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1 **On the seasonal predictability of the 2020 North Atlantic tropical cyclone**
2 **season**

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10 ABSTRACT: The 2020 Atlantic tropical cyclone (TC) season was unprecedented, producing a
11 record storm count that exceeded the projected ranges of seasonal forecasts, despite their anticipa-
12 tion of above-normal activity. Here we assess the predictability of the extreme 2020 season from
13 sea surface temperature (SST) forcing by examining a hierarchy of statistical, dynamical, and deep
14 learning (DL) modeling frameworks, all constrained by observed SSTs. Although with observed
15 SST forcing, physics-based models anticipated a relatively active season, the observed outcome lay
16 beyond their ensemble ranges. To investigate this discrepancy, we first evaluate whether the large-
17 scale environmental conditions of 2020 were conducive to the record-breaking activity. Although
18 observed large-scale conditions in 2020 were generally favorable, they were not exceptionally so
19 relative to other highly active seasons and did not indicate that a record-breaking season should have
20 occurred. We next construct a 1,000-member ensemble using a DL model forced with observed
21 SSTs to quantify the distribution of plausible outcomes. Using this large ensemble, the observed
22 season emerges as a 0.5% event: highly unlikely, but not implausible. These results suggest that
23 SST forcing provided only moderate predictive constraint, and that internal atmospheric variability
24 could have played a role in the observed hyperactivity on top of the moderately favorable large-scale
25 conditions. The inability of physics-based model ensembles to encompass the observed outcome
26 does not indicate model failure or missing predictors, but reflects both limited predictability of the
27 season from SST forcing and the inability of small ensembles to sample extreme tail risk.

28 SIGNIFICANCE STATEMENT: The 2020 Atlantic tropical cyclone (TC) season was the most
29 active on record. Although seasonal forecasts anticipated above-normal activity, they did not
30 encompass the observed hyperactivity. Even state-of-the-art physics-based, statistical, and deep
31 learning (DL) models, when given “perfect” observed sea surface temperature (SST) inputs, did
32 not capture the observed outcome within their simulated ranges. We assess how predictable this
33 extreme season was from the observed SSTs. We show that while the large-scale conditions
34 in 2020 were generally favorable for TCs, they were not exceptionally conducive to a record-
35 breaking season. Using a 1,000-member DL ensemble forced with observed SSTs, we find that
36 the 2020 season was an extremely rare but plausible outcome that a small ensemble could not have
37 represented.

38 1. Introduction

39 Tropical cyclones (TCs) that form in the North Atlantic Ocean pose severe hazards to communities
40 across the United States and the Caribbean (Young and Hsiang 2024; Weinkle et al. 2018; Klotzbach
41 et al. 2018; Pielke et al. 2008). Accurate forecasts are therefore essential for preparedness and risk
42 mitigation. Short-term weather forecasts, issued several days in advance, provide critical guidance
43 for evacuation decisions and emergency response. In contrast, seasonal forecasts, issued before the
44 onset of the hurricane season, aim to anticipate the overall characteristics of the upcoming season,
45 particularly the range in possible storms that will form throughout the course of the season.

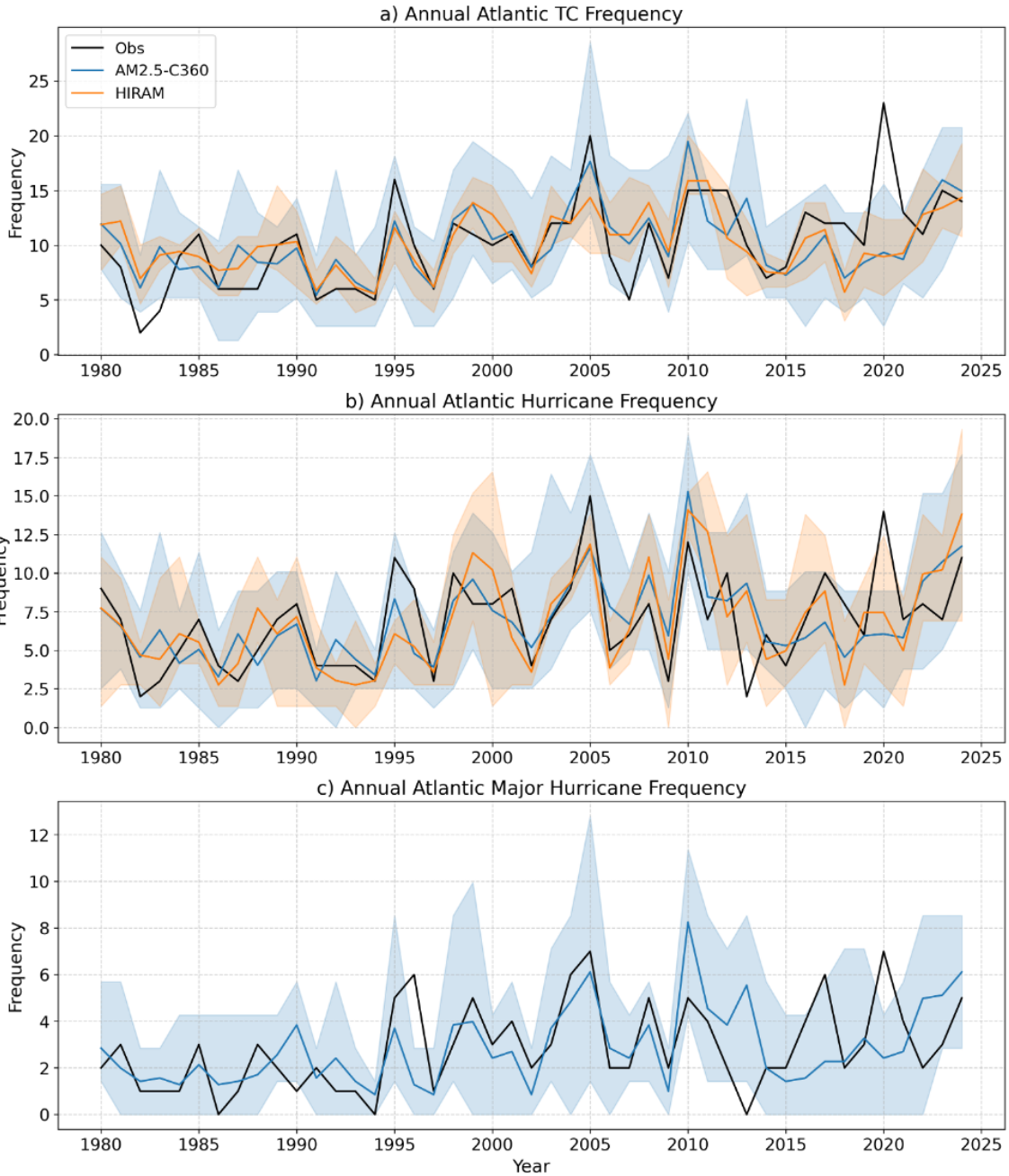
46 Many seasonal modeling and prediction systems rely on sea surface temperatures (SSTs) as a
47 primary source of predictability. Atmosphere-only dynamical models, including both physics-
48 based and deep learning (DL) approaches, can generate seasonal TC forecasts using predicted
49 SSTs or extensions of pre-season SST anomalies as their boundary conditions (e.g., Chen and
50 Lin 2013, 2011; Gao et al. 2019; Zhang et al. 2025a; Zhao et al. 2010). Similarly, retrospective
51 experiments commonly force these models with observed SSTs to isolate the component of seasonal
52 predictability attributable to ocean boundary forcing (e.g., Chan et al. 2021; Chien et al. 2025;
53 Kortum et al. 2024; Levin et al. 2026; Zhao et al. 2009). Statistical and statistical-dynamical
54 frameworks likewise use SSTs as predictors of seasonal TC activity (e.g., Eusebi et al. 2025a;
55 Vecchi et al. 2011; Villarini et al. 2010). Although these approaches do not aim to reproduce the

56 exact storms that occurred, they provide physically plausible realizations of seasons consistent with
57 the specified SST forcing.

58 The 2020 Atlantic TC season presents a striking test of this SST-based predictability frame-
59 work. By many measures, the season was the most active on record, producing the high-
60 est number of long-duration TCs and tying the record for major hurricanes, while ranking
61 second in hurricane count (Fig. 1). Although seasonal forecasts correctly anticipated above-
62 normal activity, they substantially underestimated the magnitude of the hyperactivity. For exam-
63 ple, NOAA’s May 2020 outlook ([https://www.cpc.ncep.noaa.gov/products/outlooks/
64 hurricane2020/May/hurricane.shtml](https://www.cpc.ncep.noaa.gov/products/outlooks/hurricane2020/May/hurricane.shtml)) projected 13–19 named storms, 6–10 hurricanes, and
65 3–6 major hurricanes, compared with the observed 30, 14, and 7, respectively.

66 Atmosphere-only dynamical models forced with observed monthly SSTs also failed to encompass
67 the observed hyperactivity across their ensemble ranges. Two state-of-the-art models developed
68 at the Geophysical Fluid Dynamics Laboratory (GFDL), HiRAM and AM2.5-C360, have demon-
69 strated skill in simulating historical interannual and multidecadal Atlantic TC variability under
70 observed SST forcing (Levin et al. 2026). Yet for 2020, despite simulating a moderately active
71 season relative to the historical record, neither model’s ensemble range (ten members for AM2.5-
72 C360, and five members for HiRAM), captured the observed storm counts across any intensity
73 category (Fig. 1). While observations occasionally fall outside ensemble spread in individual
74 years (e.g., 1982 and 2013), as expected from internal variability (Vecchi and Villarini 2014), the
75 2020 season represents the most extreme discrepancy from the ensemble range in the historical
76 record for long-duration TCs, and it was the only year where the observed count lay outside the
77 ensemble range across all intensity bins. Although the 2020 season’s activity was remarkable, such
78 hyperactivity is not without precedent. Even the similarly hyperactive 2005 season fell within the
79 simulated range of AM2.5-C360, with both models indicating strongly enhanced activity.

