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Space-time susceptibility modeling of hydro-morphological processes at the Chinese national scale

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Abstract

Hydro-morphological processes (HMP; any process in the spectrum between debris flows and flash floods) threaten human infrastructures and lives; and their effects are only expected to worsen in the context of climate change. One of the ways to limit the potential damage of HMP is to take preventive or remedial actions probabilistically knowing where and how frequently they may occur. The expected information on where and how frequently a given earth surface process may manifest itself is referred to as susceptibility. And, for the whole Chinese territory, a susceptibility model for HMP is currently not available.

To address this issue, we propose a yearly space-time model consisting of a Generalized Linear Model of the binomial family. The target variable of such model is the annual presence/absence information of HMP per catchment across China, from 1985 to 2015. This information has been accessed via the Chinese catalogue of HMP, a data repository the Chinese government has activated in 195X and which is still currently in use. This binary spatio-temporal information is regressed against a set of time-invariant (catchment shape indices and terrain attributes) and time-variant (urban coverage, rainfall, vegetation density and land use) covariates. Furthermore, we include a regression constant for each of the 31 years under consideration and also a three-years aggregated information on previously occurred (and not-occurred) HMP. We consider two versions of our modeling approach, an explanatory benchmark where we fit the whole space-time HMP data, including a multiple intercept per year. Furthermore, we also extend this explanatory model into a predictive

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21 one, by considering four temporal cross-validation schemes (Forward-All, Forward-Sequence,
22 Backward-All, and Backward-Sequence), removing the yearly multiple intercept. In the first
23 of 31 temporal replicates, Forward-All is calibrated for 1985 and then used to predict from
24 1986 to 2015. In the second step, a model is calibrated for 1985 and 1986 combined and
25 used to validate the rest of the space-time series. This is replicated up to the last model
26 where the combined data from 1985 to 2014 is calibrated to predict the last year of the
27 HMP presence/absence data. Forward-Sequence also moves in the same temporal direction
28 but the sampling scheme sequentially extracts two years at a time, one for calibration and
29 one for validation. For instance, the first step is trained for 1985 and used to predict 1986;
30 then the second step is trained for 1986 and used to predict 1987. As for Backward-All, and
31 Backward-Sequence, their structure is the same but the temporal direction goes from 2015
32 to 1985.

33 Our explanatory model suggests that the overall number of HMP events per year has
34 increased in the last decade and that the annual susceptibility has subsequently followed
35 the same trend. As for the four cross-validation routines, Forward-Sequence shows excellent
36 performance with an average AUC of 0.83, slightly better than Forward-All, Backward-
37 All, and Backward-Sequence. From an interpretative standpoint, this implies that the best
38 spatio-temporal prediction we obtained is associated with short-term variations of the HMP
39 distribution and that such variations should be considered in a forward temporal direction.

40 Furthermore, we portrayed the annual susceptibility models into 30 maps, where the
41 south-east of China is shown to exhibit the largest variation in the spatio-temporal proba-
42 bility of HMP occurrence. Also, we compressed the whole spatio-temporal prediction into
43 three summary maps. These report the mean, maximum and 95% confidence interval of the
44 spatio-temporal susceptibility distribution per catchment, per year.

45 The information we present has a dual value. On the one hand, we provide a platform
46 to interpret environmental effects on HMP at a very large scale, both spatially (the whole
47 Chinese country) and temporally (31 years of records). On the other hand, we provide
48 information on which catchments are more prone to experience a HMP-driven disaster.
49 Hence, a step further would be to select the more susceptible catchment for detailed analysis
50 where physically-based models could be tested to estimate the potentially impacted areas.

51 **Keywords:** Hydro-morphological processes; Historical hazard archives; Susceptibility; Spa-
52 tiotemporal predictive models.

53

1 Introduction

In this work, the term hydro-morphological process (HMP) was used to address a class of earth surface phenomena where solid and fluid phases of a gravitationally-driven moving mass are not well determined. Thus, in this class we refer to a broad spectrum of processes in between debris flows and flash floods. The reasons behind such initial disclaimer are due to the nature of the dataset we used and further explanations will be provided later in the text.

This class of HMPs includes some of the most frequent and damaging natural disasters, and their occurrence shows a close relationship with climatic changes (Blöschl *et al.*, 2020; Prein *et al.*, 2017; Westra *et al.*, 2014). Because of this, HMPs have increasingly been reported to threaten human lives and infrastructure in recent years (e.g., Špitalar *et al.*, 2014). To prevent or limit the losses, it is crucial to estimate where and when these processes may occur. In turn, this enables administrations to plan ahead and mitigate future risks (Lombardo *et al.*, 2020; Rossi *et al.*, 2019).

HMPs are extremely rapid phenomena. Just few hours are needed between the triggering heavy rain and their manifestation (Borga *et al.*, 2007; Marchi *et al.*, 2010). They can also be generated by snow-melt but it is generally the intensity and duration of precipitation that control the process through the water discharged over a given area. Then, the overland flows follow the river network, entrains all sorts of debris and leaves it strewn especially when the runoff intersects urban areas (Norbiato *et al.*, 2008). In such cases, roads may be blocked, drainage systems clogged, cars trapped, lives lost and property destroyed (Karagiorgos *et al.*, 2016; Mahmood *et al.*, 2017). For this reason, HMP prediction models are primarily implemented in a physically-based framework where one can reliably introduce the rainfall input and simulate the process by accounting for topography and soil hydrological characteristics (Tramblay *et al.*, 2010). This is usually performed specifically for small areas (Rozalis *et al.*, 2010) but recent advancement have led to develop similar applications on much wider regions, simulating different types of HMPs from catchment (Javelle *et al.*, 2010; Bout *et al.*, 2018) to country-wise scales (Gourley *et al.*, 2017), and even up to continental scales (Paprotny *et al.*, 2017). These different levels of details all share a common structure where a design storm is used as the input. The design storm can be either inferred from long time-series of rainfall data via extreme value statistics (Li *et al.*, 2019a). Or, it can be directly plugged in by using near-real-time rainfall data obtained from meteorological forecasts (Collier, 2007). As for the remaining information, terrain characteristics are commonly derived from global DEM data (Adnan *et al.*, 2019) or from site-specific LiDAR surveys (Crema *et al.*, 2018). Besides, soil parameters are required to describe the hydrological characteristics and the associated ability to retain water or to convert it into runoff (Norbiato *et al.*, 2008). This can be obtained via in-situ tests whenever the area is relatively small (Cenci *et al.*, 2016) and from global estimates such as ISRIC, for large scale assessments (Ragettli *et al.*, 2017). These methods have the inherited ability to produce HMP runout estimates, such as total impacted spatial extent, flow heights, kinetic energy, volumes and more, which are crucial

94 information for engineering design and master plans (Li *et al.*, 2019b). However, the appli-
95 cability of physically-based models inevitably suffers from considerable limitations whenever
96 the study target involves continental to global scales (Bout *et al.*, 2018; Glade and Crozier,
97 2005), with very few exceptions to this rule (Liao *et al.*, 2012). In fact, for large areas, the
98 required input information is typically quite smooth, assuming it is even accessible. And,
99 collecting suitable geotechnical data is difficult if not impossible (Gaume *et al.*, 2009) over
100 large regions. As a result, a complementary branch in the natural hazard community has
101 developed statistically-based models during the last decades. This methods do not offer the
102 same breath of results produced from the physically-based counterpart (e.g., they do not
103 spatially predict runout-impacted areas nor flow-heights, etc.). However, they provide useful
104 information on areas potentially subjected to HMPs, learning from past events from which
105 spatio-temporal projections are made (Gourley *et al.*, 2013).

106 The present work fits in the second category. Specifically, the Chinese government has
107 recently completed a long lasting project where all the available information on historical
108 HMPs has been collated for the whole Chinese territory. We use the term HMP specifically
109 because the Chinese catalogue reports a wide spectrum of earth surface processes without
110 explicitly attributing a class. This catalogue starts from reports gathered even from ancient
111 China and it covers the period until 2015. Because of this wide temporal coverage, the
112 data differs in quality across space and time and the Chinese government has decided to use
113 a more general classification, consistent through time. More specifically, the data collated
114 until 1949 is relatively poor and the situation improves substantially from 1950 onward as
115 the current Chinese government was established. Nevertheless, even from 1950 up to 1980,
116 the data may still have some positional issues because the digital system did not exist (Li
117 *et al.*, 2018). The Chinese HMP report system became standardized after the 1980ies, with
118 more available technologies being used to record the location (latitude and longitude), date
119 and time as well as the losses expressed either in the number of victims or economical value
120 (Guo *et al.*, 2018). In light of this considerations, we subset the Chinese HMP catalogue
121 extracting all the available information from 1985 to 2015. We note here that since 1985
122 we also have access to meteorological digital data collected and aggregated daily from the
123 Chinese rain gauge network.

124 We use this data to build a space-time HMP susceptibility model. A susceptibility
125 model essentially estimates the probability of occurrence of a given natural process within
126 a specific mapping unit and temporal unit. Mapping units constitute the spatial structure
127 under which a given study area is subdivided. They can consist of a regular lattice (usually
128 grid-cells or rarely hexagons) or they can represent geographic features such as catchments
129 or administrative units (Carrara *et al.*, 1991, 1995; Reichenbach *et al.*, 2018; Lombardo
130 *et al.*, 2019). Irrespective of the specific geometry, a mapping unit represent the object upon
131 which a statistical model estimates the probability of occurrence of the target hazard. As
132 for temporal units, this represent the time span upon which the selected model makes a
133 prediction. For physically-based models, this is typically expressed in hours or days whereas

134 for statistical models this may involve a much larger time span. In this work we opted for
135 a catchment partition, having accessed the most updated watershed delineation of China
136 (Shen *et al.*, 2017). As for the temporal partition, we selected an annual unit of time.

137 As for the method, we chose a binomial Generalized Linear Model (GLM) assuming
138 that the spatio-temporal population of HMPs across China behaves according to a Bernoulli
139 probability distribution. This procedure is quite common and actually represents the most
140 common practice in the geomorphological literature (e.g., Budimir *et al.*, 2015; Lombardo
141 *et al.*, 2015; Reichenbach *et al.*, 2018).

142 We stress here that the susceptibility to any surface process is not stationary or time-
143 invariant (Lombardo *et al.*, 2020). It actually varies through time as the environmental
144 conditions change. For instance, landscape evolution processes may modify the terrain,
145 hence changing the hydrology of a given area. Similarly, settlement growth and urbanization
146 experienced a dynamic expansion and the urbanization itself has become denser through
147 time, especially in China. This may have changed the distribution of permeable surfaces
148 in favor of concreted and sealed land covers (see, Gong *et al.*, 2019). Also, climate changes
149 may contribute to vary the HMP triggering conditions through space and time, especially
150 because rainfall regimes have become less diluted during wet seasons and they have become
151 more concentrated in narrow time windows. All these contributing/triggering factors can be
152 accounted for in statistical models. For instance, if climate change and accelerated urbaniza-
153 tion control the HMP occurrence distribution, then a space-time statistical model should be
154 able to capture their influences and show a potential increase in HMP occurrences in recent
155 years.

156 The present manuscript is organized as follows: Section 2 introduces the study area
157 and Section 3 describes the material and methodology framework used in susceptibility
158 modelling. This is followed by a detailed description of the model performance and the
159 resulting susceptibility maps in Section 4. Finally, Section 5 discusses the supporting and
160 opposing arguments on this work. And Section 6 summarizes our final remarks.

161 2 Study Area

162 China approximately covers the area between latitudes 18° and 54° N, and longitudes 73°
163 and 135° E. It is characterized by a vast territory and a complex landscape. Based on ge-
164 omorphological characteristics, China can be divided into six homogeneous regions (Wang
165 *et al.*, 2020): eastern plains, southeastern hills, southwestern mountains, north-central plains,
166 northwestern basins and Tibetan Plateau. About two-thirds of China is covered by moun-
167 tainous areas (Liu *et al.*, 2018). The southern China consists of hilly and mountainous
168 terrains, while the western and northern China is dominated by plains and basins. The an-
169 nual rainfall records are strongly controlled by the distance to the coastline and precipitation
170 amounts gradually decrease from the southeast to northwest of China. The eastern plains
171 and southern coasts are severely influenced by the East Asian Summer Monsoon, where most

172 of China’s agricultural land and settlements are located. In this context, only the northwest
173 China has a predominantly arid climate and a lower population density.

174 **3 Material and Methods**

175 **3.1 Hydro-morphological processes in China**

176 As previously introduced, the Chinese catalogue of HMPs is a digital collection of events,
177 describing a spectrum of phenomena where a fast moving mass – consisting of a ill-defined
178 proportion of solid and fluid – propagates across the landscape, potentially causing destruc-
179 tion in its path. As a result, the above mentioned spectrum encompasses processes from
180 debris flows (where the solid and liquid phases are almost equally represented) to flash
181 floods (where the fluid phase is much larger than the solid one). Each HMP record in the
182 database contains information on geographic coordinates, date and time as well as (but not
183 always) two loss estimates, expressed as victims and costs.

184 Because of this rich information, it would be theoretically possible to extract HMPs that
185 have actually resulted in a disaster (i.e., life losses > 0 OR economical losses > 0). However,
186 not all the HMPs contain the loss information. For this reason, instead of modeling a subset
187 of the whole database, we opted for the entirety of the available information, including
188 “innocuous” and disastrous HMPs. This information is geographically summarised in Figure
189 1 where we highlight the spatio-temporal distribution of HMPs upon which we have built
190 our modeling routine.

191 Overall, the Chinese database reports 24,956 HMPs in the time span of 31 years (1985-
192 2015) with a substantially varying concentration across space and time, with the exception
193 of the western arid to semi-arid sector where essentially no events have been recorded.

