

A systematic approach and software for the analysis of point patterns on river networks

Wolfgang Schwanghart^{1*}, Christian Molkenhain², Dirk Scherler^{3,4}

¹ Institute of Environmental Science and Geography, University of Potsdam, 14476 Potsdam-Golm,
Germany

² Institute of Mathematics, University of Potsdam, 14476 Potsdam-Golm, Germany

³ German Research Centre for Geosciences (GFZ), Earth Surface Geochemistry, 14473 Potsdam,
Germany

⁴ Institute of Geological Sciences, Freie Universität Berlin, 12249 Berlin, Germany

* Corresponding author: w.schwanghart@geo.uni-potsdam.de

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Abstract

Many geomorphic phenomena such as bank failures, landslide dams, riffle-pool sequences and knickpoints can be modelled as spatial point processes. However, as the locations of these phenomena are constrained to lie on or alongside rivers, their analysis must account for the geometry and topology of river networks. Here, we introduce a new numeric class in TopoToolbox called Point Pattern on Stream networks (PPS), which supports exploratory analysis, statistical modelling, simulation and visualization of point processes. We present two case studies that aim at inferring processes and factors that control the spatial density of geomorphic phenomena along river networks: the analysis of knickpoints in river profiles, and modelling spatial locations of beaver dams based on topographic metrics. The case studies rely on exploratory analysis and statistical inference using inhomogeneous Poisson point processes. Thereby, statistical and probabilistic procedures implemented in PPS provide a systematic approach for treating of uncertainties. PPS provides a consistent numeric framework for modelling point processes on river networks with a wide range of applications in fluvial geomorphology, but also other disciplines such as ecology.

31 Introduction

32 Many geomorphic phenomena along rivers can be represented as spatial point processes. For
33 example, bank failures (Fonstad and Marcus, 2003; Liang et al., 2015), landslide dams (Fan et al.,
34 2020; Korup, 2006; Tacconi Stefanelli et al., 2015), riffle-pool sequences (Golly et al., 2019), wood
35 jams (Scott et al., 2019; Wohl, 2013), and knickpoints (Berlin and Anderson, 2007; Gailleton et al.,
36 2019; Phillips and Lutz, 2008; Schwanghart and Scherler, 2020) are phenomena that occur at specific
37 locations along rivers and that – at particular spatial scales of analysis – can be represented as point
38 features. Many questions about these processes are inherently linked to their spatial arrangement. For
39 example: Do these phenomena occur randomly in space, or are there mechanisms that cause these
40 phenomena to cluster spatially? Are there interactions between these phenomena that generate some
41 characteristic spacing between them or do additional factors exist that promote their spatial density? A
42 spatial point process is a stochastic mechanism that generates patterns of points in space. The analysis
43 of point patterns – a major subject within the field of spatial statistics – is concerned with
44 understanding and modelling the stochastic and deterministic mechanisms that generate the patterns
45 (Baddeley et al., 2015). While point pattern analysis has pervaded many geoscientific disciplines,
46 applications in geomorphology are relatively rare (Bishop, 2007b, 2007a; Clark et al., 2018; Kandakji
47 et al., 2020; Kraft et al., 2011; Lombardo et al., 2018, 2019; Oeppen and Ongley, 1975; Sochan et al.,
48 2019; Tarboton et al., 1989).

49 The aim of this study is to explore the opportunities that the analysis of spatial point patterns offers in
50 geomorphology. In particular, we are interested in point patterns that occur along river networks. The
51 network-led spatial configuration makes this kind of analysis challenging. Statistical techniques
52 designed for point patterns in two-dimensional space are usually based on the Euclidean distance
53 between points which can be very different from distances along networks (Ang et al., 2012; Okabe et
54 al., 2009). While methodological developments in geostatistics have established a mature set of tools
55 to tackle interpolation along stream networks (Cressie et al., 2006; Ganio et al., 2005; Skoien et al.,
56 2006; Ver Hoef et al., 2006), point pattern analysis on networks is a relatively young and active field
57 of research (Baddeley et al., 2015; Okabe and Sugihara, 2012).

58 Here, we present an extension to the MATLAB-based terrain analysis software TopoToolbox
59 (Schwanghart and Kuhn, 2010; Schwanghart and Scherler, 2014) called PPS (Point Pattern on Stream
60 networks), which implements the statistical principles and techniques of point pattern analysis on
61 linear networks. PPS complements other tools for point pattern analysis. The R-package spatstat is
62 among the most comprehensive software packages that also handles point patterns on networks
63 (Baddeley et al., 2015) and has strongly influenced the design of PPS. In addition, SANET (Okabe et
64 al., 2018) is a toolbox for ArcGIS for analyzing events that occur on networks or alongside networks.
65 Incorporating PPS in TopoToolbox offers seamless workflows including data import, analysis,
66 modelling and visualization in the MATLAB programming environment. The ease of working in one

67 computational programming environment and the availability of computational tools for working with
68 river network data was a major motivation to develop PPS alongside TopoToolbox.

69 In the following text, we provide a brief summary of the theory, computational methods, and
70 implementation of PPS. We furthermore present two applications in which point pattern analysis
71 serves as an approach to investigating and modelling the occurrence of geomorphic forms and
72 processes along river networks.

