1	UNSEEN trends: Detecting decadal changes in
2	100-year precipitation extremes
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17	Sample sizes of observed climate extremes are typically too small to reliably constrain non-
18	stationary behaviour. To facilitate detection of non-stationarities in 100-year precipitation values
19	over a short period of 35 years (1981-2015), we apply the UNprecedented Simulated Extreme
20	ENsemble (UNSEEN) approach, by pooling ensemble members and lead times from the ECMWF
21	seasonal prediction system SEAS5. We generate a 3500-year UNSEEN dataset of autumn 3-day
22	extreme precipitation events across Western Norway and Svalbard. The UNSEEN ensemble shows
23	that an event of 1.5 times the magnitude of the most severe flood episode recorded in Western
24	Norway can arise with a return period of ~2000 years. Applying the novel UNSEEN-trends
25	approach, we demonstrate that for Svalbard the 100-year event in 1981 could be expected to
26	occur with a return period of around 40 years in 2015. These new insights have important
27	implications for current design-level practices and for understanding the underlying causes of non-
28	stationarities.
29	Handling the non-stationarity of climate extremes is an active area of research ^{1–3} that is confounded
30	by the brevity and sparsity of observational records ^{4–6} . Non-stationary precipitation analyses
31	typically focus on detecting multidecadal to centennial changes in annual precipitation maxima ^{7–9} .

- 32 However, annual maximum precipitation events do not necessarily cause high impacts and hence, a
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potentially more pressing research challenge is the detection of changes in larger extremes^{10,11}, such
 as the 1-in-100-year event. Furthermore, the impacts of abrupt warming in recent decades may not
 yet be detectable in short precipitation records. Therefore, robust detection of short-term (decadal,
 rather than centennial) trends in climate extremes may provide valuable and actionable information.

37 An emerging alternative to traditional observation-based extreme value analysis is to pool ensemble members from numerical weather prediction systems^{12–22} – the UNprecedented Simulated Extreme 38 ENsemble (UNSEEN) approach^{20,22}. This technique creates numerous alternative pathways of reality, 39 40 thus increasing the event sample size for statistical analysis. The larger sample size offers a broader view of present-day hazard and, therefore, has potential to improve design-levels. For example, the 41 2013/14 winter flooding in the UK had no observational precedent, but could have been anticipated 42 with the UNSEEN approach²². Similarly, estimates of storm surge levels of the River Rhine^{12,13}, global 43 ocean wind and wave extremes^{15,16,18}, and losses from extreme windstorms¹⁹ have all been improved 44 45 with the UNSEEN approach. UNSEEN can also enhance food security through better drought exposure estimates^{14,21} and can assist policy makers and contingency planners by quantifying and 46 explaining the most severe events possible in the current climate, such as heatwaves in China²⁰. 47 48 However, validating the UNSEEN method is a well-recognised difficulty in existing studies, and 49 UNSEEN has not yet been used to facilitate detection of non-stationarity in climate extremes over 50 short periods of several decades.

51 Here, we provide a framework to systematically evaluate the robustness of the UNSEEN approach 52 and we present a novel UNSEEN-trends approach, where we aim to provide confident short-term 53 trend estimates by using the larger event sample to better constrain changes in climate extremes. We do this in a storyline context²³, where we take observed flood episodes as a starting point for our 54 analysis. We select the west coast of Norway and the Svalbard Archipelago as study regions; two 55 56 contrasting areas in terms of precipitation extremes. Western Norway faces the highest extremes within Europe²⁴ and has a dense station network^{25,26}, whereas Svalbard is a semi-desert with only a 57 few observation stations²⁷. Both regions have faced severe damages from recent extreme events, 58 such as the September 2005²⁸ and October 2014²⁹ floods over Western Norway and the slush-59 avalanche inducing extreme precipitation event over the Svalbard Archipelago in 2012³⁰. The 60 extreme events were driven by atmospheric rivers^{27–29} (ARs), which cause heavy precipitation over a 61 62 prolonged period. As AR-related floods predominantly occur in autumn and frequently strengthen over a period of several days^{28,29}, we select autumn (September to November) spatial averaged 63 64 (Supplementary fig. 1) three-day extreme precipitation (SON-3DP) as target events.