194 **3.2 Mapping unit**

195 The nature of the Chinese HMP catalogue implies that the various processes included may
196 act on different spatial scales. For instance, debris flows usually have a more limited spatial
197 extent, thus slope- to catchment- based models are the most suitable to represent the physical
198 expression of these phenomena. Conversely, on the other side of the spectrum, flash floods
199 can travel much longer distances, therefore covering larger geographic scales and associated
200 models, from slope to regional ones. Because of this, choosing the most appropriate mapping
201 unit becomes a crucial step to handle the spatio-temporal dimension of the HMP data. We
202 recall here that a mapping unit, in its most basic form, represents the geographic object upon
203 which the landscape is partitioned. In case of relatively small study areas, examples exist
204 where HMPs are modeled along specific streamlined and neighboring areas by adopting a fine
205 squared lattice. This type of resolution and characterization of the HMPs cannot be used in
206 our case, where the size of the Chinese territory would result in billions of grid-cells or data
207 points. Therefore, in case of such large geographic context, a common spatial partition choice

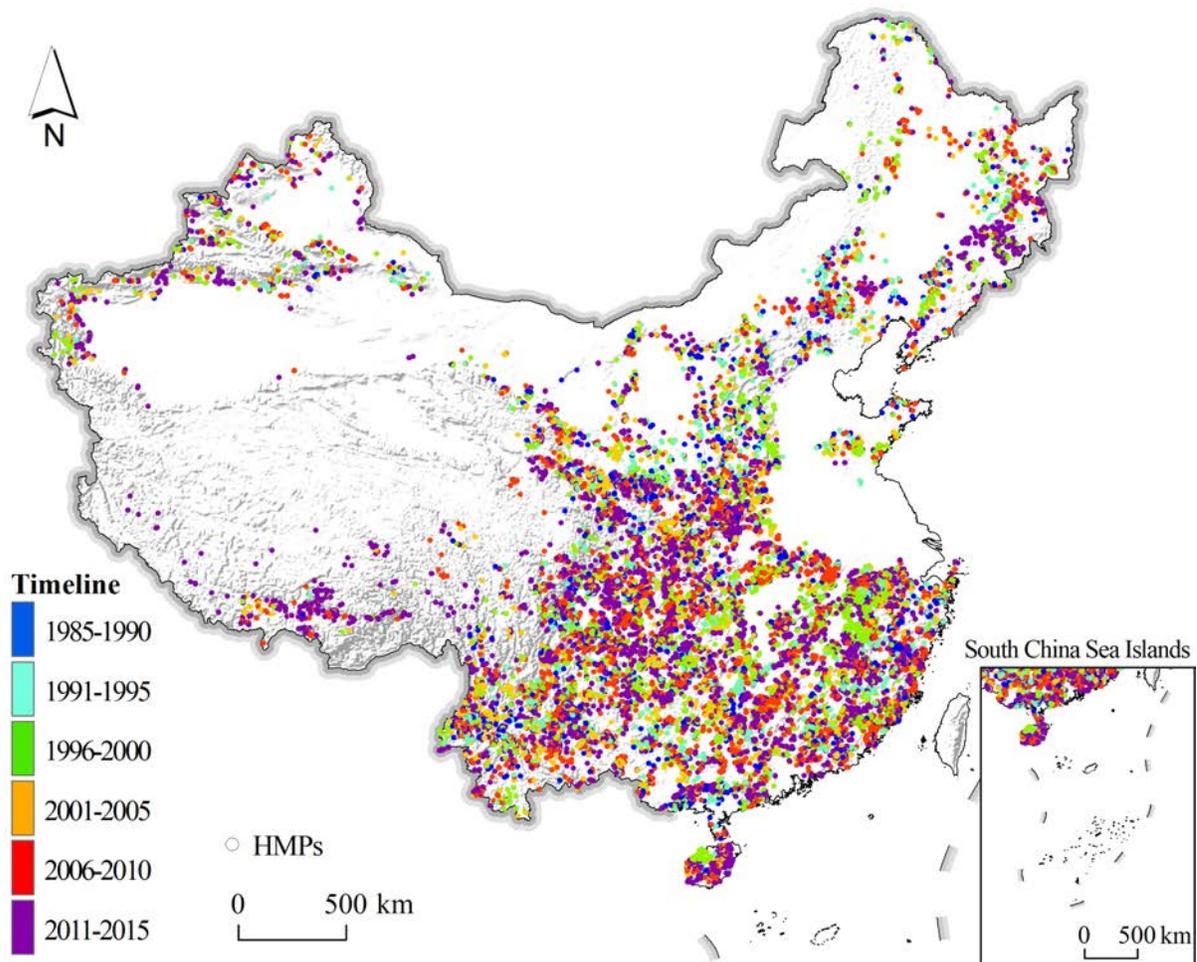


Figure 1: The multi-temporal HMPs in China from 1985 to 2015.

208 could be represented by administrative boundaries, upon which estimating the probability of
 209 HMP occurrences. However, the resulting susceptibility model would neglect the hydrology
 210 behind the natural process. In fact, administrative boundaries do not necessarily follow
 211 streams or catchment divides, where HMP occurrences can be considered independent or
 212 nearly-independent from each other. Therefore, a good solution to represent the spatial
 213 scale of HMPs, while respecting the hydrological realization of the natural phenomena, is
 214 to consider a catchment partition of the Chinese territory. To support the analyses in this
 215 work, we selected the 12th level catchment delineation published by Shen *et al.* (2017),
 216 which partitions the whole Chinese territory into 73,587 catchments. The corresponding
 217 distribution of catchment sizes is bimodal (see Figure 2) and it spans from 0.1 km² to
 218 667 km², with average area of 130 km² and a 95% confidence interval – measured as the
 difference between the 97.5 and the 2.5 percentiles of the distribution – of 231 km².

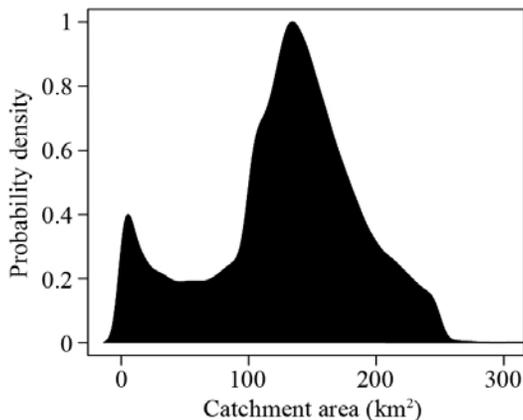


Figure 2: Probability density distribution of catchments sizes in China, computed from the 12th level published in Shen *et al.* (2017).

219

220 As any other mapping unit partition used the context of susceptibility modeling, a pre-
 221 processing step is required. The presence/absence information of HMPs is to be assigned
 222 to each catchment. To do so, we assign a presence (1) and absence (0) label to catchments
 223 where at least one HMP record is contained within a specific temporal unit (see details
 224 below).

225 3.3 Temporal unit

226 As much as the mapping unit choice aggregates HMP occurrences over space, whenever a
 227 dataset has a temporal connotation one should also choose a temporal unit. A temporal
 228 unit is the time interval through which we aggregate HMP occurrences and assign a suitable
 229 presence/absence conditions. In our case, the HMP dataset has very fine resolution with date
 230 and time available. However, the properties or covariates we will use in the model (see Section
 231 3.4) do not share the same temporal resolution. For instance, rainfall and temperature are

232 available with a daily resolution across China, vegetation cover and urban development are
233 available on a yearly basis while terrain properties do not exhibit any temporal changes.
234 Therefore, choosing a timescale that allows for meaningful interpretation and suitable data
235 is also crucial. In this context, the coarse temporal resolution of the covariates inhibits our
236 ability to build a finely resolved space-time model. And, in any case, choosing a fine temporal
237 resolution would inevitably increase the computational burden. Thus, choosing a reasonable
238 trade-off is required. Due to the characteristics of some crucial covariates, we chose a yearly
239 temporal unit. Such temporal unit implies that we assign a presence (1) and absence (0)
240 label to catchments where at least one HMP record is contained within a year time window.

241 3.4 Covariate set

242 HMPs are the result of several interplaying factors. These primarily feature:

- 243 1. precipitation, for it represents the main trigger;
- 244 2. catchment morphology, for it controls the time of concentration and other hydro-
245 dynamic parameters;
- 246 3. terrain attributes, for they control the path of the overland flows as well as the avail-
247 ability of material to be mobilized and transported;
- 248 4. soil hydrology, for it controls the interaction of the water with the earth surface;
- 249 5. vegetation density, for it can absorb part of the rainfall discharge and interact with
250 soil through the root system;
- 251 6. temperature, for it controls evapotranspiration and hence the soil moisture;
- 252 7. urbanization, for it may change the natural hydrology both because of impermeable
253 surface placed over permeable ones, and because buildings can also reduce the hydraulic
254 section through which HMPs may flow into.

255 In the context of space-time modeling, these properties need to be considered both in
256 terms of their spatial distribution and also in terms of their temporal evolution. In fact,
257 some properties will be more stationary over time, whereas some will have a much more
258 rapid rate of change. For instance, at the scale of the Chinese territory, soil hydrology can
259 be considered quite stationary within the 31 years under consideration. Conversely, rainfall,
260 vegetation and urbanization have a much faster spatio-temporal variation. Therefore, certain
261 properties can be introduced as a single realization (or map) whereas other properties should
262 be accounted for their successive temporal realizations (or maps).

263 We also consider antecedent HMPs, calculated over a time window of three years and
264 binarized into presence/absence conditions per catchment. We do so, to carry the spatial
265 signal of the HMPs. In fact, within a relatively short time window, we expect the suscepti-
266 bility to HMPs to be quite spatio-temporally consistent or stationary. In other words, areas
267 that have experienced HMPs in the recent past are more likely to suffer from HMP events
268 in the near future (Samia *et al.*, 2017). Hence, introducing the information of previously

269 occurred HMPs should better inform the model of this short-term spatial dependence and
270 improve its overall prediction capacity (Lombardo *et al.*, 2020).

271 As much as we tried to capture some residual dependence over space via antecedent
272 HMP events per catchment, we also tried to consider the presence of residual temporal
273 dependencies. Our assumption is that if climate change has produced a increasing trend in
274 rainfall extremes and resulting HMPs, a multiple intercept should also show an increasing
275 contribution through time.

276 The modeling protocol we implemented makes use of both types of covariates, featuring
277 properties that can be safely considered time-invariant within three decades: terrain and
278 catchment characteristics as well as soil type and climatic zones. And, also by featuring
279 properties that are explicitly time-variant within the same period: climate, vegetation and
280 human activity, as well as antecedent HMP events).

281 Due to the size of the study area and the temporal connotation of the database, the
282 number of covariate is inevitably large especially because a crucial step consists of ag-
283 gregating the covariate values in space (at the catchment scale) and time (at the yearly
284 scale). Due to the numerous data sources, the spatial resolution of the covariate set
285 we chose ranges from 90 m (SRTM, <https://earthexplorer.usgs.gov/>) to 8 km (NDVI,
286 <https://climatedataguide.ucar.edu/>). To summarize the spatial signal of each covariate (per
287 catchment) we calculated its mean and standard deviation. In case of stationary covariates,
288 such as terrain attributes, the spatial mean and standard deviation is a sufficient approxi-
289 mation where the mean reflects the main bulk of the pixel distribution per catchment and
290 the standard deviation highlight the associated variability. These values are kept constant
291 through time. As for catchment morphological indices, one single value is computed per
292 catchment and even in this case, the indices are kept constant through time (they are re-
293 peated for each of the 31 years).

294 For covariates that are nonstationary over time (such as rainfall, temperature and vegeta-
295 tion) we compute the spatial mean per catchment as well as the temporal mean and standard
296 deviation in a year. As for the anthropic signal, the percent of urbanized area with respect
297 to the total catchment size is directly calculated on a yearly basis, hence it does not need any
298 spatio-temporal aggregation. To this purpose, we employed the World Settlement Footprint
299 (WSF) Evolution which outlines at 30 m spatial resolution the global settlement growth
300 from 1985 to 2015 on a yearly basis (Marconcini *et al.*, 2020a). The WSF evolution has
301 been generated by exploiting the recently released WSF2015 layer, which maps worldwide
302 the settlement extent for the year 2015 (Marconcini *et al.*, 2020b). In particular, for each
303 pixel denoted as settlement in the WSF2015, a temporal analysis has been performed by
304 means of historical Landsat-5 and Landsat-7 optical satellite imagery to identify when the
305 construction took place. Here, an iterative procedure has been implemented where - starting
306 backwards from 2015 - training samples for the settlement and non-settlement class are ex-
307 tracted out of the map obtained at time t and Random forest binary classification has been
308 employed to outline the settlement extent at time $t-1$. Ultimately, zonal statistics have been

309 computed to determine yearly for each catchment partition the corresponding total amount
310 of settlement area.

311 A summary of all the covariates we considered is provided in Table 1.

312 **3.5 Susceptibility Modeling**

313 In this work, because of the vast study area and the long time series, we opted to create a sus-
314 ceptibility model that can feature spatio-temporal characteristics. We do so by considering
315 two types of models, an explanatory one and a set of predictive ones. The explanatory model
316 is a model built by using the whole available information. In our case, it is a model where
317 the entirety of China is taken into consideration together with its 31 years observations. In
318 such a way, one can build a model that can be used for interpretation, to understand the
319 statistical role of every environmental factor with respect to HMP occurrences. However,
320 such models do not have a predictive connotation because no new data is used to test the
321 classification performance. In fact, predictive models are built by calibrating the analysis
322 over a portion of the data. And, the calibrated relations are used to make a prediction over
323 an unknown dataset.

324 We stress here that the natural hazard community – at least the part of it using statistical
325 models – usually performs calibration by randomly subsetting a percentage of the data
326 over space and test the validation performance over the complementary cases. However,
327 prediction or forecast are terms usually referred to estimates of future occurrences, hence
328 in time. Rarely, studies dedicated to susceptibility models are validated in time (or chrono-
329 validated) (Lombardo and Tanyas, 2020; Cama *et al.*, 2015), mostly because of the inherited
330 complexity of obtaining accurate multi-temporal inventories (Guzzetti *et al.*, 2012).

331 Because our dataset spans over such a large time window, we actually have the chance to
332 test whether it is possible to forecast future occurrences. Thus, we have opted to assess the
333 predictive capacity of future HMP occurrences by considering four cross-validation schemes:

- 334 1. Forward-All or MOD1: This validation procedure starts by calibrating our binomial
335 GLM (more details in Section 3.5.1) over a specific year (e.g., 1985) and testing over
336 the remaining time series (e.g., 1986-2015). In the second step, the previous reference
337 year is combined with the next (e.g., 1985 and 1986) to predict HMPs in the remaining
338 years (1987 to 2015). This moving window moves one year at a time until completion
339 of the time series.
- 340 2. Forward-Sequence or MOD2: This validation scheme iteratively calibrates over a spe-
341 cific year (e.g., 1985) and predicts only the next (e.g., 1986). In the second step, the
342 calibration aggregates the subsequent year (e.g., 1985 and 1986) and predicts only the
343 next (e.g., 1987). This is repeated until the completion of the time series in 2015.
- 344 3. Backward-All or MOD3: This validation scheme is analogous to MOD1 but it is imple-
345 mented in the opposite temporal direction. Specifically, we calibrate over the last year

Table 1: Covariates' summary: (Time-invariant variables: terrain feature, stream/catchment feature, soil type, and climatic zone; Time-variate variables: climatic indicators, NDVI, settlement area, 3-years antecedent HMP events.

