

# The relationship between lake surface area and maximum depth

B. B. Cael<sup>1</sup> and D. Seekell<sup>2,3</sup>

1. National Oceanography Centre, Southampton, UK. cael@noc.ac.uk 2. Climate Impacts Research Centre, Umeå University, Abisko, Sweden. 3. Department of Ecology and Environmental Science, Umeå University, Umeå, Sweden. david.seekell@umu.se

This paper is a non-peer reviewed preprint submitted to EarthArXiv. This paper is in review at Limnology Oceanography: Letters.

*Author Contributions:* B. B. Cael and D. Seekell conceptualized the paper. B. B. Cael conducted the analysis. B. B. Cael and D. Seekell wrote the paper.

To whom correspondence should be addressed: cael@noc.ac.uk and david.seekell@umu.se.

*Article Type:* Letter.

*Running Header:* Fractal scaling of lake maximum depth.

*Keywords:* Lake Morphometry | Maximum Depth | Fractal Brownian Motion | Hurst Exponent | Factal Scaling | Global Limnology

*Data Availability:* Data are publicly available at the reference cited in the text; code will be given a Zenodo DOI before publication should this manuscript be accepted, including an example of how to upscale methane fluxes.

*Abstract:* Maximum depth varies among lakes from <1 to 1741 meters, but attempts to explain this variation have achieved little predictive power. In this paper, we describe the probability distribution of maximum depths based on recent developments in the theory of fractal Brownian motions. The theoretical distribution is right-tailed and adequately captures variations in maximum depth in a dataset of 8,164 lakes (maximum depths 0.1 to 135 meters) from the northeastern United States. Maximum depth increases with surface area, but with substantial random variation - the 95% prediction interval spans more than an order of magnitude for lakes with any specific surface area. Our results explain the observed variability in lake maximum depths, capture the link between topographic characteristics and lake bathymetry, and provide a means to upscale maximum-depth-dependent processes, which we illustrate by upscaling the diffusive flux of methane from northern lakes to the atmosphere.

*Significance Statement:* Many foundational ecosystem characteristics in lakes, such as thermal stratification, are constrained by maximum depth. Aquatic scientists have struggled to both explain this variation and to predict maximum depths for lakes without detailed bathymetric surveys. The morphometry of collections of lakes has previously been shown to follow theoretical predictions that assume Earth's topography approximates a fractal Brownian motion. In this study, we demonstrate how these theories relate maximum depth to surface area. Maximum depth increases with surface area, but massive lake-to-lake variation is an inherent characteristic of this pattern. This is because the maximum depth is a single random displacement on a topographic profile. The probability distribution of maximum depths can be used to upscale the patterns and processes for many lakes at the regional or global scale. We demonstrate this for diffusive flux of methane from temperate lakes to the atmosphere - a process that correlates strongly with maximum depth but not surface area and cannot be up-scaled without the results presented in this paper.

---

## 1 Introduction

2 Maximum depth varies among lakes between  $\sim 0.1$  and 1741 meters [27, 13]. This variation engen-  
3 ders patterns of diverse ecosystem characteristics including mixing and thermal stratification [15], the  
4 relative sizes of littoral and pelagic habitats [22], and carbon cycling including methane evasion [16].  
5 However, there is a paucity of bathymetric data relative to the global abundance of lakes, and em-  
6 pirical relationships that relate maximum depth to lake and landscape characteristics consistently fail  
7 to develop sufficient predictive power to estimate maximum depth in lakes that have not been depth  
8 sounded [26]. In particular, there is need to develop scaling relationships that relate maximum depth  
9 to lake characteristics that are easily measured across broad geographic regions, such as surface area  
10 (e.g., [2, 23]). Such relationships provide the simple rules used to generalize understanding of aquatic  
11 ecosystem patterns and processes at regional to global scales [7].

12 Bathymetric surveys are time consuming, and prohibitively expensive for large numbers of lakes [26,  
13 14]. While global perspectives have come to dominate the aquatic sciences over the last twenty years,  
14 the proportion of lakes depth sounded has remained low, preventing up-scaling of empirical results  
15 from local to global scales [26, 21, 8]. This has spurred a series of empirical studies seeking to  
16 predict lake-specific maximum depth based on surface area and other easily mapped characteristic,  
17 typically some metric of vertical relief within a buffer zone around each lake (e.g., [21, 14, 27, 19, 18,  
18 12]). Typically, these studies assume that larger lakes should be deeper than smaller lakes, and that  
19 integrative measures of topography (i.e. variance, slope) should relate to lake bathymetry. However,  
20 these might not be reliable assumptions. For example, while the deepest lakes all have large surface  
21 areas ( $\geq 500 \text{ km}^2$ ), many lakes with large surface areas are remarkably shallow, often only a few meters  
22 deep [13]. Further, correlations between surface area and maximum depth are often weak. For example,  
23 the Pearson correlation between the logarithms of surface area and maximum depth for Canadian lakes  
24 is  $r = 0.46$  [19]. Based on this correlation, the probability of a larger lake being deeper than a smaller  
25 lake is only slightly better ( $p = 0.65$ ) than a fair coin flip ( $p = 0.5$ ), if the two lakes are selected at  
26 random [9]. Additionally, it is not clear that topography should predict maximum depth. Specifically,  
27 many lakes have relatively sudden and catastrophic origins that may not reflect the processes shaping  
28 topography [29]. Additionally, maximum depth is essentially a single random perturbation along a  
29 combined topographic-bathymetric profile [26]. While integrative measures of topography may be able  
30 to predict integrative measures of bathymetry, there is no clear reason to believe that they should  
31 accurately predict the value of a single random displacement [26]. The consequence of this is highly  
32 uncertain predictions that preclude the application of these equations for upscaling. For example, the  
33 error ratio for maximum depth predictions for Canadian lakes is approximately two [3, 19]. This means  
34 that a lake predicted to have a maximum depth of 10 meters will have a maximum depth between 5 and  
35 20 meters [3]. Error ratios are even higher in other studies (e.g., [12]). This magnitude of prediction  
36 error is material across the entire size spectra of lakes.