65 Previous UNSEEN studies have used the Hadley Centre global climate model, HadGEM3-GC2^{14,20-22} 66 and the European Centre for Medium-ranged Weather Forecasts (ECMWF) ensemble prediction systems^{15–18} and earlier version of the seasonal prediction system^{12,13,19}. Here, we are the first to use 67 the latest ECMWF seasonal prediction system SEAS5³¹ for its high-resolution, large ensemble, long 68 69 homogeneous hindcast period (1981-2015) and open access. The ECMWF atmospheric model has 70 shown skill in simulating atmospheric rivers for Northern Europe³², giving confidence in the realism 71 of these extreme events in SEAS5, hence is a good candidate for the UNSEEN method. We use the 25 72 ensemble members across lead times of 2-5 months, resulting in a sample of 100 members (called 73 the UNSEEN ensemble) and evaluate the independence and stability of the pooled sample for SON-74 3DP events across Western Norway and Svalbard. We then use the UNSEEN-trends approach to 75 identify unprecedented extreme precipitation events and to detect trends in 100-year precipitation 76 events over the last 35 years. These findings will help understanding the robustness of current 77 design levels and may improve our understanding of physical processes driving climate extremes and

78 their non-stationarity.

79 Ensemble member independence and model stability

The independence of ensemble members is an important requirement for the UNSEEN approach, as
dependent members would artificially inflate the sample size, without adding new information.
Previous studies have assessed the independence of ensemble members for lead times 9-10
days^{15,16,18}, but to the best of our knowledge, no independence test has yet been performed in
UNSEEN studies of seasonal prediction systems.

For the regions studied here, the ensemble members from lead times beyond one month are not
dependent on atmospheric initial conditions, because the synoptic patterns related to ARs are
known not to be predictable beyond two weeks^{32,33}. However, predictability on a seasonal timescale
may be found through slowly varying components of the ocean-atmosphere system. Therefore,
while the ensemble members might represent unique weather events because of the independency
to the atmospheric initial conditions, the weather events could have a conditional bias induced by
favourable conditions in the slowly varying components of the ocean-atmosphere system.

To test the seasonal dependence of SON-3DP, we first select the seasonal maximum event for each forecast then concatenating these events to create a 35-year timeseries (Fig. 1a,b,c). To robustly assess the independence between each of the ensemble members, we calculate the Spearman rank correlation coefficient (ρ) for every distinct pair of ensemble members (Fig. 1d), resulting in 300 ρ values for each lead time. The value of ρ ranges from ca. -0.6 to 0.6, and the median correlation is close to zero for all lead times for both Western Norway and Svalbard (Fig. 1e,f). The range in ρ

98 values is expected due to the large number of correlation tests, and none of the lead times fall 99 outside the range that would be expected for uncorrelated data for the West Coast of Norway (Fig. 100 1e). For Svalbard, slightly higher ρ values are found, with the median correlation still within the 101 expected range, but the interquartile range just exceeding the upper boundary of the confidence 102 intervals for the first two lead times (Fig. 1f). The small correlations found for Svalbard might be 103 driven by the trend that we detect for this region (UNSEEN-trends section), and thus, the UNSEEN 104 ensemble members represent unique events that follow the slowly evolving climate signal, as 105 desired.

106 A second potential issue for generating the UNSEEN ensemble could be a drift in the simulated climatology^{34,35}, which may alter precipitation extremes over longer lead times. Therefore, model 107 108 stability is a requirement for pooling lead times. Model stability is assessed by comparing the 109 distribution of predicted SON-3DP events across different lead times. For both regions, the probability density functions of the pooled SON-3DP events for the considered lead times are 110 111 remarkably similar (Fig. 2a,b). Moreover, the empirical extreme value distributions of the individual 112 lead times fall within the uncertainty range of the distribution of all lead times pooled together and 113 thus, the model can be considered stable over lead times (Fig. 2c,d).

114 Fidelity of UNSEEN extremes for Western Norway

115 Confidence in simulated `unprecedented extremes' in large ensembles is complicated by the inability 116 to validate extremes, given the limited sample sizes of observations. Here, we evaluate the UNSEEN ensemble with 1) rank histograms, commonly applied in ensemble forecast verification³⁶ and 2) by 117 bootstrapping the ensemble into datasets of 35 years and assessing whether observations fall within 118 the range of the bootstrapped distribution, following previous UNSEEN studies^{20,22} (see Methods). 119 We perform this analysis for the SEAS5 UNSEEN SON-3DP ensemble over Western Norway, because 120 the dense station network of the country^{25,26} facilitates model evaluation (unlike in Svalbard). For a 121 comprehensive global model validation of SEAS5, see Johnson et al.³¹. 122