Category	Indicator	Definition		
Terrain feature	EIV_μ	Mean of elevation.		
	ELV_σ	Standard deviation of elevation.		
	SLP_μ	Mean of slope.		
	SLP_σ	Standard deviation of slope.		
	PLC_μ	Mean of plan curvature.		
	PLC_σ	Standard deviation of plan curvature.		
	PRC_μ	Mean of profile curvature.		
	PRC_σ	Standard deviation of profile curvature.		
Stream / catchment feature	Wandering ratio (Chorley, 1957)	$R_w = \frac{L_{MF}}{L_B}$	Drainage density (Strahler, 1952)	$D_d = \frac{L_u}{A}$
	Form factor (Horton, 1932)	$F_f = \frac{A}{L_B}$	Relief ratio (Schumm, 1956)	$R_r = \frac{R_B}{L_B}$
	Elongation ratio (Schumm, 1956)	$R_e = \frac{2}{L_B \times (A/\pi)^{0.5}}$		
	<p>A is the drainage area; L_{MF} is the length along the longest watercourse from the mouth to the head of the channel; L_B is the maximal length of the line from a basin mouth to a point on the perimeter equidistant from the basin mouth in either direction around the perimeter; R_B is the elevation difference between the highest point on the drainage divide and the mouth; L_u is the order-wise total stream length based on Strahler stream order.</p>			
Soil type	The area percentage of each kind of soil in each catchment. The soil types include Clay, ClayLoam, Loam, LoamSand, Sand, SandyClay, SandyClayLoam, SandyLoam, Silt, SiltClay, SiltClayLoam, SiltLoam.			
Climatic zone	The area percentage of each climatic zone in each catchment. The climatic zones include north temperate zone, central temperate zone, south temperate zone, north subtropics zone, central subtropics zone, south subtropics zone, north tropics zone, central tropics zone, highland climatic zone.			
Climatic indices	RAIN_Tμ_Sμ	The mean of each catchment (Sμ) with the mean daily rainfall in each year (Tμ).		
	RAIN_Tσ_Sμ	The mean of each catchment (Sμ) with the standard deviation of the daily rainfall in each year (Tσ).		
	RAIN_TA_SA	The maximum of each catchment (SA) with the maximum daily rainfall in each year (TA).		
	Annual RAIN_Sμ	The mean of each catchment (Sμ) with the annual rainfall in each year.		
	TEM_Tμ_Sμ	The mean of each catchment (Sμ) with the mean daily temperature in each year (Tμ).		
	TEM_Tσ_Sμ	The mean of each catchment (Sμ) with the standard deviation of the daily temperature in each year (Tσ).		
	TEM_TA_SA	The maximum of each catchment (SA) with the maximum of the daily temperature in each year (TA).		
NDVI	NDVI_Tμ_Sμ	The mean of each catchment (Sμ) with the mean NDVI in each year (Tμ).		
	NDVI_Tσ_Sμ	The mean of each catchment (Sμ) with the standard deviation of the NDVI in each year (Tσ).		
Settlement area	The estimated settlement area per polygon in km ² in each year.			
Antecedent disasters	The cumulative quantity of flash flood disasters occurred in the antecedent 3 years.			

(2015) and predict the whole time series backward (from 1985 to 2014). In the next step we then calibrate aggregating the information of the previous year (e.g., 2015 and 2014) to predict the remaining time series (1985 to 2013). This operation is backwardly repeated until the completion of the time series in 1985.

4. Backward-Sequence or MOD4: this model is analogous to MOD2 but again it is implemented in the opposite temporal direction. This means that the calibration starts in 2015 and it is used to predict the previous year only (2014). Then the calibration integrates the information from the previous year (2015 and 2014) to predict only one step back in time (2013). This operation is repeated backwardly until the time series is completed in 1985.

Notably, each of these validation schemes inevitably produces 30 testing outputs, whereas the explanatory model only produces one training output.

3.5.1 Generalized Linear Models

The vast majority of statistically-based susceptibility models are carried out by using Generalized Linear Models (Budimir *et al.*, 2015; Reichenbach *et al.*, 2018). This class of models assumes that the response variable follows an exponential family distribution such as Gaussian, Poisson, Bernoulli and more. Among those, the Bernoulli case, also referred to as Binary Logistic Regression, corresponds to a model where the target variable can take on only two values. Therefore, a binomial GLM estimates the probability that a given mapping unit belongs to one of the two classes (by standard, this is the class 1, or the class conveying the presence of HMPs, rather than 0). More specifically, a binomial GLM can be denoted as follows:

$$\text{logit}(\pi) = \frac{\pi}{1 - \pi} = \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \dots + \beta_n X_n \quad (1)$$

where, the target variable Y is assumed to be Binomial with a probability π of a given catchment to experience a HMP. The β_0 term is the global intercept and β_n are the regression coefficients estimated for X_n covariates. The *logit*, or the natural logarithm of the odds, allows for the conversion of the odds into probabilities.

This framework allows for continuous and discrete covariates. Each class of a discrete covariate is modeled independently from the other classes, or technically it is assumed to be *independent and identically distributed* (iid). More specifically, the model will assign a different regression constant to each class separately from the others. Notably, in this work we make use of iid covariates for a multiple yearly intercept for the explanatory reference model. The remaining covariates are all continuous in nature and used as linear properties both in the explanatory and predictive models.

379 3.5.2 Estimates of confidence intervals

380 In statistics, any model should allow for inference on a distribution of estimates rather than
381 a single estimate. In other words, obtaining a mean prediction is as important as measuring
382 the uncertainty around that mean value. Therefore, in this work we sought to retrieve both
383 the mean behavior of every regression coefficient and performance metric as well as the
384 estimated variability associated with them.

385 To do so, we present two schemes, one for the explanatory model and one for the validation
386 routines (MOD1 to MOD4). When implementing the explanatory model (we recall here that
387 it is fitted using the whole available information), we have also added a bootstrap simulation
388 step (Efron and Tibshirani, 1994). This step essentially re-samples with replacement the
389 whole dataset and re-fits the same model structure to the simulated dataset. We do this
390 over 100 bootstrap replicates to estimate the sampling distribution of each parameter we
391 store during the explanatory analyses. Besides, we implement the 10-fold cross validation
392 to evaluate the overall performance on the whole dataset. As for the validation routines in
393 MOD1 to MOD4, the variability of the tests is summarized via the 30 estimates, one for
394 each of the 30 years under consideration.

395 3.5.3 Model evaluation

396 The primary tool to assess the performance of our HMP susceptibility model consists of the
397 Receiver Operating Characteristic curves (ROC, Hosmer and Lemeshow, 2000) and their
398 integral or Area Under the Curve (AUC, Hosmer and Lemeshow, 2000). The former is
399 the most common threshold independent metric used in classification problems (Rahmati
400 *et al.*, 2019). It is constructed by slicing the probability spectrum at various cutoff, and by
401 computing the confusion matrix at each step. As a result, it is possible to calculate the False
402 Positive Rate or FPR ($FP / [FP+TN]$) and the True Positive Rate or TPR ($TP / [TP+FN]$)
403 for each cutoff. The integral of the curve defined by the FPR and TPR pairs calculates from
404 different cutoffs can be then used as an index of performance. Specifically, $AUC = 1$ indicate
405 a perfect classification, $0.9 < AUC < 1$ refers to outstanding performance, $0.8 < AUC < 0.9$
406 marks excellent performance whereas $0.7 < AUC < 0.8$ are acceptable results. Any AUC
407 value from 0.7 to 0.5 indicates a range of poor performance down to results comparable to
408 a random classification.

409 We make use of the AUC throughout the manuscript. We also implement a Jackknife
410 test in the validation scheme (Lombardo and Mai, 2018; O'Banion and Olsen, 2014). A
411 Jackknife test is essentially divided into two steps. The first one runs single (j^{th}) variable
412 models whereas the second runs all-but-one-variable ($(j - 1)$) models. In both cases, the AUC
413 is calculated to offer a comprehensive summary of covariates contributions. Single variable
414 models highlight stand-alone performance of specific covariates in explaining HMP occur-
415 rences. All-but-one-variable models highlight performance drop resulting from the removal
416 of one single covariate at a time, with respect to a full model using them all at once.

417 Notably, the validation scheme in this work includes training and testing 30 temporal
418 models. As a result, we have run 30 Jackknife tests, one for each year from 1985 to 2015.

419 4 Results

420 4.1 Explanatory Model and its cross-validation

421 In this section, we reported the regression coefficients obtained from a susceptibility model
422 built by using all the available spatio-temporal information. These estimates were used to
423 interpret the relation between HMP occurrences and environmental conditions (or covari-
424 ates). Firstly, each regression coefficient is characterized by a distribution of values which
425 have been retrieved from 100 nonparametric bootstrap replicates. Figure 3 summarizes
426 each model component. Among the continuous covariates (see Fig.3), climatic indices (e.g.
427 $RAIN_{T\sigma_S\mu}$, $AnnualRAIN_S\mu$), terrain attributes (e.g. $PLC_σ$, $SLP_σ$), catchment
428 morphology (e.g. form factor) present notable positive regression coefficients. In addition,
429 catchments located in Central temperate and South temperate zones also suffer more from
430 the HMPs. More details on the interpretation of this covariate effects will be provided in
431 Section 5.

432 Besides, we made use of an iid effect for each *year*, whose result is shown in Figure 4. The
433 year-specific regression constants show an interesting pattern. For each year from 2002 to
434 2014, all regression coefficients are significantly positive and the whole distribution is quite
435 distant from the zero line (between 0.5 and 1) with an exception of 2004. As for each year in
436 the period between 1985 and 2001, the regression constants are also estimated with a positive
437 median coefficient, although some of them appear to be not significant (the distribution
438 of regression constants also show negative values). Besides, the regression coefficients vary
439 around the zero. More details on the interpretation of this temporal iid effect will be provided
440 in Section 5.

441 to complete the analyses on the whole spatio-temporal domain, we also run a 10-fold
442 cross validation. We recall here that a 10-fold cross validation implies randomly partitioning
443 the whole data population into ten complementary subsets, each time extracting 90% and
444 10% for calibration and validation, respectively. Figure 5 presents the performance of the
445 10-fold cross-validation scheme. Specifically, panel 5a reports 10 ROC curves obtained by
446 using 90% of the spatiotemporal HMP data; and panel 5b reports ROC curves obtained by
447 testing over 10% of the spatiotemporal HMP data. The respective mean AUC values do not
448 significantly change, as they both returned 0.84. This attest both for excellent goodness-
449 of-fit and prediction-skill according to Hosmer and Lemeshow (2000) as well as a indicating
450 robust results with differences that can be distinguished only at the third decimal place.

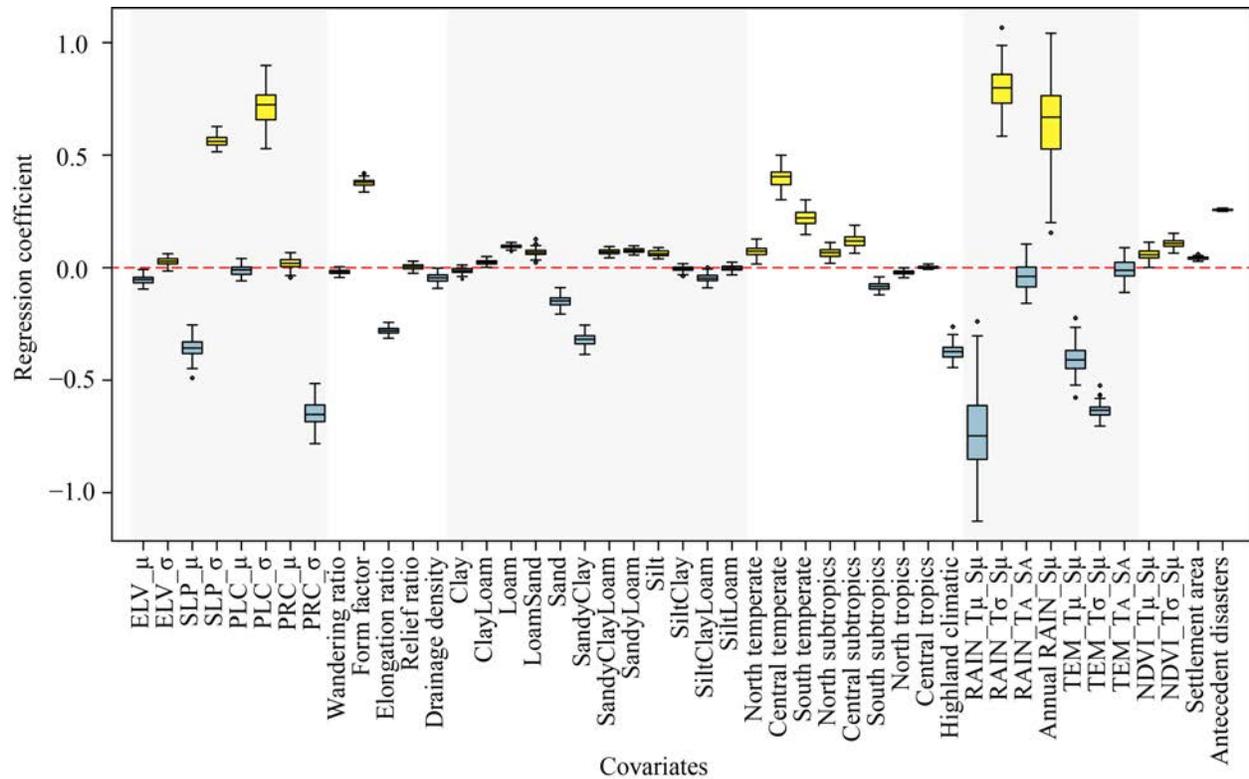


Figure 3: Regression coefficients estimated through the explanatory model built by using the whole HMP spatio-temporal information across China. The covariates shown in this figure are continuous in nature. The red dash line corresponds to zero or no-contribution to the model. Boxplots shown in blue indicate a median negative correlation to HMPs while yellow indicates a median positive one.