37 In this paper, we describe the theoretical basis for predicting maximum depth from surface area.  
38 This theory is based on the characteristics of fractional Brownian motions, which approximate key  
39 features of Earth's topography and are the basis for other lake scaling relationships including mean  
40 depth and volume (e.g., [11, 23, 2, 25]). Specifically, we describe how recent developments generalizing  
41 the arcsine laws from Brownian motions to fractional Brownian motions relate to the problem of  
42 maximum depths [4, 5]. We test goodness-of-fit of the theoretical distribution derived based on these

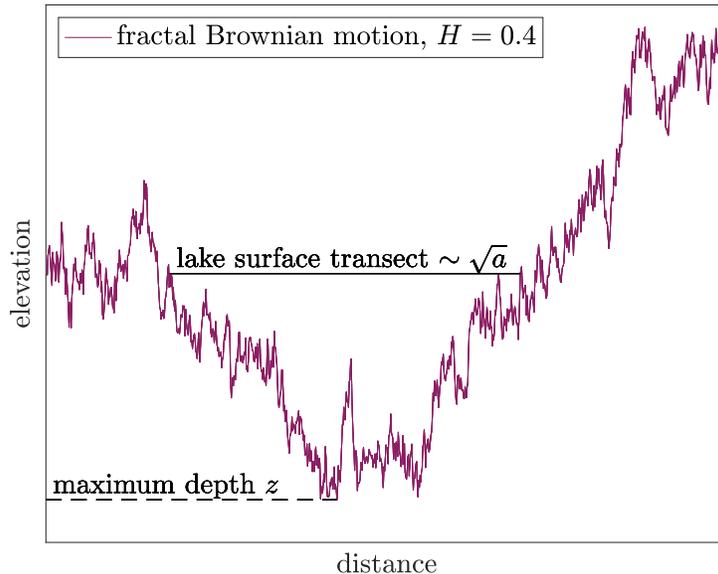


Figure 1: An example fractal Brownian motion path, with Hurst exponent  $H = 0.4$ . This path is taken to be an analog of a transect along a landscape; lakes fill in sections along this transect, which thus sets their maximum depth.

43 theoretical developments to a bathymetry database with more than 8000 lakes. Finally, demonstrate  
 44 how this distribution can be applied to advance understanding of global lake characteristics by up-  
 45 scaling diffusive methane flux from temperate lakes to the atmosphere, a process that has previously  
 46 been shown to correlate with maximum depth but not surface area, and hence a process that is difficult  
 47 to upscale without the results presented in this paper [16]. Collectively, these results both advance  
 48 fundamental understanding of patterns of lake morphometry, and provide tools for upscaling depth  
 49 dependent processes when seeking to place lakes within a global context.

## 50 Theory

51 Earth's topography is approximately scale-invariant [30]. This is self-evident for many landforms  
 52 including lakes. For example, it would be difficult or impossible to determine if a lake is 1 ha or 1,000  
 53 ha on a map or image without a scale reference (e.g., a scale bar on a map or a house on an aerial  
 54 image) because the key characteristics of lake geometry are similar across wide ranges of scales [25,  
 55 26, 23]. In particular, Earth's topography is self-affine, which implies different scaling in the vertical  
 56 and horizontal directions [30]. This is also self-evident; when walking or driving away from a mountain

57 range, the profile flattens more rapidly than it compresses horizontally [26]. Specifically, the Hurst  
 58 exponent  $H$  captures the relationship between horizontal and vertical scales. For a surface  $Z(\vec{X})$ ,  
 59 rescaling by any coefficient  $b$  conforms, in a statistical sense, to  $b^{-H}Z(b\vec{X}) = Z(\vec{X})$ .  $H$  takes values in  
 60 the range  $(0,1)$ . When  $H = \frac{1}{2}$ , transects across  $Z$  have the statistical properties of a one-dimensional  
 61 Brownian motion; when  $H > \frac{1}{2}$  ( $< \frac{1}{2}$ ), surfaces are smoother (rougher) than Brownian trajectories  
 62 (e.g., [11]). These characteristics engender the derivation of empirically robust and theoretically sound  
 63 power-law relationships between lake surface area and various aspects of lake morphometry including  
 64 abundance, volume, mean depth, perimeter, and hydrological connectivity [2, 23, 25, 23]. However,  
 65 to our knowledge, the relationship between maximum depth and surface area has not been described  
 66 [26].