The rank histograms clearly indicate an under-forecasting bias of the absolute SON-3DP values within the UNSEEN ensemble (Supplementary Fig. 2). This is confirmed by the bootstrapping test, that shows that the observed mean and standard deviation fall outside the 95% confidence intervals of the UNSEEN ensemble (Supplementary Fig. 3). The UNSEEN SON-3DP anomalies and standardized anomalies do show rank uniformity, and thus are suggested to be reliable (Supplementary Fig. 2). Such under-forecasting biases precipitation extremes are not uncommon in global Earth System Models³⁷, especially for a mountainous region like Western Norway. 130 As the UNSEEN SON-3DP deviations from the mean show good agreement to the observed values 131 (Supplementary Fig. 2), the ratio between the mean observed extremes and the mean simulated 132 extremes (1.74) is applied as a constant bias correction to generate the bias corrected UNSEEN 133 ensemble (henceforth referred to as UNSEEN-BC). Note that we found little sensitivity to using the 134 median (1.72), 5-year (1.69) or 20-year (1.70) values in the bias correction procedure and, hence 135 chose a constant value to avoid extrapolations beyond the quantile range. The bootstrapping test shows that the statistics derived from the observed precipitation fall within the 95% intervals of 136 137 UNSEEN-BC for timeseries of 35 years (Supplementary Fig. 4), i.e. the precipitation of the single realization of reality is one of the plausible realizations of UNSEEN-BC and, therefore, UNSEEN-BC is 138 139 indistinguishable from the observed values.

140 We then fit the GEV distribution to the observations, the UNSEEN and the UNSEEN-BC ensemble (see 141 Methods and Fig. 3). Interestingly, the fitted distributions show that the UNSEEN-BC ensemble 142 diverges from the observed values for return periods above ~35 years. To evaluate the discrepancy, 143 we test the sensitivity of the results to the choice of extreme value distribution (Supplementary Fig. 144 5). Whilst the Gumbel distribution (shape parameter $\xi = 0$) shows a relatively good fit to the 145 observations and a similar distribution to the UNSEEN ensemble, the fit is not as good as using a full 146 GEV with fitted shape parameter, as suggested by Supplementary Fig. 5 and confirmed by the 147 likelihood ratio test (p-value = 0.03 for the observed and p-value = 1.54×10^{-7} for the UNSEEN 148 ensemble). In addition, results are also very sensitive to outliers, as can be seen when the observed 149 extreme value distribution is fitted on a sample where the largest value is increased by 10% 150 (Supplementary Fig. 5). This confirms the challenge associated with estimating the magnitude of 151 events of long return periods (greater than 20 years) from an observed time series of only 35 years, 152 with more trust in estimations resulting from the larger UNSEEN sample.

153 We find that the 2005 and 2014 observed extreme events (two largest events in Fig. 3) are similar in 154 magnitude and represent events with return periods of 21 years (Cl of 19-24 years) when compared 155 with the extreme value distribution of UNSEEN-BC. Based on the observed values, the return period 156 estimate of 60 years for the events would be very uncertain, with the lower confidence interval 157 never reaching the event magnitude (Cl of 18 - ∞ years). Moreover, the highest UNSEEN-BC event is 158 1.5 times higher than the highest observed event, with an estimated return period of ~2000 years (CI 159 of 1150-4800 years). The estimated return period of this event based on the observations is completely dominated by the uncertainties (~5000, 600 - ∞ years) and can only be statistically 160 161 modelled, while for the UNSEEN estimate, it is a physically simulated `empirical' event within 3500

162 years of data. The observed flood episode caused flooding and landslides with severe damage²⁸ and

163 UNSEEN-BC indicates what kind of events beyond the observed record are plausible in the present164 climate.

165 UNSEEN-trends in 100-year precipitation over last 35 years

Climate models can be used to detect changes^{38–41} or to attribute extreme events to human causes⁴², 166 167 but are less suited to detecting trends over the recent past such as the last 35 years. By design, 168 climate model simulations are initialized once at the beginning of a centennial run. Contrastingly, 169 here we use seasonal forecasts that are initialized every month, and thus are more constrained by 170 real-world climate variability than climate model simulations. Consequently, seasonal forecasts 171 sample a smaller range of climate conditions but are closer to reality than climate model 172 simulations. This means that their use is consistent with analysing trends over the recent past 173 described by the available forecast period (for SEAS5, currently 35 years). Furthermore, the model 174 setup and version are the same for the entire hindcast simulation, ensuring that, with respect to the 175 models and initialization, SEAS5 is a homogeneous dataset and thus suitable for climate analysis and 176 detection of UNSEEN-trends.