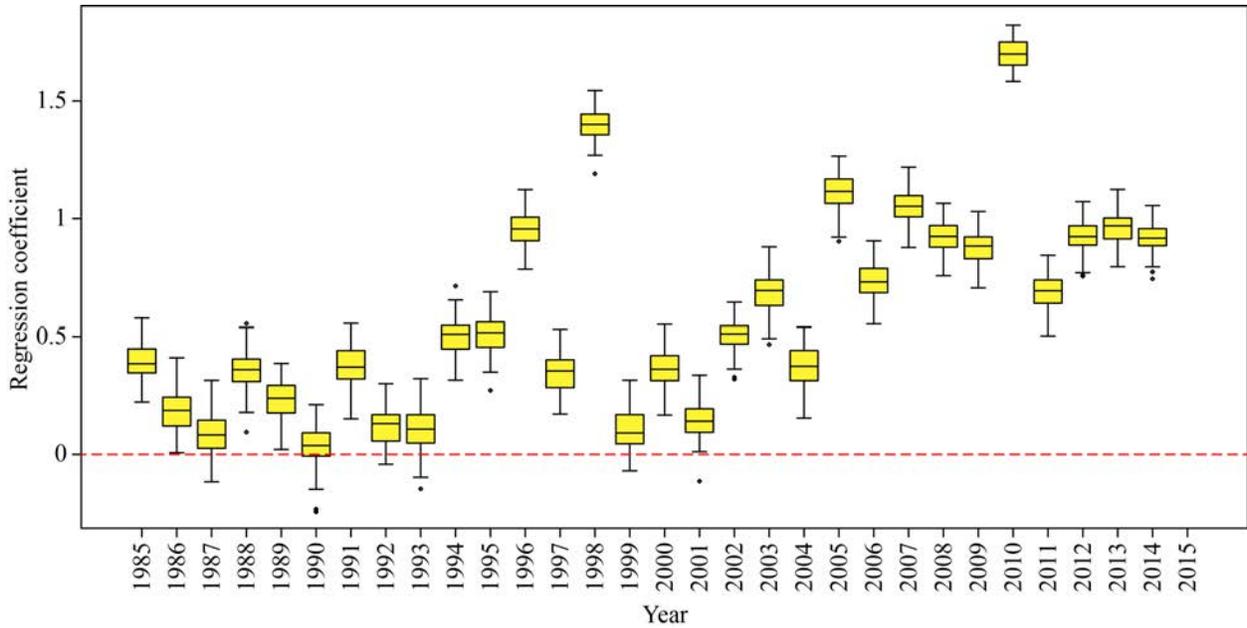


Figure 4: Regression coefficients estimated through the explanatory model built by using the whole HMP spatio-temporal information across China. The covariates shown in this figure are categorical in nature and correspond to the yearly contribution to the model. The red dash line corresponds to zero or no-contribution to the model.

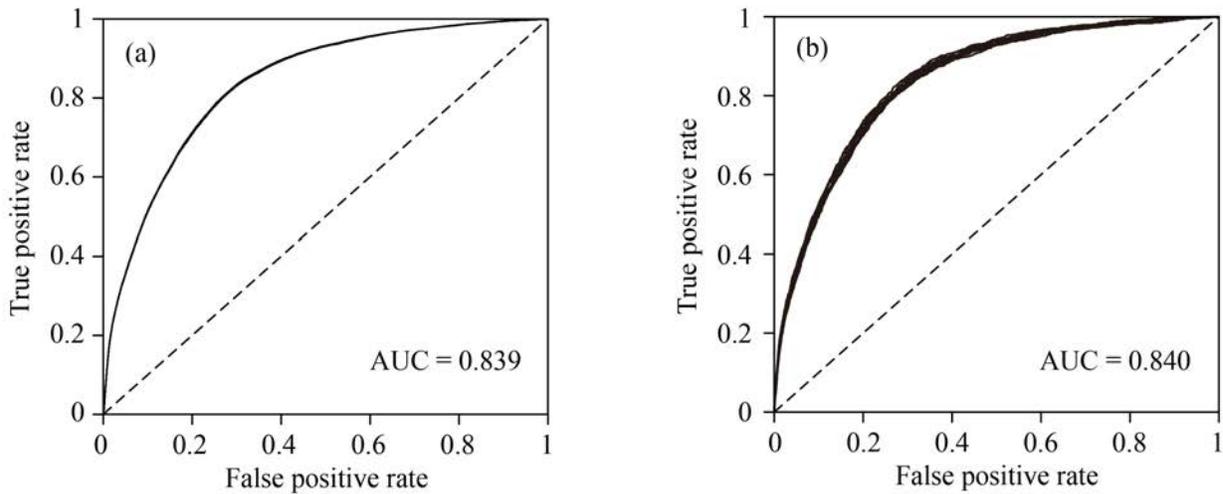


Figure 5: ROC curves obtained via 10-fold cross-validation. (a) Ten calibration models (90%), (b) Ten validation models (10%). The AUC reported in both panels corresponds to the mean of the ten replicates, respectively.

4.2 Temporal Validation Routines

Here we present the four temporal validation schemes described in Section 3.5. For each temporal validation scheme, we summarize the model performance in Figure 6. All models are reported with a mean temporal AUC greater than 0.82. We recall here that this value corresponds to excellent performance according to the AUC classification system proposed by Hosmer and Lemeshow (2000). However, two distinct patterns arise in the four temporal validation routines. The AUCs obtained for each year in MOD1 and MOD3 appear quite smooth. In MOD1, this is also associated with a downward shift in AUC when comparing calibration and validation performances (Figure 6a). As for MOD3, calibration and validation performance largely overlap, with the exception of the period in between 2009 and 2015 where the validation routine shows a significant drop in predictive capacity (Figure 6c). In case of MOD2 and MOD4, the AUC values estimated for each year present a much rougher temporal variation. Between these two validation schemes, MOD4 less accurately predicts the HMPs in the last years of our AUC time series (Figure 6d). As for MOD2, a similar difference in performance between calibration and validation is shown for the initial years of our HMP time series (Figure 6b). However, the initial years from 1986 to 1989 contain less HMP occurrences, thus a relatively low performance in this period is much more acceptable than a relatively low performance in the latest years. In light of these considerations, and because of a slightly better performance overall, we consider MOD2 (or Forward-Sequence) as the best validation scheme compared to the other three.

We stress again that a close look at MOD2 in Figure 6b highlights some fluctuations in the AUC time series for the validation whereas the calibration appears much more stable through time in terms of estimated performance. This is better presented in Figure 7 where we show 30 ROC curves, one for each year. The panel (a) corresponds to the training ROC curves and aside for a few years, they consistently overlap through time. As for the validation shown in panel (b) a marked spread can be seen in the curves spanning from 1986 to 2015. We note here that the relatively poorer performance registered at the start and end of the time series also correspond to two years with a relatively lower number of observations. Conversely, the other relatively low AUC values between the two endpoints always appear in the following year containing very large numbers of HMP occurrences. This may be due to the fact that an abrupt increase in HMPs, may induce some variations in the estimated correlations between HMPs and covariates. This in turn, may also induce variations in the susceptibility patterns, which may end up not matching the HMPs of the subsequent year (likely to be representative or closer to the average Chinese susceptibility pattern). This explanation fits well the year 1998. That year was characterized by an exceptionally large number of HMPs in China, and the temporal validation of 1999 returned the poorest performance we observed across the whole temporal sequence. Notably, such temporal variations in performance has been similarly shown in other studies, where the authors reported that effect of climate change may be responsible for large uncertainties in the prediction of HMPs (e.g., Collier, 2007).

To provide a comprehensive overview of the model structure and covariates' role in the

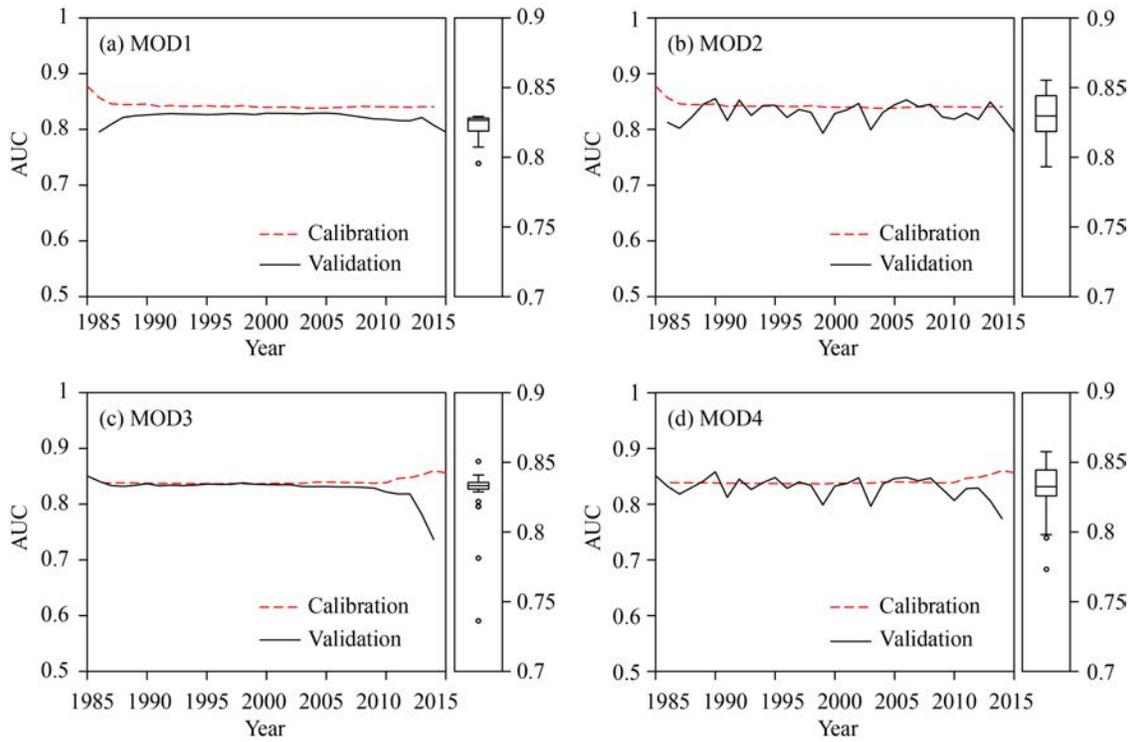


Figure 6: Each panel corresponds to one of the four temporal validations we tested. The line plots report the AUC time series from 1985 to 2015. The boxplots summarize the AUC variation over the thirty years.

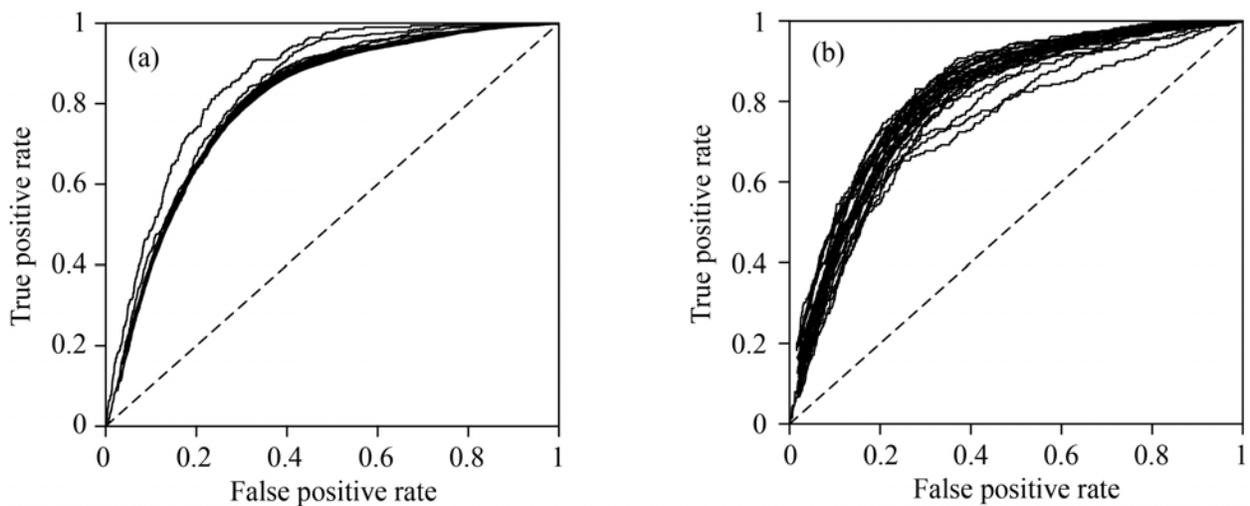


Figure 7: The ROC curves obtained by MOD2. (a) Calibration model, (b) Validation model

491 temporal validation, we performed a suite of Jackknife tests (Jiao *et al.*, 2019). We recall
492 here that Jackknife tests are essentially replicates of a reference model whose structure
493 is perturbed by either building single-variable (*only-one-variable*) models for each of the
494 covariate in the reference structure. Or, by removing one covariate at a time (*all-but-one-*
495 *variable*) from the whole set of covariates. Many example of Jackknife tests exist in the
496 susceptibility literature, but they have been limited so far to a pure spatial domain (see
497 for instance, Park, 2015; Lombardo *et al.*, 2016; Ramos-Bernal *et al.*, 2019). Here, because
498 we consider both spatial and temporal dimensions, we iterated the *only-one-variable* and
499 *all-but-one-variable* models thirty times, once per year from 1985 to 2015.

500 Figure 8a presents the AUC obtained via *only-one-variable* models. It indicates that
501 most of the terrain attributes, climatic indices, and antecedent disasters could contribute
502 to a model with an AUC greater than 0.6. At the same time, the *all-but-one-variable*
503 models in Figure 8b indicates that removing either of SLP_{σ} , form factor, elongation ratio,
504 $RAIN_{T\sigma}S\mu$, and antecedent disasters from the model will induce an obvious AUC drop.

505 4.3 Regression Coefficients

506 In addition to assessing model performance, another crucial step in any modeling protocol is
507 to evaluate how reasonable regression coefficients may be from an interpretative standpoint.
508 In this work, we already summarized a similar information for our benchmark fit. Never-
509 theless, regression coefficients estimated for the temporal validation scheme could shed light
510 on the variability that each covariate effect may exhibit through time. Here, we assigned
511 the yellow color for a positive β value, which indicates the probability of HMP occurrence
512 will increase by a factor equal to the exponential of the β value. Conversely, the blue color
513 indicates a decrease.

514 Among the terrain attributes, the standard deviation of slope (SLP_{σ}) and plan cur-
515 vature (PLC_{σ}) play an important role in controlling the estimated probability of HMP
516 occurrences (Figure 9). In terms of catchment morphology, form factor and elongation ratio
517 show a positive effect. Most soil types present non significant and negligible contributions
518 to the thirty cross validation schemes, with the exception of Sandy Clay which appears to
519 be negatively correlated to HMPs, although with a slight negative influence. Furthermore,
520 catchments located in Central temperate, South temperate, and Central subtropics zones
521 appear to be more prone to HMPs than the others.

522 The summary presented above reports the role of time-invariant properties. As for time-
523 variant covariates, $AnnualRain_{S\mu}$ showed the largest significant and positive effect out of
524 all the climatic indices, followed by $RAIN_{T\sigma}\mu$ (the intra-annual rainfall variance within
525 a given catchment). The 3-years antecedent disasters in a given catchment also appeared to
526 be significant and to increase the susceptibility estimates.

