

1 **Title:** Testing for the effects of pre-season temperature and winter-chilling on land-surface  
2 phenology of coniferous and broadleaved forests in Central Europe

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9 Landsat

10

11 **Abstract:** Phenology is an important indicator of climate change impacts on vegetated  
12 ecosystems. Changes in leaf unfolding dates in response to changing temperatures have been  
13 well documented from in-situ phenological measurements across Central Europe. However, it  
14 is unclear whether those processes can be scaled to the landscape scale, which is important to  
15 accurately represent phenology in (global) vegetation models. Moderate resolution remote  
16 sensing time series, which measure land surface phenology instead of species specific  
17 phenophases, can help answering this question. We here test for the effect of pre-season  
18 temperature and winter-chilling on the inter-annual variation in start of season derived from  
19 Landsat time series for a forest landscape in southern Germany. The landscape is comprised of  
20 broadleaved and coniferous forests and thus typical for montane forest landscapes in Central  
21 Europe. We find strong evidence for average pre-season mean daily temperature driving inter-  
22 annual variation in start of season, with a  $-3.74 \text{ d } ^\circ\text{C}^{-1}$  earlier start of season for broadleaved  
23 forests and a  $-2.68 \text{ d } ^\circ\text{C}^{-1}$  earlier start of season for coniferous forests over the time period 1985  
24 to 2015. This relationship, however, was modulated by the number of chilling days during  
25 winter, with a decreasing effect of pre-season temperature with decreasing number of chilling  
26 days. The inter-annual variation in start of season predicted from our model – i.e., calibrated  
27 solely from Landsat satellite time series – showed good agreement with in-situ observations of  
28 bud-break (Pearson's  $r = 0.79/\text{RMSE} = 4.88 \text{ d}$  for broadleaved forests and  $r = 0.62/\text{RMSE} =$   
29  $6.57 \text{ d}$  for coniferous forests). We conclude that in-situ based processes are also detectable at  
30 the landscape-scale and that considering winter-chilling is important for accurately predicting  
31 phenology, which should be recognized in (global) vegetation models.

32

## 33 **1. Introduction**

34 Phenology is a key property of ecosystems (Forrest & Miller-Rushing, 2010). Due to its  
35 sensitivity to climate, phenology is also a valuable indicator of climate change (Cleland,  
36 Chuine, Menzel, Mooney, & Schwartz, 2007). Climate change effects on vegetation  
37 phenology have been documented in numerous experimental and observational studies,  
38 finding shifts in the start of the growing season for many plant species as a result of climate  
39 change (Cleland et al., 2007; Menzel et al., 2006; Parmesan & Yohe, 2003). Those changing  
40 phenological patterns have profound impacts on ecosystem functions, such as carbon uptake  
41 (Keenan et al., 2014). In order to better understand the impacts of changing phenology on  
42 ecosystem function, it is fundamental to monitor, model and ultimately predict vegetation  
43 phenological dynamics at varying spatial and temporal scales (Pau et al., 2011; Tang et al.,  
44 2016).

45 A key data source for monitoring and modeling phenological dynamics are in-situ  
46 phenological measurements. Those measurements are generally taken at the level of  
47 individuals (e.g., single trees or plants), and the resulting information is used to calibrate  
48 species-specific phenological models that predict phenological phases (e.g., leaf unfolding)  
49 based on a set of aggregated meteorological variables (Basler, 2016; Chuine, 2000; Hufkens,  
50 Basler, Milliman, Melaas, & Richardson, 2018). However, such species-specific models are  
51 often difficult to regionalise to the landscape or ecosystem scale (Y. Fu, Zhang, Dong, &  
52 Yuan, 2014; Richardson et al., 2012; Tang et al., 2016), particularly if the exact species or  
53 community composition are unknown (Jeremy I. Fisher, Richardson, & Mustard, 2007; Liang  
54 & Schwartz, 2009). Further, calibrating species-specific models across large geographic  
55 extents is challenging, as in-situ phenological observations are lacking in many regions  
56 worldwide, or are difficult to compare due to varying measurement protocols and/or target  
57 species (Fitchett, Grab, & Thompson, 2015). An improved understand of phenological  
58 dynamics at the landscape and ecosystem scale is, however, required for further developing

59 the representation of phenology in global vegetation models (Richardson et al., 2012; Yang,  
60 Mustard, Tang, & Xu, 2012).

