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1 **Space-time data-driven modeling of**
2 **precipitation-induced shallow landslides in South**
3 **Tyrol, Italy**

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6 **Abstract**

7 Shallow landslides represent potentially damaging processes in mountain areas worldwide.
8 These geomorphic processes are usually caused by a combination of predisposing, prepara-
9 tory, and triggering environmental factors. At regional scales, data-driven methods have
10 been used to model shallow landslides by addressing the spatial and temporal components
11 separately. So far, few studies have explored the integration of space and time for land-
12 slide prediction. This research leverages generalized additive mixed models to develop an
13 integrated approach to model shallow landslides in space and time. We built upon data
14 on precipitation-induced landslide records from 2000 to 2020 in South Tyrol, Italy (7,400
15 km²). The Slope Unit-based model predicts landslide occurrence as a function of static
16 and dynamic factors while seasonal effects are incorporated. The model also accounts for
17 spatial and temporal biases inherent in the underlying landslide data. We validated the
18 resulting predictions through a suite of cross-validation techniques and tested potential ap-
19 plications. The analysis revealed that the best-performing model combines static ground
20 conditions and two precipitation time windows: short-term cumulative precipitation prior
21 to the landslide event and medium-term cumulative precipitation. We tested the model’s
22 predictive capabilities by predicting the dynamic landslide probabilities over hypothetical
23 non-spatially explicit precipitation scenarios and historical precipitation associated with a
24 heavy precipitation event on August 5th, 2016. The novel approach shows the potential to
25 integrate static and dynamic landslide factors for large areas, accounting for the underlying
26 data structure and data limitations.

27 **Keywords:** Space-time modeling; GAMMs, Dynamic landslide modeling; Rainfall-induced
28 landslides

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1 Introduction

Landslides are potentially damage-causing geomorphic processes in hilly and mountain areas, which yearly cause thousands of people to lose their lives and belongings (Petley, 2010; Froude and Petley, 2018).

The term “shallow landslide” is generally used to refer to landslides (Cruden and Varnes, 1996; Varnes, 1978; Hungr et al., 2014) in which the sliding surface is located within a depth from a few decimeters up to generally less than two meters (Raetzo et al., 2002). Despite their limited size, shallow landslides can be particularly destructive due to their swift propagation and rapid formation, seemingly lacking pre-event geomorphic evidence (Persichillo et al., 2017).

The occurrence of shallow landslides is driven by a combination of static and dynamic environmental controls. The predisposing factors indicate a location potentially prone to landsliding (e.g., topography, underlying lithology), whereas the preparatory factors (e.g., land cover changes) and triggering factors (e.g., intense precipitation, seismic load) may cause the actual slope failure (Glade et al., 2012; Crozier, 1986). Hence, achieving reliable predictions of landslide occurrence should require a comprehensive consideration of both static and dynamic controls (Corominas et al., 2014).

The selection of an appropriate modeling approach is conditioned by several factors, including the scale of analysis, data quality, and data availability (van Westen et al., 2008; Guzzetti et al., 1999). Specifically, for large areas, data-driven models have been extensively utilized to model shallow landslide occurrences (Reichenbach et al., 2018). However, such assessments have traditionally addressed the spatial component (i.e., *where landslides are likely to occur*) and the temporal component (i.e., *when or under which dynamic conditions landslides are likely to occur*) separately.

In a purely spatial context, data-driven models derive statistical relationships between information on past landslide occurrence (i.e., landslide inventory) and static environmental controls (i.e., predisposing factors) to estimate the spatial likelihood or landslide susceptibility (Brabb, 1984; Guzzetti et al., 2005). The quality of such assessments is strongly influenced by the completeness and quality of the underlying input data. Regional landslide inventories rarely provide a spatially representative sample of past slope instabilities due to the lack of consistent mapping practices throughout an area. This often leads to underrepresented landslide populations (e.g., in areas distant from infrastructure), and consequently, this limitation may introduce bias into the resulting models (Steger et al., 2017; Lima et al., 2021). Such biased models possess restricted practical applicability because they may reflect the methods and underlying assumptions employed during the collection of the landslide inventory. As an illustration, the landslide susceptibility would exhibit probability patterns that correlate with the underlying landslide mapping strategy and the effectively surveyed areas (Bornaetxea et al., 2018; Knevels et al., 2020, i.e., areas explicitly surveyed during the generation of the landslide inventory). To address these challenges, Steger et al. (2021a) proposed novel strategies using generalized additive mixed models (GAMMs) to account for

70 the data collection effects within landslide models.

