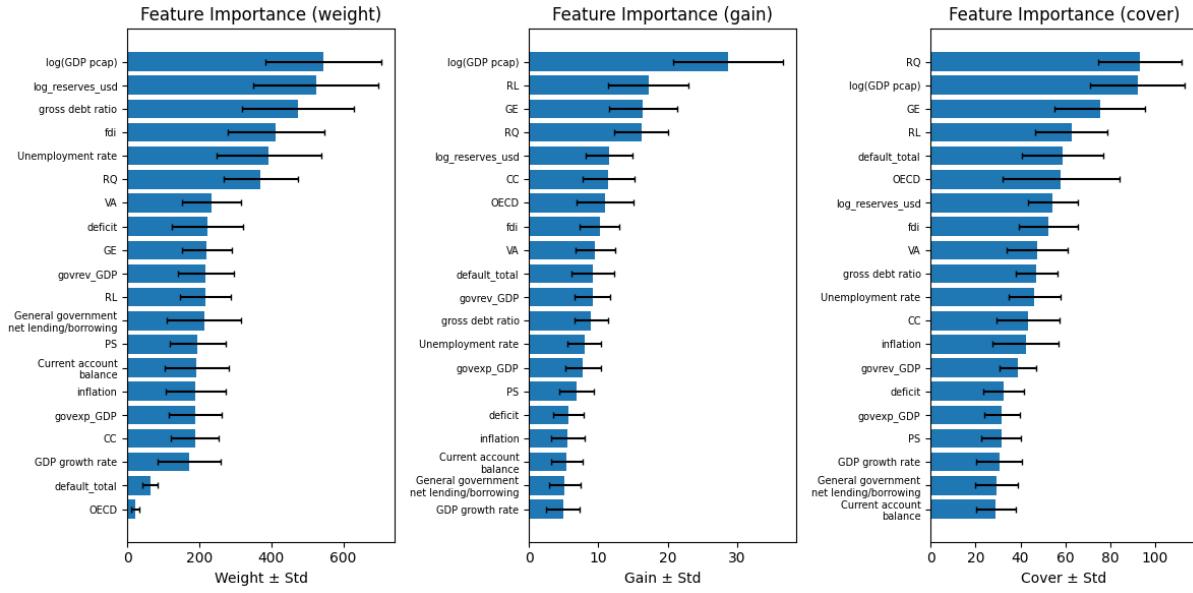


Figure S14. Feature importance rankings from the XGBoost credit rating prediction model. "Weight" describes how often a feature is used; "Gain" describes how useful a feature is in minimizing loss; "Cover" describes how much a feature reduces uncertainty. The error bars show the range of 500 random seeds.