87 The inability of these dynamical models’ ensemble range to capture the 2020 TC season when
88 forced with observed SSTs raises a central question: to what extent was the hyperactive 2020
89 season predictable from the observed SST forcing? The inability of SST-forced models to capture
90 the observed outcome across their ensemble spread could indicate deficiencies in the models
91 or observational inputs, or it could instead reflect the intrinsic limits of predictability, with 2020



80 FIG. 1. Annual frequency of (a) TCs, (b) hurricanes, and (c) major hurricanes from 1980–2024. The
 81 observational record (black) is computed using the methodology and duration thresholds of Landsea et al.
 82 (2010a) to estimate the observed TC count in panel (a). Model simulations include AM2.5-C360 forced with
 83 observed and bias-corrected SSTs from Chan et al. (2021) (blue) and HiRAM forced with the same SSTs.
 84 Shading denotes the full ensemble range (minimum–maximum) for each model (ten members for AM2.5-C360;
 85 five members for HiRAM), and solid lines indicate ensemble means. Panel (c) shows only AM2.5-C360 results,
 86 as HiRAM does not produce storms of sufficient intensity to be classified as major hurricanes.

92 representing a rare but dynamically plausible realization that a small ensemble size failed to sample.
93 We evaluate these competing explanations through two broad hypothesis families:

94 The first family of hypotheses considers whether deficiencies in models or observational datasets
95 contributed to the poor simulation of 2020 Atlantic TC counts. Within this framework, we assess
96 the following possibilities:

- 97 1. Atmosphere-only models may exhibit structural limitations in their representation of TC
98 formation, reducing their reliability in simulating interannual Atlantic TC variability.
- 99 2. Observed storm counts in recent decades, as shown in Fig. 1 (black lines), may be inaccurate.
- 100 3. Errors may be present in the observed monthly SST datasets used to force the models in 2020.
- 101 4. SST-forced models may not adequately represent the influence of observed aerosol changes on
102 TC activity. Two major events in 2020 likely altered aerosol concentrations over the Atlantic
103 basin. First, International Maritime Organization (IMO) regulations on the sulfur content of
104 shipping fuel took effect on 1 January 2020, reducing maximum sulfur content from 3.5% to
105 0.5%. Implemented to improve air quality, this policy reduced aerosol loading over the oceans
106 and altered regional radiative fluxes (Diamond 2023; Jordan and Henry 2024; Zhang et al.
107 2025b). Second, the onset of the COVID-19 pandemic led to substantial reductions in global
108 emissions. Previous work suggests that decreased aerosol concentrations over North America
109 and Europe can be associated with enhanced Atlantic TC activity (Murakami 2022, 2024).
110 Because the SST-forced simulations analyzed here employ prescribed CMIP5 aerosol forcings
111 rather than observed 2020 aerosol fields, the combined radiative and dynamical effects of these
112 aerosol perturbations may not be represented.

113 The second hypothesis family considers the possibility that the hyperactive 2020 season was
114 a highly unlikely, but dynamically plausible, realization of the internal variability of the climate
115 system that could not have been captured by the small ensemble sizes of the GFDL models.
116 Under this interpretation, the large-scale environmental conditions in 2020 may have permitted
117 a wide range of outcomes, including a low-probability hyperactive season, but there would have
118 been limited predictability from the monthly mean observed SST forcing. Subseasonal variability
119 and internal atmospheric fluctuations could have amplified activity beyond the ensemble mean
120 response, leading to an outcome that was not well captured by the small ensemble size of the

121 models. To fully address this second family of hypotheses, this paper seeks to answer the following
122 research questions:

- 123 1. Was the 2020 season an extreme tail event in the small AM2.5-C360 and HiRAM ensembles
124 whose likelihood could only be quantified by a very large calibrated ensemble?
- 125 2. To what extent could the hyperactive 2020 TC season have been predictable from the observed
126 monthly SST forcing?

127 In the following sections we apply statistical analyses, dynamical diagnostics, DL-based climate
128 simulations, and a theoretical framework to evaluate the two competing hypothesis families and the
129 posed research questions. The evidence presented in this paper suggests that the 2020 season was
130 a low-probability outcome, rather than a structural failure in the models or observational record.
131 Our results suggest that the observed 2020 SST forcing provided limited predictability for the
132 extreme hyperactivity that occurred, with the observed season representing a tail outcome that was
133 not sampled by the relatively small GFDL ensembles. At the same time, we cannot exclude the
134 possibility that additional mechanisms not examined in this paper influenced the 2020 season.

135 **2. Methods and data**

136 *a. Dynamical Models*

137 To place the 2020 Atlantic TC season in the context of recent historical variability, we generate
138 a multi-ensemble record of TC, hurricane, and major hurricane activity using two global TC-
139 permitting atmospheric models developed at GFDL: AM2.5-C360 and HiRAM (Zhao et al. 2009).
140 Following the methods of Levin et al. (2026), we conducted multi-ensemble historical experiments
141 with both models. We generated ten (five) ensemble members for AM2.5-C360 (HiRAM), each
142 forced with bias-corrected observed monthly SSTs from the HadISST dataset (Chan et al. 2021)
143 over the period 1980–2024. From these simulations, we compute annual counts of TCs, hurricanes,
144 and major hurricanes to construct a modeled range of historical Atlantic TC activity.

145 To evaluate hypothesis 4 from Section 1, which proposes that the 2020 reduction in atmospheric
146 aerosols was not represented in the SST-forced simulations, we conduct additional experiments.
147 We generate a five-member ensemble of AM2.5-C180 (a slightly lower-resolution configuration of
148 AM2.5-C360 with approximately 50-km grid spacing) spanning 1980–2024. In these simulations,

149 the model is forced with observed aerosol fields from the MERRA-2 reanalysis (Gelaro et al. 2017),
150 rather than the CMIP5 aerosol reconstructions used in the historical integrations. This experiment
151 isolates the effect of realistic interannual aerosol variability on modeled TC counts.

152 As a second approach, we perform an idealized modeling experiment to obtain additional samples
153 and better quantify the relationship between sulfate and black carbon aerosols and Atlantic TC
154 activity. Using the HiRAM model, we perform four simulations to isolate the effects of sulfate
155 aerosols, which declined markedly in 2020 following the implementation of the IMO shipping
156 regulation, and black carbon, which also decreased in 2020 due to reduced industrial activity and
157 travel during the COVID-19 pandemic.

158 We perform a control experiment (cntl) in which the model is forced by the annual cycle of
159 SSTs averaged over the period 1986–2005, following the methodology of Vecchi et al. (2019).
160 This climatological SST cycle is repeated for 200 years following a 10-year spin-up, such that the
161 200-year mean represents the model’s steady-state climatological baseline. We then perform three
162 additional experiments using the same SST forcing as the control. In the first experiment, sulfate
163 aerosols are removed over the North Atlantic (zeroSO4NA), and the model is integrated for 50
164 years. In the second experiment, sulfate aerosols are removed globally (zeroSO4global), and the
165 model is again integrated for 50 years. In the final experiment black carbon emissions are removed
166 globally (zeroBCglobal), and the model is integrated over 50 years. This serves as an exaggerated
167 analogue to the global reduction in black carbon brought on by the COVID-19 pandemic.

168 To identify TCs and seed disturbances in both models, we follow the tracking procedure described
169 in Levin et al. (2026). Hurricanes are defined as TCs that attain maximum winds exceeding 33 m s^{-1}
170 at least once during their lifetime. Major hurricanes are identified in AM2.5-C360 as storms whose
171 maximum winds exceed 45 m s^{-1} , a threshold slightly lower than the observed 50 m s^{-1} standard
172 because the model rarely produces storms at the observed major hurricane intensity. HiRAM does
173 not reliably simulate storms of sufficient intensity to meet this major hurricane threshold, and
174 therefore major hurricane statistics are reported only for AM2.5-C360.

175 *b. Statistical-Dynamical Model*

176 To characterize the range of plausible seasonal Atlantic TC counts over the recent historical
177 period (1980–2024), we employ the statistical–dynamical model of Vecchi et al. (2011) that uses

178 monthly SST data as an input and source of seasonal TC predictability. This framework estimates
179 the expected annual Atlantic TC count, λ , using a Poisson regression in which λ depends solely
180 on SST. λ is modeled as a function of the mean seasonal SST anomaly in the Atlantic main
181 development region (MDR) and the mean seasonal SST anomaly across the tropics, since these
182 two predictors together have been shown to be a robust indicator of Atlantic TC activity in the
183 present climate (Vecchi and Soden 2007; Villarini et al. 2011; Eusebi et al. 2025b). Specifically,
184 the logarithm of λ is modeled as the following linear function:

$$\lambda = \exp(1.707 + 1.388, SST_{MDR} - 1.521, SST_{TROP}), \quad (1)$$

185 where SST_{MDR} and SST_{TROP} are defined relative to the 1982–2005 climatology. The coefficients
186 of Equation 1 have not been updated from those of Vecchi et al. (2011). The MDR is defined as
187 10° – 25° N, 80° – 20° W, and the tropical mean spans 30° S– 30° N.

188 This formulation provides a full probabilistic distribution of seasonal outcomes conditional on
189 the SST state. Thus, for each year in 1980–2024, the model yields a distribution of possible TC
190 counts. We use this distribution to quantify the percentile rank of the observed 2020 season relative
191 to SST-forced expectations.

192 We further use this model to evaluate hypothesis 3 from Section 1, which posits that discrepancies
193 between observed and simulated 2020 TC counts may arise from inconsistencies in the observed
194 SST datasets used to force atmosphere-only models. The AM2.5-C360 and HiRAM simulations
195 analyzed here (Section a) are forced with a corrected version of the HadISST dataset (Chan et al.
196 2021). To assess the sensitivity of the inferred TC distribution to the choice of SST product,
197 we apply the statistical–dynamical model separately using three independent SST datasets: the
198 Optimum Interpolation Sea Surface Temperature dataset (OISST; Huang et al. (2021)), the Hadley
199 Centre Global Sea Ice and Sea Surface Temperature dataset (HadISST; Schneider et al. (2013)),
200 and the Extended Reconstructed Sea Surface Temperature dataset (ERSST; Huang et al. (2017)).

201 *c. DL-based Model*

202 Although the multi-ensemble historical integrations with AM2.5-C360 and HiRAM provide a
203 dynamical range of Atlantic TC outcomes and require substantial computational resources, this

204 study questions whether their ensemble sizes of ten and five members, respectively, are= too small
205 to fully characterize the probabilistic distribution of plausible 2020 seasonal outcomes.