527 Notably, the summary of the covariates' effects shown above is quite static as it overlooks
528 the temporal variation that each model component exhibit as the temporal-validation is
529 performed. To complement this information, in Figure 10 we show the temporal evolution of

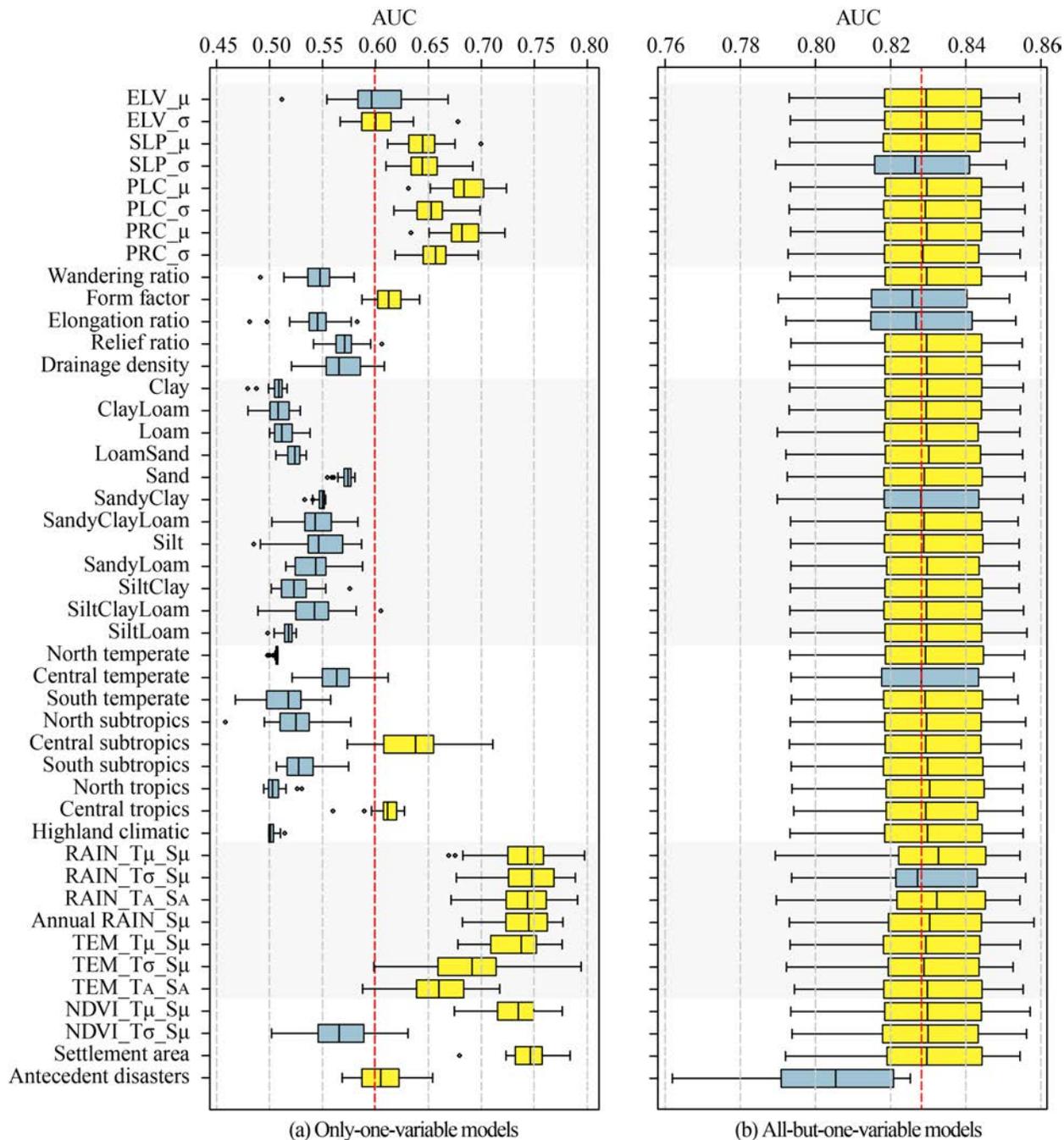


Figure 8: Jackknife test for covariates. Only-one variable models are shown in the left panel and all-but-one variable models in the right panel. Red line indicates the corresponding mean value of all combinations. Blue boxplots indicate a covariate-specific median AUC lower than the mean AUC computed for all covariates. Yellow boxplots correspond to higher covariate-specific median AUC.

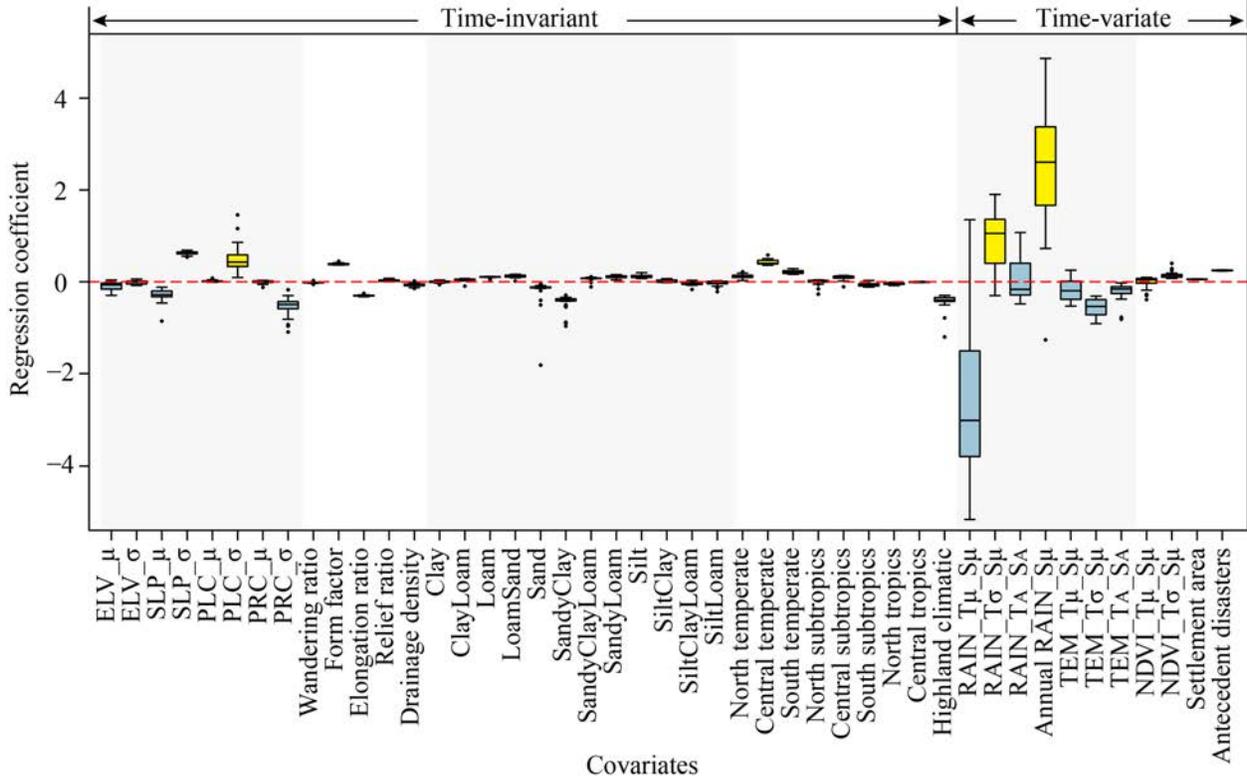


Figure 9: Regression coefficient obtained by MOD2.

530 the regression coefficients belonging to covariates that appeared to be significant in Figure
 531 9.

532 More specifically, to better distinguish the variance of the covariates' effects through time,
 533 we split Figure 10 in two panels, according to the magnitude of the regression coefficients.
 534 Panel (a) summarizes β coefficients whose magnitude through time ranges from -0.5 to 0.5,
 535 whereas panel (b) presents the same information for β coefficients whose magnitude through
 536 time ranges from -5 to 5. Most of the covariates in both panels indicated a constantly similar
 537 effect on HMP occurrence, whereas, few covariates showed a large variation through time.
 538 For instance, the annual rainfall (*AnnualRAIN_Sμ*) indicated an increasing positive influ-
 539 ence from 1985 to 2014. However, the variance of NDVI (*NDVI_Tσ_Sμ*) within each year
 540 showed a decreasing effect before 1990, after which the trend flattened until the end of the
 541 time series. Overall, the covariates which exhibited the largest variation through time all cor-
 542 respond to climatic indices, especially those associated with rainfall (see *AnnualRAIN_Sμ*
 543 and *RAIN_Tσ_Sμ* in Figure 10b).

544 4.4 Susceptibility Mapping

545 HMPs susceptibility maps generated via MOD2 are drawn in Figure 11 from 1996 to 2015.
 546 These have been classified into very low (VL), low (L), low to medium (LM), medium to

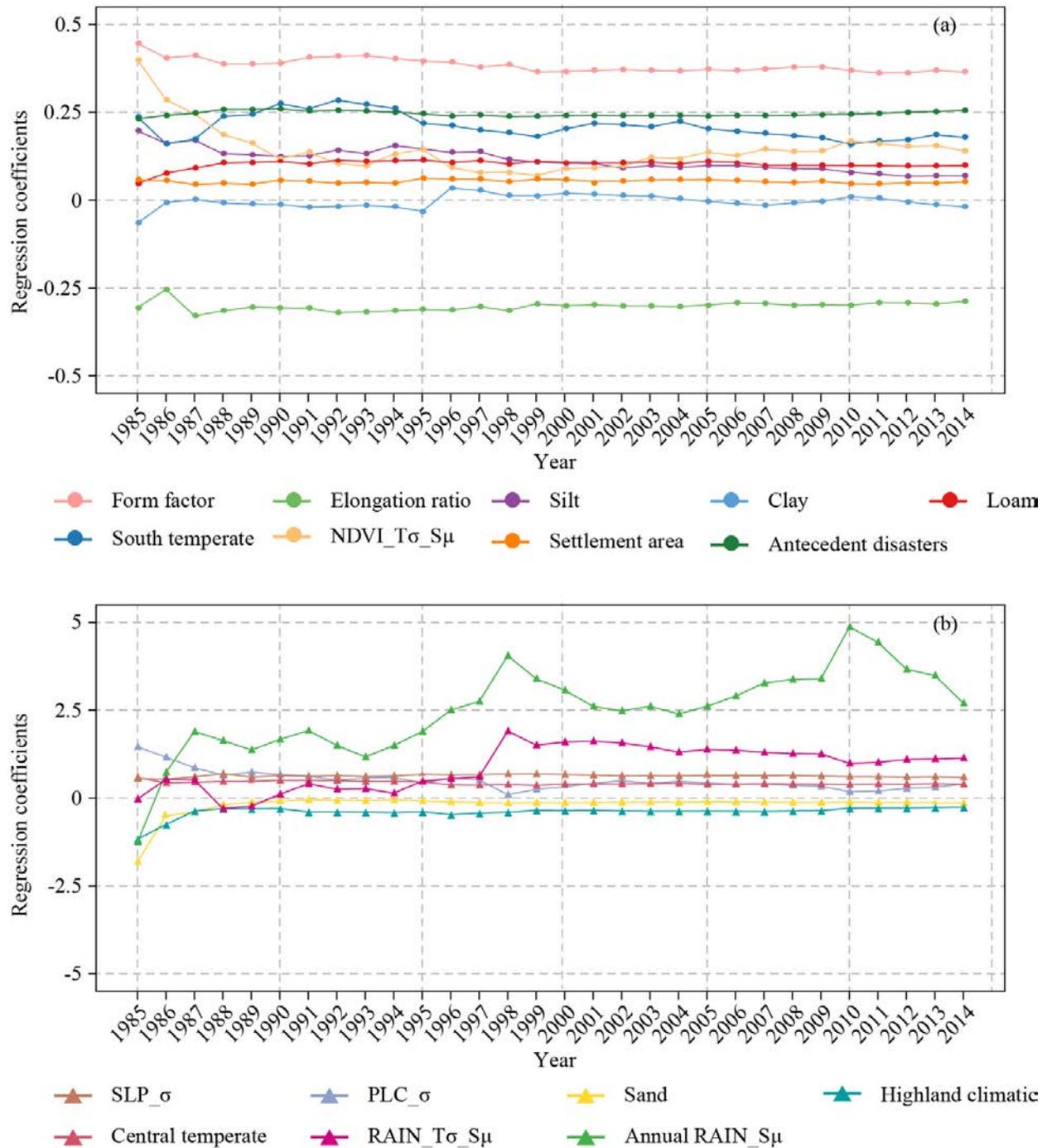


Figure 10: The regression coefficient varying across the of 1986-2015 obtained by MOD2.

547 high (MH), high (H), and very high (VH) according to break points that have been set as the
548 2.5%, 25%, 50%, 75%, and 97.5% percentiles of the whole probability spectrum combined.
549 In other words, to reclassify the numerical susceptibility into classes, we have concatenated
550 all the space-time HMP probability estimates into a single vector, from which five percentiles
551 have been extracted to ensure a common color palette among the 30 maps.

552 Looking at the time series of susceptibility maps (Figure 11), at the beginning of the study
553 period probabilities are generally lower, especially in the western sector. Besides, as the time
554 series evolves towards recent years, the probability spectrum appears to shift towards higher
555 susceptibility classes. More specifically, catchments with very low probabilities of HMP
556 occurrences mainly appear from 1986 to 1988; whereas catchments presenting very high
557 probability of HMP occurrences characterize the south-east sector of China since 1997.

558 To summarize the space-time susceptibility information depicted in Figure 11, we further
559 generated three maps aimed at delivering the mean and the maximum susceptibility together
560 with the 95% confidence interval (see Figures 12a, 12b and 12c respectively).

561 Looking at the susceptibility patterns depicted in the mean and maximum maps, two
562 macro-regions stand out the most. The western sector of China appears to be consistently
563 estimated as non susceptible. There, the susceptibility appears to be generally confined
564 within the first 10% of the national distribution. On the contrary, the south-eastern sector
565 appears to be generally the most susceptible. There, most of the catchments are associated
566 with susceptibilities estimated above 70% of the national probability distribution. Notably,
567 few exceptions exist to this observation due to the existence of large plains, where catchments
568 are generally gentler in topography. Other catchments with high HMP susceptibility, albeit
569 lower than the south-east, can still be found in central, north-east and southern China.

570 Interestingly, the 95% confidence interval map – we recall here to be computed as the
571 difference between the 97.5th and 2.5th percentiles of the spatio-temporal probability spec-
572 trum shown in Figure 11 – marks analogous geographic patterns to the mean and maximum
573 maps. This is an indication of the robustness of our model. In fact, this means that areas
574 with low susceptibilities are estimated with similar values through time. Conversely, areas
575 with high susceptibility exhibit a much larger degree of variation through time, as expected
576 just by looking at the raw data in Figure 1.