61 Remote sensing, which measures so-called land surface phenology and thus delivers  
62 an integrated view on phenology independent of individual species (Morisette et al., 2009),  
63 might overcome limitations of field-based phenological data. Classical approaches tracking  
64 phenology from remote sensing data made use of the synoptic view of the Advanced Very  
65 High Resolution Radiometer (AVHRR) sensor family (M. A. White et al., 2009). AVHRR,  
66 however, has a spatial resolution that is too coarse (1 km) to infer phenological dynamics in  
67 spatially heterogeneous landscapes. The MODerate Imaging Spectroradiometer (MODIS),  
68 which has a higher spatial resolution of 250 m, might improve the estimation of spatial  
69 drivers (Friedl et al., 2014), yet its limited temporal depth (starting in 2001) makes it difficult  
70 to infer temporal trends in phenological dynamics. Dense Landsat time series were recently  
71 suggested to overcome the limitations of MODIS and AVHRR in monitoring vegetation  
72 phenology (J. I. Fisher & Mustard, 2007; J. I. Fisher, Mustard, & Vadeboncoeur, 2006;  
73 Melaas, Friedl, & Zhu, 2013; Melaas, Sulla-Menashe, & Friedl, 2018; Nijland, Bolton,  
74 Coops, & Stenhouse, 2016; Senf, Pflugmacher, Heurich, & Krueger, 2017), since they have a  
75 spatial resolution of 30 m and cover a time-span of more than 30 years (from 1984 onwards).  
76 Landsat time series might thus be well suited to characterize vegetation phenology at the  
77 landscape scale, and thus gain additional insights supporting field-based phenological studies.

78 For Central Europe, a substantial shift in leaf unfolding was identified over the past  
79 four decades (Menzel et al., 2006). It was hypothesized that warmer spring temperatures  
80 caused by climate change are the main driver explaining this observation. Yet, more recent  
81 research highlights the importance of winter chilling for phenological dynamics (Y. H. Fu,  
82 Piao, et al., 2015; Y. H. Fu, Zhao, et al., 2015; Laube et al., 2014). In particular, the effect of  
83 warmer spring temperatures on leaf unfolding might be compensated by reduced winter-  
84 chilling with increasing winter temperatures. While the modulating effect of winter chilling is

85 increasingly well understood based on lab experiments and field studies, it is unknown  
86 whether it also reproduces at the landscape scale, that is, whether similar effects can be found  
87 for land surface phenology derived from satellite time series.

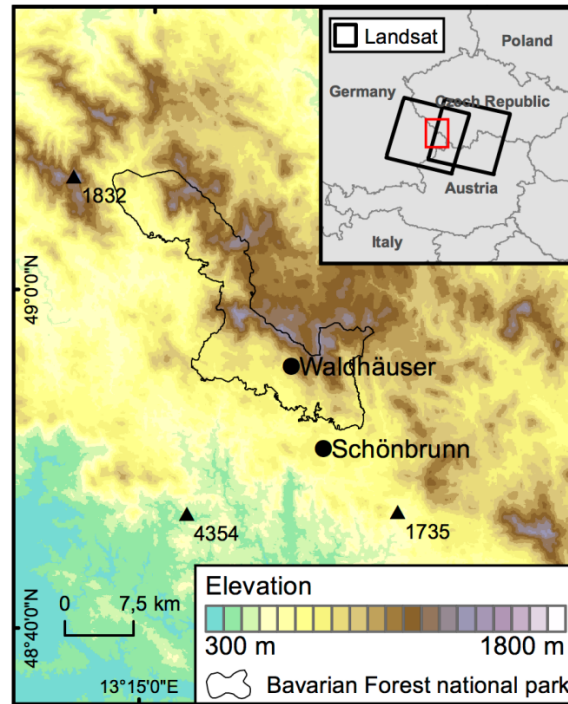
88         We here test for the interacting effects of pre-season temperature and winter-chilling  
89 on the inter-annual variation in start of season derived from Landsat time series. Our  
90 hypothesis was that while warmer pre-season temperatures would lead to earlier start of  
91 seasons, the effect will be modulated by the number of chilling days during winter. For testing  
92 this hypothesis, we integrate Landsat time series and meteorological observations into a  
93 combined phenological model. We calibrate the model for a landscape in southern Germany,  
94 comprised of broadleaved and coniferous forests, thus representing typical forest types in  
95 Central European (sub-) montane landscapes. We finally compare predictions from our model  
96 to in-situ phenological observations of leaf unfolding, determining whether pre-season  
97 temperature and winter-chilling effects identified from land surface phenology reproduce  
98 field-based observations of leaf-unfolding.