With a 36 km resolution and 25 members, the ECMWF SEAS5 reforecast set used here is based on a
modelling system of high resolution and associated with a large ensemble compared to current highresolution global climate models⁴³. SEAS5 greenhouse gas radiative forcing captures the long-term
trends in emissions³¹, and we show that the global mean temperature trend in SEAS5 follows ERA5⁴⁴
(Supplementary Fig.6). Whilst regionally, we find a cold bias over the Norwegian study domain, the
trend is consistent with ERA5 for both Western Norway and Svalbard (Supplementary Fig. 6),
confirming the capacity of SEAS5 to detect recent trends.

184 To illustrate the added value of UNSEEN-trends, we extend the GEV distribution to include a time 185 covariate and fit this distribution to the observed and UNSEEN SON-3DP (see Methods). Using the 186 observations, we find an increase in 100-year SON-3DP of 4% over 1981-2015 in Western Norway, 187 but associated with large uncertainties ranging from -27% to 34% (Fig. 4 a,b). The UNSEEN-trend 188 estimate of 2% is more constrained due the larger sample size, with confidence intervals ranging 189 from -3% to 7%. A negative trend is thus statistically possible, indicating that the trend over Western 190 Norway is not significant. For Svalbard, we find a significant positive UNSEEN-trend of 8%, with 191 uncertainty bounds ranging between 4-12%.

In addition to the trend in 100-year SON-3DP events, we illustrate the change in all return values by plotting the GEV distribution with the covariates 1981 and 2015 (Fig. 4 c,d). The likelihood ratio test shows that the GEV distribution including a time covariate improves the model fit for Svalbard (p-

- value = 2.7e-07). We find that the frequency of the event that used to be a 100-year event in 1981
- has an expected return period of 41 years in 2015 (Fig. 4 c,d). For Western Norway, the GEV
- distribution including a time covariate does not improve the model fit for either the observed (p-
- value = 0.58) or the UNSEEN-ensemble (p-value = 0.65), and thus, the stationary GEV distribution, as
- 199 presented in Fig. 3, is most appropriate.

200 Discussion and Conclusion

- In this study, we test the robustness of the UNSEEN approach and we use the large sample to
 constrain short-term UNSEEN-trends in high-impact precipitation events for Western Norway and
 Svalbard. We show that with SEAS5, the effective sample size of autumn 3-day precipitation (SON3DP) events in Western Norway and Svalbard can be increased by a factor of 100 compared to
 observations, because ensemble members are independent and the model is stable over lead times.
 Validating UNSEEN events and trends is a complex task, but our approach reproduces observed
 extremes well after bias correction for Western Norway, a region with extensive records²⁶.
- 208 The insights presented in this study are specific to Western Norway and Svalbard SON-3DP but the 209 independence, model stability and model fidelity tests applied to the UNSEEN approach could be 210 transferred to other regions, temporal resolutions and spatial extent of the events, seasons and 211 climate variables. Global validation of the UNSEEN ensemble will highlight in which regions the 212 approach may enhance the robustness of design level estimation, with a potentially high value in 213 supporting data scarce regions⁴⁵. Furthermore, the large sample size may allow estimation of extremes using empirical approaches that avoid assumptions about underlying distributions and 214 their non-stationarity, thereby offering the possibility of improved design estimates¹⁰ and empirical 215 216 attribution of physical mechanisms. A wide range of scientific disciplines might benefit from the 217 UNSEEN method by forcing seasonal prediction systems into impact models to assess 218 unprecedented impacts and improve understanding of the physical mechanisms leading to these 219 events.
- 220 The results from the two study areas highlight the value of both the UNSEEN and the UNSEEN-trends 221 approach. For the well-monitored Norwegian domain, we are able to bias correct the UNSEEN 222 ensemble (UNSEEN-BC) and therefore we can better estimate the return period of the 2005 and 223 2014 flood episodes. We find that the flood episodes are not rare exceptions; rather they might be expected to occur once in 20 years under a stationary climate. Furthermore, the UNSEEN-BC 224 225 ensemble shows that an event of 1.5 times the magnitude of the highest observed event could arise. 226 The September 2005 and October 2014 flood episodes were identified as high-impact events in 227 previous end-user engagement sessions within the Translating Weather Extremes into the Future

(TWEX) project, and thus, the results found from the UNSEEN-BC ensemble are of high relevance to
decision makers and end-users. This application of the UNSEEN approach is similar to previous
research on the 2013/14 winter floods in the UK²² and for the 1990 windstorm losses over Germany
and the UK¹⁹. A difference to the previous studies is that we run the analysis on a three-day
resolution, whereas monthly averages have been used so far. The observed record and the UNSEENtrend show that there is no significant trend over Western Norway between 1981-2015, and
therefore justify using the stationary GEV distribution.