577 The last panel of Figure 12 depicts seven cluster drawn from the maximum susceptibility
578 in the same figure. These have been manually interpreted on the basis of expert opinion.
579 Clusters *I* to *V* correspond to regions are affected by monsoon. The reason to split *I* and *II*
580 are due to the difference of terrain and annual rainfall whereas the reason to split *I* and *III*
581 into two parts is due to the mountain range that acts as a strong topographic barrier. More
582 specifically:

- 583 • Cluster *I*: the region with the most severe erosion due to the topographic control;
- 584 • Cluster *II*: the region mostly affected by monsoon;
- 585 • Cluster *III*: less annual rainfall, Loess Plateau affected by widespread gully incisions;

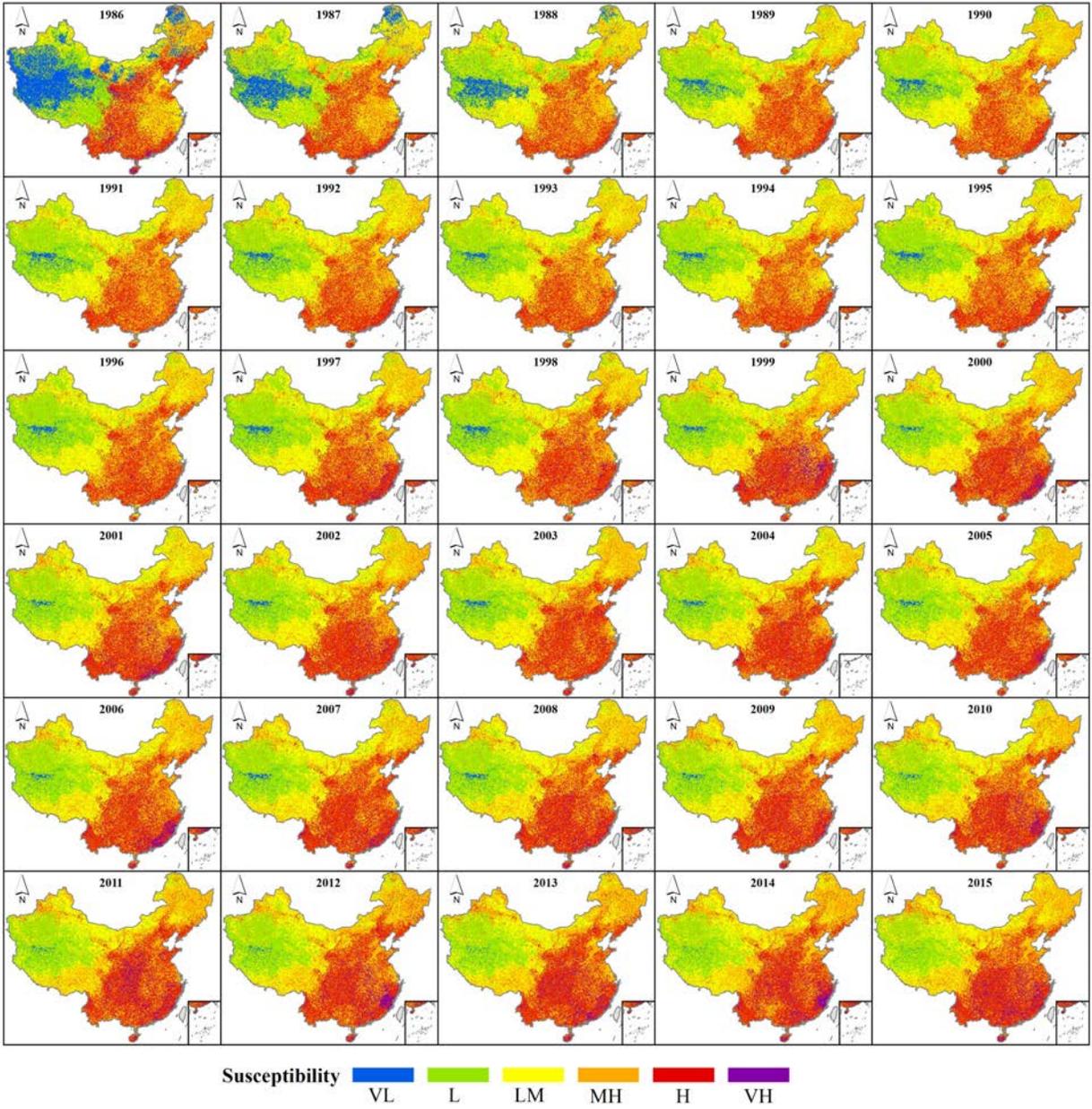


Figure 11: Susceptibility of HMPs in China during 1986-2015 detected via MOD2.

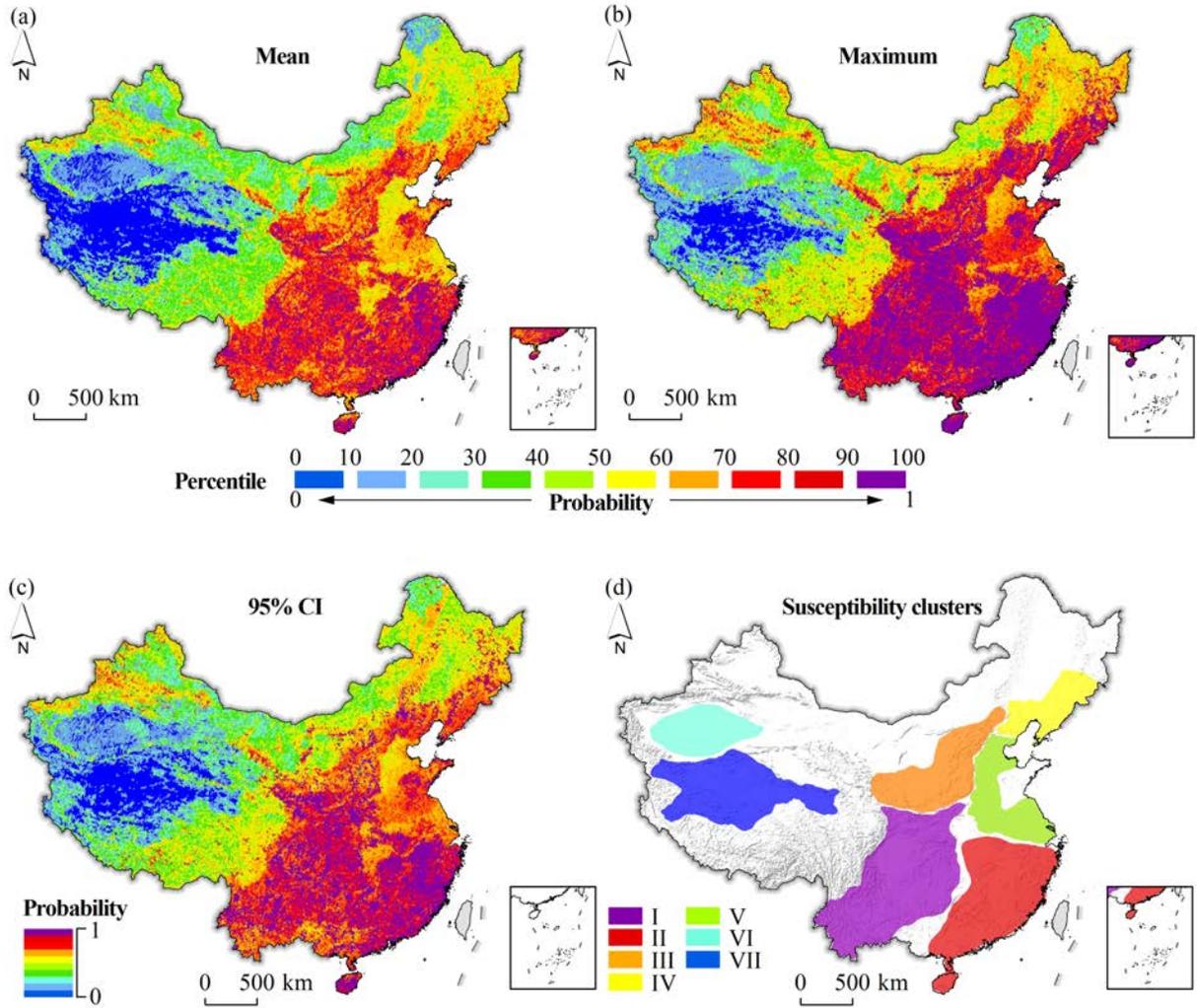


Figure 12: Summary of HMP susceptibility estimated for China from 1986 to 2015 via MOD2: (a) Mean susceptibility, (b) Maximum susceptibility, (c) 95 % CI susceptibility. Panel (d) shows seven interpreted clusters from panel (b).

- 586 • Cluster *IV*: this sector of China shows a relatively large proneness towards HMP
587 although the rainfall intensity due to the incoming monsoons in this area is much
588 lower than the precipitation discharged to the south. This is primarily due to the local
589 rough topography which contributes to increase the probability of HMP occurrence;
- 590 • Cluster *V*: plains with widespread flat terrains;
- 591 • Cluster *VI*: distinct characteristics attributable to the Taklamakan Desert and the
592 Tarim Basin;
- 593 • Cluster *VII*: sparsely populated area corresponding to the Changtang Plateau and
594 Qinghai Hoh Xil Plateau.

595 Figure 12 is meant to compress the spatio-temporal susceptibility information in the
596 geographic space. To do the same for the temporal case, we went back to Figure 11 and
597 computed the for each year the areas assigned with one of the six susceptibility class. From
598 these, we generated a stacked barplot (see Figure 13) reporting the proportion of China
599 associated with one of the six classes, showing the evolution through time from 1986 to
600 2015. What stands out the most is that the areal percentage of catchments with very low
601 (VL) susceptibility decreased sharply in the first three years. This effect is mostly due to
602 the fact that as the time series progressed, more HMP have been recorded, which generally
603 leads to a higher probability of HMP. On the opposite side of the probability spectrum, the
604 proportion of China associated with very high (VH) HMP susceptibility can be seen to have
605 increased from 1997 onward. We remind here the reader that despite these changes may
606 appear small in a simple graphical summary such as Figure 13, in reality a variation of even
607 just 1% of the total Chinese territory involves several hundreds thousands of km² and several
608 hundreds actual catchments.

609 5 Discussion

610 5.1 Supporting arguments

611 This work estimates and investigates the spatio-temporal variation of HMP susceptibility
612 patterns over China. Because of the vast space-time domain, many options exists on how to
613 build and validate a space-time predictive model (Lombardo *et al.*, 2020).

614 We chose a binomial GLM, which we calibrated and validated through different strategies.
615 The first strategy we used exploited the whole space-time domain, from which catchments
616 with high variations in slope steepness and planar curvature appear to increase the overall
617 susceptibility to HMPs. The influence of slope with respect to HMPs is widely acknowledged
618 in the literature. However, for analyses expressed at the catchment scale, the effect of the
619 terrain steepness may be lost. This may be the reason why in our model, the positive
620 role of the slope steepness is expressed in terms of standard variation, a common proxy for

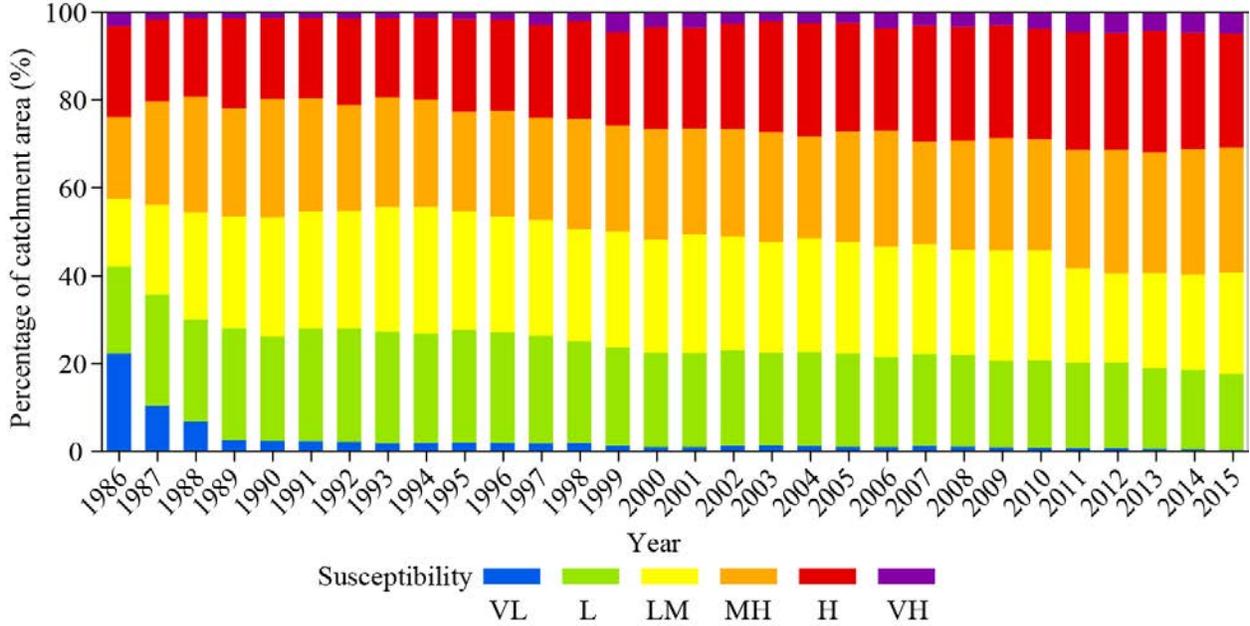


Figure 13: Proportion of the Chinese territory estimated to be HMP-susceptible to HMP, from 1986 to 2015, via MOD2.

621 topographic roughness. A similar reasoning can be inferred for the standard deviation of the
 622 planar curvature.

623 Unsurprisingly, another positive contribution is carried by the rainfall patterns, expressed
 624 through the $RAIN_T\sigma_S\mu$ and the $AnnualRAIN_S\mu$ (see Figure 3). It should be noted
 625 that the spatio-temporal rainfall signal is carried in the model via four summary statistics of
 626 the precipitation over the mapping (catchment) and over the temporal (year) units. This is
 627 certainly the reason behind the overall negative contribution estimated for $RAIN_T\mu_S\mu$. In
 628 fact, in any multivariate analysis, whenever slightly dependent covariates interact with each
 629 other, the estimation of their sign and amplitude can also depend on each other presence
 630 within the model. Because the four rainfall aggregation measures stem from the same spatio-
 631 temporal information, it is safe to assume that some degree of dependence can exist among
 632 the four we computed. Therefore, the overall influence of rainfall on HMP occurrences
 633 should be interpreted as the combined effect of the four covariates and their estimated
 634 regression coefficients, which returns an overall increasing effect of the HMP susceptibility
 635 as the rainfall increases (see Figure 3 and note the following median values: $\beta_{RAIN_T\mu_S\mu} =$
 636 -0.75 , $\beta_{RAIN_T\sigma_S\mu} = 0.80$, $\beta_{RAIN_TA_SA} = -0.04$, $\beta_{AnnualRAIN_S\mu} = 0.67$).