67 When topography approximates a Brownian motion ( $H = \frac{1}{2}$ ), the maximum depth over a given interval  
 68 is described by the third arcsine law [4]. Essentially, because the maximum depth is a single random  
 69 displacement on a topographic profile, the arcsine law shows that maximum depth converges as a  
 70 probability distribution based on surface area. Recently, this result has been generalized to cases where  
 71  $H \neq \frac{1}{2}$  [4, 5]. This generalization allows application of the arcsine law to predict the distribution of lake  
 72 maximum depth  $p(z)$  based on surface area and the Hurst coefficient. Unfortunately, the mathematical  
 73 form of  $p(z)$  is almost comically complex. First, define the normalized maximum depth

$$y = \frac{z}{\sqrt{2}L^H}$$

74 and define the probability distribution in terms of  $y$ , i.e.  $p(y)$  (later we recast this in terms of  $z$  as  
 75 well). Doing so,

$$p(z) = \frac{1}{\sqrt{2}L^H} f(y)$$

76 where  $\frac{1}{\sqrt{2}L^H}$  is a normalization constant that ensures the total probability integrates to one, and  $f(y)$   
 77 has the form:

$$f(y) = \sqrt{\frac{2}{\pi}} e^{-y/2} y^{(1/H)-2} e^{(H-\frac{1}{2})(4 \ln y + \mathcal{G}(y))}$$

78 and  $\mathcal{G}(y)$  is the complex expression:

$$\mathcal{G}(y) = \frac{y^4}{6} {}_2F_2\left(1, 1; \frac{5}{2}, 3; \frac{y^2}{2}\right) - 3y^2 + \pi(1 - y^2)\operatorname{erfi}\left(\frac{y}{\sqrt{2}}\right) + \sqrt{2\pi}e^{y^2/2}y + (y^2 - 2)(\gamma_E + \ln(2y^2))$$

79 where  $F$  is the hypergeometric function,  $\operatorname{erfi}$  is the imaginary error function, and  $\gamma_E$  is the Euler-  
80 Mascheroni constant [4, 5].

81 To date, the arcsine laws for fractal Brownian motions are only proven for the one-dimensional case.  
82 Therefore, lake areas must be recast as one-dimensional lengthscales  $L$  to empirically test the theo-  
83 retical depth distribution. The lengthscale  $L$  is in a sense the length of a randomly chosen horizontal  
84 transect along a lake's bathymetry, which includes the lake's maximum depth (solid black line in Figure  
85 1). As surface area  $a$  [ $\text{m}^2$ ] is the fundamental horizontal metric for lakes, it would be preferable to  
86 cast  $L$  in terms of  $a$ ; from unit considerations one necessarily must have  $L \propto \sqrt{a}$ , but the coefficient  
87 relating the two is not easily derived from first principles because lake surfaces often have complex  
88 shapes, and their bathymetry can also be complex, including multiple basins and/or having maximum  
89 depths far from their centers. We therefore introduce the free parameter  $\ell$  which allows us to relate  
90  $L$  to  $a$  according to  $L = \ell\sqrt{a}$ .  $\ell$  should be no larger than what it would be for a circular lake with  
91 its maximum depth at the center, i.e.  $\ell = \sqrt{2/\pi} \approx 0.8$ , but is likely much smaller as lakes take  
92 myriad shapes that are often very far from circular, and their maximum depths of course do not have  
93 to be in their center [23]. To compare the theoretical distribution to observed depths, it is further  
94 useful to use a normalized maximum depth  $y = z/\sqrt{2}L^H$ , and define the probability distribution in  
95 terms of  $y$ , i.e.  $p(y)$  (which can later be recast in terms of  $z$ ). Thus in terms of maximum depth and  
96 area, normalized maximum depth is:  $y = z/\sqrt{2}(\ell\sqrt{a})^H$ . This allows an overall test of goodness-of-fit  
97 which is needed because of difficulty finding large numbers of lakes with identical surface areas on real  
98 landscapes.

## 99 Empirical Test

100 We tested goodness-of-fit of normalized maximum depths to the theoretical distribution using the  
101 publicly available LAGOS database, which includes maximum depths for  $N = 8164$  lakes in the  
102 northeastern United States [20, 21] (Accessed 28 August 2021). We selected this database because it  
103 has been extensively documented and has previously been utilized for developing and testing predictive  
104 models for lake maximum depth [21]. Maximum depth spans more than three orders of magnitude

105 among these lakes, from  $z = 0.1\text{m}$  to  $z = 135\text{m}$  while surface area spans from  $a = 4\text{ha}$  to  $a = 989\text{ha}$ .