206 Recent advances in DL weather models have shown substantial promise for simulating TCs.
207 Once trained, these models can be integrated efficiently with significantly lower computational
208 costs compared to traditional dynamical models like AM2.5-C360 and HiRAM (e.g. (Bi et al.
209 2023; Chen et al. 2023; Lam et al. 2023; Lang et al. 2024; Kochkov et al. 2024)). In this study,
210 we employ the Ai2 Climate Emulator version 2 (ACE2; Watt-Meyer et al. (2025)) to perform
211 multi-ensemble annual simulations of historical Atlantic TC seasons, including 2020. We use a
212 version of ACE2 trained on the ERA5 reanalysis (Hersbach et al. 2020). Previous studies have
213 demonstrated that ACE2 is numerically stable over simulations spanning hundreds to thousands of
214 years (Watt-Meyer et al. 2025) and can realistically reproduce TC characteristics across subseasonal
215 to interannual timescales (Chien et al. 2025). These properties make ACE2 a robust platform for
216 generating large ensembles of Atlantic TC seasons and for investigating the range of plausible
217 outcomes for extreme events such as the 2020 season.

218 We generate a 1,000-member SST-forced ensemble for each year during 2005–2020 (including
219 the anomalously active 2005 and 2010 seasons) and for the anomalously inactive 1982 season.
220 Following Chien et al. (2025), ensemble members for a given year are generated by repeatedly
221 integrating the ACE2 model under identically fixed SST boundary forcing. For a given target year
222 (e.g., 2020), observed SSTs from that year are prescribed and repeated cyclically. The model is
223 then integrated autoregressively for 100 consecutive years under these repeating SST boundary
224 conditions. Each simulated year represents one dynamically distinct realization of the same SST-
225 forced year and is treated as an independent ensemble member.

226 To obtain 1,000 members efficiently, we perform ten parallel integrations (“chunks”), each
227 producing 100 ensemble members via the 100-year repeated-SST integration described above.
228 Distinct initial atmospheric states are used for each chunk. Specifically, we initialize ACE2
229 from publicly available initial conditions and integrate the model forward for one year using SST
230 boundary conditions from the year preceding the target year (e.g., 2019 SSTs when constructing the
231 2020 ensemble). The final model state from this spin-up integration is then used to initialize one
232 100-year repeated-SST chunk. This procedure is repeated to generate ten distinct initial conditions,
233 yielding $10 \times 100 = 1,000$ ensemble members. Because SST boundary conditions are held fixed

234 within each ensemble while atmospheric initial conditions differ, the resulting spread reflects
 235 internal atmospheric variability under identical SST forcing. This design allows us to quantify the
 236 distribution of plausible seasonal TC outcomes conditioned on the observed SST state.

237 To detect TCs in ACE2, we follow the methodology of Chien et al. (2025) and Watt-Meyer et al.
 238 (2025), utilizing the TempestExtremes framework (Ullrich and Zarzycki 2017). By generating a
 239 1,000-member ensemble for the 2020 season, we can robustly estimate the percentile rank of the
 240 observed 2020 storm count within the modeled distribution and to characterize the tail behavior
 241 of seasonal variability. In doing so, we can assess whether the observed hyperactivity represents
 242 a statistically rare but dynamically plausible realization, consistent with the second family of
 243 hypotheses outlined in Section 1.

244 *d. Theoretical Framework*

245 To assess whether the large-scale environmental conditions in 2020 were favorable for a hyperac-
 246 tive Atlantic season, we apply the theoretical framework of Hsieh et al. (2020) to ERA5 reanalysis
 247 data (Hersbach et al. 2020) over the period 1980–2024.

248 The framework conceptualizes TC formation as a two-stage process involving precursor seed
 249 disturbances and subsequent development into fully formed TCs. Under this formulation, the
 250 annual number of North Atlantic TCs, N_{TC} , is approximated as

$$N_{TC} \approx SPI \times P(\Lambda), \quad (2)$$

251 where SPI is the seed propensity index, a proxy for the number of precursor seed disturbances,
 252 and $P(\Lambda)$ is a proxy for the nondimensional probability that a seed disturbance develops into a TC.
 253 The seed propensity index is defined as

$$SPI = (-\omega) \cdot \frac{1}{1 + Z^{-1/\sigma}}, \quad (3)$$

254 where

$$Z = \frac{f + \zeta}{\sqrt{|\beta + \partial_y \zeta| U}}. \quad (4)$$

255 The variable ω is the 500 hPa pressure coordinate vertical velocity ($\omega < 0$ for upward motion),
 256 $\sigma = 0.69$ is a nondimensional fitting parameter, and Z is nondimensional variable that represents
 257 low level vorticity spinup. f is the Coriolis parameter, β is its meridional gradient, ζ is 850 hPa
 258 relative vorticity, and U is a constant wind speed of 20 m/s which is empirically fit in Hsieh et al.
 259 (2020).

260 Next, $P(\Lambda)$ is defined as

$$P(\Lambda) = \frac{1}{1 + (\Lambda_0/\Lambda)^{1/\gamma}}, \quad (5)$$

261 where

$$\Lambda = \frac{\nu_s \cdot \chi}{PI}. \quad (6)$$

262 The variables $\Lambda_0 = 0.014$ and $\gamma = -0.9$ are dimensionless fitting parameters, and Λ denotes the
 263 ventilation index (Tang and Emanuel (2010); Tang and Emanuel (2012)). ν_s is vertical wind
 264 shear between 250 hPa and 850 hPa, PI is potential intensity, and χ is a dimensionless variable
 265 representing entropy deficit.

266 In this study, we extend the existing framework by incorporating monthly mean data from
 267 ERA5, averaged over the tropical Atlantic (10° – 30° N) during June through November, to derive a
 268 proxy for TC activity. Basin-wide means are calculated over a region extending beyond the Main
 269 Development Region (10° – 25° N, 80° – 20° W) to include the Gulf of Mexico, in order to account
 270 for the high concentration of storms that occurred there in 2020 (Fig. S4). Including the Gulf of
 271 Mexico ensures that differences in TC track locations among years are appropriately represented
 272 in the analysis.

273 We also construct an alternative proxy for annual Atlantic TC counts that explicitly incorporates
 274 tracked precursor seeds. In this formulation, annual TC activity is approximated as

$$N_{TC} \approx N_s \times P(\Lambda), \quad (7)$$

275 where N_s denotes the observed annual count of explicitly tracked TC seeds. Here, seed frequency
 276 is diagnosed directly from reanalysis data, allowing us to more explicitly separate variability in
 277 seed frequency from variability in development probability.

278 **3. Results**

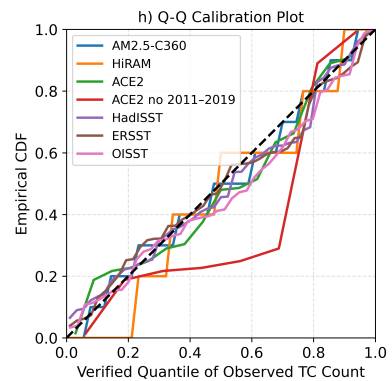
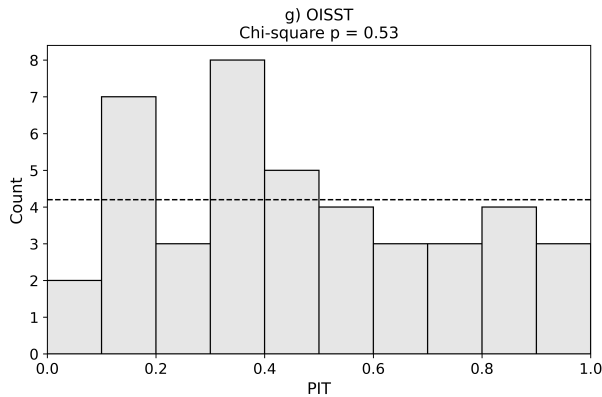
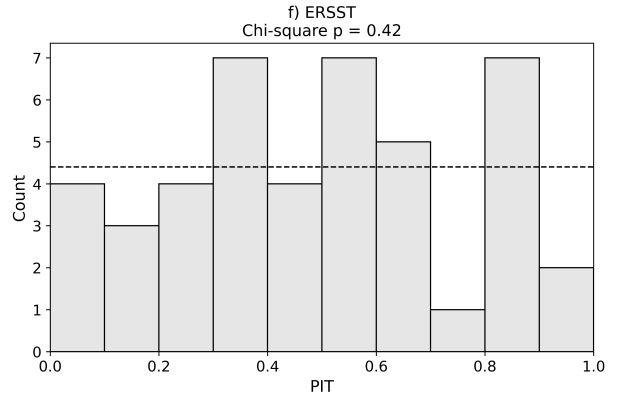
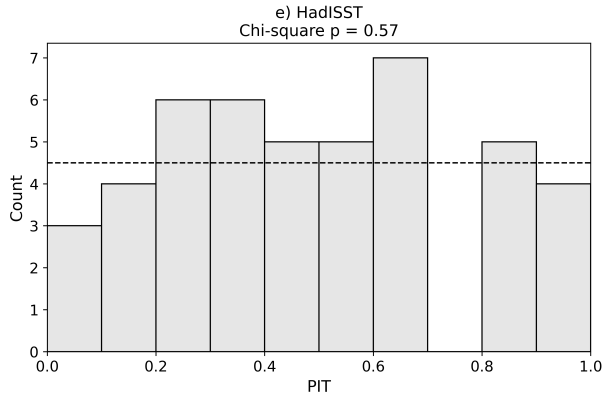
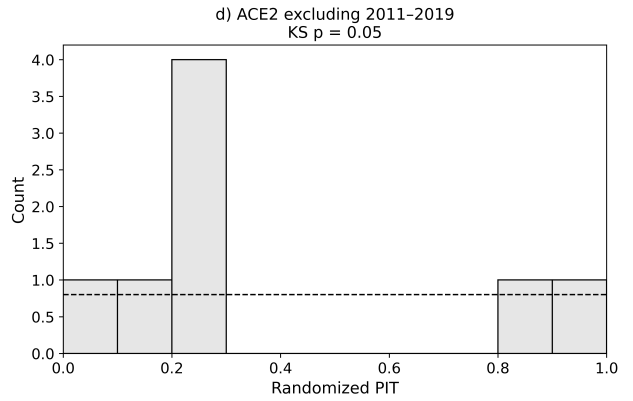
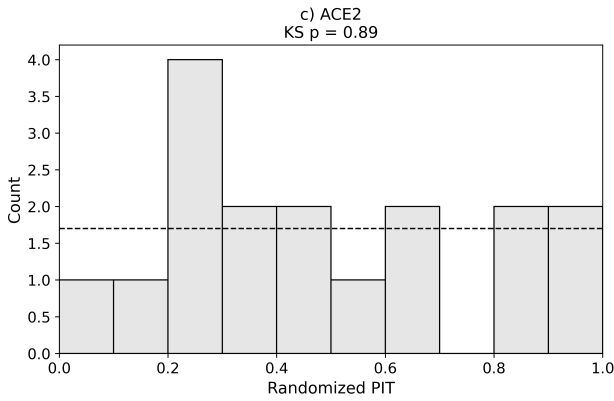
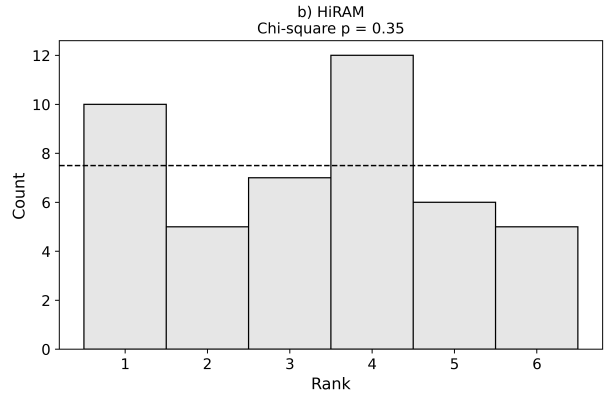
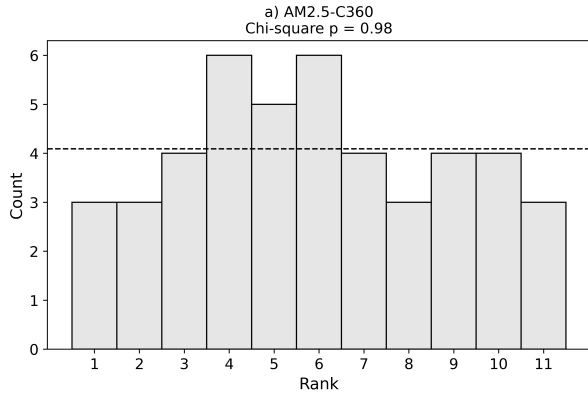
279 *a. Testing the first hypothesis family*