637 As for the temperature, the effect is much clearer there, as all the three summary statistics
 638 derived from the original spatio-temporal temperature signal appear to have a negative con-
 639 tribution to the model. This means that at increasing temperatures the probability of HMP
 640 occurrences consistently decreases in space and time, irrespective of the three components
 641 at hand.

642 We also stress here the relevance of antecedent 3-years disasters. This idea stems from the

643 fact that certain types of hazard persist or cluster both in time and space, hence by featuring
644 antecedent occurrences in the model can help predicting future HMPs. This concept is not
645 new in landslide studies (Samia *et al.*, 2018, 2020), although a similar approach has not been
646 tested yet when modeling HMPs statistically.

647 An additional and equally interesting contribution to the model was carried by human
648 interference. Other researchers have already pointed out a similar consideration (Bronstert,
649 2003; Plate, 2002), which we tested in this work by including the presence of build-up area
650 per catchment and per year (Marconcini *et al.*, 2020b). The expansion of human settlements
651 has a dual effect in our model. On the one hand, it undeniably constitutes an interference
652 with the environment, potentially leading to HMP occurrences (Duncan, 2013). On the
653 other hand, human expansion also means that a larger number of people are being exposed
654 to disasters (Cutter *et al.*, 2018). This in turns may bring some degree of bias in the HMP
655 inventory because events that occur in non-inhabited areas may not be recorded, especially
656 due to the size of the study area. Conversely, events that occur in inhabited areas are much
657 more likely to be recorded.

658 As regards the temporal validation scheme we tested, it is important to justify why we
659 chose Mod2 as our best and further presented it to the readers. When looking at performance,
660 not only the central tendency (mean or median) but also the variation around it constitute a
661 relevant criterion. The variation is essentially described as the difference between the highest
662 and lowest performance. Among the two terms, we chose the lowest performance, together
663 with the median AUC, to be our primary mean of selecting the best temporal validation
664 scheme. In fact, in an ideal situation one should avoid selecting a model that can poorly
665 perform even as rare as this may occur. Therefore, MOD2 has become our best validation
666 scheme for it both provides a median value comparable to MOD1, MOD3 and MOD4. And,
667 it returned a minimum AUC much higher than the other temporal validation routines.

668 In terms of covariates' influence on HMP susceptibility, MOD2 offers a slightly different
669 perspective than the first exploratory tests. The morphological characteristics of the catch-
670 ments largely contribute to the HMP susceptibility (see form factor and elongation ratio
671 in Figure 10. And even more interestingly, $RAIN-T_{\sigma}-S_{\mu}$ and $AnnualRAIN-S_{\mu}$ not only
672 dominate the probability estimates to a much larger extent than any other covariate. But,
673 they also show a quite distinctive increasing trend through time.

674 Ultimately, we decided to remind the reader that the susceptibility we estimated for
675 the whole Chinese territory is actually much finer in resolution than what it looks like in
676 the previous figures. To do this, we have selected eight important and large catchments.
677 In Figure 14 we show the HMP susceptibility estimated for each of those catchments via
678 MOD2, and specifically through the maximum probability of HMP occurrence shown in
679 Figure 12. Looking at Figure 14 becomes evident that our model is built on a very detailed
680 spatial partition of the Chinese landscape. And, within each of the eight selected major
681 catchments, it is possible to further distinguish susceptible sub-catchments that upon which
682 local administrations can made decisions to ensure the safety of local inhabitants.

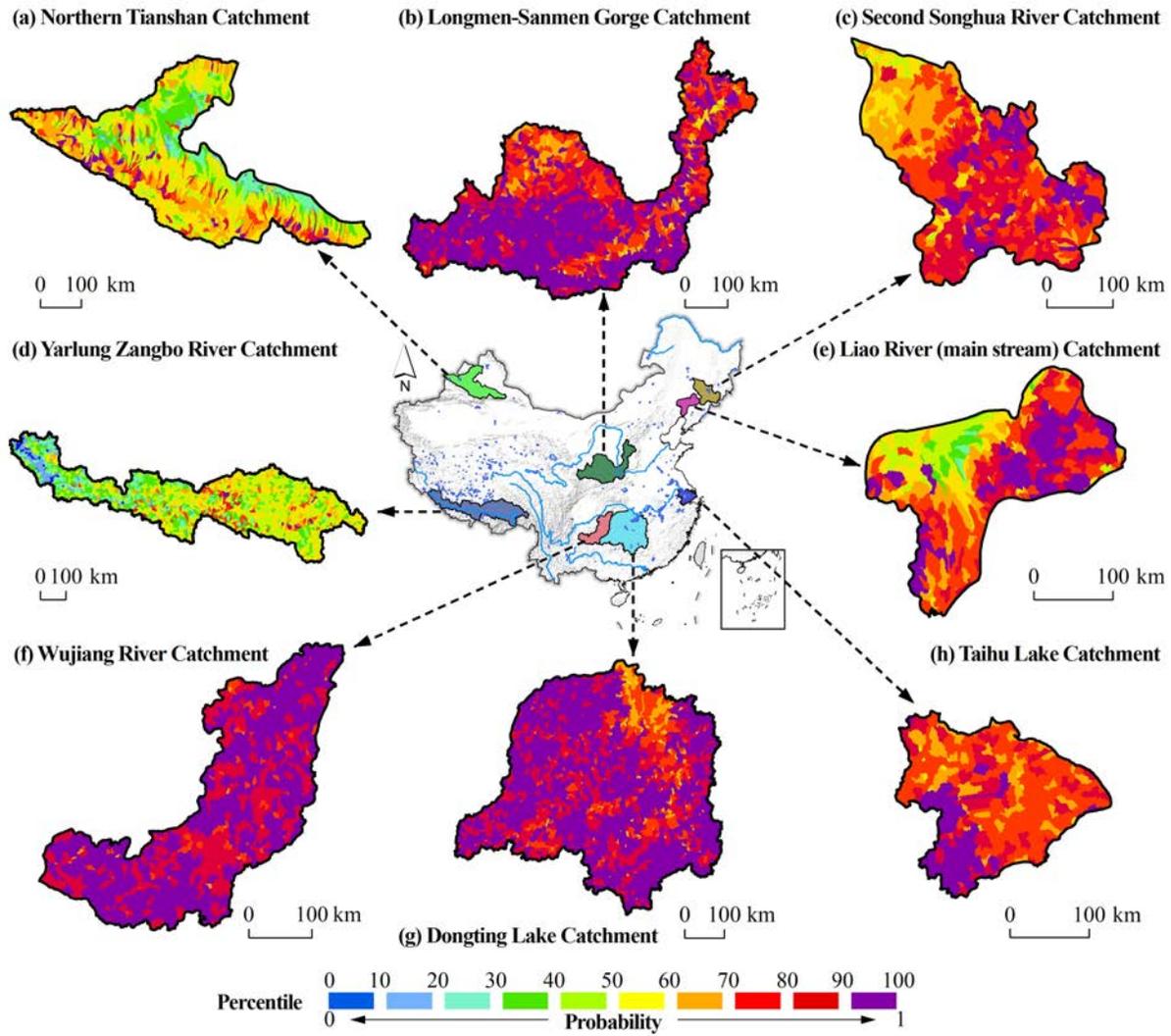


Figure 14: Details of specific large catchments across the Chinese territory. The HMP susceptibility corresponds to the maximum probability estimated via MOD2 between 1986 and 2015, this being shown in Figure 12. The catchments we report here for graphical purposes are: (a) Northern Tianshan Catchment; (b) Longmen-Sanmen Gorge Catchment; (c) Second Songhua River Catchment; (d) Yarlung Zangbo River Catchment; (e) Liao River (main stream) Catchment; (f) Wujiang River Catchment; (g) Dongting Lake Catchment; (h) Taihu Lake Catchment.

683 5.2 Opposing arguments

684 The model we present is both spatial and temporal in nature. Among the suitable space-
685 time models we have chosen a relatively simple one, a binomial GLM. Instead of this, we
686 could have opted for a binomial GAM (Generalized Additive Model) extension to account
687 for possible nonlinear covariates' effects. And, to include potential variables acting at a
688 latent level, hence requiring complex CAR or SPDE components to be featured as well.
689 We maintain that our choice has proven to be a valid option, for both our spatio-temporal
690 cross-validation and temporal validation schemes returned AUC values well above 0.8, the
691 threshold for excellent binary classifiers according to [Hosmer and Lemeshow \(2000\)](#).

692 Some may argue that on a 30-year long record, the accuracy of the inventory may have
693 drastically changed in recent times. As a result, the inventory may be biased towards a
694 larger number of HMP records at the end of the time series. We maintain that the HMP
695 inventory is reliable and should not be affected by this type of bias. In fact, the Chinese
696 government has supported the initiative of creating this inventory long before the 1980'ies.
697 And, by the starting year of our time series (1985), the recording protocol had already been
698 standardised at the whole Chinese territory scale. Surely, we cannot entirely disregard the
699 possibility of some sort of bias due to the size of the study area. We know for a fact (and
700 already shared the information with the readers in [Section 4](#)) that the western sector of the
701 Chinese territory is devoid of large settlements. This may imply that the lack of HMP record
702 in the region and the subsequent low susceptibility estimated there (see [Figures 11 and 12](#))
703 could be due (to some extent) to a lack of interest rather than a real absence of HMPs.
704 We have tried to investigate this potential issue, by checking local news and other source
705 of information. But, we have not found records of HMPs in the region. Thus, we can only
706 assume the inventory to be reliable.

707 The spatio-temporal dataset we have built features a much larger number of catchments
708 without HMP records rather than catchments recorded with actual HMP events. In other
709 words, we used an unbalanced presence/absence dataset. In turn, this affects the estimated
710 probability spectrum, resulting in a positively skewed susceptibility distribution. We main-
711 tain that this is a realistic pattern for such a vast spatiotemporal domain where probabilities
712 of HMP are generally very low, with the exception of few catchments that are very suscep-
713 tible (see [Frattini *et al.*, 2010](#)). However, it is worth mentioning that the geomorphological
714 community often operates with a balanced dataset of presence and absence cases. This
715 in turn makes the probability spectrum much more normally-distributed and centered at
716 around 0.5 ([Rossi *et al.*, 2010](#)). Both approaches are valid, although creating a balanced
717 presence/absence dataset distorts the global intercept ([Lombardo and Mai, 2018](#)) making
718 the interpretation of the probabilities valid only in a relative sense rather than the common
719 notion of probability available in any other statistical application ([Petschko *et al.*, 2014](#)).
720 Therefore, we maintain that our unbalanced choice is valid and suitable to treat such a large
721 spatiotemporal domain.

6 Conclusions

The Chinese territory hosts a vast and diverse landscape that in the last thirty years has been struck by thousands of hydro-morphological processes. Such processes, spanning from debris flows to debris floods and floods have been monitored and recorded in a multi-temporal digital archive thanks to a Chinese program centrally coordinated but enacted by local administrations. In this work, we explore and exploit this archive to produce the first catchment-scale-based HMP susceptibility model of China, from 1985 to 2015.

We distinguished seven macro-regions where the estimated probability of HMP occurrence can be interpreted and explained. The south-eastern regions are the most susceptible to HMPs, primarily because of the monsoon control on precipitation regimes. This observation of a spatial patterns is strictly entwined with the temporal observation that the susceptibility estimates tend to increase in recent years. This may be due to the fact that climatic changes have narrowed the duration of storms and increased their intensity. This is literally a recipe for disasters, especially in the case of HMP.

For transparency, we are sharing the shapefile containing the susceptibility estimates for each year under consideration. Although we cannot share the raw data, the method we propose is certainly reproducible in any other geographic context. For this reason, we consider our work an example of continental-decadal scale HMP susceptibility. We stress here that other examples currently present in the literature have all been built by using a grid-cell partition of space, where each grid-cell has a resolution in the order of kilometers. Therefore, their actual use is hindered by the fact that over several kilometers, the landscape can contextually feature floodplains as well as mountain ridges. Here we presented a multi-temporal HMP susceptibility model built at the catchment scale. Therefore, the information we provide is expressed at a scale that respects the geomorphology and hydrology of the phenomena under consideration. We consider this work a first order indicator of catchments under threat though. And, we expect the operational use of this information to be exploited in a second stage, where physically-based models will be run for catchments with a large probability of HMP occurrence.

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References

- 755
- 756 Adnan, M. S. G., Dewan, A., Zannat, K. E. and Abdullah, A. Y. M. (2019) The use of water-
757 shed geomorphic data in flash flood susceptibility zoning: a case study of the Karnaphuli
758 and Sangu river basins of Bangladesh. *Natural Hazards* **99**(1), 425–448.
- 759 Blöschl, G., Kiss, A., Viglione, A., Barriendos, M., Böhm, O., Brázdil, R., Coeur, D.,
760 Demarée, G., Llasat, M. C., Macdonald, N. *et al.* (2020) Current European flood-rich
761 period exceptional compared with past 500 years. *Nature* **583**(7817), 560–566.
- 762 Borga, M., Boscolo, P., Zanoni, F. and Sangati, M. (2007) Hydrometeorological analysis of
763 the 29 August 2003 flash flood in the Eastern Italian Alps. *Journal of Hydrometeorology*
764 **8**(5), 1049–1067.
- 765 Bout, B., Lombardo, L., van Westen, C. and Jetten, V. (2018) Integration of two-phase
766 solid fluid equations in a catchment model for flashfloods, debris flows and shallow slope
767 failures. *Environmental Modelling & Software* **105**, 1–16.
- 768 Bronstert, A. (2003) Floods and climate change: interactions and impacts. *Risk Analysis:*
769 *An International Journal* **23**(3), 545–557.
- 770 Budimir, M., Atkinson, P. and Lewis, H. (2015) A systematic review of landslide probability
771 mapping using logistic regression. *Landslides* **12**(3), 419–436.
- 772 Cama, M., Lombardo, L., Conoscenti, C., Agnesi, V. and Rotigliano, E. (2015) Predicting
773 storm-triggered debris flow events: application to the 2009 Ionian Peloritan disaster (Sicily,
774 Italy). *Nat Hazards Earth Syst Sci* **15**(8), 1785–1806.
- 775 Carrara, A., Cardinali, M., Detti, R., Guzzetti, F., Pasqui, V. and Reichenbach, P. (1991)
776 GIS techniques and statistical models in evaluating landslide hazard. *Earth Surface Pro-*
777 *cesses and Landforms* **16**(5), 427–445.
- 778 Carrara, A., Cardinali, M., Guzzetti, F. and Reichenbach, P. (1995) Gis technology in map-
779 ping landslide hazard. In *Geographical Information Systems in Assessing Natural Haz-*
780 *ards*, Advances in Natural and Technological Hazards Research, pp. 135–175. Dordrecht:
781 Kluwer, Springer. ISBN 978-90-481-4561-4 978-94-015-8404-3.
- 782 Cenci, L., Laiolo, P., Gabellani, S., Campo, L., Silvestro, F., Delogu, F., Boni, G. and Rudari,
783 R. (2016) Assimilation of H-SAF soil moisture products for flash flood early warning
784 systems. Case study: Mediterranean catchments. *IEEE Journal of Selected Topics in*
785 *Applied Earth Observations and Remote Sensing* **9**(12), 5634–5646.
- 786 Collier, C. (2007) Flash flood forecasting: What are the limits of predictability? *Quarterly*
787 *Journal of the Royal Meteorological Society: A journal of the atmospheric sciences, applied*
788 *meteorology and physical oceanography* **133**(622), 3–23.