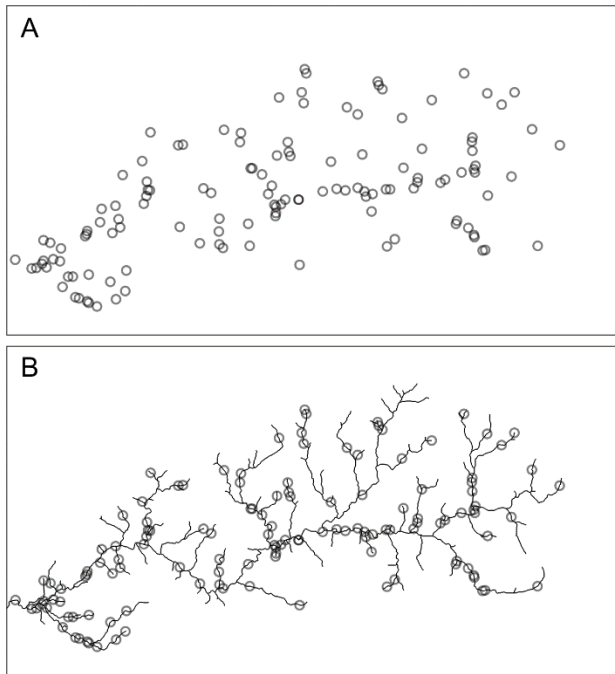
73 Point processes on networks

74 Spatial analysis of point patterns is predicated on the concept of first and second order effects or
75 variations. First order variations arise from spatial trends or other covariates that control the spatial
76 density of points. For example, the spatial density of bank collapses along a river is a function of the
77 type of rocks or sediments, but may additionally be controlled by spatial trends in water level
78 fluctuations, river gradient and planform geometry (Fonstad and Marcus, 2003; Liang et al., 2015).
79 Bank collapses can also impact the occurrence of other events of bank failures. Once a bank has failed
80 it may change patterns of river flow and/or make adjacent banks susceptible to failure due to
81 debuitressing. Close to an existing bank failure we might thus expect even more bank failures. In this
82 case, we hypothesize a second order effect due to direct physical interactions that cause bank collapses
83 to be more frequent close to other failures. Another example for a second order variation is the effect
84 of seed dispersal on the spatial density of plants, but we may also think of processes that inhibit small
85 distances between adjacent points such as the competition for nutrients, light and water. In fluvial
86 geomorphology, riffle-pool and step-pool sequences are phenomena that exhibit regular distances
87 (Golly et al., 2019; Knighton, 1998; Tarboton et al., 1989). A major goal of point pattern analysis
88 pertains to the analysis and modelling of first and second order variations from point data (Baddeley et
89 al., 2015). Although this might appear straightforward at first glance, separating the two effects from
90 each other is often challenging.

91 Commonly, spatial point processes are analyzed in two or three spatial dimensions and time.
92 Frequently, however, the events (entities, points, locations) occur on or alongside networks. Car
93 accidents, for example, are events on a road network whereas supermarkets are locations alongside the
94 road network. Whether on or alongside, the coordinates of these points are constrained by a spatial
95 network (network-constrained events or, in short, network events (Okabe and Sugihara, 2012)). Paths
96 between points follow the network's edges and thus distances rarely follow direct Euclidean distances.
97 Instead, standard practice is to measure distances in networks by the length of the shortest path, least-
98 cost or resistance distances (Rakshit et al., 2017). To this end, many existing methods in point pattern
99 analysis rely on the Euclidean distance which may be inappropriate or fallacious if applied to network
100 events (Okabe and Sugihara, 2012; Rakshit et al., 2017) (Figure 1). It may seem straightforward that
101 distances in river networks ought to be calculated in metric units from the outlet or channelheads, but

102 we may also weight these distances by stream flow (Ver Hoef et al., 2006) or elevation (Foltête et al.,
103 2008), or use metrics such as χ -transformed distance (Harkins et al., 2007; Perron and Royden, 2013)
104 which are increasingly used in the analysis of river profiles and network topology. The choice of
105 distance metric depends on the application and should be guided by additional information (Rakshit et
106 al., 2017). Hence, not all network-constrained points must be analyzed using network-derived
107 distances. In an analysis of the spatial patterns of river junctions, for example, Oeppen and Ongley
108 (1975) relied on the planar Euclidean distance.

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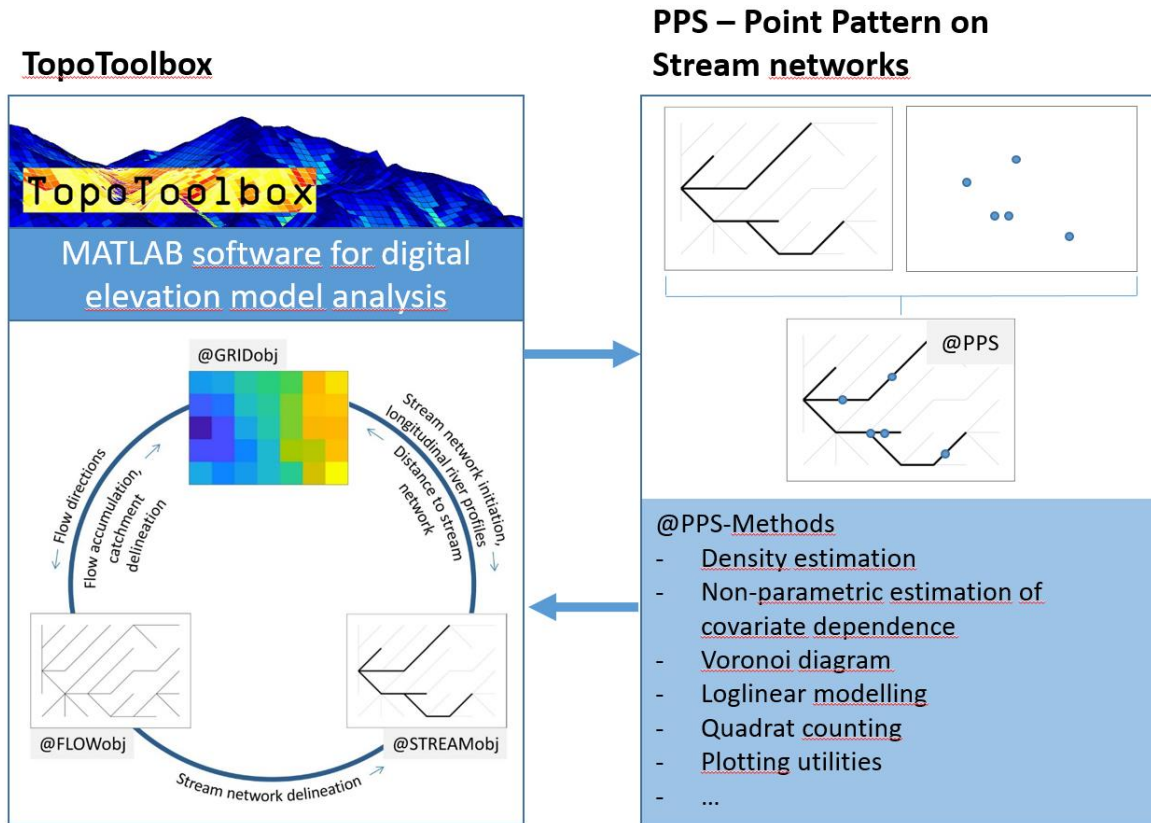
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111 *Figure 1: Spatial point processes clearly lack a completely random pattern (A) if we ignore that their locations are*
112 *constrained by a network. If we take this constraint into account (B), it is more difficult to decide if the observed point pattern*
113 *is completely random or not.*

114 TopoToolbox as the basis for PPS

115 PPS is based on TopoToolbox, a MATLAB software for topographic analysis (Schwanghart and
116 Scherler, 2014). TopoToolbox pursues an object-oriented programming approach that simplifies
117 programming tasks which involve gridded digital elevation models (DEMs) and topographic
118 derivatives (Figure 2). A DEM is stored as an object of the class GRIDObj which includes the matrix
119 of elevation values and information on extent, resolution, and coordinate reference system. Flow
120 directions are derived from DEMs and are stored as an instance of the class FLOWObj. Using
121 topological sorting of the flow network (Braun and Willett, 2013; Hergarten and Neugebauer, 2001),
122 this computational object enables the derivation of drainage basins or computations such as flow
123 accumulation (Schwanghart and Scherler, 2014). Moreover, FLOWObj is the basis for the delineation
124 of stream networks which are stored as an object of the class STREAMObj. Any computation with
125 stream networks adopts highly efficient algorithms from graph theory (Heckmann et al., 2015). PPS

126 takes advantage of the algorithms that are readily available in TopoToolbox and extends their
 127 capabilities to numerous new applications that enable the analysis of point patterns on stream networks
 128 (Figure 2).