280 In this section, we present the results to evaluate the first family of hypotheses as to why the
281 GFDL models poorly represented the hyperactive 2020 Atlantic TC season within their ensemble
282 ranges: whether there were fundamental deficiencies in the models or observational datasets.

283 1) HYPOTHESIS NUMBER ONE

284 This hypothesis poses that the dynamical atmosphere-only models used in this study to generate
285 the historical TC record in Figure 1 (AM2.5-C360 and HiRAM) may be limited in their ability to
286 represent TC genesis, and thus, are unreliable to assess Atlantic TC interannual variability.

287 This explanation appears unlikely. In addition to the demonstrated skill of AM2.5-C360 and
288 HiRAM in reproducing the long-term historical variability of Atlantic TC activity (Levin et al.
289 2026), the influence of SSTs on TC frequency and intensity (Chan et al. 2021), and the annual TC
290 cycle (Yang et al. 2021), it is notable that the anomalous 2020 season also lies at or beyond the
291 ensemble range in other independent modeling systems. Seasonal forecasts from physics-based
292 coupled models such as GFDL SPEAR and FLOR (Murakami et al. 2025), as well as from DL-
293 based models such as NeuralGCM (Zhang et al. 2025a), similarly place 2020 at the edge or outside
294 their ensemble distributions. The fact that multiple structurally distinct modeling frameworks
295 fail to encompass the observed 2020 hyperactivity within their ensemble spread suggests that a
296 model-specific deficiency is unlikely to be the sole explanation.



297 FIG. 2. Reliability diagnostics for the modeling systems used in this study against the observed seasonal
298 Atlantic TC counts computed using the methodology of Landsea et al. (2010a). Panels (a) and (b) show rank
299 histograms for AM2.5-C360 (10-member ensemble) and HiRAM (5-member ensemble), respectively, forced
300 with observed SSTs over 1980–2024 ($n=45$ years). For these relatively small discrete ensembles, calibration
301 is assessed using rank histograms, with the p -values from Chi-square goodness-of-fit tests against the uniform
302 distribution shown in the subplot titles. Panels (c) and (d) show randomized probability integral transform
303 (PIT) histograms for the 1,000-member ACE2 DL ensemble for 1982 and 2005–2020 ($n=17$) and for the same
304 years excluding the model training period (2011–2019; $n=8$), respectively. Because ACE2 produces a very
305 large ensemble and the outcome variable is discrete count data, randomized PIT values are used instead of
306 rank histograms. Given the limited sample sizes for ACE2, calibration is assessed using Kolmogorov–Smirnov
307 (KS) tests against the uniform distribution. Panels (e)–(g) show randomized PIT histograms for the statistical-
308 dynamical model of Vecchi et al. (2011) forced with HadISST (1980–2024, $n=45$), ERSST (1980–2023, $n=44$),
309 and OISST (1981–2022, $n=42$), respectively. Because these models produce full probabilistic distributions, PIT
310 diagnostics are used; Chi-square goodness-of-fit tests against the uniform distribution are shown in the subplot
311 titles. In panels (a)–(g), the dashed horizontal black line denotes the expected histogram frequency under perfect
312 calibration, for which the ranks or verification quantiles of the observed TC counts are uniformly distributed.
313 Higher p -values indicate weaker evidence against uniformity. Panel (h) shows a Q–Q calibration diagnostic
314 for all models in panels (a)–(g). For each verification year, the percentile at which the observed TC count
315 falls within the corresponding forecast distribution is computed (using randomized PIT for discrete ensemble
316 forecasts), yielding one verification quantile per year. The empirical distribution of these verification quantiles
317 is then compared against the standard uniform distribution. A well-calibrated forecast system should produce
318 verification quantiles that lie close to the 1:1 dashed black line.

319 Additionally, we assess the reliability of the historical AM2.5-C360 and HiRAM ensembles
320 against observed seasonal Atlantic TC counts using rank histograms (Figures 2a,b; Hamill (2001);
321 Gneiting et al. (2007)). While both histograms show a slight visual concentration near the center,
322 which could qualitatively suggest mild underdispersion, chi-square goodness-of-fit tests against
323 the uniform distribution yield high p -values ($p = 0.98$ for AM2.5-C360; $p = 0.35$ for HiRAM),
324 providing no evidence to reject the null hypothesis of uniformity. This indicates that, over the
325 historical record, the ensemble spread is broadly consistent with the observed distribution of
326 seasonal TC counts. The Q–Q calibration diagnostic (Figure 2h) provides additional support, with

327 both models lying close to the 1:1 line, indicating that the observed TC counts fall at approximately
328 uniformly distributed quantiles within the forecast distributions. Together, these diagnostics suggest
329 that the GFDL ensembles are reasonably well calibrated over the historical period, despite failing
330 to encompass the extreme 2020 season within their ensemble ranges.

331 2) HYPOTHESIS NUMBERS TWO AND THREE

332 The second hypothesis regarding the reliability of the TC observations appears unlikely, as
333 the post-1980 Atlantic TC record is considered highly reliable given modern satellite, aircraft
334 reconnaissance, and remote sensing observations, with confidence further enhanced in recent years
335 by newer geostationary satellites. Additionally, this study uses the observed TC record computed
336 using the methods of Landsea et al. (2010b), which excludes short-duration storms that may have
337 been missed in earlier decades due to less advanced observing capabilities, thereby improving
338 temporal homogeneity in the historical record. Moreover, the anomalous 2020 activity extended to
339 hurricanes and major hurricanes, which are even less susceptible to detection ambiguity, making
340 observational error an unlikely explanation for the model–observation discrepancy.

341 The third hypothesis suggests there are errors in the SST dataset used to force the atmosphere-
342 only models. To test this, we apply the statistical–dynamical model of Vecchi et al. (2011), which
343 predicts seasonal Atlantic TC activity from SST, using three independent SST products: OISST,
344 HadISST, and ERSST. The PIT histograms for the resulting seasonal TC distributions (Fig. 2) show
345 no statistical evidence of deviation from uniformity based on chi-square tests, indicating that these
346 statistical-dynamical reconstructions are well calibrated. The corresponding time series of the
347 statistical-dynamical models (Fig. S1) capture substantial interannual variability, with correlations
348 of $r = 0.66$ – 0.70 against observations. Despite known differences among SST datasets in their
349 long-term trends (Menemenlis et al. 2025), the predicted 2020 TC distributions are nearly identical,
350 with ensemble means ranging from 10.9 to 11.8 storms. This provides no evidence that errors in
351 HadISST, which forces the dynamical simulations, explain the discrepancy.

352 3) HYPOTHESIS NUMBER FOUR

353 Our fourth hypothesis suggests that the failure of the observed 2020 TC count to fall within
354 the AM2.5-C360 and HiRAM ensemble ranges reflects limitations of SST-forced models in rep-
355 resenting the impacts of abrupt aerosol perturbations. To test this, we use two approaches. First,

356 we perform a five-member AM2.5-C360 ensemble simulation over 1980–2024 using observed
357 reanalysis-based aerosol concentrations, rather than the prescribed CMIP5 aerosol forcing used in
358 the historical simulations shown in Fig. 1. Even when forced with observed aerosol concentrations,
359 the model simulates 2020 as only a moderately active season (Fig. S2). The absence of a change
360 in simulated Atlantic TC activity when including observed aerosols suggests that any radiative or
361 dynamical effects associated with the 2020 aerosol reduction were either small or already reflected
362 indirectly in the observed SST boundary conditions.

363 Additionally, we conduct idealized sensitivity experiments in which sulfate and black carbon
364 aerosols are substantially reduced to emulate an exaggerated analogue of the 2020 aerosol decline.
365 The resulting annual Atlantic TC count distributions are similar to the control, with all experi-
366 ments producing approximately 12 TCs per year (Fig. S3). Kolmogorov–Smirnov tests show no
367 statistically significant differences from the control in any experiment ($p = 0.28–0.99$), suggest-
368 ing that, within this modeling framework, the direct radiative effects of prescribed low-frequency
369 (>monthly) aerosol variations do not exert a robust influence on Atlantic TC activity. This does not
370 preclude aerosol effects operating through SST-mediated pathways or synoptic-scale interactions,
371 which are not represented in these experiments.