## 99 **2. Data and methods**

### 100 ***2.1 Study landscape***

101 The study landscape is located in the Bavarian Forest National park, with moderate  
102 topography ranging from approximately 300 to 1,800 meters in elevation (Figure 1). The  
103 ecosystem is characterized by mountain beech forests with *Fagus sylvatica* (European beech)  
104 being the leading species. Other broad-leaved tree species include *Acer pseudoplatanus*  
105 (Sycamore maple) or *Quercus robur* (English oak), but *F. sylvatica* is by far the most  
106 abundant. The higher elevation areas of the ecosystem, in turn, are covered by mountain  
107 spruce forests, consisting mainly of *Picea abies* (Norway spruce), with some *Abies alba*  
108 (European silver fir) intermixed.

109



110

111 **Figure 1:** Study landscape with national park boundaries, the three meteorological stations  
 112 (triangles), and the two in-situ phenological observation sites (points) from the International  
 113 Phenological Gardens of Europe network.

114 **2.2 Landsat data**

115 Landsat has been used to create wall-to-wall maps of phenological parameters (e.g., start of  
 116 season), which were subsequently used to summarize the inter-annual variability in each  
 117 parameter for a specific ecosystem or region (Melaas et al., 2013; Nijland et al., 2016).  
 118 However, creating wall-to-wall maps is computationally intensive and might be too  
 119 cumbersome if one is primarily interested in estimating the temporal dynamics of phenology  
 120 for a certain landscape or ecosystem. Alternatively, a sample-based approach can be used,  
 121 where only a sample of the Landsat time series is used to estimate the temporal dynamics in  
 122 phenological metrics (Melaas et al., 2018; Senf et al., 2017). We identified suitable sampling  
 123 locations for both broadleaved forests (i.e., mountain beech forests) and coniferous forests  
 124 (i.e., mountain spruce forests) using an existing land cover map generated by National Park  
 125 authorities from aerial imagery. To exclude young and/or disturbed forests we only sampled  
 126 mature stands, that is, stands with a minimum age of 60 years. The sample size was set to n =

127 250, but we also tested alternative sample sizes with no substantial differences found (data not  
128 shown).

129 We created cloud-, snow-, and shadow-free Enhanced Vegetation Index (EVI) time  
130 series for each sample location using all data from the Landsat archive, accessed via the  
131 Google Earth Engine (Gorelick et al., 2017). We preferred EVI over other vegetation indices  
132 as it has been shown to allow better estimation of key phenological dates from remote sensing  
133 data (Klosterman et al., 2014). Finally, as we were primarily interested in spring phenology,  
134 we separated the time series into spring/summer and autumn/winter observations. We  
135 followed a method suggested by Melaas et al. (2013), which excludes all observations before  
136 day of year 80 (winter observations) and then uses an iterative algorithm identifying the  
137 transition from summer to autumn. In essence, the algorithm fits a linear model to a running  
138 window of 21 observations and identifies the day of year with the first negative slope as the  
139 summer-autumn transition date. We dropped all autumn/winter observations from further  
140 analysis.