235 Contrastingly, for Svalbard, the UNSEEN-trends approach shows that what was a 100-year event in 236 1981 is to be expected to return once in 41 years in 2015. The trend in extreme precipitation over Svalbard could not be detected from observation-based studies due to the sparse observation 237 238 network in this area²⁷. Despite very few precipitation extremes being recorded in the Svalbard 239 Archipelago, it is assumed that their frequency and magnitude are increasing in a warming 240 climate^{27,30,46}, which is confirmed by our UNSEEN-trends analysis. Those precipitation extremes are 241 connected to the inflow of relatively warm air and, thus, can cause severe landslides and so-called rain-on-ice events³⁰. Both could have significant impacts on people living in the Arctic and on the 242 243 local ecosystem.

244 In due course, the drivers of changes in climate extremes could be investigated with the UNSEEN-245 trends approach. For example, to assess the non-stationarity of extreme precipitation, covariates 246 other than time could be selected, such as ocean temperatures, modes of climate variability, or 247 indicators of large-scale synoptic weather systems. This may improve our physical understanding of 248 the non-stationary processes and could provide insight into potential model biases, thereby 249 improving confidence in detected trends. Century-long seasonal hindcasts, such as the ASF-20C global atmospheric seasonal hindcasts⁴⁷, might prove useful in assessing the sensitivity of UNSEEN-250 251 trends to different time windows over a longer time period.

252 Our results for Western Norway highlight the strength of UNSEEN in estimating design-levels and present-day climate hazards, backed by a growing body of literature^{12,13,22,14–21}, and the results for 253 254 Svalbard emphasise the significance of our novel UNSEEN-trends approach in estimating non-255 stationarities in climate extremes. Both underline the need to rethink current design-level estimates 256 based upon observations alone. We think further applications can 1) help estimate design values, 257 especially relevant for data scarce regions; 2) improve risk estimation of natural hazards by coupling 258 UNSEEN to impact models; 3) detect trends in rare climate extremes, including variables other than 259 precipitation; and 4) increase our physical understanding of the drivers of non-stationary climate 260 extremes, through the possible attribution of detected trends.

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411		
	Mathada	

412 Methods

- 413 **Data.** We use the fifth generation of the ECMWF seasonal forecasting system SEAS5 to generate the
- 414 UNSEEN ensemble. SEAS5 is a global coupled ocean, sea-ice, and atmosphere model, which has been
- 415 introduced in fall 2017³¹. The atmospheric component is based on cycle 43r1 of the ECMWF
- 416 Integrated Forecast System. The spatial horizontal resolution is 36 km and it has 91 vertical levels.
- 417 The ocean (Nucleus for European Modelling of the Ocean, NEMO⁴⁸) and sea-ice (Louvain-la-Neuve
- 418 Sea Ice Model, LIM2⁴⁹) models run on a 0.25-degree resolution. The atmosphere is initialized by ERA-
- 419 Interim⁵⁰ and the ocean and sea-ice components are initialized by the OCEAN5 reanalysis⁵¹. ECMWF
- 420 provides a re-forecast (also known as hindcast) dataset for calibration of the operational forecasting
- 421 system SEAS5. The data are initialized monthly with 25 ensemble members, each with 7-month

forecast length on a daily resolution, covering the years 1981-2016³¹. The ensemble members are
 generated from perturbations to the ocean and atmosphere initial conditions and from stochastic
 model perturbations.

425

426 In the UNSEEN approach, ensemble members and initialization dates are pooled to increase the 427 sample size of the variable of interest. Here, we generate an UNSEEN ensemble for the west coast of 428 Norway and for the Svalbard Archipelago to focus on recent atmospheric river (AR) related severe 429 events^{27–29}. ARs have been connected to precipitation extremes in the observed records for both Norway^{52,53} and Svalbard²⁷ and occur in September to March. AR-related floods mostly occur in 430 431 autumn, because snowfall during winter precipitation events results in storage rather than runoff. One-day and five-day precipitation are a common diagnostic for extreme analysis^{6,54}. ARs frequently 432 433 strengthen over a period of several days^{28,29} and therefore multi-day diagnostics prevent splitting events. Following the 2014 flood episode²⁹, we have chosen three-day total precipitation in this 434 435 study. We thus select autumn (September to November) 3-day extreme precipitation (SON-3DP) as 436 target events.