129
 130 *Figure 2: Numerical classes in TopoToolbox and the new PPS class.*

131
 132 **Numeric implementation and methods of PPS**

133 Computational representations of networks can rely on either vector or raster representations (Okabe
 134 and Sugihara, 2012). Being built on the STREAMObj class, PPS uses a hybrid approach. An object of
 135 class STREAMObj is derived from a DEM. Thus, the nodes of the PPS stream network refer to cell
 136 centers of the DEM. The topology of the network is determined by edges that link the cell centers in
 137 cardinal and diagonal directions (8-connectivity). Each node in the network can have attribute values
 138 which we refer to as node-attribute list. An instance of PPS is created by combining a stream network
 139 with a point dataset represented by a set of coordinates. If the points are not located on the stream
 140 network, they are snapped to the nearest nodes of the stream network, and their distance to the stream
 141 can be an attribute of the points. Formally, PPS thus adopts a fine-pixel approximation of a point
 142 pattern (Baddeley et al., 2015).

143 *Table 1: Overview on PPS functions.*

Function	Description
<i>Creating an instance of PPS</i>	

PPS	Constructor function that creates an instance of class PPS from a stream network (STREAMObj) and a set of points. Alternatively, the function can generate randomly distributed points on stream networks, or calculate intersections with a network of lines.
<i>Explorative analysis</i>	
density	Kernel density estimator on stream networks
ecdf	Empirical cumulative density function
intensity	Intensity (points per unit distance)
histogram	Histogram of point pattern on stream network
rhoht	Nonparametric estimation of covariate dependence
cluster	Hierarchical spatial clustering of points
<i>Inference and simulation</i>	
fitloglinear	Fitting a loglinear intensity model
bayesloglinear	Bayesian analysis of a loglinear intensity model
quadratcount	Quadrat counting
random	Simulation of points using a loglinear intensity model
simulate	Simulation of points using random thinning
ploteffects	Plot effect of a single predictor variable in a model
roc	Receiver-operating characteristics curve
<i>Other utilities</i>	
as	Utility to convert PPS object to other formats
pointdistances	Pairwise distances between points in PPS
voronoi	Voronoi tessalation of the river network based on points in PPS
hasduplicates	Determine if PPS has duplicate points
removeduplicates	Remove duplicate points in PPS
convhull	Calculate convex hull of points
aggregate	Merge labelled points to a new object of PPS
idw	Inverse distance weighted interpolation on stream networks
shapewrite	Export PPS as shapefile
<i>Visualization</i>	
plot	Plot stream network with points
plotc	Plot colored stream network with points
ploteffects	Plot effect of covariate in a loglinear model
plotdz	Plot longitudinal profile with points
plotpoints	Plot points only
wmplot	Plot stream network with points in a webmap

144

145 A PPS object is created using an instance of STREAMObj and a set of coordinates of points, line
 146 features (e.g. fault traces) that intersect the stream network, or a model that randomly generates points
 147 (Figure 2). Supported models are the binomial and the homogeneous Poisson point process that
 148 randomly distribute points on the network given a specified total number of points and intensity
 149 (average number of points per unit length), respectively. For example, the pattern in Figure 1b was
 150 generated by a Poisson process with an intensity of $5 \times 10^{-4} \text{ m}^{-1}$. Once initiated, an object of PPS can
 151 access numerous functions (or methods) which are summarized in Table 1. The functions are broadly
 152 categorized into tools for explorative analysis, inference and simulation, and visualization. In addition,
 153 there are a number of conversion tools and other utilities such as interpolation tools.

154 Explorative analysis of point patterns often begins with kernel density estimates to highlight spatially
 155 varying densities of points. While kernel density estimates are straightforward in 1D, 2D or higher

156 dimensions, they are not directly applicable to networks. Conventional 2D kernel density estimators
157 applied to points on river networks may easily overestimate densities along adjacent rivers albeit the
158 rivers may be disconnected. Applying 1D kernel density estimators to networks, however, is also
159 fallacious because it fails to conserve mass where networks branch (McSwiggan et al., 2017; Okabe
160 and Sugihara, 2012). The function *density* adopts the approach of McSwiggan et al. (2017) who
161 implement Gaussian kernel density estimation on networks using an approach that perceives Gaussian
162 kernels as heat kernels and the variable densities along the network as Brownian diffusion
163 (McSwiggan et al., 2017).

164 Clustering is a technique that groups similar objects to classes. In spatial point pattern analysis this
165 technique is used to detect spatial clusters of points, and to merge them eventually to a set of new
166 points. The function *cluster* uses hierarchical clustering based on the shortest-path distances of all
167 points (Okabe and Sugihara, 2012). The resulting spatial clusters can subsequently be merged using
168 the function *aggregate*, which computes cluster centers by finding the network node that minimizes
169 the sum of squared shortest distances from each point in the cluster.

170 An important question in the analysis of point patterns is whether the intensity of points depends on
171 spatial covariates. Parametric models describing this dependence have a long tradition in point pattern
172 analysis. These models require that the dependence structure of the model is known. Yet, often we do
173 not know the form of the model, or the form is too complicated to be fitted by a parametric model.
174 Thus, nonparametric estimation provides an important exploratory approach, since it determines the
175 model structure from the data. While nonparametric models do not completely lack parameters, they
176 model the relationship between variables with fewer assumptions, and are thus particularly suitable for
177 explorative analysis (Baddeley et al., 2012). We implemented this nonparametric technique in PPS
178 with the function *rho*, which also calculates confidence intervals using bootstrapping.