### 141 **2.3 Meteorological data**

142 We used meteorological data to create annual estimates of pre-season temperature and  
143 winter chilling for the study landscape. Pre-season temperature ( $T_{pre}$ ) was defined as the  
144 average mean daily temperature  $T_{mean}$  in the months April and May, being the approximate  
145 months of vegetation green-up in our study region (Senf et al., 2017). Winter chilling ( $D_{chil}$ )  
146 was defined as the number of days with  $0 < T_{mean} \leq 5 \text{ } ^\circ \text{C}$  as suggested by previous studies  
147 (Y. H. Fu, Zhao, et al., 2015). Both variables were averaged from three meteorological  
148 stations operated by the German Weather Service (station-ids: 1735, 1832, 4354; Figure 1;  
149 [ftp://ftp-cdc.dwd.de/pub/CDC/observations\\_germany/climate/daily/kl/historical/](ftp://ftp-cdc.dwd.de/pub/CDC/observations_germany/climate/daily/kl/historical/)). We  
150 restricted the time-series to our study period (1985 – 2015) and z-transformed the data to  
151 represent anomalies in units of standard deviation.

### 152 **2.4 Landsat phenological model**

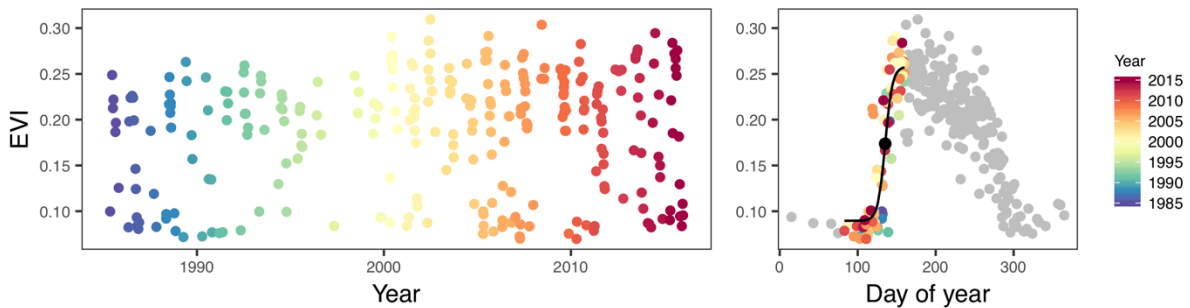
153 To estimate average phenological parameters for each pixel sampled, all available EVI  
 154 observations are sorted by observation day of year  $t$ , independent of the year of observation  
 155 (see Figure 2). Following previous research (Elmore, Guinn, Minsley, & Richardson, 2012;  
 156 K. White, Pontius, & Schaberg, 2014), we model this average phenological curve by a logistic  
 157 function with four parameters and normally distributed errors (see Figure 2 and Table 1):

$$EVI_i = \beta_{1[i]} + \frac{\beta_{2[i]}}{\left(1 + e^{-\beta_{3[i]}*(t-\beta_{4[i]})}\right)} + \varepsilon_i \quad (1)$$

$$\varepsilon_i \sim N(0, \sigma^2)$$

158 The parameters ( $\beta_i$ ) of the logistic function were modelled as a multivariate normal  
 159 distribution, thus allowing for correlation between them. We re-parameterized the  
 160 multivariate normal distribution to improve model convergence following recommendations  
 161 in Senf et al. (2017).

162



163

164 **Figure 2:** Example pixel time series showing the annual observations (left panel) and the  
 165 pooled time series (right panel). Shown here is the Enhanced Vegetation Index (EVI). The  
 166 pooled time series is truncated to spring and summer observations as described in Section 2.3,  
 167 and only those observations are colored. A phenological model as described in Section 2.4  
 168 was fit to the data. The black dot represents the start of season estimated from the model.