437

438 Since the forecasts are initialized every month on the first of the month and run over 7-months 439 length, there are five initialization months (May-September) available to forecast the entire target 440 autumn season (September-November). The first month is removed to avoid potentially dependent 441 events. In the end, 100 forecasts, based on 25 ensemble members with 4 initialization dates are 442 used to forecast the autumn season of each year (Fig. 1a-c). The window of 35 years between 1981 443 and 2016 leads to a total of 3500 forecasts of autumn weather conditions that could have occurred. 444 We extract the maximum 3-day cumulative precipitation within autumn from the 3500 forecasts (SON-3DP), using the xarray package⁵⁵ in Python. To focus on the large-scale systems as experienced 445 446 in recent severe events, we use only the large-scale precipitation output of the model. The west 447 coast of Norway is mountainous and characterised by large topographic variations. Catchment-scale 448 processes in these mountainous areas cannot be resolved by a global model with 36 km resolution. 449 Therefore, the precipitation timeseries presented in this study are spatial averages where the 200-450 year precipitation exceeds 90 mm for the west coast of Norway (4-7° E, 58-63° N) and 35 mm for 451 Svalbard (8-30° E, 76-80° N) (Supplementary Fig. 1).

452

To evaluate the precipitation extremes simulated by SEAS5, we use a 1x1 km gridded station-based precipitation product for Norway²⁵. The data have recently been corrected for underestimation caused by wind-induced under catch and uses more information in the interpolation scheme for

data-scarce areas, resulting in higher precipitation in data scarce areas²⁶. We upscale this gridded
dataset to the same resolution as SEAS5 and extract SON-3DP values for the same spatial domain
over 1981-2016. Note, for the Svalbard Archipelago no gridded precipitation dataset is available as a
reference dataset. We use ERA5⁴⁴ for the global and regional temperature evaluation of SEAS5.

461 **Ensemble member independence testing.** The method for independence testing applied in this 462 study is inspired by previous research on potential predictability: the ability of the model to predict itself^{32,36}. The potential predictability of a model is calculated by using one of the forecast ensemble 463 464 members as the observations and the mean of the other ensemble members as the forecast. The 465 correlation between the 'observed' ensemble member and the mean of the other ensemble 466 members is calculated for every ensemble member and this range gives an estimate of the ability of 467 the model to forecast itself. Because this method assesses the correlation between ensemble 468 members, it can be used to find the degree of ensemble members' dependence. In seasonal 469 forecasting, this method is used to identify any predictability in the seasonal prediction system. In 470 contrast, here we seek to demonstrate that there is no potential predictability in the system for the 471 ensemble members to represent independent, unique events.

An illustration of our method to test for independence is shown in Fig. 1. A potential predictability
test is performed but instead of correlating an ensemble member to the mean of the other
ensemble members, a pairwise correlation test is applied between all ensemble members to
robustly assess the individual ensemble member dependence. Indeed, we concatenate the seasons
together member by member, even though they do not necessarily originate from the same run.
This approach was chosen because the underlying initialization method remains the same for each
member over different seasons.

479 For the 25 ensemble members, there are 300 distinct pairings in the correlation matrix for each of 480 the four lead times being analysed (may-August). We calculate the spearman ρ statistics on the 481 standardized SON-3DP anomalies (deviation from mean divided by the standard deviation) for each 482 distinct pair. From the 300 ρ values for each lead time, boxplot statistics are calculated: the 483 whiskers, the interquartile range and the median. When testing for significance of the 300 ρ values, 484 care must be taken not to falsely detect significant correlations because of the large number of tests. 485 For example, with a confidence interval of 5%, 15 out of the 300 correlations would be expected to 486 be significant by chance alone. To avoid these problems, a permutation test is performed. The 487 dataset, which previously consisted of 25 timeseries (members) of 35 datapoints (years) for four 488 initializations months (lead times), is resampled into 100 timeseries of 35 datapoints, with

datapoints randomly picked from all members, years and lead times to remove potential
correlations. This randomized dataset is split into four pseudo lead times of 25 timeseries, in order
to calculate the boxplot statistics from the same amount of correlation coefficients (300) as before.
The data are resampled 1000 times (without replacement), resulting in 4000 boxplot statistics (4

- 493 pseudo lead times * 1000 resampled series), from which the confidence intervals are calculated
- 494 based on a 5% significance level (the 2.5 and 97.5 percentiles).
- 495

496 Model stability. The extreme precipitation distribution must be similar over lead times in order to 497 generate the UNSEEN ensemble. We use four initialization months (May-August) forecasting the 498 target autumn season with lead times 2-5 months. For each lead time, 25 ensemble members over 499 35 years result into an 875-year long dataset and the pooled ensemble into 3500 years. To compare 500 the distributions, we first plot the probability density function for each of the lead times using 501 ggplot2⁵⁶. Secondly, we plot the extreme value distributions, focussing more on the tails of the 502 distribution. We calculate empirical quantiles of the extreme precipitation ensemble without 503 assuming any distribution a priori, to avoid problems regarding statistical modelling of the 504 extremes^{10,57}. The quantile (Q) of a distribution is the inverse of the distribution function (F(x)):