106 This database does not distinguish between lakes and reservoirs.

107 We compare the theoretical and empirical depth distributions using the Kolmogorov-Smirnov statistic

108  $D = \max(P(y) - E(y))$ , where  $P$  ( $E$ ) is the theoretical (empirical) cumulative distribution function

109 for  $y$  [28]. For a given value of  $H$  and  $\ell$ , we calculate  $y$  for each lake, then  $E(y)$  from the empirical

110 distribution of  $y$  values, then  $p(y)$  based on that  $H$  value, then  $P(y)$  by integrating  $p(y)$ , then finally

111  $D$  by comparing  $E(y)$  and  $P(y)$ . First, we set  $H = 0.4$  as an estimated value for Earth's topography

112 [17, 6], which has also been shown to capture global scale relationships between mean depth and

113 surface area [2]. We then select  $\ell$  by systematically varying  $\ell$  and repeating this procedure until  $D$  is

114 minimized. We also test what the best-fit value of  $H$  is by systematically varying both  $H$  and  $\ell$  until

115  $D$  is minimized. Finally, because small lakes are often not accurately represented in lake databases,

116 we repeat this process but only considering the upper half of the distribution, i.e. using a modified

117  $D_{alt} = \max(P(y > \text{median}(y)) - E(y > \text{median}(y)))$  [26]. This tests the goodness-of-fit of the larger

118 well mapped lakes to the upper tail of the theoretical distribution.

## 119 Results

120 The theoretical distribution  $p(z)$  adequately captures the empirical distribution of maximum depth

121 when  $H = 0.4$  (Figure 2;  $D = 0.033$ ). The distribution is right-skewed and unimodal, with few lakes

122 having a  $y \approx 0$ , many lakes having a  $y \approx 0.5$ , and a heavy tail of relatively deep lakes having a

123  $y \in (1, 4)$ . The best-fitting  $\ell$  value was 0.17, substantially lower than the theoretical limit of  $\sqrt{2/\pi}$ ;

124 this presumably reflects the extent to which these lakes are not circular and their maximum depths are

125 not located in their centers. Furthermore, when  $H$  is left as a free parameter rather than externally

126 set by topographic considerations, the best-fitting  $H = 0.38$  is negligibly different to  $H = 0.4$  (Figure

127 A1;  $D = 0.027$ ;  $\ell$  also changes only slightly, from 0.17 to 0.18).

128 This distribution  $p(y)$  can be recast as a distribution for maximum depth  $z$  for lakes of a given area  $a$

129 (Figure 3); for a lake with  $a = 1\text{ha}$  this corresponds to a median maximum depth of 3.9m and a 95%

130 confidence interval of (0.4m, 10.5m). This wide range of predicted depths arises naturally from random

131 topographic variations, and is consistent with the ranges of variability observed in empirical datasets

132 (Figure 3). In this context of a particular  $a$  value, the difference between the theoretical and empirical

133 distributions can be most easily understood by comparing the percentiles of each, which underscores

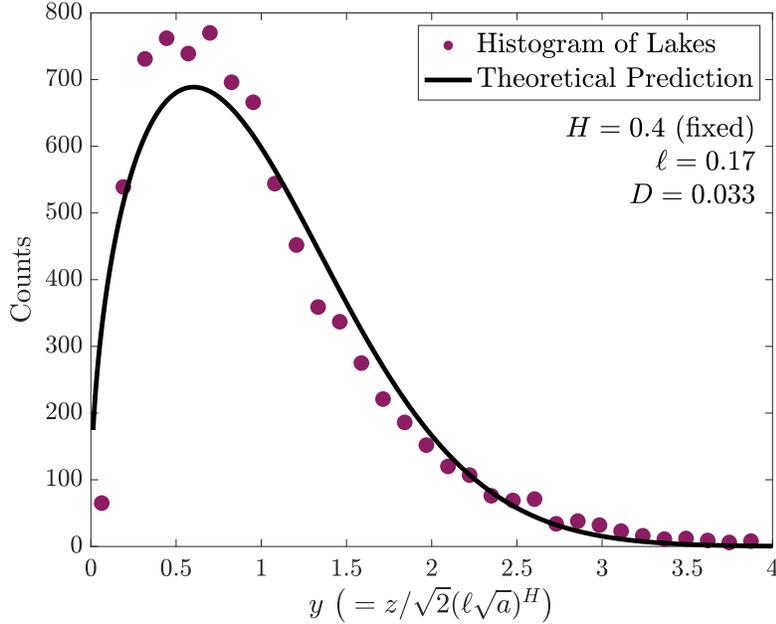


Figure 2: Theoretical probability distribution versus histogram of normalized maximum depth ( $y$ ) values for 8164 lakes. The Hurst exponent is fixed at  $H = 0.4$  and the area-lengthscale coefficient  $\ell$  ( $= 0.17$ ) is left as a free parameter.  $D = 0.033$  is the Kolmogorov-Smirnov statistic.

134 that the correspondence is indeed quite good. If we treat the empirical and theoretical  $y$  distributions  
 135 as if they were maximum depth distributions for lakes with  $a = 1\text{ha}$ , the 1st-91st percentiles of the  
 136 theoretical distribution for maximum depth are all within 8cm of those of the empirical distribution,  
 137 which is remarkable given that errors in predicting maximum depth are primarily material for shallow  
 138 and medium depth lakes. Lakes in the 92-99th percentile have larger absolute deviations from their  
 139 expected values (10-56cm), but as these expected values are larger these correspond to small relative  
 140 deviations (5-17%).