179 Nonparametric analysis of covariate dependence makes no assumptions about the shape of the
180 functional relationship between point density and an explanatory variable. However, if the type of
181 relationship is known or hypothesized, then parametric techniques are a more powerful way to analyze
182 the data (Baddeley et al., 2015). The most common model in point pattern analysis is the
183 inhomogeneous Poisson point process model with an intensity which is a loglinear function of the
184 covariates (Baddeley et al., 2015)

$$\lambda(u) = e^{B(u) + \theta^T Z(u)} \quad (1)$$

185 where λ is the intensity of points at locations u , B is a known baseline intensity, and θ is a vector of p
186 parameters for a vector-valued function $Z(u) = [Z_1(u) \dots Z_p(u)]$. Loglinear models assume that the
187 intensity is intrinsically positive-valued and enables to model the dependence of intensity on numeric
188 and categorical variables. The model assumes no interactions between points and thus has the
189 advantage that parameter estimation can rely on standard techniques such as logistic regression or
190 Poisson regression. PPS implements Poisson models using the function *fitloglinear*. The function

191 accesses the function *fitglm*, which is part of the MATLAB Statistics and Machine Learning Toolbox
192 and fits generalized linear least squares problems. PPS also features a Bayesian approach to analyze
193 loglinear models. The function *bayesloglinear* interfaces with the BayesReg Toolbox (Makalic and
194 Schmidt, 2011, 2016) which provides highly efficient and numerically stable implementations of
195 penalized regression techniques.

196 PPS features tools to study first and second order effects in point processes. However, current
197 inferential methods in PPS are based on models that assume that point patterns do not exhibit second
198 order effects. Variable densities of points in space are assumed to relate to some factor or covariate.
199 Models exist that can be used to explain clustering or regular patterns and include Cox, Neyman-Scott,
200 Gibbs or Hawkes models. These models are currently not supported in PPS.

201 Case studies

202 Applying the techniques and tools outlined in the previous section, we present two case studies in
203 which the analysis of point patterns is used to extract information about geomorphological processes
204 that take place on or alongside rivers. In the first case study, we demonstrate how explorative analysis
205 of knickpoints in river profiles of the Big Tujunga catchment in California can help reveal two phases
206 of landscape rejuvenation. In the second case study, we investigate the spatial distribution of beaver
207 dams in the Tualatin basin, Oregon, and model their geomorphometric constraints. For brevity, some
208 of the data and methods of the case studies are summarized in Table 2. All data are open and freely
209 available.

210 *Table 2: Data used in the case studies.*

Case study	Knickpoints in the Big Tujunga catchment	Beaver dams in the Tualatin basin
Location	California, USA, 34.2°N, 118.2°W	Oregon, USA, 45.4°N, 122.8°W
Catchment area	293 km ²	1803 km ²
DEM (spatial resolution)	SRTM-1 (30 m)	NED (10 m)
Point pattern	52 knickpoints detected by knickpointfinder	510 beaver dams from Smith (2019)
Additional data	Vector data with faults from (USGS and NMBMMR, 2019)	Stream network vector data from Nagel et al. (2017)

211

212 [Knickpoints in the Big Tujunga basin](#)

213 Rivers in the Big Tujunga catchment in the San Gabriel Mountains feature numerous knickpoints
214 along their longitudinal profiles. These knickpoints are unrelated to lithological boundaries and they
215 are found in relatively narrow elevation bands (Wobus et al., 2006), which suggests that they formed

216 at the range front due to acceleration in slip rate of the Sierra Madre Fault Zone, and the concomitant
217 adjustment of the stream network to the higher uplift rate (DiBiase et al., 2015). The aim of this
218 example is to illustrate how an explorative analysis of knickpoint patterns helps in assessing a model
219 of landscape response times to changes in tectonic uplift.

220 The most widely used model of fluvial incision and knickpoint migration is the stream power incision
221 model (SPIM) (Lague, 2014), which states that the rate at which elevations z along a river change over
222 time t is a function of uplift U , erosional efficiency K , upslope area A and local river gradient

$$\frac{\partial z(x)}{\partial t} = U(x, t)K(x, t)A(x, t)^m \left| \frac{dz}{dx} \right|^n \quad (2)$$

223

224 where x is the distance from the river outlet along the flow network, and the exponents m and n are
225 empirical constants. Assuming that U and K do not vary in time and space, and that drainage
226 configurations remain unchanged, the steady state channel slope is calculated with

$$\left| \frac{dz}{dx} \right| = \left(\frac{U}{K} \right)^{\frac{1}{n}} A(x)^{-\frac{m}{n}} \quad (3)$$

227

228 a relation between channel slope and area that predicts an upward concave river profile (Hack, 1957).
229 Based on Eq. 3, Harkins et al. (2007) and Perron and Royden (2013) introduced a coordinate
230 transformation which linearizes the power-law relation. The linearization takes the integral of the left
231 and right term in Eq. 3 so that elevation becomes a linear function