505
$$Q(p) = F^{-1}(p) = \inf \{x: F(x) \ge p\}, \quad 0 (1)$$

506 Where the return value is associated with the quantile of percentile (*p*):

507
$$p = 1 - \frac{1}{T}$$
 (2)

508 With T being the return period. We use the quantile function in R⁵⁸ to compute the empirical return 509 values and we refer to Hyndman & Fan⁵⁹ for more specifics.

510

511 Fidelity of the UNSEEN ensemble for Western Norway. We first evaluate the UNSEEN ensemble and 512 then compare UNSEEN design-levels to observation-based design-levels. As a first assessment of the 513 biases within the SON-3DP UNSEEN ensemble, we use rank histograms. Rank histograms indicate 514 over-dispersion or under-dispersion and over-forecasting or under-forecasting bias³⁶. Here, we have 515 100 members (4 lead times and 25 ensemble members) for each year over 1981-2015. The rank of 516 the observations within the 100 ensembles is calculated for each year and the resulting 35 ranks are 517 plotted as a histogram over the range 1-100. If the observations are mostly in the upper (lower) 518 ranks, this indicates that the observed values are higher (lower) than the forecasted values and 519 therefore the forecasts are under-forecasting (over-forecasting). Similarly, when the observations

520 are mostly in the outer (inner) ranks, this indicates that the observed values show more (less)

521 variability and thus the forecasts are under-dispersed (over-dispersed). We create rank histograms

522 for the raw SON-3DP UNSEEN ensemble, for the anomalies from the mean and for the standardized

anomalies, where the anomalies are divided by the standard deviation.

To compare UNSEEN to the observed record in more detail, we apply a bootstrap test presented in previous studies^{20,22}. We bootstrap 10,000 timeseries of 35 years with replacement from all ensembles (100 x 35 years) and calculate the mean, standard deviation, skewness and kurtosis for each. We test whether the four distribution statistics derived from the observed precipitation time series over the period 1981-2015 fall within the 95% confidence intervals for the statistics derived from the bootstrapped timeseries.

530 We then fit the Generalized Extreme Value (GEV) distribution, described by a location ($-\infty < \mu < 531 \quad \infty$), scale ($\sigma > 0$) and shape ($-\infty < \xi < \infty$) parameter⁶⁰:

532
$$F(x) = \exp\left[-\left(1 + \xi\left(\frac{x-\mu}{\sigma}\right)\right)^{-\frac{1}{\xi}}\right], \qquad \left(1 + \xi\left(\frac{x-\mu}{\sigma}\right)\right) > 0 \tag{3}$$

533 And we test the sensitivity to using the Gumbel distribution with $\xi = 0$, simplifying the distribution 534 to:

535
$$F(x) = \exp\left[-\exp\left(-\left(\frac{x-\mu}{\sigma}\right)\right)\right], \qquad -\infty < x < \infty$$
(4)

536 The quantiles of the distribution can again be obtained by inverting the distribution:

537
$$x_{p} = \begin{cases} \mu - \frac{\sigma}{\xi} [1 - \{-\log(1-p)\}^{-\xi}], & for \xi \neq 0\\ \mu - \sigma \log\{-\log(1-p)\}, & for \xi = 0 \end{cases}$$
(5)

538 Where the return value x_p corresponds to the return period 1/probability (p). For all statistical 539 model fits in this study (including non-stationary fits described in the next section), we apply 540 Maximum Likelihood Estimation (MLE) to estimate the parameters of the distributions, utilizing the 541 extRemes package⁶¹ in R⁵⁸. The 95% confidence intervals of the distributions are calculated based on 542 the normal approximation, which is the default of the extRemes package.