## 141 Discussion

142 Scaling relationships provide simple rules for understanding hydrographic patterns at regional and  
 143 global scales. Our study contributes to this understanding by describing the relationship between sur-  
 144 face area and maximum depth. Prior work has primarily focused on developing empirical relationships  
 145 with multiple linear regression or similar methods (e.g., [27, 12, 21]), whereas our study provides a  
 146 rigorous theoretical perspective which is accurate and generalizable among regions. Lake area and the  
 147 Hurst coefficient are the key factors controlling lake depth, with a large stochastic component. These

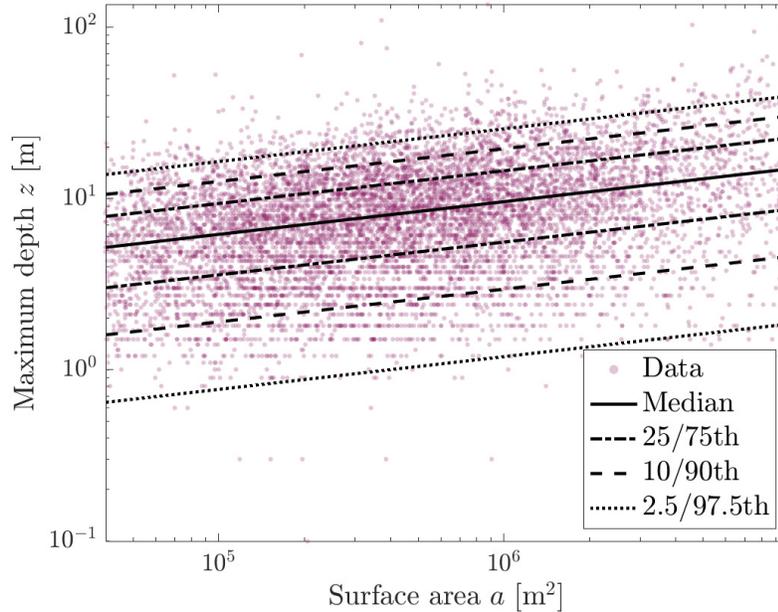


Figure 3: Lake surface areas  $a$  and maximum depths  $z$ , with percentiles of the probability distribution for  $z$  (from figure 2, for a given  $a$ ) superimposed.

148 factors can be measured without bathymetric surveys, and hence our results can be applied to existing  
 149 hydrographic and topographic data sets to estimate characteristics of lake across broad geographic  
 150 extents including littoral habitat size, carbon burial, or greenhouse gas evasion [22, 16, 1].

151 Theoretical lake morphometry, including our study, primarily derives from the assumption that Earth's  
 152 topography approximates a fractal Brownian surface [11, 2, 25]. This assumption is imperfect because  
 153 1) it implies that the landscape is a static surface whereas real landscapes are dynamic and evolve over  
 154 time, 2) it implies scale-invariance whereas real landscapes are shaped by scale dependent processes,  
 155 and 3) it implies that lakes are formed by flooding preexisting landscape depressions, which is often  
 156 not the case [11, 29]. Within this context, linear regression analyses based on landscape characteristics  
 157 essentially are trying to leverage the differences between real and random surfaces to improve maximum  
 158 depth predictions. However, the low explanatory power of such analyses, and the general consistency  
 159 between theoretical predictions and empirical patterns, indicates that the imperfect assumptions of  
 160 the fractal Brownian motion do not materially detract from their application to lake morphometry.  
 161 Hence, although somewhat stylized, the fractional Brownian motion is an effective starting point for  
 162 developing predictions about the global characteristics of lakes.

163 In our analysis, the number of medium-depth lakes is slightly overestimated and number of small  
164 lakes slightly underestimated by the theoretical distribution, but the overall fit is remarkably good.  
165 However, whether  $H$  is set by external topographic considerations ( $H = 0.4$ ) or estimated from the  
166 lake  $y$  distribution ( $H = 0.38$ ), there are some discrepancies between the theoretical depth distribution  
167 and observations is not perfect. Specifically, at the lower end of the distribution, with the theory  
168 slightly overestimating the number of medium-depth lakes ( $y \in (1, 2)$ ), underestimating the number of  
169 shallow lakes ( $y \in (\frac{1}{4}, 1)$ ), and overestimating the number of very shallow lakes ( $y < \frac{1}{4}$ ). The general  
170 ontogeny of lakes is decreased depth overtime due to sedimentation [22]. Empirical deviations from  
171 the theoretical depth distribution are consistent with this ontogeny, specifically very shallow lakes  
172 are rarer than expected, probably because they have transitioned from lake to wetland or terrestrial  
173 ecosystem. Additionally, the accentuated peak for shallow lakes is consistent with patterns expected  
174 due sedimentation. There are several other factors that can contribute to discrepancies between theory  
175 and empirical patterns, primarily related to the collection of bathymetric data. First, large lake  
176 databases typically contain samples strongly biased to certain lake characteristics, and the values in  
177 the data we analyzed may not be completely representative of the true maximum depth distribution [26,  
178 22]. Second, maximum depths are often only reported to one decimal place, which can lead to significant  
179 rounding errors for shallow lakes. This is clearly visible in Figure 3, where there is a regular patterning  
180 of shallow maximum depths, but not deep maximum depths (i.e. equal spacing among points along the  
181 ordinate). Finally, very shallow lakes may be mis-classified as wetlands and therefore not included in  
182 bathymetric databases. Collectively, these factors most strongly impact small and shallow lakes, and  
183 this observation is consistent with the excellent fit between theoretical and empirical distributions for  
184 deep lakes and somewhat weaker fit across the whole depth distribution. Specifically, when we fit only  
185 the upper-half of the  $y$  distribution (i.e. using  $D_{alt}$  above; Figure A2); the goodness-of-fit was excellent  
186 ( $D_{alt} = 0.012$ ). Advancing global limnology relies on developing robust probability distributions  
187 for lake characteristics [7]. The variety of factors causing discrepancies between the theoretical and  
188 observed distributions highlight the challenges faced in characterizing these distributions, even in ideal  
189 cases like this where theory provides clear guidance on the appropriate distribution.