$$z(x) = z(x_b) + \left(\frac{U}{KA_0^m} \right)^{\frac{1}{n}} \chi \quad (4)$$

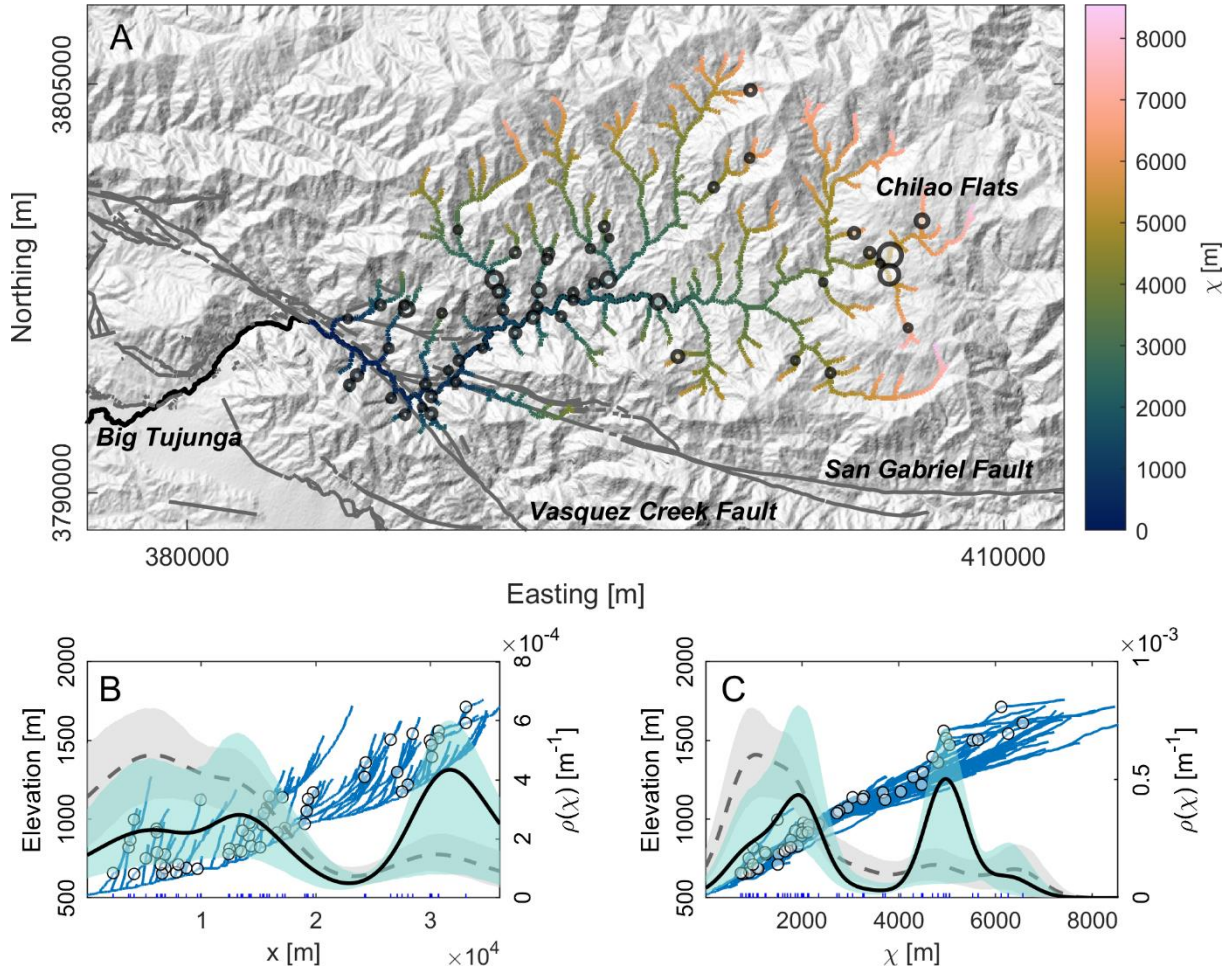
232 where

$$\chi = \int_{x_b}^x \left(\frac{A_0}{A(x)} \right)^{\frac{m}{n}} dx \quad (5)$$

233 with A_0 (which we set to 10^6 m^2) and x_b is the location of the base level (Perron and Royden, 2013).
234 The linear form of the SPIM (with $n=1$) predicts that perturbations to river elevations, for example by
235 base level change, migrate upstream as a function of upstream area (Berlin and Anderson, 2007). χ -
236 transformation normalizes for upstream area so that any base level change at x_b in the past, should
237 result in knickpoints that cluster at a specific value of χ , irrespective of whether the perturbation has
238 travelled upstream the trunk river, or any of its tributaries (Perron and Royden, 2013; Schwanghart
239 and Scherler, 2020). χ thus serves as a metric for distances travelled by perturbations upstream in the
240 river network (Fox et al., 2014).

241 In order to test the knickpoint celerity model in the Big Tujunga catchment, we derived a stream
242 network with a minimum supporting upslope area of 0.9 km^2 . Locations of knickpoints were

243 calculated using the function knickpointfinder, an automated method of knickpoint identification
 244 based on iterative fitting of strictly concave stream profiles that is implemented in TopoToolbox and
 245 described in Stolle et al. (2019). Applying a tolerance of 20 m – which is about the maximum
 246 elevation error recorded along streams of the SRTM-1 (Schwanghart and Scherler, 2017) – yields 52
 247 knickpoints (Figure 3A). Knickpoint height – the elevation difference between the fitted profile and a
 248 knickpoint, and a measure taken here for the prominence of each knickpoint – ranges between 22 and
 249 216 m.

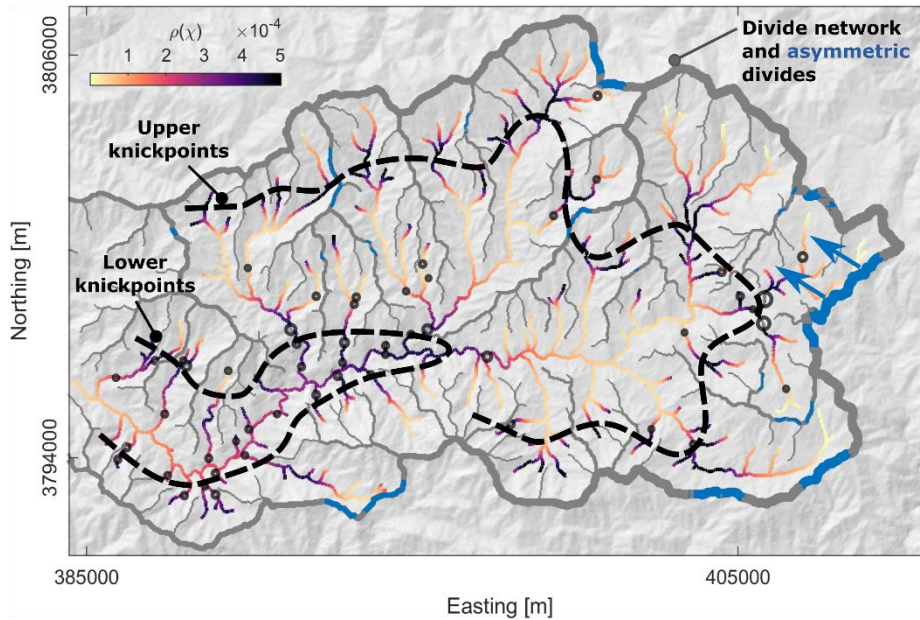


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251 *Figure 3: Knickpoint patterns in the Big Tujunga catchment. A) Hillshade map of the catchment and faults (gray lines; after*
 252 *Morton and Miller, 2006), knickpoints and χ -values of the river network. The size of the knickpoint symbols linearly scales*
 253 *with knickpoint heights, which range between 22 and 216 m. B) Distribution of knickpoints along river profiles (blue lines).*
 254 *Gray dashed line shows the nonparametric dependence of knickpoint locations (with gray envelopes indicating bootstrapped*
 255 *95% confidence intervals) as a function of distance from the range-bounding fault. The black line shows the dependence*
 256 *estimate weighted by the knickpoint height. The bandwidth for both estimates is 3000 m. C) Same as B), but with the*
 257 *covariate being χ and bandwidth being 400 m.*

258 The majority of knickpoints are located in the lower part of the catchment (Figure 3A), which is also
 259 reflected by the nonparametric estimate (function *rho*) which shows how knickpoint locations
 260 depend on the distance to the range-bounding fault (Figure 3B, dashed gray line). Weighting
 261 knickpoints by their squared heights (black line) the occurrence of few but prominent knickpoints in
 262 the upper part of the basin is accentuated. We calculated χ with an m/n ratio of 0.4 which has
 263 previously been used by Perron and Royden (2013) for the same catchment. Figure 3C is similar to B,

264 but depicts density estimates as a function of χ . Again, a non-weighted density estimation highlights
265 the knickpoints in the vicinity to the catchment outlet, whereas weighting them reveals two
266 pronounced peaks at χ values around 2000 and 5000 m. However, uncertainty intervals (based on
267 bootstrapping) of the density estimates of the second peak are high and reflect the scarcity of
268 knickpoints in the upper part of the catchment.