71 The assessment of landslide occurrence timing is frequently hindered by the absence of
72 either multi-temporal landslide data or historical records pertaining to triggering events such
73 as rainfall or earthquakes ([van Westen et al., 2006](#); [Guzzetti et al., 2012](#)). For large areas, em-
74 pirical approaches have traditionally been the primary choice for determining critical rainfall
75 conditions for precipitation-induced landslide occurrence ([Guzzetti et al., 2007](#); [Glade et al.,](#)
76 [2000](#)). These approaches rely on rainfall thresholds derived from past landslide observations
77 ([Guzzetti et al., 2007](#); [Kirschbaum and Stanley, 2018](#)) and often serve as the foundation for
78 landslide early warning systems ([Segoni et al., 2018a](#); [Gariano et al., 2015](#)). This is achieved
79 by using several representative parameters such as rainfall intensity-duration ([Guzzetti et al.,](#)
80 [2007](#)), cumulative rainfall event-duration ([Peruccacci et al., 2017](#)), and antecedent rainfall
81 conditions ([Monsieus et al., 2019](#)). Conventionally, these approaches adopt a presence-only
82 framework, focusing solely on the rainfall that leads to landslides while disregarding rainfall
83 events that do not trigger landslides. So far, few studies have adopted presence-absence
84 frameworks, which are commonly observed in landslide susceptibility studies, with the dis-
85 tinction that susceptibility solely accounts for static ground-related conditions (e.g., slope
86 steepness, lithology, land cover), while rainfall thresholds are dedicated to the meteorological
87 aspects ([Steger et al., 2023](#); [Segoni et al., 2018a](#)). This situation depicts the current separa-
88 tion in the geoscientific literature between two components of the landslide hazard definition
89 (i.e., the "where" and "when" landslides may occur) ([Corominas et al., 2014](#)).

90 The integration of both space and time in data-driven modeling techniques is rarely done
91 ([Lombardo and Tanyas, 2020](#); [Samia et al., 2020](#); [Lombardo and Tanyas, 2021](#); [Bajni et al.,](#)
92 [2023](#)). Nonetheless, there have been notable endeavors to incorporate dynamic controls into
93 the modeling process by aggregating meteorological factors over specific time periods (e.g.,
94 mean annual rainfall; maximum daily rainfall per inventoried period [Wang et al., 2022](#); [Dahal](#)
95 [et al., 2022](#)). These approaches are better equipped to capture climate variability or long-
96 term meteorological predisposition rather than short-term dynamics. However, by focusing
97 on such long time windows, the impact of extreme events may be diluted or even lost ([Camera](#)
98 [et al., 2021](#)). The incorporation of intense rainfall into data-driven modeling mainly revolves
99 around event-based assessments and early warning systems ([Knevels et al., 2020](#); [Segoni](#)
100 [et al., 2018b](#); [Kirschbaum and Stanley, 2018](#)). For instance, [Knevels et al. \(2020\)](#) integrates
101 rainfall by using the maximum rainfall intensity observed within an hourly time window as a
102 predictor within a susceptibility model. In contrast, [Segoni et al. \(2018b\)](#) couples the output
103 of a purely spatial susceptibility model with the results of rainfall threshold exceedance
104 analyses in a heuristic approach to predict landslide occurrence.

105 To summarize, the integrated modeling of landslides in space and time remains a challenge
106 and is seldom explored in the majority of the geoscientific literature. Its development and
107 application would enhance the understanding of the critical conditions (e.g., rainfall) that
108 lead to slope failure and the quality and reliability of landslide predictions. This research
109 aims to integrate shallow landslide modeling in space and time using data-driven techniques.

110 Specifically, a binary GAMM is used to account for precipitation as a dynamic predictor
111 at different temporal scales while integrating static landscape characteristics and seasonal
112 effects. We conducted the analysis for a space-time domain covering the territory of South
113 Tyrol (Italy) for 21 years (2000 - 2020).

114 2 Materials and methods

115 2.1 Study area

116 The province of South Tyrol spans approximately 7,400 km² and is located in the Eastern
117 Alps, constituting the northernmost part of Italy, as shown in Figure 1. It shares borders
118 with Austria to the north and northeast, Switzerland to the northwest, and the rest of
119 Italy to the south. According to Provincial Statistics Institute (<https://astat.provinz.bz.it/>)
120 South Tyrol is home to approximately 535,000 people with Bozen, Meran, and Brixen being
121 the most populated municipalities.