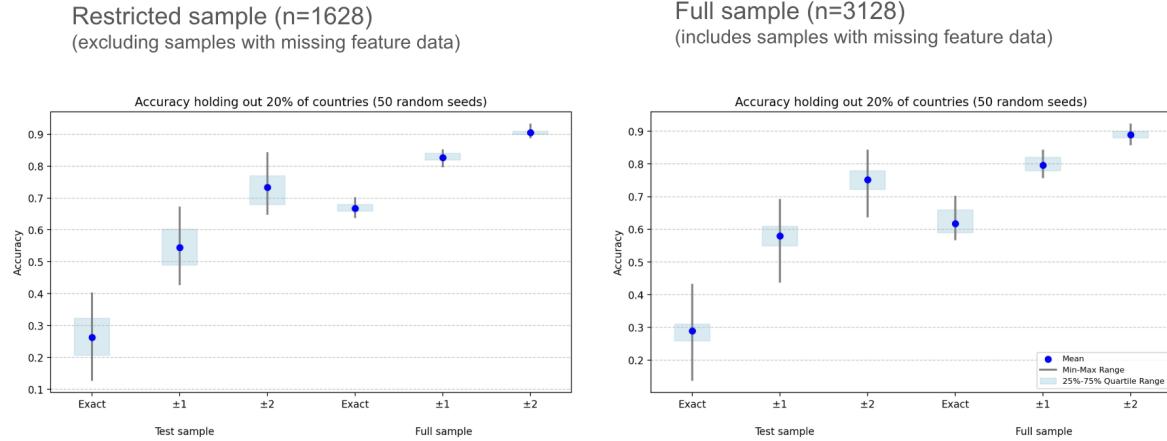
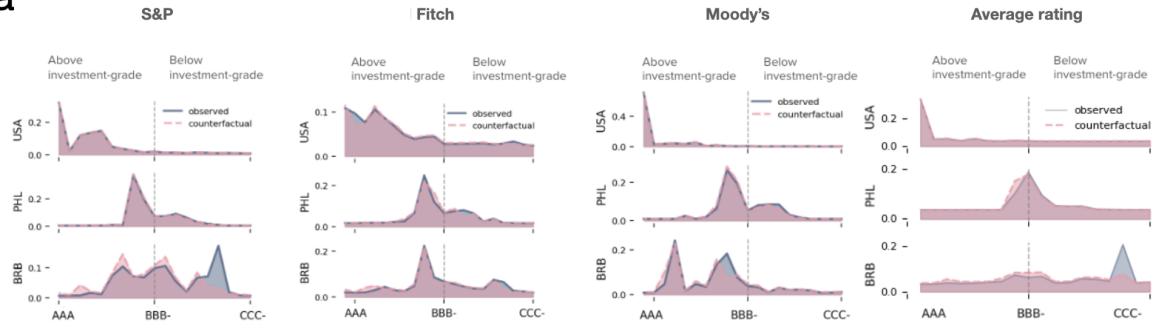
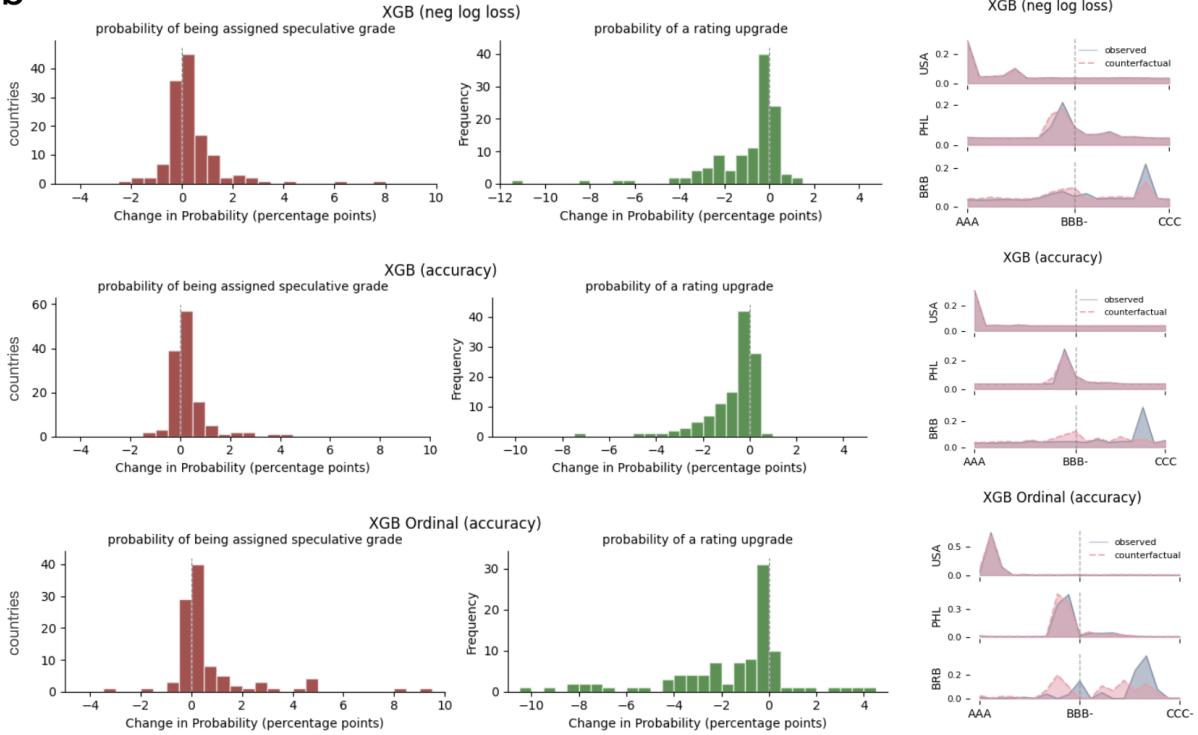