269

270 *Figure 4: Actual and expected spatial patterns of knickpoints in the Big Tujunga basin. The two dashed lines are manually*
271 *drawn to highlight the two generations of upstream migrating knickpoints and their expected locations. The gray lines depict*
272 *the drainage divide network (Scherler and Schwanghart, 2020), with blue sections showing asymmetric divides and the*
273 *inferred movement is indicated by the blue arrows.*

274 Mapping the patterns of knickpoint density obtained from the weighted nonparametric dependence
275 model in Figure 3C back to spatial coordinates (Figure 4) reveals the expected spatial locations of
276 knickpoints. Clearly, as the model was obtained from actual knickpoint locations, both must be
277 consistent to a certain degree. Notwithstanding, actual and expected knickpoint patterns show notable
278 differences in many locations that require explanation. These differences are particularly obvious for
279 the older wave of knickpoints that mark the transition to the Chilao Flats and that are expected to be
280 present high up in other tributaries to the Big Tujunga as well. However, most headwater channels are
281 devoid of knickpoints. There are several explanations for a lack of consistency between expected and
282 actual knickpoint patterns. First, variations in bedrock erodibility manifest themselves in a series of
283 waterfalls in the overstepped knickzone straddling the Chilao Flats. These waterfalls have been
284 previously found to have slowed down knickpoint retreat by at least an order of magnitude (DiBiase et al.,
285 2015). Other tributaries may lack such resistant layers and thus knickpoints may have already left
286 the system. Second, headwater channels may be dominated by debris-flow processes (Hergarten et al.,
287 2016; Stock and Dietrich, 2003) which may result in faster incision and possibly smearing of
288 knickpoints in the channels. Third, inconsistencies between expected and observed knickpoint patterns
289 may arise from drainage reorganization. Our analysis weighted the most prominent knickpoints, yet

290 these knickpoints may be those that have been particularly affected by divide migration. The margins
291 of the Chilao Flats show highly asymmetric divides (Scherler and Schwanghart, 2020) (Figure 4)
292 which suggest possibly past and ongoing drainage reorganization. Such reorganization may
293 significantly alter drainage areas and discharge, and thus impact on knickpoint celerities which in
294 return will result in more scattered knickpoint locations (Schwanghart and Scherler, 2020).

295 **Beaver dams in the Tualatin basin, Oregon**

296 Beavers are ecosystem engineers that build dams across and alongside rivers. These wood
297 accumulations increase floodplain storage of water, sediment, organic matter and nutrients, and thus
298 have several ecological benefits (Bouwes et al., 2016; Macfarlane et al., 2017; Wohl, 2013). As beaver
299 dams impound water upstream, they also raise the possibility of beaver dam outburst floods. Although
300 such outburst floods are rare, there were cases where such events greatly exceeded discharges of
301 meteorological floods (O'Connor et al., 2013). Given both ecological benefits and outburst hazard,
302 potential beaver dam locations should thus be known for managing river restoration and flood risk.

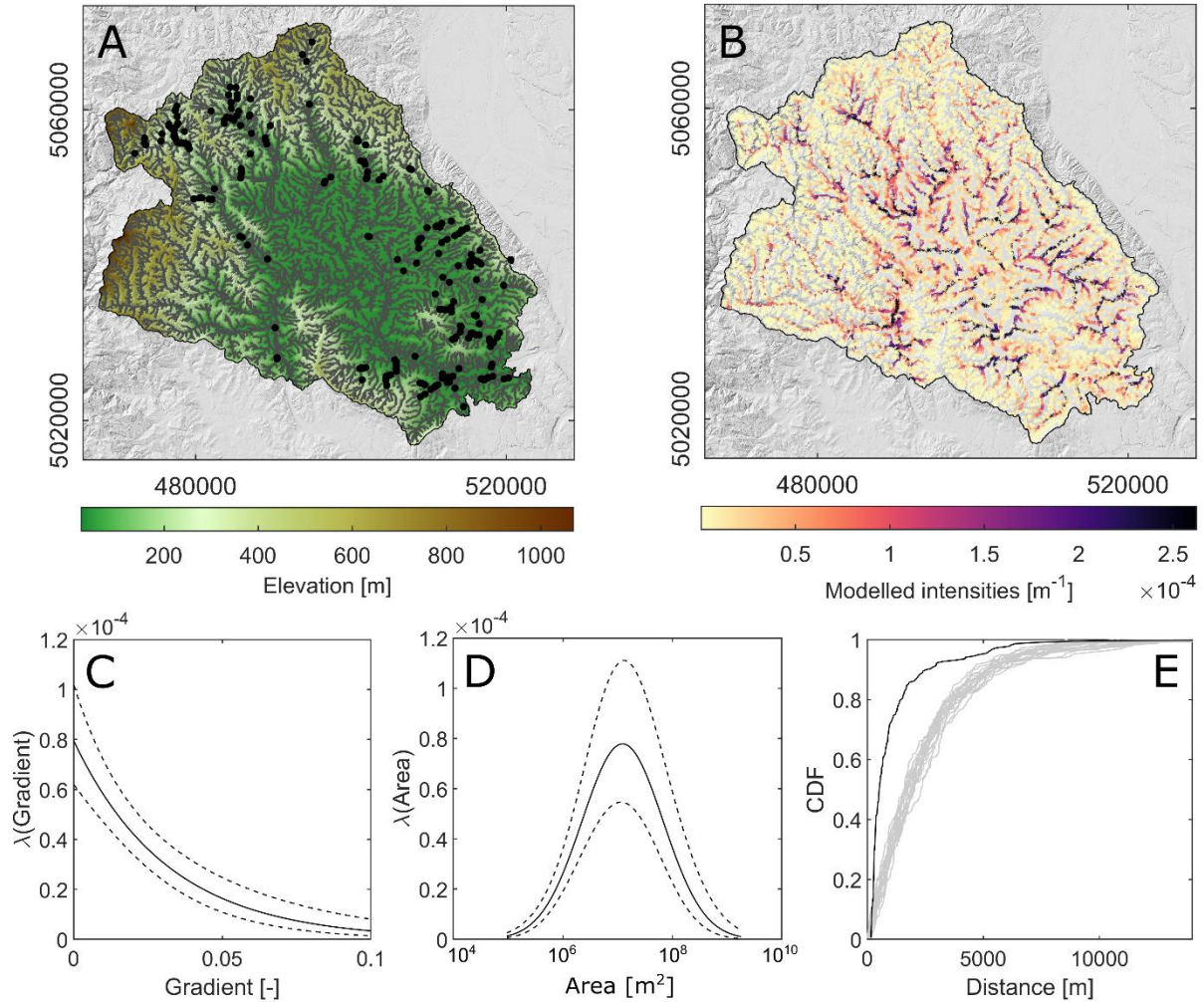
303 In this case study, our analysis focuses on topographic controls on the occurrence of beaver dams that
304 can be derived solely from catchment-scale digital elevation data. Several properties determine the
305 degree to which beavers colonize and sustain a population (Gurnell, 1998), and we hypothesize that
306 beaver habitats are primarily a function of stream flow and stream gradient. Beavers require sufficient
307 stream flow as a reliable water source. Yet, rivers should neither be too wide nor too deep to inhibit
308 building and persistence of dams (Collen and Gibson, 2000; Gurnell, 1998; Macfarlane et al., 2017).
309 At the same time, river gradient should be relatively low to impound sufficiently large areas.
310 Therefore, steep and rocky rivers are generally less favored by beavers as dams in such streams are
311 susceptible to damage during high-magnitude discharges and have low impounding efficiency
312 (Gurnell, 1998).