169

170 **Table 1:** Phenological parameters of the model.

Parameters $\beta_i$	Name	Description
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$\beta_{1[i]}$	Minimum	The minimum spectral value.
$\beta_{2[i]}$	Magnitude	The magnitude of spectral change during the year.
$\beta_{3[i]}$	Change rate	The change rate in the inflection point.
$\beta_{4[i]}$	Start of season	The day of year of the inflection point.

171

172 To account for temporal variation in the model parameters  $\beta_i$  among years  $j$ , we  
173 follow Senf et al. (2017) and replace  $\beta_i$  by a hierarchical formulation:  $\beta'_{ij} = \beta_i + \phi_j$ ,  
174 allowing them to simultaneously vary in space and time. Thus, the model estimates average  
175 phenological parameters for each pixel, and simultaneously estimates how each parameters  
176 varies – on average – over time. While the variability in the minimum, maximum and change  
177 rate is assumed to be purely random (i.e., expressed by a multivariate normal distribution  
178 centered on zero), we specifically model the mean temporal variation in the start of season by  
179 a linear combination of pre-season temperature and winter-chilling (and their interaction):  
180  $\phi_{4[j]} \sim N(\rho_1 * T_{pre} + \rho_2 * D_{chil} + \rho_3 * T_{pre} * D_{chil}, \sigma_\rho^2)$ . Hence, the model estimates the  
181 direction and strength of influence of pre-season temperature, winter chilling and their  
182 interaction on the inter-annual variability in the start of season.

### 183 **2.5 Parameter estimation**

184 Full Bayesian inference for each model parameter was made by sampling the joint posterior  
185 distribution using Markov Chain Monte Carlo (MCMC) methods implemented in the free  
186 software Stan (Carpenter et al., 2017). To sample joint posterior distributions, we needed to  
187 assign prior distributions to each parameter. We used weakly-informative, penalizing priors,  
188 which put most probability mass on values around zero. There must thus be high evidence in  
189 the data to estimate a posterior effect that is substantially different from zero (similar to lasso-  
190 regression). Further, we centered priors of the four parameters of the logistic function to their  
191 approximate location by averaging over the estimates of 10 non-linear least square fits with  
192 standard settings. While this step is not necessary to run the model, it helps making the model

193 converge faster and is thus similar to setting initial starting values in other (frequentist)  
194 iterative fitting routines.

195         The joint posterior distributions were finally sampled using four chains à 4,000  
196 iterations, of which 2,000 were later dropped as warm-up iterations. Convergence of chains  
197 was tested by comparing the between- and within-chain variance using the  $\hat{R}$  statistic  
198 described in Gelman, Carlin, Stern, and Rubin (2014). In essence, the  $\hat{R}$  statistic compares  
199 the variance within each chain to the variance between chains, with values approaching 1.00  
200 if chains converge to a similar solution. While modern MCMC samplers greatly improved in  
201 terms of speed, sampling might still take a long time with very large samples of pixels. To  
202 overcome this inherent limitation of spatial Bayesian statistics, we implemented an ensemble  
203 pixel sampling strategy that draws  $m$  samples of  $n$  pixels, and later average over the  $m$  joint  
204 posterior probability distributions of each model parameter. The associated cost of greater  
205 imprecision of parameter estimates due to the relatively small pixel sample size and the  
206 averaging of the posterior distributions was mitigated by a sufficient number of ensemble  
207 members (see Figure S1).

208         We finally tested the fit of the ensemble model by means of posterior predictive  
209 checking. Posterior predictive checking generates replicated data drawn randomly under the  
210 model, which can then be compared to the observed data (Gelman et al., 2014). If the model  
211 fits the data well, there should be no systematic differences between the replicated data and  
212 the observed vegetation index time series.