Figure S15. Accuracy of the XGBoost credit rating prediction model for held out test samples and the full sample. The shaded area of the box plots show the inter-quartile range of accuracy after randomly assigning out 20% of countries for the held out test sample. Left panel shows the results for a restricted sample where observations with missing feature data are excluded. Right panel shows the results where observations with missing feature data are included (XGBoost treats the missing data as information, which allows us to take advantage of greater volume of training data).

a

(a)

b

(b)

Figure S16. Counterfactual credit ratings. Panel a shows the probability distribution of sovereign credit ratings for three selected countries (The United States, The Philippines, Barbados) using observed vs counterfactual debt-to-GDP ratio and GDP in 2019, based on rating data from S&P, Fitch, and Moody's. Our main model uses the average score across three rating agencies. Panel b plots the distribution of how much more likely countries are to be assigned a below-investment-grade rating (left) or the probability of receiving a rating upgrade (right) in the counterfactual scenario. Results from different training models are shown, for XGB classification using log loss as the scoring method (our main model), using accuracy as the scoring method, and XGB Ordinal.

Tables

Table 1. Robustness Checks

Category	Specification
Model specifications	<ul style="list-style-type: none"> • Testing TCs and temperature shocks separately • Testing each of the TC metrics separately • Time trends (no trend, linear, quadratic) • Alternative metric of TC exposure (using average wind speed, as in Hsiang 2014)
Time period	<ul style="list-style-type: none"> • Sample from 1980–2019 (removing COVID years; main model) • Sample from 1990–2022 (longest period preserving balanced sample, with exposure shocks starting in 1980)
Testing subsets + removing outliers	<ul style="list-style-type: none"> • Removing large TC-impacted economies (China, US) • Removing debt ratio outliers (Venezuela, Argentina) • Testing subsets of countries (rated vs. non-rated countries)
Selecting control variables	<ul style="list-style-type: none"> • Lags of country GDP growth • Without temperature controls • Lags of TC, temperature shocks
Standard error treatments	<ul style="list-style-type: none"> • Wild bootstrap (jittering the residuals) • Empirical bootstrap (country block) • Driscoll-Kraay standard errors (adjust for both serial correlation over time and cross-sectional dependence across units)
Different GDP debt datasets	<ul style="list-style-type: none"> • Global Debt Database (main model; based on IMF data, Mbaye et al. 2018) • World Economic Outlook / World Bank (Kose et al. 2022) • Global Macro Database (Müller et al. 2025)

category	ordinal scale	Moody's	S&P	Fitch
investment-grade	21.0	Aaa	AAA	AAA
	20.5	Aaa*/Aa1**	AAA*/AA+**	AAA*/AA+**
	20.0	Aa1	AA+	AA+
	19.5	Aa1*/Aa2**	AA+*/AA**	AA+*/AA**
	19.0	Aa2	AA	AA
	18.5	Aa2*/Aa3**	AA*/AA-**	AA*/AA-**
	18.0	Aa3	AA-	AA-
	17.5	Aa3*/A1**	AA-*/A+**	AA-*/A+**
	17.0	A1	A+	A+
	16.5	A1*/A2**	A+*/A**	A+*/A**
	16.0	A2	A	A
	15.5	A2*/A3**	A*/A-**	A*/A-**
	15.0	A3	A-	A-
	14.5	A3*/Baa1**	A+*/BBB+**	A+*/BBB+**
	14.0	Baa1	BBB+	BBB+
	13.5	Baa1*/Baa2**	BBB+*/BBB++	BBB+*/BBB++
	13.0	Baa2	BBB	BBB
	12.5	Baa2*/Baa3**	BBB*/BBB-**	BBB*/BBB-**
	12.0	Baa3	BBB-	BBB-
speculative-grade	11.5	Baa3*/Ba1**	BBB-*/BB+**	BBB-*/BB+**
	11.0	Ba1	BB+	BB+
	10.5	Ba1*/a2**	BB+*/BB**	BB+*/BB**
	10.0	Ba2	BB	BB
	9.5	Ba2*/Ba3**	BB*/BB-**	BB*/BB-**
	9.0	Ba3	BB-	BB-
	8.5	Ba3*/B1**	BB-*/B+**	BB-*/B+**
	8.0	B1	B+	B+
	7.5	B1*/B2**	B+*/B**	B+*/B**
	7.0	B2	B	B
	6.5	B2*/B3**	B*/B-**	B*/B-**
	6.0	B3	B-	B-
	5.5	B3*/Caa1**	B-*/CCC+**	B-*/CCC+**
	5.0	Caa1	CCC+	CCC+
	4.5	Caa1*/Caa2**	CCC+*/CCC++	CCC+*/CCC++
	4.0	Caa2	CCC	CCC
	3.5	Caa2*/Caa3**	CCC*/CCC-**	CCC*/CCC-**
	3.0	Caa3	CCC-	CCC-
	2.5	Caa3*/Ca**	CCC*/CC**	CCC-*/CC**
	2.0	Ca	CC	CC
	1.5	Ca*	CC*	CC*
	1.0	C	C	C