- 789 Crema, S., Marra, F., Andreoli, A., Scorpio, V., Kofler, C., Cavalli, M., Marchi, L., Borga,
790 M. and Comiti, F. (2018) Integrating high-resolution hydrology and geomorphometry for
791 flash flood characterization. *EGUGA* p. 13142.
- 792 Cutter, S. L., Emrich, C. T., Gall, M. and Reeves, R. (2018) Flash flood risk and the paradox
793 of urban development. *Natural Hazards Review* **19**(1), 05017005.
- 794 Duncan, J. M. (2013) Slope stability then and now. In *Geo-Congress 2013: Stability and*
795 *Performance of Slopes and Embankments III*, pp. 2184–2203.
- 796 Efron, B. and Tibshirani, R. J. (1994) *An introduction to the bootstrap*. CRC press.
- 797 Frattini, P., Crosta, G. and Carrara, A. (2010) Techniques for evaluating the performance
798 of landslide susceptibility models. *Engineering Geology* **111**(1), 62–72.
- 799 Gaume, E., Bain, V., Bernardara, P., Newinger, O., Barbuc, M., Bateman, A., Blaškovičová,
800 L., Blöschl, G., Borga, M., Dumitrescu, A. *et al.* (2009) A compilation of data on European
801 flash floods. *Journal of Hydrology* **367**(1-2), 70–78.
- 802 Glade, T. and Crozier, M. J. (2005) Landslide hazard and risk: concluding comment and
803 perspectives. *Landslide hazard and risk*. Wiley, Chichester pp. 767–774.
- 804 Gong, P., Li, X. and Zhang, W. (2019) 40-Year (1978–2017) human settlement changes in
805 China reflected by impervious surfaces from satellite remote sensing. *Science Bulletin*
806 **64**(11), 756–763.
- 807 Gourley, J. J., Flamig, Z. L., Vergara, H., Kirstetter, P.-E., Clark III, R. A., Argyle, E.,
808 Arthur, A., Martinaitis, S., Terti, G., Erlingis, J. M. *et al.* (2017) The FLASH Project:
809 improving the tools for flash flood monitoring and prediction across the United States.
810 *Bulletin of the American Meteorological Society* **98**(2), 361–372.
- 811 Gourley, J. J., Hong, Y., Flamig, Z. L., Arthur, A., Clark, R., Calianno, M., Ruin, I., Ortel,
812 T., Wiczorek, M. E., Kirstetter, P.-E. *et al.* (2013) A unified flash flood database across
813 the United States. *Bulletin of the American Meteorological Society* **94**(6), 799–805.
- 814 Guo, L., He, B., Ma, M., Chang, Q., Li, Q., Zhang, K. and Hong, Y. (2018) A comprehensive
815 flash flood defense system in China: overview, achievements, and outlook. *Natural Hazards*
816 **92**(2), 727–740.
- 817 Guzzetti, F., Mondini, A. C., Cardinali, M., Fiorucci, F., Santangelo, M. and Chang, K.-T.
818 (2012) Landslide inventory maps: New tools for an old problem. *Earth-Science Reviews*
819 **112**(1-2), 42–66.
- 820 Hosmer, D. W. and Lemeshow, S. (2000) *Applied Logistic Regression*. Second edition. New
821 York: Wiley.

- 822 Javelle, P., Fouchier, C., Arnaud, P. and Lavabre, J. (2010) Flash flood warning at un-
823 gauged locations using radar rainfall and antecedent soil moisture estimations. *Journal of*
824 *Hydrology* **394**(1-2), 267–274.
- 825 Jiao, Y., Zhao, D., Ding, Y., Liu, Y., Xu, Q., Qiu, Y., Liu, C., Liu, Z., Zha, Z. and Li, R.
826 (2019) Performance evaluation for four GIS-based models purposed to predict and map
827 landslide susceptibility: A case study at a world heritage site in Southwest China. *Catena*
828 **183**, 104221.
- 829 Karagiorgos, K., Thaler, T., Heiser, M., Hübl, J. and Fuchs, S. (2016) Integrated flash flood
830 vulnerability assessment: insights from East Attica, Greece. *Journal of Hydrology* **541**,
831 553–562.
- 832 Li, H., Lei, X., Shang, Y. and Qin, T. (2018) Flash flood early warning research in China.
833 *International Journal of Water Resources Development* **34**(3), 369–385.
- 834 Li, H., Wang, D., Singh, V. P., Wang, Y., Wu, J., Wu, J., Liu, J., Zou, Y., He, R. and
835 Zhang, J. (2019a) Non-stationary frequency analysis of annual extreme rainfall volume
836 and intensity using Archimedean copulas: A case study in eastern China. *Journal of*
837 *Hydrology* **571**, 114–131.
- 838 Li, W., Lin, K., Zhao, T., Lan, T., Chen, X., Du, H. and Chen, H. (2019b) Risk assessment
839 and sensitivity analysis of flash floods in ungauged basins using coupled hydrologic and
840 hydrodynamic models. *Journal of Hydrology* **572**, 108–120.
- 841 Liao, Z., Hong, Y., Kirschbaum, D. and Liu, C. (2012) Assessment of shallow landslides from
842 Hurricane Mitch in central America using a physically based model. *Environmental Earth*
843 *Sciences* **66**(6), 1697–1705.
- 844 Liu, C., Guo, L., Ye, L., Zhang, S., Zhao, Y. and Song, T. (2018) A review of advances in
845 China’s flash flood early-warning system. *Natural Hazards* **92**(2), 619–634.
- 846 Lombardo, L., Bakka, H., Tanyas, H., van Westen, C., Mai, P. M. and Huser, R. (2019) Geo-
847 statistical modeling to capture seismic-shaking patterns from earthquake-induced land-
848 slides. *Journal of Geophysical Research: Earth Surface* **124**(7), 1958–1980.
- 849 Lombardo, L., Cama, M., Conoscenti, C., Märker, M. and Rotigliano, E. (2015) Binary
850 logistic regression versus stochastic gradient boosted decision trees in assessing landslide
851 susceptibility for multiple-occurring landslide events: application to the 2009 storm event
852 in Messina (Sicily, southern Italy). *Natural Hazards* **79**(3), 1621–1648.
- 853 Lombardo, L., Fubelli, G., Amato, G. and Bonasera, M. (2016) Presence-only approach to
854 assess landslide triggering-thickness susceptibility: a test for the Mili catchment (north-
855 eastern Sicily, Italy). *Natural Hazards* **84**(1), 565–588.

- 856 Lombardo, L. and Mai, P. M. (2018) Presenting logistic regression-based landslide suscepti-
857 bility results. *Engineering geology* **244**, 14–24.
- 858 Lombardo, L., Opitz, T., Ardizzone, F., Guzzetti, F. and Huser, R. (2020) Space-time
859 landslide predictive modelling. *Earth-Science Reviews* p. 103318.
- 860 Lombardo, L. and Tanyas, H. (2020) Chrono-validation of near-real-time landslide suscepti-
861 bility models via plug-in statistical simulations. *Engineering Geology* p. 105818.
- 862 Mahmood, M. I., Elagib, N. A., Horn, F. and Saad, S. A. (2017) Lessons learned from Khar-
863 toum flash flood impacts: An integrated assessment. *Science of the Total Environment*
864 **601**, 1031–1045.
- 865 Marchi, L., Borga, M., Preciso, E. and Gaume, E. (2010) Characterisation of selected extreme
866 flash floods in Europe and implications for flood risk management. *Journal of Hydrology*
867 **394**(1-2), 118–133.
- 868 Marconcini, M., Gorelick, N., Metz-Marconcini, A. and Esch, T. (2020a) Accurately mon-
869 itoring urbanization at global scale—the world settlement footprint. In *IOP Conference*
870 *Series: Earth and Environmental Science*, volume 509, p. 012036.
- 871 Marconcini, M., Metz-Marconcini, A., Üreyen, S., Palacios-Lopez, D., Hanke, W., Bachofer,
872 F., Zeidler, J., Esch, T., Gorelick, N., Kakarla, A. *et al.* (2020b) Outlining where humans
873 live, the World Settlement Footprint 2015. *Scientific Data* **7**(1), 1–14.
- 874 Norbiato, D., Borga, M., Degli Esposti, S., Gaume, E. and Anquetin, S. (2008) Flash flood
875 warning based on rainfall thresholds and soil moisture conditions: An assessment for
876 gauged and ungauged basins. *Journal of Hydrology* **362**(3-4), 274–290.
- 877 O’Banion, M. and Olsen, M. (2014) Predictive seismically-induced landslide hazard mapping
878 in oregon using a maximum entropy model (MaxEnt). In *Proceedings of the 10th national*
879 *conference in earthquake engineering, Earthquake Engineering Research Institute, Anchor-*
880 *age*.
- 881 Paprotny, D., Morales-Nápoles, O. and Jonkman, S. N. (2017) Efficient pan-European river
882 flood hazard modelling through a combination of statistical and physical models. *Natural*
883 *Hazards and Earth System Sciences* **17**(7), 1267.
- 884 Park, N.-W. (2015) Using maximum entropy modeling for landslide susceptibility mapping
885 with multiple geoenvironmental data sets. *Environmental Earth Sciences* **73**(3), 937–949.
- 886 Petschko, H., Brenning, A., Bell, R., Goetz, J. and Glade, T. (2014) Assessing the quality
887 of landslide susceptibility maps—case study lower austria. *Natural Hazards and Earth*
888 *System Sciences* **14**(1), 95–118.

- 889 Plate, E. J. (2002) Flood risk and flood management. *Journal of Hydrology* **267**(1-2), 2–11.
- 890 Prein, A. F., Rasmussen, R. M., Ikeda, K., Liu, C., Clark, M. P. and Holland, G. J. (2017)
891 The future intensification of hourly precipitation extremes. *Nature Climate Change* **7**(1),
892 48–52.
- 893 Ragettli, S., Zhou, J., Wang, H., Liu, C. and Guo, L. (2017) Modeling flash floods in
894 ungauged mountain catchments of China: A decision tree learning approach for parameter
895 regionalization. *Journal of Hydrology* **555**, 330–346.
- 896 Rahmati, O., Kornejady, A., Samadi, M., Deo, R. C., Conoscenti, C., Lombardo, L., Dayal,
897 K., Taghizadeh-Mehrjardi, R., Pourghasemi, H. R., Kumar, S. *et al.* (2019) PMT: New an-
898 alytical framework for automated evaluation of geo-environmental modelling approaches.
899 *Science of the total environment* **664**, 296–311.
- 900 Ramos-Bernal, R. N., Vázquez-Jiménez, R., Tizapa, S. S. and Matus, R. A. (2019) Char-
901 acterization of Susceptible Landslide Zones by an Accumulated Index. In *Landslides*.
902 IntechOpen.
- 903 Reichenbach, P., Rossi, M., Malamud, B. D., Mihir, M. and Guzzetti, F. (2018) A review of
904 statistically-based landslide susceptibility models. *Earth-Science Reviews* **180**, 60–91.
- 905 Rossi, M., Guzzetti, F., Reichenbach, P., Mondini, A. C. and Peruccacci, S. (2010) Optimal
906 landslide susceptibility zonation based on multiple forecasts. *Geomorphology* **114**(3), 129–
907 142.
- 908 Rossi, M., Guzzetti, F., Salvati, P., Donnini, M., Napolitano, E. and Bianchi, C. (2019) A
909 predictive model of societal landslide risk in Italy. *Earth-Science Reviews* **196**, 102849.
- 910 Rozalis, S., Morin, E., Yair, Y. and Price, C. (2010) Flash flood prediction using an uncal-
911 ibrated hydrological model and radar rainfall data in a Mediterranean watershed under
912 changing hydrological conditions. *Journal of Hydrology* **394**(1-2), 245–255.
- 913 Samia, J., Temme, A., Bregt, A., Wallinga, J., Guzzetti, F. and Ardizzone, F. (2020) Dy-
914 namic path-dependent landslide susceptibility modelling. *Natural Hazards and Earth Sys-*
915 *tem Sciences* **20**(1), 271–285.
- 916 Samia, J., Temme, A., Bregt, A. K., Wallinga, J., Stuijver, J., Guzzetti, F., Ardizzone, F.
917 and Rossi, M. (2018) Implementing landslide path dependency in landslide susceptibility
918 modelling. *Landslides* **15**(11), 2129–2144.
- 919 Samia, J., Temme, A. J., Bregt, A., Wallinga, J., Guzzetti, F., Ardizzone, F. and Rossi,
920 M. (2017) Do Landslides Follow Landslides? Insights in Path Dependency from a Multi-
921 Temporal Landslide Inventory. *Landslides* **14**, 547–558.

- 922 Shen, X., Anagnostou, E. N., Mei, Y. and Hong, Y. (2017) A global distributed basin
923 morphometric dataset. *Scientific Data* **4**(1), 1–8.
- 924 Špitalar, M., Gourley, J. J., Lutoff, C., Kirstetter, P.-E., Brilly, M. and Carr, N. (2014)
925 Analysis of flash flood parameters and human impacts in the US from 2006 to 2012.
926 *Journal of Hydrology* **519**, 863–870.
- 927 Tramblay, Y., Bouvier, C., Martin, C., Didon-Lescot, J.-F., Todorovik, D. and Domergue,
928 J.-M. (2010) Assessment of initial soil moisture conditions for event-based rainfall–runoff
929 modelling. *Journal of Hydrology* **387**(3-4), 176–187.
- 930 Wang, N., Cheng, W., Wang, B., Liu, Q. and Zhou, C. (2020) Geomorphological regionaliza-
931 tion theory system and division methodology of China. *Journal of Geographical Sciences*
932 **30**(2), 212–232.
- 933 Westra, S., Fowler, H., Evans, J., Alexander, L., Berg, P., Johnson, F., Kendon, E.,
934 Lenderink, G. and Roberts, N. (2014) Future changes to the intensity and frequency
935 of short-duration extreme rainfall. *Reviews of Geophysics* **52**(3), 522–555.