## 213 ***2.6 Comparison to in-situ measurements***

214         After sampling the joint posterior distribution, we compared the predicted inter-annual  
215 variability in the start of season against in-situ observations of leaf unfolding recorded for  
216 three *F. sylvatica* trees and three *P. abies* trees at two phenological gardens within and in  
217 close proximity to the National Park, respectively (see Figure 1). The phenological

218 observations were acquired from the International Phenological Gardens of Europe network  
219 (stations: Waldhäuser and Schönbrunn; Figure 1; <http://ipg.hu-berlin.de/>).

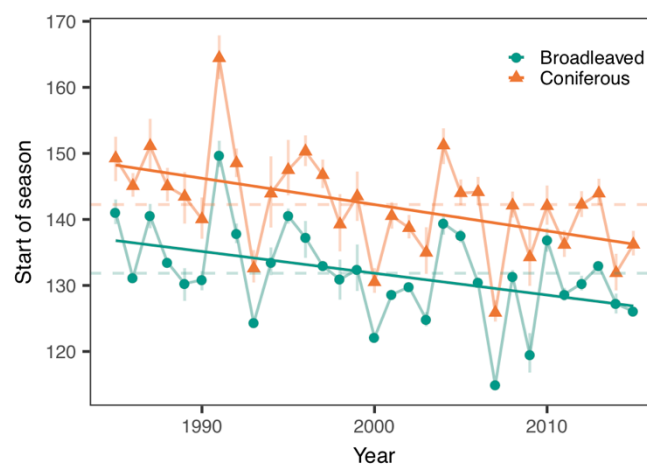
### 220 3. Results

#### 221 3.1 Model evaluation

222 All models in the ensemble showed good convergence with  $\hat{R}$  smaller 1.01 for 99% of the  
223 model parameters. Posterior predictive checks indicated no systematic deviation between  
224 simulated and raw values (Figure S2), and 96% of the raw values were within the 95%  
225 credible interval of the posterior simulations. Hence, there was little evidence to assume a  
226 substantial mismatch between model and data.

#### 227 3.2 Inter-annual variability in start of season

228 We identified substantial inter-annual variability in the start of season for broadleaved and  
229 coniferous forests (Figure 3), with an average change of  $-0.3 \text{ d yr}^{-1}$  for broadleaved and of  $-$   
230  $0.4 \text{ d yr}^{-1}$  for coniferous forests. This result indicates that the start of season was  
231 approximately 9 and 12 days earlier in 2015, compared to 1985, for broadleaved and  
232 coniferous forests, respectively. There were also some very late years (e.g., 1991) and some  
233 very early years (e.g., 2007 and 2009) identifiable for both broadleaved and coniferous  
234 forests, with overall very similar inter-annual variability for both forest types.

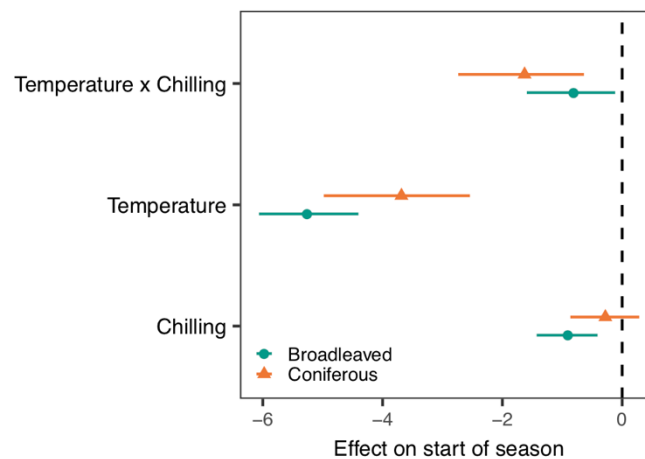


235  
236 **Figure 3:** Inter-annual variability in start of season estimated from Landsat time series for  
237 two forest types. Dots represent the median of the posterior distribution and error-bars extent

238 from the 2.5% to the 97.5% quantile of the posterior (95% credible interval). The dashed  
239 horizontal lines indicate the long-term average, whereas the solid lines indicate the trend line  
240 (OLS-estimate of the posterior medians).