372 *b. Testing the second hypothesis family*

373 In this section, we evaluate the second hypothesis family, which considers the possibility that the
374 hyperactive 2020 season was a highly unlikely but dynamically plausible outcome that could not
375 have been captured by the small ensemble sizes of the AM2.5-C360 and HiRAM models.

376 1) ANALYSIS OF ENSEMBLE SIZE

377 We first examine the historical ensembles from the dynamical AM2.5-C360 and HiRAM models
378 shown in Fig. 1. Rather than treating 2020 in isolation, we quantify how often the observed storm
379 count falls outside the modeled ensemble spread across the entire 45-year record.

380 Because the dynamical ensembles appear well calibrated (Fig. 2), the observed record can be
381 interpreted as an additional equally likely realization drawn from the same distribution as the
382 model ensemble members. Under this assumption, AM2.5-C360 (HiRAM) contains eleven (six)
383 realizations per year, comprising ten (five) model members plus the observational record. The

384 probability that the observation lies outside the modeled ensemble range, is either the minimum or
385 maximum of the combined set, is therefore

$$P(\text{outside ensemble range}) = \frac{2}{n}.$$

386 Therefore, the resulting annual exceedance probabilities are $2/11 \approx 0.18$ for AM2.5-C360 and $2/6 =$
387 0.33 for HiRAM, corresponding to expected exceedance counts over 45 years of approximately
388 eight and fifteen, respectively.

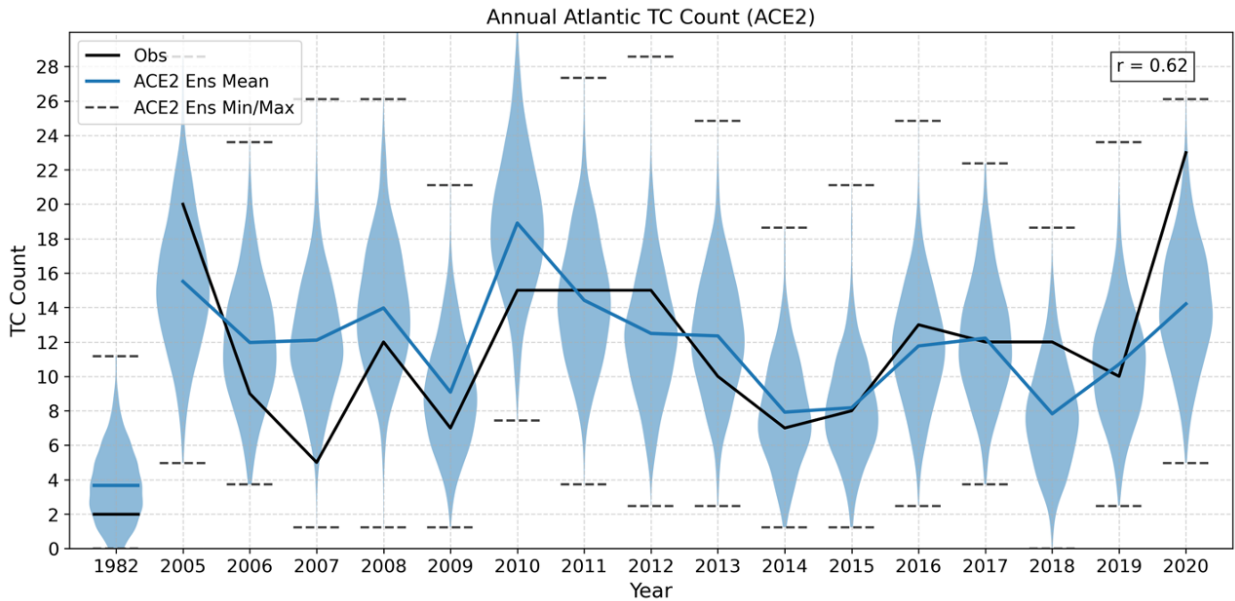
389 In Fig. 1, the observed record lies outside the AM2.5-C360 ensemble spread six times for total
390 TCs, three times for hurricanes, and five times for major hurricanes. These outcomes are well
391 within sampling variability and are therefore consistent with a calibrated system. Similarly, for
392 HiRAM, the observed record lies outside the ensemble spread fifteen times for TCs and twelve
393 times for hurricanes, both close or equal to the expectation of fifteen exceedances.

394 Thus, even in a well calibrated system with only five to ten ensemble members, it is expected
395 that the observation will fall outside the ensemble range multiple times over a 45-year period.
396 Although 2020 stands out visually because it lies outside the ensemble spread simultaneously for
397 TCs, hurricanes, and major hurricanes, similar exceedances occur elsewhere in the historical record
398 (e.g., 2013 and 1996). Therefore, this analysis supports the conclusion that the small ensemble
399 sizes of AM2.5-C360 and HiRAM could have been inadequate to capture the extremely hyperactive
400 2020 season.

401 At the same time, the limited ensemble sizes of AM2.5-C360 and HiRAM constrain our ability
402 to precisely estimate tail probabilities of extreme events like the 2020 season. This limitation
403 motivates the use of the 1,000 member ACE2 DL-based model ensemble and the statistical-
404 dynamical model from Vecchi et al. (2011) to generate a full distribution of seasonal TC counts
405 and more robustly characterize the distribution's tail behavior. The ACE2 1,000 member ensemble
406 spans 1982 and 2005-2020, encompassing both active and inactive seasons that are captured within
407 the ensemble spread of dynamical models, including 1982, 2005, 2009, and 2010 (Fig. 1). Our
408 comparison across years in the recent record allows us to assess whether the behavior of the
409 DL-based model diverges substantially from that of physics-based atmosphere-only models.

410 The observed annual TC counts and the distribution of simulated TC counts from the 1,000-
411 member historical ACE2 ensemble are shown in Fig. 3. ACE2 exhibits a moderate positive

412 correlation between the observed TC counts and the ensemble-mean simulated counts ($r = 0.62$),
413 indicating skill in distinguishing between relatively active and inactive seasons due to SST and
414 carbon dioxide external forcing alone. The ensemble mean peaks during active seasons such as
415 2005 and 2010, and shows local minima during inactive seasons such as 2009 and 2014, consistent
416 with the observed variability.



417 FIG. 3. Annual Atlantic TC counts for 1982 and 2005–2020. The adjusted observational record computed
418 using the methods from Landsea et al. (2010a) is shown in black. Blue violin plots depict the distribution of
419 simulated TC counts from the 1,000-member ACE2 ensemble forced with observed SSTs, with the ensemble
420 mean indicated by the solid blue line, and ensemble minimum and maximum indicated by the dashed dark
421 gray lines. The correlation between the ensemble-mean TC frequency and the observed counts is shown in the
422 upper-right corner.

423 Beyond reproducing the observed mean-state variability, the ensemble exhibits substantial in-
424 terannual spread. For each of the seventeen seasons shown, the simulated range of TC counts
425 (the difference between the ensemble minimum and maximum) spans roughly ten to 26 storms.
426 For instance, in 2005 the ensemble indicates that between approximately five and 28 TCs were
427 dynamically plausible, whereas in the relatively inactive 2014 season the simulated range extends
428 from one to nineteen TCs. Despite this considerable variability, the observed TC count falls within

429 the ensemble distribution for most seasons in the historical record. Furthermore, the randomized
430 PIT histogram for the full ACE2 record (Fig. 2c) is not statistically distinguishable from uniform
431 according to the K–S test, and the corresponding Q–Q plot lies close to the 1:1 line (Fig. 2h),
432 indicating that the ensemble dispersion provides a realistic representation of seasonal uncertainty.
433 In other words, the verifying observational quantile is approximately uniformly distributed across
434 years, suggesting that the ensemble dispersion realistically represents seasonal uncertainty.

435 ACE2’s performance is not uniform across all years, however. Visual inspection of Fig. 3
436 suggests that discrepancies between the ensemble mean and observations are largest prior to 2010
437 and in 2020, while agreement improves during 2010–2019. Additionally, the randomized PIT
438 histogram for the subset of out-of-sample years (Fig. 2d,h) appears U-shaped, with a marginal
439 K–S test p-value of 0.05 and the largest departures from the 1:1 line in the Q–Q plot, suggestive of
440 reduced calibration. However, this subset contains only eight years, limiting statistical confidence.
441 These years with the weaker reliability span the out-of-sample validation period, since the model’s
442 training includes 2011–2019 (Watt-Meyer et al. 2025). These results suggest stronger performance
443 during in-sample years than in true out-of-sample evaluation.

444 There are several years in which the observed TC count falls near the extreme tails of the ACE2
445 ensemble distribution. One such case is 2007, an anomalously inactive season with only five
446 observed TCs. The 1,000-member ACE2 ensemble, forced with observed 2007 SSTs, simulated a
447 plausible range of approximately one to 26 storms, with an ensemble mean of twelve. Only fifteen
448 of the 1,000 members produced five or fewer TCs, indicating that the realized season corresponds
449 to a 1.5% lower-tail event within the modeled distribution. Although the ACE2 ensemble mean
450 represented 2020 as a relatively active season, with the ensemble mean being third highest across
451 the years in 3, below only 2005 and 2010, the observed 2020 hyperactivity verified at the extreme
452 tail of the ACE2 ensemble distribution. The ensemble simulates a mean of approximately fourteen
453 TCs (Table 1), compared to the 23 storms observed. Only 37 ensemble members produce at
454 least 20 storms (3.7%), and just five members simulate 23 or more storms (0.5%). Thus, under
455 the observed 2020 SST forcing, ACE2 characterizes the realized season as a rare but physically
456 plausible outcome, verifying at the 99.5th percentile of the modeled distribution.

457 Table 1 compares the estimated likelihood of the 2020 hyperactive season from the 1,000-
458 member ACE2 ensemble distribution with that derived from the statistical–dynamical Poisson

466 TABLE 1. Modeled distribution of 2020 seasonal TC counts using a statistical-dynamical model with three
 467 SST products as inputs and the DL-based ACE2 1,000-member ensemble model. For each modelled TC count
 468 distribution we show the distribution’s mean, quantile at which the observed 23 TC counts in 2020 verified
 469 against the modeled distribution, and the corresponding percent of such an event.