313 To test the above hypothesis, we studied the distribution of beaver dams in the Tualatin basin, Oregon
314 (Table 2, Figure 5A). In our analysis, we used upstream area as proxy for stream flow, which we
315 derived from the DEM using flow accumulation. Anthropogenic features such as bridges and culverts
316 produced some artifacts when computing the stream network from the original DEM. Thus, we used
317 hydrographic data from Nagel et al. (2017) and preprocessed the DEM using stream burning (Reuter et
318 al., 2009). We extracted the stream network based on an area threshold of 0.1 km², and smoothed the
319 profiles using the CSR (constrained regularized smoothing) algorithm (Schwanghart and Scherler,
320 2017) with a smoothing factor of $K = 10$. The smoothed elevations are subsequently used to calculate
321 the local stream gradient. Commonly, stream gradients derived from DEMs fluctuate strongly as they
322 are highly sensitive to errors in the elevation data (Wobus et al., 2006). Our approach of smoothing the
323 profiles created local gradients that mimic those obtained from a moving window approach with a
324 kernel size of ~200 m.

325 Beaver dam locations were obtained from the version 2.0 of the data released by Smith (2019). The
326 data was compiled by the U.S. Geological Survey (USGS) and comprises information on 510 beaver
327 dams. Some recorded locations are very close to each other and likely correspond to the same beaver
328 populations. Thus, we merged locations using the function *cluster* (Table 1). The function implements
329 hierarchical clustering based on the shortest-path distance matrix of all points using an average linkage
330 method. We chose a cutoff of 160 m and obtained 217 unique locations, which we used for the
331 subsequent analysis.

332 The pattern of beaver dams (Figure 5A) suggests that their intensity is spatially inhomogeneous. This
333 hypothesis can be tested using techniques such as quadrat counting (function *quadratcount*). Quadrat
334 counting subdivides the network into roughly equal sized subnetworks and then counts the number of
335 locations within each subnetwork. Under the assumption of complete spatial randomness, the
336 distribution of points in each subnetwork should follow a Poisson distribution with homogeneous
337 intensity, a hypothesis that we investigate with a χ^2 -test (note that χ^2 has nothing in common with the
338 χ -transformation in the previous case study). The χ^2 -test underscores ($p < 0.0001$) the visual
339 impression that spatial locations of beaver dams in the Tualatin Basin are not completely random.

340 To test whether drainage area and stream gradient can be used to explain spatial variations in beaver-
341 dam density, we fit a loglinear model with stream gradient and the decadic logarithm of upslope area
342 as independent variables. The loglinear model has an intercept and a first-degree polynomial for
343 gradient and second-degree polynomial for upslope area. Moreover, we add an interaction term
344 (product of both predictors) to investigate whether the interrelationship of stream gradient and upslope
345 area determines spatial beaver-dam densities.



346

347 *Figure 5: Modelling the locations of beaver dams in the Tualatin basin, Oregon, US. A) Hillshade map of the basin, stream*
 348 *network, and the locations of beaver dams (black dots). B) Modelled intensities of beaver dams using an inhomogeneous*
 349 *Poisson point pattern. C+D) Fitted responses to a single predictor: C) Stream gradient and D) drainage area. E) Empirical*
 350 *nearest-neighbor distance distribution function for actual beaver dam locations (solid black line) compared to distribution*
 351 *functions of 20 simulated point patterns derived from the inhomogeneous Poisson model (gray lines).*

352

353 We fit the model using stepwise regression which removes parameters or terms that fail to improve the
 354 model fit measured by the Akaike-Information Criterion (AIC). Stepwise regression removes the
 355 interaction term so that the final model is

$$\hat{\lambda}(u) = e^{\beta_0 + \beta_1 g(u) + \beta_2 a(u) + \beta_3 a^2(u)} \quad (6)$$

356 where $\hat{\lambda}$ is the estimated density of beaver dams (Figure 5B), β_0 is an offset, and β_{1-3} are the
 357 parameters for stream gradient g and the decadic logarithm of upslope area a and its quadratic form,
 358 respectively. Overall, the model is highly significant compared to a model with a pure offset ($p =$
 359 7.28×10^{-82}) and the area under the ROC (receiver-operating characteristic) curve, a measure of
 360 aggregated classification performance, is 0.85 (0.83-0.86 simulation confidence intervals). The values
 361 for the parameters, their uncertainties and individual p-values are listed in Table 3 and the fitted
 362 responses to the single variables are shown in Figure 5C and D.