122 The study area is characterized by its diversified geomorphology, geology, and climatic
123 conditions. The landscape is dominated by pronounced variations in altitude (i.e., elevation
124 ranging between ~200 m and ~3,900 m a.s.l.) with narrow valleys encompassed by steep
125 slopes. The three main rivers comprise the Rienz River, the Eisack River, and the Etsch
126 River, the latter being Italy’s second longest river, flowing eastward into the Adriatic Sea.

127 Geologically, the area is characterized by the presence of the Periadriatic Line, a major
128 tectonic structure that subdivides the region into two distinct sections: the Southalpine and
129 the Austroalpine (Piacentini et al., 2012). The latter constitutes the western and northern
130 parts of South Tyrol and is mainly composed of metamorphic rocks such as mica schists,
131 amphibolites, orthogneisses, and paragneisses. On the other hand, the Southalpine section,
132 located in the southern and southeastern parts of the study area, is dominated by volcanic
133 rocks, including porphyries, which are covered by carbonatic successions of dolomites and
134 limestones (Stingl and Mair, 2005). Furthermore, the northeastern areas correspond to the
135 Tauren Window, mainly formed by medium-grade calcium-rich metasediments, serpentinites,
136 and metagranitoids as described by Oxburgh (1968).

137 The climatic conditions exhibit considerable seasonal and spatial variations, predomi-
138 nantly characterized by a continental climate. Airflows of western Atlantic and Mediter-
139 ranean currents influence the monthly precipitation patterns, as indicated in Crespi et al.
140 (2021); Marra et al. (2014); Adler et al. (2015). The western inner valleys, particularly the
141 Vinschgau Valley, represent the driest areas with an average annual precipitation of ~500
142 mm. Conversely, the mountainous areas in the north and northeast experience the highest
143 levels of precipitation, reaching ~1,500 mm annually. Seasonally, the maximum precipitation
144 is observed during the summer months, with an average of ~120 mm per month, followed
145 by autumn with ~80 mm per month. In contrast, winter tends to be the driest period, with
146 an average monthly precipitation of ~30 mm.

147 The terrain and climatic settings render this landscape highly prone to landsliding, with

148 predominant failure types belonging to falls, topples, and slides ([Schlögel et al., 2020](#)). In
149 South Tyrol, researchers have explored several aspects related to shallow landslides, such as
150 the effects of land use on landslides ([Tasser et al., 2003](#)), modeling shallow landslide suscep-
151 tibility ([Piacentini et al., 2012](#)), inventory-based exploratory analysis of landslides ([Steger](#)
152 [et al., 2021b](#)), the development of strategies to address bias in shallow landslide susceptibil-
153 ity modeling ([Steger et al., 2021a](#)), and recent investigations into understanding seasonally
154 dynamic precipitation conditions responsible for shallow landslide occurrence ([Steger et al.,](#)
155 [2023](#)). However, no regional-scale approach has yet been developed to integrate static and
156 dynamic factors for explaining shallow landslide occurrence in South Tyrol.