Table 2. Conversion table of ratings from Moody's, S&P, Fitch to a descending linear ordinal scale. Ratings greater than BBB-/Baa3 are considered investment-grade ratings.

	Ordered Probit	Ordered Forest	XGB Classification	XGBoost Ordinal
What	Maximum likelihood estimation (MLE) to find parameter values that best fit the observed data	Random forest bagging + imposing ordinal nature of outcome variables	Gradient boosted decision trees; trees are built sequentially to minimize loss	Gradient boosted decision trees; trees are built sequentially to minimize loss
Pros	Naturally preserves ordinal nature of outcome variables	Account for nonlinear combinations of predictors.	Account for nonlinear combinations of predictors, and missing data also counts as information. Model is able to learn the ordered nature of outcome variables.	Ordinal nature of outcomes are preserved through a series of binary predictions.
Cons	Requires parallel regression assumption	Performs poorly in replicating the ordinal nature of outcomes 'Probabilities' for predictions are assigned based on the 'voting' behavior of trees	Need to be careful with overfitting.	Need to be careful with overfitting. On rare occasions, violates the Kolmogorov axioms of probability

Table 3. Table listing different rating prediction models that were tested in this study, their pros and cons.

Variable	Definition	Source
Current account balance	The difference between a country's savings and investments, reflecting net trade in goods and services plus net income and transfers (% of GDP)	World Economic Outlook
General government net lending/borrowing	Government fiscal balance; surplus or deficit as a percentage of GDP	World Economic Outlook
Gross debt ratio	Total government debt as a percentage of GDP	World Economic Outlook
log(GDP pcap)	Log of GDP per capita	World Economic Outlook
GDP growth rate	Annual percentage change in real GDP	World Economic Outlook
Deficit	Government budget deficit, typically expenses minus revenues (% of GDP)	World Economic Outlook
log_reserves_usd	Natural logarithm of foreign exchange reserves in USD; indicates external liquidity buffer	World Economic Outlook
Foreign Direct Investment (FDI)	Cross-border investment flows into a country (% of GDP)	World Economic Outlook
Inflation	Annual percentage change in consumer price level	World Economic Outlook
Unemployment rate	Percentage of the labor force without jobs but actively seeking work	World Economic Outlook
default_total	Indicator of whether a country is in some form of sovereign debt default in a given year	BoC-BoE Sovereign Default Database
govexp_GDP	Government expenditure as a percentage of GDP	World Economic Outlook
govrev_GDP	Government revenue as a percentage of GDP	World Economic Outlook
Voice and Accountability (VA)	Captures perceptions of citizens' ability to participate in government selection and freedom of expression	World Governance Indicators
Regulatory Quality (RQ)	Captures perceptions of the government's ability to formulate and implement sound policies that promote private sector development	World Governance Indicators
Government Effectiveness (GE)	Captures perceptions of public service quality and policy implementation credibility	World Governance Indicators
Rule of Law (RL)	Captures perceptions of contract enforcement, property rights, police, and court quality	World Governance Indicators
Control of Corruption (CC)	Captures perceptions of the extent to which public power is exercised for private gain	World Governance Indicators

Table 4. Table listing the prediction variables, definitions, and data source used in the XGB model.