190 Our approach to understanding the relationship between maximum depth and surface area is useful  
191 for predicting characteristics of collections of lakes, but not individual lakes. The only accurate way to  
192 measure morphometry for individual lakes remains the detailed bathymetric survey. However, many  
193 urgent questions in the aquatic sciences are framed in a global perspective where the characteristics

194 of large populations are of specific interest [7, 8]. Our approach is well suited for application to these  
195 questions. In particular, the characteristics of small numbers of lakes are typically up-scaled based  
196 on abundance-area distributions to estimate population level characteristics [7, 24]. This poses a  
197 problem for ecosystem characteristics that are closely tied to maximum depth but not surface area.  
198 Diffusive methane flux from temperate lakes to the atmosphere is one example of these characteristics  
199 [16]. Methane is a potent greenhouse gas and estimating evasion from lakes to the atmosphere is a  
200 priority for understanding contributions of lakes to the global carbon cycle. A recent synthesis of  
201 diffusive flux measurements for temperate lakes revealed a statistically significant inverse relationship  
202 with maximum depth ( $p < 0.001$ ), but no significant relationship with surface area ( $p > 0.05$ ) [16]. We  
203 applied the empirical relationship for diffusive methane flux from [16] to the full database of lakes  
204 from 17 northeastern United States for which surface areas are available ( $N = 141,265$ ) by randomly  
205 simulating maximum depths for these according to the distribution above (with  $H = 0.4$ , and  $\ell = 0.17$   
206 estimated from the 8,164 lakes from these same 17 states for which maximum depth measurements were  
207 available; Figure 2) and then calculating diffusive methane flux for each lake as a function of surface  
208 area and simulated maximum depth. We repeat this process many times to estimate uncertainty due  
209 to the randomness of simulated maximum depths. This results in an estimated  $0.5 \pm 0.04$  Tg  $\text{CH}_4/\text{year}$   
210 from these lakes. This is a substantially different estimate than if one just takes an average rate  
211 per unit area across all lakes, because the relationship between methane flux and maximum depth is  
212 nonlinear [16] and this average will be dominated by lakes with large surfaces, deep maximum depths,  
213 and low rates of methane flux per unit area; one may then substantially underestimate overall methane  
214 flux as a result. For instance, multiplying the average of  $0.9 \text{ g CH}_4/\text{m}^2/\text{year}$  from [16] by the total  
215 surface area of the 141,265 lakes above results in an estimate of  $0.03 \text{ Tg CH}_4/\text{year}$ . Hence, our results  
216 represent an important methodological advancement for up-scaling lake characteristics correlated with  
217 depth but not surface area. For illustrative purposes (as the parameter  $\ell$  has been fit here to lakes  
218 from the northeastern United States rather than lakes globally), applying the relationship from [16]  
219 globally to the lake areas from the Hydrolakes database (<https://hydrosheds.org>) [18] via simulating  
220 maximum depths as we describe here results in a global diffusive methane flux of  $27 \pm 3$  Tg  $\text{CH}_4/\text{year}$ .  
221 Matlab code demonstrating the application of the maximum depth distribution for up-scaling is in the  
222 supplemental materials.

223 Our study highlights the far reaching influence of the Hurst coefficient on global-scale lake charac-  
224 teristics. Specifically, the differences between horizontal and vertical scaling described by the Hurst

225 coefficient underlie differences in characteristics between large and small lakes - small lakes are typi-  
 226 cally deeper relative to their surface area compared to large lakes which has implications for energy  
 227 and carbon budgets across the lake size spectra [1]. Additionally, the Hurst coefficient is involved  
 228 in most other lake scaling relationships, including for abundance, perimeter, volume, and hydrologic  
 229 connectivity [11, 25, 2, 23]. Despite this, empirical measurements of the Hurst coefficient for Earth's  
 230 topography and bathymetry are relatively rare and highly variable ( $H = 0.4-0.7$ )[11, 10, 31]. Devel-  
 231 oping such measurements should be an important priority for advancing global scale understanding of  
 232 lakes. These measurements could explain variations in scaling relationships among regions, as well as  
 233 improve the precision of predictions by reducing uncertainty in parameterization.