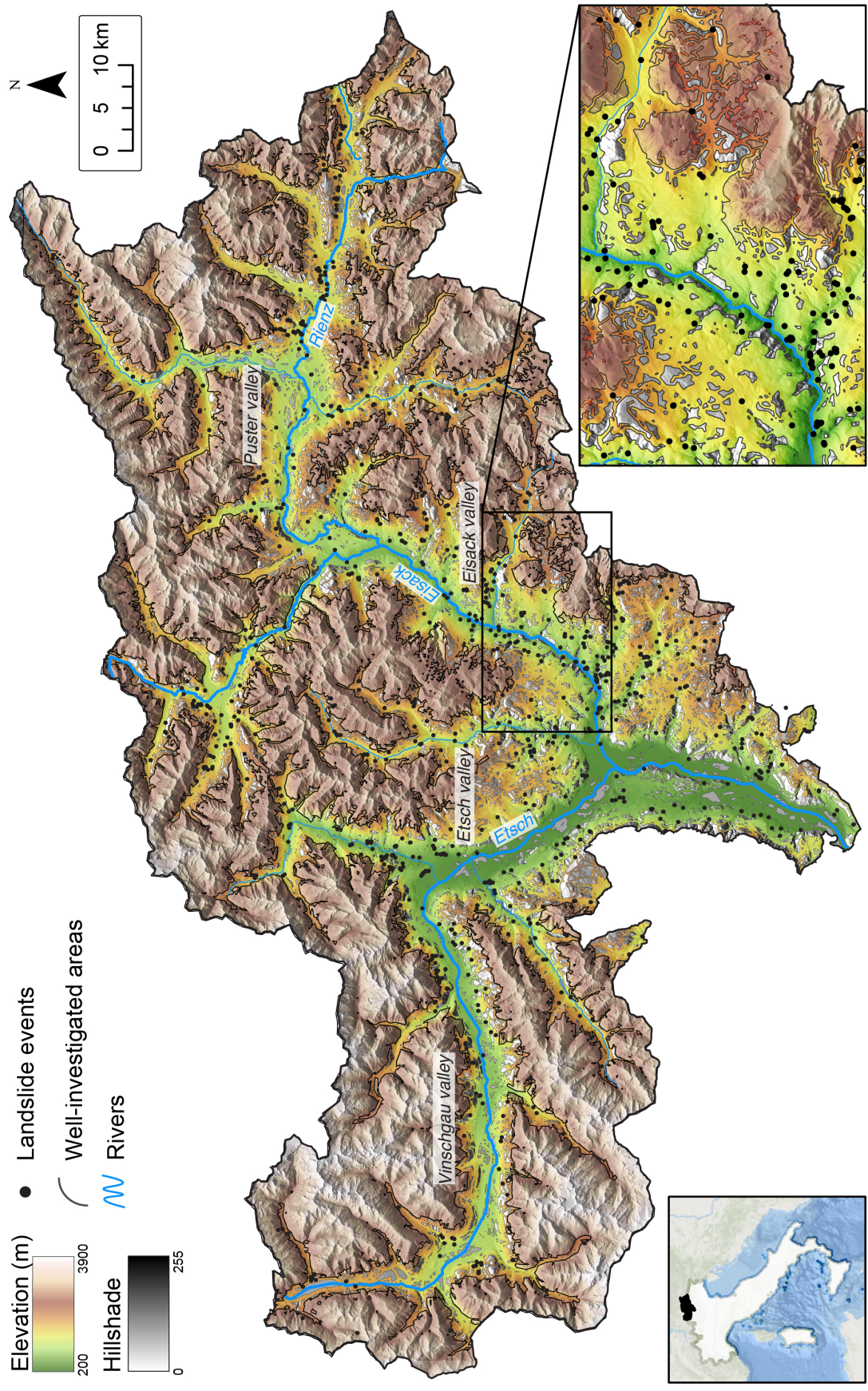


Figure 1: Study area showing the elevation and the filtered landslide scarp locations (n=1006). The dotted line covering the more contrasted zones represents the well-investigated or effectively surveyed areas (i.e., areas with certainty that a landslide is mapped).

2.2 Data

2.2.1 Landslide inventory

The landslide information is extracted from the Italian landslide inventory (*Inventario dei Fenomeni Franosi in Italia*; IFFI), which can be consulted in the IdroGeo platform (<https://idrogeo.isprambiente.it/app/>). The IFFI project has been coordinated by the Institute for Environmental Protection and Research (ISPRA) since the early years 2000 (Trigila et al., 2007). Each region and autonomous province in Italy is responsible for collecting landslide data for their respective areas of jurisdiction.

In the Autonomous Province of Bolzano, the available point-based information depicts landslide scarp locations which are frequently mapped in the field using Global Positioning Systems (GPS) (Trigila et al., 2010). Landslides that triggered an intervention by the provincial authorities are systematically inventoried, whilst landslides that did not cause any damage (e.g., landslides that occur far from infrastructure) are usually not documented (Steger et al., 2021a). This can be seen in Figure 1, where the well-investigated areas or areas in which landslides are rigorously mapped are highlighted. The inventory shows that up to February 2022, there were 11,416 landslide events and the majority fall into the categories of falls/topples, slides, and flows. Each of these landslides is documented with various attributes that provide detailed information on the movement type, the material involved, the cause, and the occurrence date.

As the primary objective of this study is to dynamically model the probability of landslide occurrences over space and time, specific analyses were supported by selectively extracting from the inventory only those landslide events that were accompanied by the occurrence date information. Further elaboration on this methodical detail is provided in 2.3.