543

544 UNSEEN-trends. In this study, we present the idea of performing trend analysis on seasonal

545 hindcast, as the seasonal hindcasts provide a larger sample than observations and a higher

- resolution than climate models (see the UNSEEN-trends section for more details). We apply well-
- 547 established extreme value theory^{60,62,63}, by allowing the location (μ) and scale (σ) parameters of the

548 GEV distribution (given in equation 3) to vary linearly with time (*t*). Because the scale parameter 549 needs to be positive, a log-link function is used:

550
$$\mu(t) = \mu_0 + \mu_1 t$$
 (6)

$$\ln \sigma \left(t \right) = \phi_0 + \phi_1 t \tag{7}$$

552

551

This approach selects one block maximum per year, leading to 35 data points over the years 1981-553 554 2015 based on observed records. With UNSEEN-trends, we have 100 times more values for each 555 year and thus increase confidence in the regression analysis (see Fig.4a,b for illustration). As for the 556 stationary method, we use MLE to estimate the parameters of the distributions and the normal 557 approximation to find the 95% confidence intervals of return values. We focus on the changes in the 558 100-year quantiles, because these are associated with the design-levels mostly used in flood 559 defence⁶⁴. The trend in the 100-year return value is defined as the percentual change between 1981 560 and 2015:

$$\Delta x_T = 100 * \left(\frac{x_T(\mu_{2015}, \ln \sigma_{2015}, \xi) - x_T(\mu_{1981}, \ln \sigma_{1981}, \xi)}{x_T(\mu_{1981}, \ln \sigma_{1981}, \xi)} \right)$$

562 Where x_T is defined by equation (5).

563 The robustness of the trends to experiment decisions like the block size and the regression method 564 can be further investigated but are beyond the scope of this research. For example, 6-month blocks can be selected at the expense of the ensemble size. This will result in 25 realizations, in comparison 565 566 with 3-month blocks, which contain 100 realizations. A block size of three months (September-567 November) is chosen in this study. A linear trend in time is assumed in this study. With the large 568 amount of data, more complex regression methods can be explored. The ECMWF SEAS5 seasonal 569 prediction system is used in this study, but other seasonal prediction systems with available 570 hindcasts could also be assessed to test the model sensitivity to return value and trend estimation.

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576 Data and code availability

- 577 SEAS5 re-forecast data and ERA-5 data are openly available on the Copernicus Climate Change
- 578 Service (C3S) Climate DataStore (<u>https://cds.climate.copernicus.eu/</u>) and the SeNorge daily total
- 579 precipitation data are available at DOI: <u>https://doi.org/10.5281/zenodo.2082320</u>. The extracted
- 580 SON-3DP UNSEEN ensembles as well as the extracted SON-3DP observations will be made available
- on GitHub, along with all code to reproduce the analysis in this paper. Code is available to reviewers
- 582 on request.

583 Author contributions

- 584 T.K. and M.M. conceived and T.K, M.M, L.J.S., T.M., R.L.W., C.P., P.B. designed the study. T.K. drafted
- the paper with extensive contributions from L.J.S., T.M., R.L.W., C.P. and M.M.. T.K. analysed the
- 586 data with input from all authors. M.M. acquired the data. T.K. produced the figures; L.F. produced
- 587 Supplementary Fig. 6.

588 Competing interests

589 The authors declare no competing interests.

591 Figures





593 Fig. 1 | A workflow for analysing ensemble member dependence. a, August 2014 initialized 25-594 member seasonal forecasts of 3-day precipitation time series over the SON forecast horizon. Ensemble members 0 and 1 are shown in blue and orange, respectively. b, From the forecast 595 596 members 0 and 1, the September-November (SON) maximum value for the 2014 season is selected. 597 c, A series of the maximum 3-day precipitation values for the SON season for each year in the 598 hindcast record is created for member 0 and member 1. The 2014 maximum, as illustrated in b, is 599 encircled. d, The standardized anomaly of the maximum 3-day precipitation series for the two 600 members are correlated. Spearman's rho correlation is shown. This process is repeated for the 300 distinct ensemble member pairings for each of the four lead times (May-August). e,f, Boxplots of the 601 602 resulting 300 Spearman's rho correlations for each lead time over Norway (e) and Svalbard (f). Grey 603 shading shows the confidence intervals of the boxplot statistics (whiskers, interquartile range and 604 median), based on a permutation test with 5% significance level.

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- 609





Fig. 2 | SEAS5 model stability of extreme precipitation over Western Norway and Svalbard. The empirical probability density (**a**,**b**) and extreme value (**c**,**d**) distribution of SON-3DP for each lead time and for all lead times together (in black), for the West Coast (**a**,**c**) and Svalbard (**b**,**d**) domains. Grey shading in **c**,**d**, illustrates the 95% confidence intervals of the distribution of the pooled lead times, bootstrapped to timeseries of similar length to the individual lead times with n = 10,000.





Fig. 3 | The extreme precipitation distribution for UNSEEN and UNSEEN-BC, as compared to the

precipitation record over Western Norway. The data points show the SON-3DP events and the solid 620 lines show the GEV fitted to the data, including 95% confidence intervals.