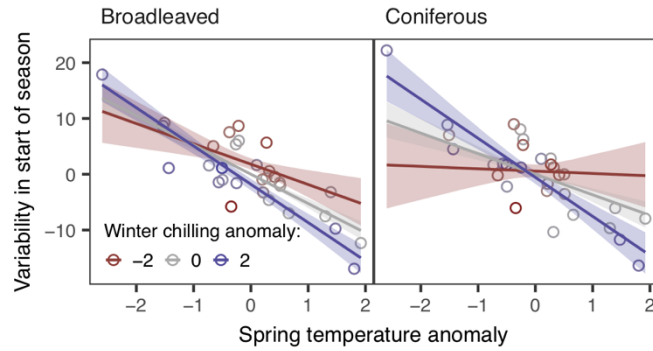
### 241 3.3 Driver analysis

242 The driver analysis revealed that the inter-annual variability in the start of season is highly  
243 sensitive to pre-season temperature (Figure 4), with a  $-3.7 \text{ d } ^\circ\text{C}^{-1}$  earlier start of season for  
244 broadleaved and a  $-2.7 \text{ d } ^\circ\text{C}^{-1}$  earlier start of season for coniferous forests, respectively. The  
245 effect of pre-season temperature on inter-annual variability in start of season was, however,  
246 modulated by the total number of chilling days in the preceding winter (Figure 4), with a  
247 decreasing effect of pre-season temperature on inter-annual variability in the start of season  
248 with a decreasing number of chilling days (Figure 5). Hence, the advance in spring phenology  
249 following a warmer than average spring were less pronounced when the preceding winter had  
250 a below-average number of chilling days.



251  
252 **Figure 4:** Effects of pre-season temperature and winter-chilling on start of season. All  
253 variables are z-transformed and effects are thus on the standard deviation scale. Dots  
254 represent the median of the posterior distribution and error-bars extent from the 2.5% to the  
255 97.5% quantile of the posterior (95% credible interval).

256

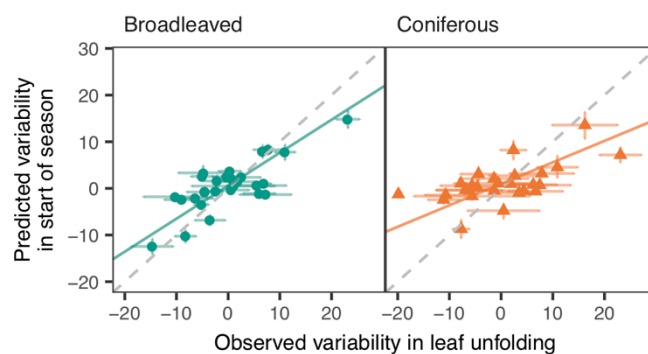


257

258 **Figure 5:** Variability in start of season with changing pre-season temperature and changing  
 259 number of winter-chilling days as estimated from the model.

260 **3.4 Comparison to in-situ measurements**

261 Comparing predictions from the model (i.e., based on the regression relationship shown in  
 262 Figure 5) to in-situ observations of bud-break (Figure 6), we found strong agreement for  
 263 broadleaved forests (Pearson's  $r = 0.79$  [0.74 – 0.80] and RMSE = 4.88 [4.60 – 5.21] days)  
 264 and moderate agreement for coniferous forests (Pearson's  $r = 0.69$  [0.63 – 0.75] and RMSE =  
 265 6.57 [6.00 – 7.13] days). Hence, our model – based only on pre-season temperature and  
 266 winter chilling and calibrated solely from Landsat time series – was able to predict the general  
 267 inter-annual variability in spring phenology as recorded in in-situ measurements.



268

269 **Figure 6:** Comparing the inter-annual variability in start of season predicted from our model  
 270 to in-situ observations of leaf unfolding derived from two phenological gardens in and around  
 271 the national park (see Figure 1). Dots represent the median of the posterior (prediction;  $y$ -axis)  
 272 and the arithmetic mean (leaf unfolding;  $x$ -axis). Uncertainty in the prediction is expressed by

273 the 95% credible interval, whereas error-bars for the leaf unfolding represent 95% confidence  
274 intervals.