234 **Appendix: alternate versions of Figure 2**

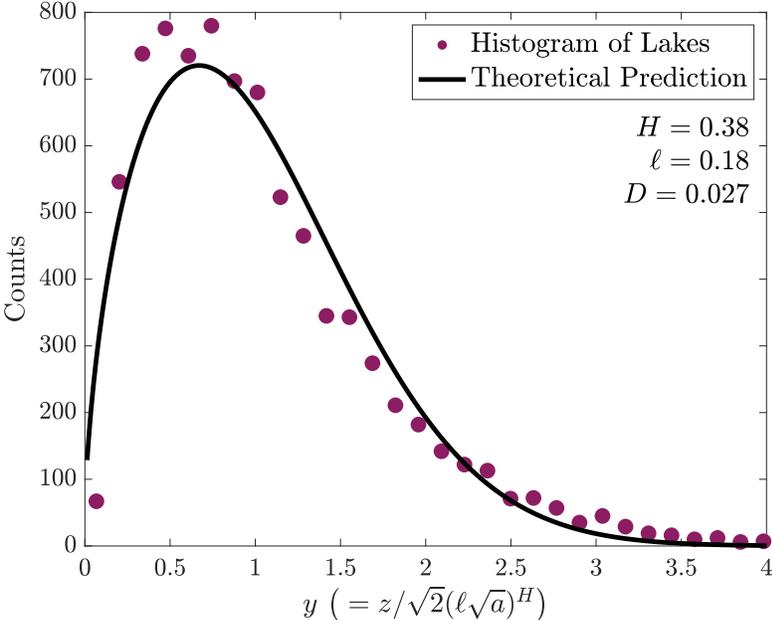


Figure A1: As in Figure 2, but with  $H$  left as a free parameter. The best-fit Hurst exponent is  $H = 0.38$  (0.4 in Figure 2) and the best-fit  $\ell = 0.18$  (0.17 in Figure 2);  $D$  reduces to 0.027 (0.033 in Figure 2).

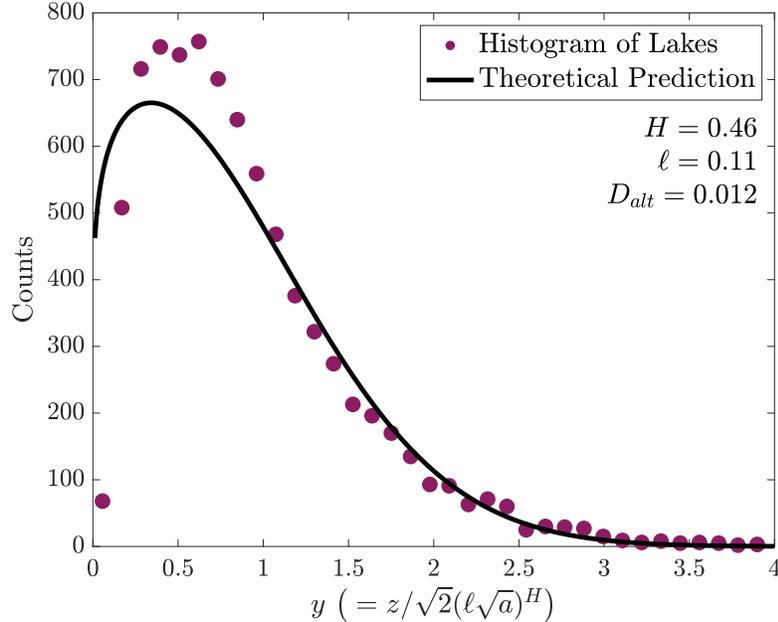


Figure A2: As in Figure 2, but with  $H$  left as a free parameter and  $D$  is only evaluated for  $y$  values above the median (i.e.  $D_{alt}$  is used). The best-fit Hurst exponent is  $H = 0.46$  (0.4 in Figure 2) and the best-fit  $\ell = 0.11$  (0.17 in Figure 2) while the fit to the upper half of the distribution is much improved ( $D_{alt} = 0.012$  as opposed to  $D = 0.033$  – albeit over the full distribution – in Figure 2).

235 **Acknowledgments:** This paper is based on research supported by the Knut and Alice Wallenberg  
 236 Foundation, the Swedish Research Council Formas, and Umeå University, and the National Environ-  
 237 mental Research Council (grants NE/N018087/1, NE/T010622/1, and NE/R015953/1).

## 238 References

- 239 [1] B. B. Cael and D. Seekell. “Morphometry constrains carbon burial in boreal lakes”. In: *Limnol.*  
 240 *Oceanogr.* (2021), In Review.
- 241 [2] BB Cael, AJ Heathcote, and DA Seekell. “The volume and mean depth of Earth’s lakes”. In:  
 242 *Geophysical Research Letters* 44.1 (2017), pp. 209–218.
- 243 [3] S. R. Carpenter et al. In: *Comparative Analyses of Ecosystems: Patterns, Mechanisms, and*  
 244 *Theories*. Ed. by J. Cole, G. Lovett, and S. Findlay. New York: Springer-Verlag, 1991, pp. 67–96.
- 245 [4] Mathieu Delorme and Kay Jörg Wiese. “Maximum of a fractional Brownian motion: Analytic  
 246 results from perturbation theory”. In: *Physical review Letters* 115.21 (2015), p. 210601.
- 247 [5] Mathieu Delorme and Kay Jörg Wiese. “Perturbative expansion for the maximum of fractional  
 248 Brownian motion”. In: *Physical Review E* 94.1 (2016), p. 012134.
- 249 [6] Peter Sheridan Dodds and Daniel H Rothman. “Scaling, universality, and geomorphology”. In:  
 250 *Annual Review of Earth and Planetary Sciences* 28.1 (2000), pp. 571–610.
- 251 [7] J. A. Downing. “Global Limnology: Up-Scaling Aquatic Services and Processes to Planet Earth”.  
 252 In: *Verh. Internat. Verein. Limnol.* 30 (2009), pp. 1149–1166.
- 253 [8] J. A. Downing. “Limnology and oceanography: Two estranged twins reuniting by global change”.  
 254 In: *Inland Waters* 4 (2014), pp. 215–232.