Model	Model Mean TC count	Verified Quantile	Event Percent
Statistical-Dynamical (HadISST)	11.0	0.999	0.1%
Statistical-Dynamical (ERSST)	11.8	0.997	0.3%
Statistical-Dynamical (OISST)	10.9	0.999	0.1%
ACE2 1,000 member ensemble	14.1	0.995	0.5%

459 model described in Eq. 1, forced separately with HadISST, ERSST, and OISST SST datasets. The
 460 statistical–dynamical model is shown to be well calibrated in Fig. 2, with PIT histograms that are
 461 statistically indistinguishable from uniform and Q–Q plots that closely follow the 1:1 line. Across
 462 all modeling frameworks, the results consistently indicate that the observed 2020 SST boundary
 463 conditions did not strongly favor such an exceptional season and provided limited predictability
 464 for a hyperactive outcome, with estimated probabilities of only 0.1–0.5% under the observed SST
 465 forcing.

470 Although a 0.5% event is highly unlikely in any single year, its occurrence over a multidecadal
 471 record is less implausible. Assuming independence between seasons, the probability of observing
 472 at least one 0.5% event over a 45-year period (e.g., 1980–2024) is

$$P(\text{At least one 0.5\% event}) = 1 - (0.995)^{45} \approx 0.20. \quad (8)$$

473 Thus, an event with a 0.5% annual probability has roughly a 20% chance of occurring at least once
 474 in a 45-year historical record.

475 To evaluate how ensemble size influences the ability to capture the full range of plausible
 476 2020 Atlantic TC outcomes, we conduct a resampling experiment using the 1,000-member ACE2
 477 simulation of the 2020 season. From this ensemble, we randomly draw either ten or fifteen
 478 members without replacement and repeat this procedure 10,000 times. These sample sizes mimic
 479 the ensemble sizes typically used in physics-based seasonal forecasting systems, which often consist
 480 of only ten to fifteen members (e.g., Murakami et al. (2025)). For each resampled ensemble,

489 TABLE 2. The percentage of 10,000 resampled realizations of annual Atlantic TC counts from the ACE2 model
 490 that satisfy specific activity thresholds. Each realization is constructed by randomly selecting either 10 or 15
 491 ensemble members from the 1,000 ACE2 simulations of the 2020 season, without replacement, and repeating
 492 this sampling process 10,000 times.

Criteria across all samples (%)	10 ensembles	15 ensembles
≥ 1 ens with at least 20 TCs	32.99%	44.96%
≥ 2 ens with at least 20 TCs	5.86%	12.06%
≥ 1 ens with at least 23 TCs	8.03%	11.51%
≥ 2 ens with at least 23 TCs	0.24%	0.56%
mean ≥ 12 TCs ('active season')	97.96%	99.62%
mean < 12 TCs ('inactive season')	2.04%	0.38%
mean ≥ 17 TCs (' $+1\sigma$ season')	0.65%	0.16%
mean ≤ 7 TCs (' -1σ season')	0.00%	0.00%

481 we evaluate whether specific activity thresholds are met (Table 2). In particular, we assess the
 482 probability that at least one or at least two members simulate (i) a season with ≥ 20 TCs (comparable
 483 to other hyperactive years such as 2005) and (ii) a season with ≥ 23 TCs, matching the observed
 484 2020 count. We also classify seasons using the ensemble mean. An “active” season is defined as
 485 one with a mean ≥ 12 TCs, consistent with the 2005–2020 observed mean (approximately twelve
 486 storms). A ten-member (fifteen-member) ensemble contains at least one realization with ≥ 23 TCs
 487 in only approximately 8% (12%) of resampled cases, highlighting how unlikely it is for an outcome
 488 as extreme as the observed 2020 season to be represented in a small ensemble size.

493 2) LARGE-SCALE ENVIRONMENTAL CONDITIONS IN 2020

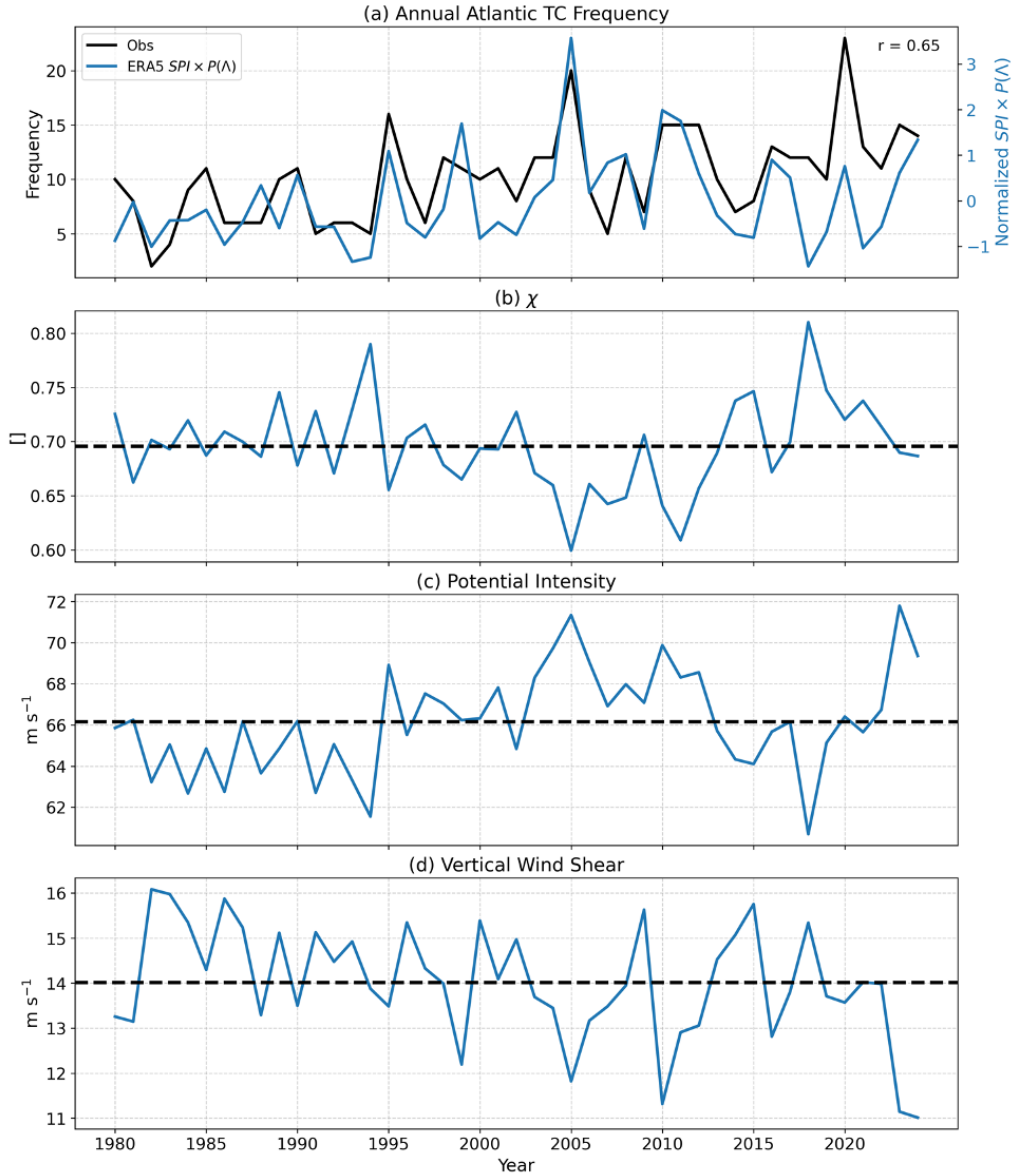
494 To evaluate if the hyperactive 2020 season was a highly unlikely but dynamically plausible real-
 495 ization of the climate system, we examine the large-scale environmental conditions that prevailed
 496 during that year. Internal atmospheric variability operating on subseasonal timescales could have
 497 amplified activity beyond what would be anticipated from seasonal-mean conditions alone.

498 To test this idea, we leverage ERA5 reanalysis data in combination with the theoretical TC proxy
 499 introduced by Hsieh et al. (2020), defined as $N_{TC} \approx \text{SPI} \times P(\Lambda)$, which approximates seasonal TC
 500 counts as a function of the large-scale environment. We compute the seasonally (June–November)
 501 and tropical Atlantic main develop region (MDR)-averaged TC proxy using monthly mean ERA5

502 inputs for the period 1980–2024 ($SPI \times P(\Lambda)$, Fig. 4a). This framework reproduces a substantial
503 portion of the observed interannual variability in Atlantic TC activity, yielding a correlation of
504 $r = 0.65$ between observed and both proxy-derived TC counts. Thus, seasonal-mean large-scale
505 environmental conditions explain much of the historical variability in Atlantic TC activity.

516 Although the TC proxy estimates the 2020 season to be moderately active across the entire
517 historical record of Fig. 4 and a local maximum in favorability in the years surrounding it, the
518 observed TC count substantially exceeds the proxy estimate. When the 2020 season is excluded,
519 the correlation between the observed counts and the seasonal $SPI \times P(\Lambda)$ (blue line) increases to
520 $r = 0.70$. Deviations between the proxy and observations are also evident in the surrounding years,
521 particularly from 2018 to 2021. These results indicate that, based on the combination of several
522 key seasonally averaged large-scale conditions alone, 2020 did not appear exceptionally favorable
523 for the record-breaking hyperactivity that occurred.

524 Additionally, we examine the individual components of $P(\Lambda)$, the proxy for the likelihood that a
525 TC will develop from a seed disturbance, averaged seasonally (June–November) over the Atlantic
526 MDR (Figs. 4b–d). In 2020, the thermodynamic variables, potential intensity and moist entropy
527 deficit, were close to their 1980–2024 climatological means, with moist entropy deficit slightly less
528 favorable than average. Similar behavior is found using ASO averages (Fig. S6), corresponding to
529 the climatological peak of the Atlantic TC season. Vertical wind shear in 2020 was more favorable
530 than average, but not exceptionally so relative to other active years. Several seasons, including
531 2005, 2010, 2011, 2012, 2016, 2023, and 2024, exhibited comparably or more favorable shear
532 conditions. While Bell et al. (2021) reported that ASO vertical wind shear in 2020, based on
533 NCEP/NCAR reanalysis, was the fifth lowest since 1980, our ERA5 analysis suggests the anomaly
534 was less extreme: in addition to the years identified by Bell et al. (2021), we find that 2006, 2012,
535 2023, and 2024 also had lower ASO vertical wind shear than 2020 (Fig. S6). Thus, although
536 the large-scale thermodynamic and dynamic environment in 2020 was generally conducive to TC
537 activity, the combined seasonal-mean values of vertical wind shear, potential intensity, and moist
538 entropy deficit do not support the expectation of a record-breaking season. This finding suggests
539 that internal atmospheric variability likely amplified TC activity beyond what would be anticipated
540 from mean conditions alone.