#### 275 **4. Discussion**

276 We here used dense Landsat time series to test for the interacting effects of pre-season  
277 temperature and winter-chilling on the inter-annual variation in start of season. We did so by  
278 explicitly integrating Landsat time series with meteorological observations; that is, instead of  
279 creating phenological products from remote sensing data that are subsequently used in  
280 analysis, we here calibrated phenological models directly from the millions of spectral  
281 observations available through the Landsat archive. The advantage is a reduced number of,  
282 and redundancy in, processing steps; as well as a better propagation of uncertainty, which is  
283 otherwise largely neglected in remotely sensed phenological studies.

284 Our model – calibrated solely from Landsat observations – was capable of predicting  
285 inter-annual variability in leaf-unfolding observed in-situ, despite the relatively simple model  
286 structure. This finding reinforces previous studies showing that more complex models do not  
287 necessarily lead to better predictions (Basler, 2016; Yang et al., 2012). In fact, RMSE values  
288 obtained by our model are in a similar range or even lower than comparable or more complex  
289 phenological models (Basler, 2016; Hufkens et al., 2018). Further, as our model is calibrated  
290 on land surface phenology (Morisette et al., 2009), it might better represent the landscape-  
291 and ecosystem-scale drivers of phenology than models based on species-specific in-situ  
292 observations (Liang & Schwartz, 2009). This is mainly due to a more aggregated view going  
293 beyond individual species and species differences (Jeremy I. Fisher et al., 2007). This  
294 difference might be particularly important for calibrating phenological models used in global  
295 vegetation models (Richardson et al., 2012; Yang et al., 2012).

296 The effects of pre-season temperature and winter-chilling on inter-annual variability in  
297 start of season identified in this study largely corroborate the current literature based on in-  
298 situ observations and lab experiments (Y. H. Fu, Piao, et al., 2015; Y. H. Fu, Zhao, et al.,

299 2015; Laube et al., 2014). Thus, recent insights on the importance of winter-chilling for  
300 phenological dynamics also reproduce at the landscape scale, highlighting the need for  
301 including those processes into (global) vegetation models. From an ecosystem-scale  
302 perspective, this might indicate that warming-related reductions in chilling days might slow  
303 down recent observed changes in spring phenology, with cascading effects on other  
304 ecosystem processes thus as carbon uptake. Yet, the exact processes are still elusive, and  
305 additional drivers such as precipitation or the previous year's phenology are discussed in the  
306 current literature (Y. H. Fu, Piao, et al., 2015; Y. S. Fu et al., 2014). Our model – in  
307 conjunction with global long-term satellite records from the Landsat archive – can be  
308 extended to include those drivers and thus offers great potential for understanding  
309 phenological responses under climate change in more detail.

## 310 **5. Conclusion**

311 We here used dense Landsat time series to show the interacting effects of pre-season  
312 temperature and winter-chilling on the inter-annual variation in start of season. We thus found  
313 additional evidence for the combined effect of temperature forcing and winter chilling in  
314 predicting the start of season, but also showing that the relationship holds true at the  
315 landscape scale. Hence, our research suggests that ecosystem models incorporating phenology  
316 should recognize the importance of winter chilling, especially under climate change scenarios.  
317 Remote sensing – and in particular Landsat remote sensing – might help calibrating those  
318 models, as we here showed Landsat-based models can reproduce field-based estimates, which  
319 are rare in many regions worldwide.

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323 Berlin for access to the phenological observations. We also thank Marius Derenthal for help  
324 with code implementation in the Google Earth Engine.

325

326 **Data and code accessibility**

327 All data and code used in this publication are available for review at

328 [https://github.com/corneliussenf/phenoBayes\\_drivers](https://github.com/corneliussenf/phenoBayes_drivers) and will be archived using

329 <https://zenodo.org/> after the paper has been accepted for publication.

330

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