- 255 [9] W. P. Dunlap. “Generalizing the common language effect size indicator to bivariate normal  
256 correlations”. In: *Psychological Bulletin* 116 (1994), pp. 509–511.
- 257 [10] J.-S. Gagnon, S. Lovejoy, and D. Schertzer. “Multifractal earth topography”. In: *Nonlin. Pro-  
258 cesses Geophys.* 13 (2006), pp. 541–570.
- 259 [11] Michael F Goodchild. “Lakes on fractal surfaces: a null hypothesis for lake-rich landscapes”. In:  
260 *Mathematical Geology* 20.6 (1988), pp. 615–630.
- 261 [12] A. J. Heatcote, P. A. del Giorgio, and Y. T. Prairie. “Predicting bathymetric features of lakes  
262 from the topography of their surrounding landscape”. In: *Canadian Journal of Fisheries and  
263 Aquatic Sciences* 72 (2015), pp. 643–650.
- 264 [13] CE Herdendorf. “Large lakes of the world”. In: *Journal of Great Lakes Research* 8.3 (1982),  
265 pp. 379–412.
- 266 [14] JW Hollister, WB Milstead, and MA Urrutia. “Predicting Maximum Lake Depth from Surround-  
267 ing Topography”. In: *Plos one* 6 (2011), e25764.
- 268 [15] WM Lewis. “A revised classification of lakes based on mixing”. In: *Canadian Journal of Fisheries  
269 and Aquatic Sciences* 40 (1983), pp. 1779–1787.
- 270 [16] M Li et al. “The significant contribution of lake depth in regulating global lake diffusive methane  
271 emissions”. In: *Water Research* 172 (2020), p. 115465.
- 272 [17] David M Mark and Peter B Aronson. “Scale-dependent fractal dimensions of topographic sur-  
273 faces: an empirical investigation, with applications in geomorphology and computer mapping”.  
274 In: *Journal of the International Association for Mathematical Geology* 16.7 (1984), pp. 671–683.
- 275 [18] M. L. Messsager et al. “Estimating the volume and age of water stored in global lakes using a  
276 geo-statistical approach”. In: *Nat. Commun.* 7 (2016), p. 13603.
- 277 [19] C. K. Minns et al. “A preliminary national analysis of some key characteristics of Canadian  
278 lakes”. In: *Can. J. Fish. Aquat. Sci.* 65 (2008), pp. 1763–1778.
- 279 [20] Samantha K Oliver et al. “Predicted and observed maximum depth values for lakes in a 17-  
280 state region of the U.S.” In: *LAGOS – Long Term Ecological Research Network* (2015). DOI:  
281 10.6073/pasta/f00a245fd9461529b8cd9d992d7e3a2.
- 282 [21] Samantha K Oliver et al. “Prediction of lake depth across a 17-state region in the United States”.  
283 In: *Inland Waters* 6.3 (2016), pp. 314–324.
- 284 [22] D Seekell et al. “Patterns and Variation of Littoral Habitat Size Among Lakes”. In: *Geophysical  
285 Research Letters* 48 (2021), e2021GL095046.
- 286 [23] D Seekell et al. “The Fractal Scaling Relationship for River Inlets to Lakes”. In: *Geophysical  
287 Research Letters* 48 (2021), e2021GL093366.
- 288 [24] D. Seekell, P. Byström, and J. Karlsson. “Lake morphometry moderates the relationship between  
289 water color and fish biomass in small boreal lakes”. In: *Limnol. Oceanogr.* 63 (2018), pp. 2171–  
290 2178.
- 291 [25] D. Seekell et al. “A Fractal-Based Approach to Lake Size-Distributions”. In: *Geophys. Res. Lett.*  
292 40 (2013), pp. 517–521.
- 293 [26] D. A. Seekell. In: *Thule: Kungl. Skytteanska Samfundets Årsbok 2018*. Ed. by R. Jacobsson.  
294 Umeå, Sweden: Kungliga Skytteanska Samfundet, 2018, pp. 109–119.
- 295 [27] S Sobek, J Nisell, and J Fölster. “Predicting the depth and volume of lakes from map-derived  
296 parameters”. In: *Inland Waters* 1.3 (2011), pp. 177–184.
- 297 [28] Michael A Stephens. “EDF statistics for goodness of fit and some comparisons”. In: *Journal of  
298 the American statistical Association* 69.347 (1974), pp. 730–737.
- 299 [29] B. V. Timms. *Lake Geomorphology*. Gleneagles Publishing, 1992.

- 300 [30] D. L. Turcotte and J. Huang. In: *Fractals in the Earth Sciences*. Ed. by C. C. Barton and P. R. La  
301 Pointe. New York: Springer, 1995, pp. 1–40.
- 302 [31] J.K. Weissel, L.F. Pratson, and A. Malinverno. “The length-scaling properties of topography”.  
303 In: *J. Geophys. Res.* 99 (1994), pp. 13997–14012.