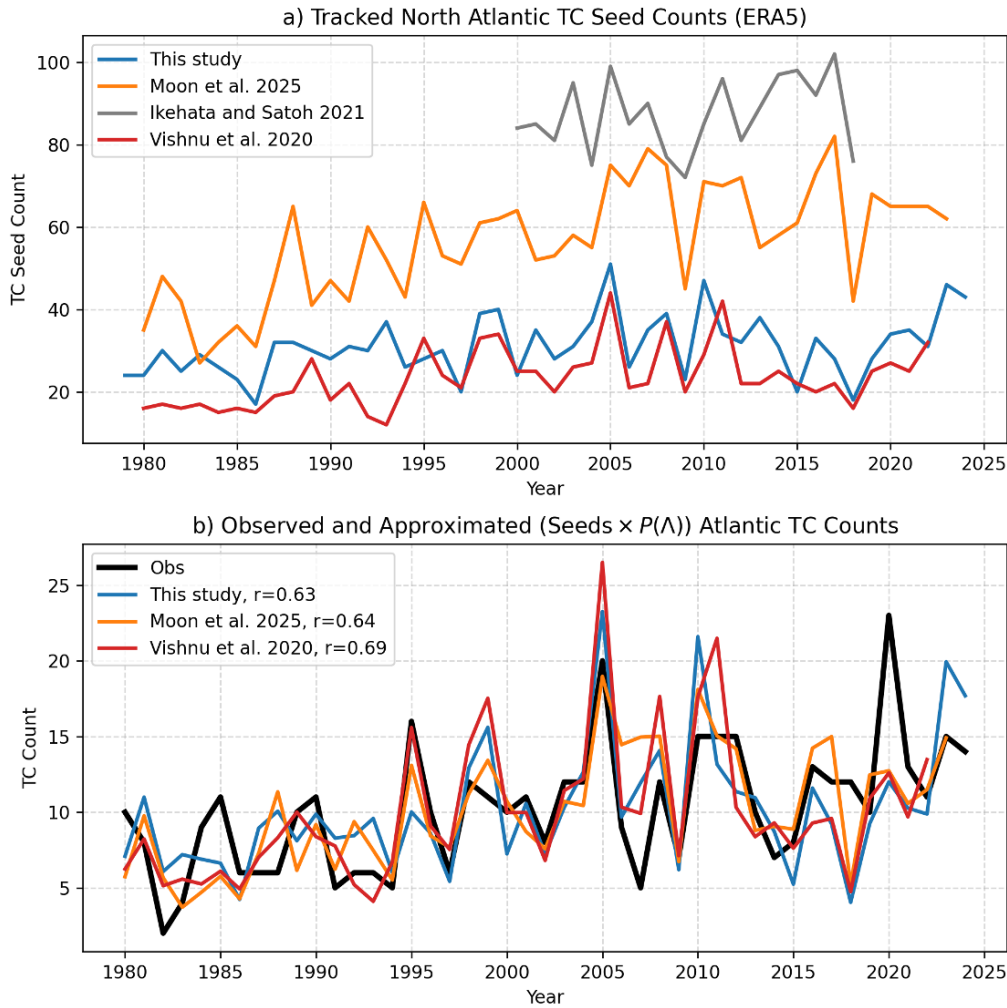
506 FIG. 4. a) Time series of observed annual Atlantic TC counts (black; left y-axis), approximated annual TC
 507 frequency ($\text{SPI} \times P(\Lambda)$); blue; right y-axis) derived from seasonally and TC main development region (MDR)-
 508 averaged monthly ERA5 data for 1980–2024. The parameterized TC frequencies are standardized using Z-score
 509 normalization, defined as $Z(x) = (x - \mu_x) / \sigma_x$, where x is the data point, μ_x is the mean over the full record,
 510 and σ_x is the standard deviation. The correlation coefficient between observed and approximated TC counts
 511 is shown in the upper-right corner of the panel. (b–d) Time series of the seasonal, MDR-averaged ERA5
 512 environmental variables that contribute to the ventilation index (Eq. 6), a key control on TC genesis: (b) moist
 513 entropy deficit (lower values are more favorable for TC development), (c) potential intensity (higher values are
 514 more favorable), and (d) vertical wind shear (lower values are more favorable). Dashed black horizontal lines
 515 indicate the 1980–2024 climatological mean of each variable.

541 In addition to examining monthly mean large-scale environmental conditions relevant for TC
542 development from a seed disturbance, we assess whether the abundance of seed disturbances
543 themselves tracked at six-hourly intervals contributed to the hyperactive 2020 TC season emerging
544 as a rare realization within the ensemble spread of multiple models. To do so, we examine explicitly
545 tracked Atlantic TC seeds from three independent ERA5-based datasets (Moon et al. (2025); Vishnu
546 et al. (2020); Ikehata and Satoh (2021)), in addition to our own tracking methodology. Each method
547 tracks storms using six-hourly June–November North Atlantic ERA5 data. The resulting seed counts
548 are shown in Fig. 5. These differing approaches yield different mean seed climatologies, yet they
549 share broad interannual patterns. Notably, none identifies 2020 as a local maximum in seed count
550 (Fig. 5). If the hyperactivity of 2020 were driven by anomalously many seeds, we would expect at
551 least one tracking methodology to identify an extreme seed year.

561 We then incorporate explicitly tracked seeds into the TC proxy, replacing SPI with observed
562 seed count to compute $N_{TC} \approx \text{Seeds} \times P(\Lambda)$, where $P(\Lambda)$ is derived from seasonal monthly ERA5
563 conditions. This modified proxy reproduces a substantial portion of observed interannual variability
564 (correlations $r = 0.63\text{--}0.69$) and successfully captures both inactive and active years such as 2005
565 and 2010. Yet across all three seed-tracking methodologies, the proxy converges on a moderately
566 active 2020 season of roughly twelve storms, below the observed 23. Thus, even when incorporating
567 explicitly tracked seed disturbances at six-hourly resolution, the observed large-scale atmospheric
568 conditions in 2020 do not appear sufficient to explain the record-breaking season, further suggesting
569 an important role for subseasonal internal atmospheric variability.

570 4. Discussion

571 This study examines the apparent unpredictability of the hyperactive 2020 Atlantic TC season
572 given the observed SST boundary conditions, during which activity across all intensity bins ex-
573 ceeded the ensemble spread of state-of-the-art GFDL atmosphere-only climate models forced with
574 observed SSTs (Fig. 1). We considered two broad classes of hypotheses to explain the discrep-
575 ancy between observations and model simulations. The first posits that model or observational
576 deficiencies prevented the ensembles from reproducing a hyperactive outcome comparable to that
577 observed. The second considers whether the 2020 season represented a dynamically plausible but



552 FIG. 5. (a) Tracked Atlantic tropical cyclone (TC) seed counts from the ERA5 dataset using the methods
 553 described in Section 2 (blue; 1979–2024), from Moon et al. (2025) (orange; 1980–2023), Ikehata and Satoh
 554 (2021) (gray; 2000–2018), and Vishnu et al. (2020) (red; 1980–2022). (b) Observed annual Atlantic TC counts
 555 (black) and approximated annual TC frequencies derived from seed counts ($\text{Seeds} \times P(\Lambda)$), where $P(\Lambda)$ represents
 556 the transition probability estimated from seasonally and MDR-averaged monthly ERA5 conditions. Seed counts
 557 are obtained following the respective tracking methods in Section 2 (blue; 1980–2024), Moon et al. (2025)
 558 (orange; 1980–2023), and Vishnu et al. (2020) (red; 1980–2022). Approximated TC counts are scaled by a
 559 multiplicative factor so that their mean matches the observed mean. Correlation coefficients between observed
 560 and approximated TC counts for each seed dataset are shown in the upper-left corner of panel (b).

578 low-probability realization under the observed SST forcing that could not have been captured by
579 the limited ensemble sizes of AM2.5-C360 and HiRAM.

580 After systematically evaluating four hypotheses within the first category using dynamical and
581 statistical–dynamical models, we find little evidence supporting model or observational failure.
582 Moreover, analysis of the large-scale environmental conditions using the theoretical framework
583 of Hsieh et al. (2020) indicates that 2020 did not exhibit exceptionally favorable environmental
584 conditions for a record-breaking season. Collectively, these results suggest that the SST-forced
585 atmosphere-only models did not fundamentally fail by simulating only moderate activity, because
586 the large-scale state itself was not strongly conducive to hyperactivity. Instead, we hypothesize that
587 an unusually favorable realization of internal atmospheric variability likely enhanced the 2020 TC
588 activity above expected levels from mean conditions alone.

589 Additional features of the 2020 season are consistent with enhanced internal variability with
590 limited predictability from observed SST boundary conditions. The season included a large number
591 of relatively weak storms compared with other hyperactive years such as 2005 and 2010 (Fig. S5).
592 These characteristics are consistent with a scenario in which weather-scale variability amplified
593 seasonal counts beyond what would be anticipated from large-scale environmental forcing alone.

594 To test whether 2020 could plausibly represent a rare but dynamically consistent outcome under
595 the observed SST forcing, we generated a 1,000-member ensemble for 2020 using the ACE2 DL
596 model forced with observed SSTs and applied the statistical–dynamical model of Vecchi et al.
597 (2011) to the same SST forcing to estimate the full distribution of possible seasonal outcomes.
598 These distributions indicate that a 23-TC season corresponds to an event with an approximate
599 probability of 0.1–0.5% under the observed 2020 SST forcing. While such an outcome is extremely
600 unlikely in any individual year, the probability of observing at least one 0.5% event over a 45-
601 year period is approximately 20%. Thus, the hyperactive 2020 season is best interpreted as a
602 rare but dynamically plausible realization rather than clear evidence of systematic model failure.
603 More broadly, it highlights both the limited predictability of extreme seasonal TC outcomes from
604 monthly SST boundary conditions alone and the difficulty of adequately sampling the observed
605 tail risk with the small (five–ten member) ensembles of AM2.50-C360 and HiRAM.