363 Table 3: Estimated parameters of a loglinear model of beaver-dam locations in the Tualatin basin, Oregon, US.

	Estimate	SE	t-statistics	p-value
β_0	-51.99	4.62	-11.26	2.18E-29
β_1	-31.60	4.99	-6.33	2.48E-10
β_2	12.97	1.36	9.55	1.35E-21
β_3	-0.91	0.10	-9.20	3.68E-20

364

365 Although the model provides a reasonable fit to the data, it may neglect other potential factors.
366 Previous studies found that stream depth, sandbar width, and anabranching (secondary rivers, sloughs)
367 as well as access to forage are important controls on the spatial distribution of beaver dams (e.g.
368 Scrafford et al., 2018). Our data and the representation of the flow network by D8 flow directions do
369 not permit us to represent these factors. In addition, beaver dams entail hydrologic (creating wetlands),
370 hydraulic (slow down runoff), geomorphic (sediment trapping), and ecological feedbacks
371 (subirrigation of downstream valley bottoms that promotes establishment and expansion of riparian
372 vegetation); all of which tend to increase stream complexity and channel-floodplain connectivity
373 (Macfarlane et al., 2017). These feedbacks may lead to spatial clustering, as beaver-engineered river
374 reaches may increase local beaver populations. Our model does not capture such clustering effects.
375 However, to test whether the data exhibits such spatial clustering after accounting for the first-order
376 effects of stream gradient and discharge, we simulated 20 realizations of beaver dam locations using
377 our model (function *simulate*), each time measuring the cumulative distribution of nearest neighbor
378 distances (the G-statistics as measured by the function *gfun* (Baddeley et al., 2015)). Figure 5E shows
379 that the actual distribution of beaver dams exhibits a much stronger clustering compared to the
380 simulated points although we declustered the original data. Whether this clustering may evolve from
381 individual beaver populations or positive feedbacks exerted by beavers on their habitats remains
382 shrouded. However, modelling such interactions may improve with more advanced point pattern
383 models, whose treatment is beyond the scope of this study and which are currently not implemented in
384 PPS.

385 Discussion

386 The two studies that we presented showcase the new TopoToolbox extension PPS, which supports the
387 analysis of point patterns on stream networks. The studies have in common that different geomorphic
388 phenomena can be conceptualized as point processes that occur on or alongside stream networks.
389 Knickpoints in bedrock rivers, for example, migrate upstream along the river network, but with no
390 apparent link between adjacent rivers. This strict constraint could be relaxed when analyzing beaver
391 populations because beavers may shortcut distances between adjacent rivers when expanding into new
392 territory. Our analysis did not take the potential movement of beavers between streams into account,
393 which may in particular affect second-order patterns of beaver dams. To this end, investigating such

394 effects would require distance metrics between points that combine distances along and aside stream
395 networks.

396 Our case study on the spatial distribution of knickpoint relied on weighting knickpoints by their
397 height. Yet, we didn't include such attributes in the analysis of beaver dams, although these
398 biogeomorphic features commonly have highly variable sizes (Turowski et al., 2013), which could be
399 used to weight observations in the models. While such attribute data was not available in this study,
400 we note that it may be useful to record this data when recording point pattern data in the field.
401 Moreover, additional attribute data could be used in the analysis of marked point patterns, a suite of
402 methods to explore and model point patterns with attribute data. However, such techniques are
403 currently not yet available in PPS.

404 PPS relies on the geographic representation of geomorphic objects or features as points, and streams as
405 lines or network of lines. It follows that the studied phenomena must be conceptualized as points,
406 although they may often have volumes associated with them and they may have vaguely defined limits
407 or be overlapping (Evans, 2012; Goodchild, 2011; Smith, 2011). As common in GIS analysis, such a
408 representation embodies spatial scale to some degree. For point pattern analysis, it is crucial to
409 remember that spacing between points may be observed if points actually represent areal
410 nonoverlapping features. Moreover, as points are constrained to lie on nodes of the stream network,
411 which are derived from the underlying DEM, the representation of network events is tightly linked to
412 the spatial resolution of the DEM. This also entails that the density of points should not be too high, as
413 it may cause points to share the same locations, a situation usually not foreseen in point pattern
414 analysis. In addition, the distance between two vertices is a lower bound of the true distance, if we
415 assume that all line vertices are located on the central line of the river (Goodchild, 2011). In
416 TopoToolbox and thus also PPS, the geometry of stream networks is determined by the Moore
417 neighborhood (8-connectivity) of the D8 flow direction algorithm. This means that cell centers are
418 rarely on the centerline of the actual stream and that river lengths can be both over- and
419 underestimated. Underestimation typically occurs for low resolution grids, while overestimation
420 occurs for high-resolution DEMs and relatively straight rivers. Relative errors in river length have
421 been estimated to range from 5-7% for distances calculated on raster data structures, and up to >30%
422 for very coarse resolution DEMs (Paz et al., 2008). In point pattern analysis, these errors will affect
423 estimates of point intensity and interpoint distances. Hence, models developed with a particular DEM,
424 cannot be easily transferred to other DEMs without analyzing how these DEMs affect distance
425 calculations.

426 Only few functions in PPS account for the directedness of stream networks. For example, the function
427 pointdistances enables to calculate nearest neighbor distances in upstream and downstream directions.
428 Most functions, however, treat the network as undirected and thus neglect that many processes on
429 stream networks have a natural direction. Sediment and nutrient transport, for example, will follow the

430 downstream flow of water, while mobile knickpoints commonly migrate upstream. Although
431 techniques of geostatistical interpolation that account for the directional dependence of dispersal in
432 river networks exist (Garreta et al., 2010), in point pattern analysis, these approaches are rare and a
433 relatively new field of research (Rasmussen and Christensen, 2019).

434 We envision numerous other potential applications of PPS. Beyond the case studies shown, potential
435 applications include the analysis of sediment tracers, the locations of oversized boulders, wood jams, or
436 landslide dams. In addition, PPS may be applied in ecology for modelling of aquatic species based on
437 sightings, for example. Finally, once point pattern models have been trained, they can be adopted in
438 simulation tools such as the TopoToolbox Landscape Evolution Model (TTLEM) (Campforts et al.,
439 2017) to study the stochastic forcing of landslides on riverscapes in long-term landscape development.

440 Conclusions

441 PPS is a new numeric class in TopoToolbox for the analysis of point patterns on stream networks. In
442 two case studies, we analyzed geomorphic phenomena whose locations are constrained to river
443 networks. Combining explorative analysis of the locations of knickpoints with χ -analysis in the Big
444 Tujunga catchment, PPS allowed us to identify two distinct generations of knickpoints. In our analysis
445 of beaver dams, we have shown that the inhomogeneous Poisson process models implemented in PPS
446 helps to infer different geomorphological factors on beaver habitats.

447 PPS focuses on exploratory data analysis and fitting of inhomogeneous Poisson point processes, which
448 both allow studying covariates that control the spatial density of points. In addition, PPS features
449 numerous tools for simulation and visualization. Incorporation into TopoToolbox enables ease of
450 access to these new functionalities from within one computational environment. Besides the presented
451 case studies, we anticipate other applications of PPS for studying processes in fluvial geomorphology
452 and landscape evolution, but it also the distribution of aquatic and riparian species or other phenomena
453 that are constrained to occur on or alongside rivers.

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