606 Seasonal forecasts from both dynamical systems (Murakami et al. 2025) and DL-based models
607 (Zhang et al. 2025a) similarly underestimated the realized 2020 storm frequency, with the observed

608 counts falling outside their ensemble ranges. Although the forecast systems in these studies per-
609 formed reasonably well in predicting seasonal accumulated cyclone energy (ACE), which reflects
610 the combined effects of storm frequency, intensity, and duration, our focus here is specifically on the
611 unprecedented storm counts across intensity categories rather than the characteristics of individual
612 storms or ACE, which fall outside the scope of this study. While seasonal outlooks generally antic-
613 ipated an active season, none suggested the degree of hyperactivity that ultimately occurred. For
614 example, NOAA's 2020 seasonal outlook assigned a 60% probability to an above-normal season,
615 indicating elevated risk but not a definitive expectation of extreme activity. Moreover, although
616 several forecasting systems predicted active conditions, the observed storm counts exceeded the
617 forecast ranges across multiple intensity bins (Table S1). Because these outlooks typically provide
618 qualitative count ranges rather than full probability distributions, they do not allow direct estimation
619 of the probability of an outcome as extreme as the observed 2020 season.

620 To address this limitation, we estimate the likelihood of the observed 2020 storm count by
621 applying May-initialized North American Multi-Model Ensemble (NMME) forecast SSTs to the
622 statistical–dynamical model of Vecchi et al. (2011), following the methodology of Villarini et al.
623 (2019). Consistent with other seasonal outlooks, these SST forecasts generally predicted a mod-
624 erately active season relative to their historical forecasts, largely reflecting favorable Atlantic-to-
625 tropical SST gradients. However, none indicated a record-breaking hyperactive outcome. This is
626 consistent with Bell et al. (2021), who similarly found that relative Atlantic SST anomalies in 2020
627 were locally elevated but not as extreme as in seasons such as 1999, 2005, or 2010. Under the
628 NMME SST forecasts, the observed 23-storm season corresponds to an event with an estimated
629 probability ranging from less than 0.05% to 4.5%, depending on the forecast model (Fig. S9).
630 The models assigning the highest probabilities (1–4.5%) exhibited warm biases in Atlantic MDR
631 SSTs relative to observations (Figs. S7 and S8), effectively creating boundary conditions more
632 favorable for TC activity than those that actually occurred. Thus, although these biased forecasts
633 made the observed hyperactivity appear more likely, even they assigned less than a 5% probability
634 to a season as extreme as 2020. In other words, even with biased, overly favorable SST conditions,
635 the models did not suggest that record-breaking hyperactivity was a likely outcome.

636 Finally, although Klotzbach et al. (2022) argued that aspects of the elevated late-season activity
637 in 2020 may have been anticipated using large-scale predictors, that analysis focused primarily

638 on intensity-related metrics such as accumulated cyclone energy and rapid intensification. To our
639 knowledge, the seasonal predictability of the realized storm frequency itself given observed SST
640 forcing has not previously been examined.

641 **5. Conclusions**

642 The evidence presented here supports the conclusion that the observed 2020 SST boundary
643 conditions did not confer strong predictability for the hyperactive outcome. Instead, the record-
644 breaking season was likely a rare realization of internal atmospheric variability superimposed on
645 a moderately favorable large-scale environmental state. Further, our results support the conclusion
646 that the hyperactive 2020 season can be understood as an unlikely outcomes that even a well-
647 calibrated system, like the modeling systems used in this study, will exhibit. Such rare but
648 dynamically plausible events may be fundamentally difficult to capture with small ensemble sizes,
649 even when the models themselves are performing appropriately. This interpretation of the 2020
650 season does not exclude the possibility that additional mechanisms not examined here contributed
651 to the observed activity. Future observations, modeling advances, or reanalysis improvements
652 could alter this assessment. New evidence could emerge to support hypotheses one through four,
653 for which we were not able to find compelling support in the results presented here. Moreover,
654 if similar large discrepancies between observations and model ensembles become more frequent,
655 reassessment of model calibration and structural assumptions would be warranted.

656 Our findings have several broader implications. First, the 2020 season occurs at the end of
657 the recent historical record and is anomalously hyperactive, which makes it disproportionately
658 influential in linear trend estimates (Fig. S10). Because trend calculations are especially sensitive
659 to extreme values near the endpoints of a time series, the inclusion of 2020 amplifies the estimated
660 1980–2024 increase in Atlantic TC, hurricane, and major hurricane activity by approximately six
661 storms per century for TCs, two storms per century for hurricanes, and one storm per century
662 for major hurricanes. Although removing 2020 does not reverse the positive sign of the TC or
663 major hurricane trends at the 95% confidence level, it substantially reduces their magnitudes. In
664 contrast, for hurricanes, excluding 2020 causes the 95% confidence interval to cross zero, implying
665 that the statistical evidence for a positive hurricane trend weakens considerably in its absence.
666 This behavior indicates that 2020 acts as a high-leverage endpoint in the regression and exerts

667 disproportionate influence on the inferred trends. This sensitivity is important because projected
668 changes in TC frequency under anthropogenic forcing remain an open and actively debated research
669 question (e.g., Knutson et al. 2020; Sobel et al. 2021). If the exceptional 2020 season primarily
670 reflects internal variability, such as subseasonal fluctuations or weather-scale noise, then its strong
671 influence on multidecadal trends complicates the interpretation of those trends as indicators of
672 externally forced climate change. These results highlight the need for caution when drawing
673 conclusions about long-term TC trends from relatively short observational records, particularly
674 when extreme endpoint seasons exert substantial leverage on estimated changes. Our findings
675 therefore stress the importance of employing statistical measures more resistant to the influence of
676 outliers, such as the median of pairwise slopes (Lazante 1996).

677 Next, this study is one of many to highlight the value of very large ensembles (e.g., Mahesh
678 et al. 2025a,b). We found that the five to ten member size from the dynamical models used in
679 this study is not sufficient to study a full distribution of potential outcomes, especially given that
680 the observed record is expected to exceed the ensemble spread across a multidecadal historical
681 record several times. Traditional modeling efforts leverage a relatively small ensemble size. For
682 example, the seasonal forecasting systems in Murakami et al. (2025) and Zhang et al. (2025a) use
683 ensembles of ten to fifteen members, and operational global meteorological agencies generally use
684 no more than a few dozen. In contrast, this study employs 1,000 ensemble members of the ACE2
685 DL model to assess the full distribution of outcomes for 2020 Atlantic TC season. By leveraging
686 the computational efficiency of DL-based weather and climate emulators, future forecast systems
687 should continue to incorporate much larger ensembles and explicitly consider tail risks, rather than
688 depending primarily on ensemble means, to prepare for low-probability but high-impact events.

689 In summary, the 2020 Atlantic TC season appears consistent with a rare realization within a
690 probabilistic framework rather than a systemic model failure. Even well-calibrated forecasting
691 systems will occasionally encounter extreme outcomes like the 2020 season. Recognizing and
692 quantifying this possibility is essential for interpreting past extremes and preparing for future risk.

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705 *Data availability statement.* Source code of the HiRAM model is available from [https://](https://www.gfdl.noaa.gov/HiRAM-quickstart)
706 www.gfdl.noaa.gov/HiRAM-quickstart, and the source code of the ACE2 model is available
707 from <https://github.com/ai2cm/ace>. The forcing data and initial conditions for the ACE2
708 model are available from [https://huggingface.co/allenai/ACE2-ERA5/tree/main/](https://huggingface.co/allenai/ACE2-ERA5/tree/main/initial_conditions)
709 [initial_conditions](https://huggingface.co/allenai/ACE2-ERA5/tree/main/initial_conditions). ERA5 data are available from [https://cds.climate.copernicus.](https://cds.climate.copernicus.eu/datasets/reanalysis-era5-pressure-levels-monthly-means?tab=overview)
710 [eu/datasets/reanalysis-era5-pressure-levels-monthly-means?tab=overview](https://cds.climate.copernicus.eu/datasets/reanalysis-era5-pressure-levels-monthly-means?tab=overview), and
711 MERRA2 data are available from [https://gmao.gsfc.nasa.gov/gmao-products/](https://gmao.gsfc.nasa.gov/gmao-products/merra-2/data-access_merra-2/)
712 [merra-2/data-access_merra-2/](https://gmao.gsfc.nasa.gov/gmao-products/merra-2/data-access_merra-2/). The OISST data are available from [https://www.ncei.](https://www.ncei.noaa.gov/products/optimum-interpolation-sst)
713 [noaa.gov/products/optimum-interpolation-sst](https://www.ncei.noaa.gov/products/optimum-interpolation-sst), the HadISST data are available from
714 <https://www.metoffice.gov.uk/hadobs/hadisst/>, and the ERSST data are available
715 from <https://www.ncei.noaa.gov/products/extended-reconstructed-sst>. The code
716 to compute the components of the tropical cyclone activity proxy are available from [https://](https://github.com/tlhsieh/tropical_cyclone_seeds)
717 github.com/tlhsieh/tropical_cyclone_seeds and [https://github.com/wy2136/](https://github.com/wy2136/wynton/tree/main/xtci/shared)
718 [wynton/tree/main/xtci/shared](https://github.com/wy2136/wynton/tree/main/xtci/shared). The TC seed dataset of Ikehata and Satoh (2021) is avail-
719 able at <https://doi.org/10.5281/zenodo.5136292>, the low pressure system dataset of
720 Vishnu et al. (2020) is available at [https://portal.nersc.gov/cfs/m3310/VishnuEtAl_](https://portal.nersc.gov/cfs/m3310/VishnuEtAl_TrackDataset/)
721 [TrackDataset/](https://portal.nersc.gov/cfs/m3310/VishnuEtAl_TrackDataset/), and the seed dataset of Moon et al. (2025) is available at [https://doi.org/](https://doi.org/10.5281/zenodo.15227119)
722 [10.5281/zenodo.15227119](https://doi.org/10.5281/zenodo.15227119). The code to run one year’s 1,000-member ensemble and process the

723 output, and track TCs is available at https://github.com/emmalevin/ACE2_1000. TC track
724 data and monthly mean environmental conditions generated from the ACE2 1,000 member ensem-
725 ble are available at the Zenodo repository: <https://doi.org/10.5281/zenodo.19456833>.

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