2.2.2 Mapping unit

The selection of the mapping unit is a crucial requirement for any landslide predictive model. Among the most commonly used mapping units, there are pixels (Lima et al., 2021), Slope Units (SUs) (Amato et al., 2019), and unique condition units (Titti et al., 2021), with SU gaining more attention in recent years (Reichenbach et al., 2018). SUs are polygons bounded by streamlines and ridges, reflecting the hydrological and geomorphological processes shaping the natural landscape (Carrara et al., 1991; Guzzetti, 2006). Alvioli et al. (2016) developed an integrated and parametrized GRASS GIS tool known as *r.slopeunits*. This software enables users to generate SU partitions that maximize the polygonal internal homogeneity and external heterogeneity of the slope aspect. To utilize *r.slopeunits*, one needs to provide the Digital Elevation Model (DEM) and specify several parameters such as the circular variance to account for the slope aspect homogeneity, the minimum SU area among others (i.e., flow accumulation threshold and cleansize; for details, see Alvioli et al., 2016). In this research, we chose an SU partition to subdivide the study area. In the process, we intentionally removed the flood plain of the Etsch-Adige River as it can be regarded as a

195 trivial terrain (i.e., easy-to-classify areas in which no landslide is expected; [Steger et al.,](#)
196 [2016](#)). In the remainder of the manuscript, we also consider other trivial areas, as described
197 in Section [2.3.2](#). Below, we report the final parameters resulting from multiple iterations in
198 *r.slopeunits* using a bilinearly resampled 30 m LiDAR-DTM ([Geokatalog, 2019](#)).

- 199 • Circular variance = 0.3
- 200 • Minimum Slope Unit area = 500,000 m²
- 201 • Flow accumulation threshold = 1,000,000
- 202 • Cleansize = 50,000 m²

203 As a result, we obtained a total of 5,379 SUs, whose size distribution has a mean of ~ 1.3
204 km² and a standard deviation of ~ 0.9 km².

205 **2.2.3 Geo-environmental data**

206 This section provides separate descriptions of the static and dynamic landslide controls used
207 to represent the predisposing, preparatory, and triggering factors as summarized in [Table 1](#).

208 **Static factors**

209 Extensive research is available on understanding the different static factors and their rela-
210 tionship to landslide occurrence, as illustrated by [Budimir et al. \(2015\)](#). In this case, we
211 used the bilinearly resampled 30 m LiDAR-DTM to derive common morphological variables
212 such as slope steepness, slope aspect, concavity, local relief, and topographic position in-
213 dex, among others. Geological information was also considered by using the "Geologische
214 Übersichtskarte Südtirol", where the lithological units were grouped into five main classes:
215 (i) crystalline, (ii) porphyry, (iii) sedimentary, (iv) plutonite, and (v) calcschist. More details
216 on the geological information can be found in [Steger et al. \(2021b\)](#). Land cover data were
217 retrieved from "Realnutzungskarte Südtirol v. 2015" and grouped into six classes: (i) agricul-
218 ture, (ii) forest, (iii) infrastructure, (iv) pasture, (v) rock, (vi) water, and glacier. Moreover,
219 the polygons of the catchment units (i.e., catchment ID) and the topographically corrected
220 SU area ([Steger et al., 2021a](#); [Moreno et al., 2023](#)) were also considered during the analysis.
221 In our modeling framework, we further accounted for the effectively surveyed area in order
222 to reduce bias stemming from a spatially uneven registration of past landslides ([Bornaetxea](#)
223 [et al., 2018](#); [Steger et al., 2021a](#)). This layer (i.e., mask) provides detailed information about
224 the locations where landslides have been meticulously mapped. By including this informa-
225 tion, we aimed to limit the inherent spatial bias that may arise from variations in the data
226 collection strategies across different areas (e.g., in areas close or far from infrastructure; see
227 [Figure 1](#)).

228 **Dynamic factors**

229 The dynamic factor that primarily influences the timing of shallow landslide occurrences in
230 Italy is precipitation ([Brunetti et al., 2010](#)). To capture this important dynamic variable,

231 we utilized the daily precipitation dataset for the Trentino-South Tyrol region, provided by
 232 [Crespi et al. \(2021\)](#), which consists of 250-m gridded daily fields. The dataset is computed
 233 by interpolating data obtained from a dense network of over 200 meteorological stations,
 234 ~ 80 of which in South Tyrol, ensuring comprehensive coverage across the study area. The
 235 preprocessing of the station data was built upon multiple methodical steps, which involved
 236 quality, consistency, and homogeneity tests, as well as gap-filling approaches to maximize
 237 series completeness. To account for the influence of the orography on the spatial distribution
 238 of precipitation, the interpolation scheme includes a local weighted linear regression with
 239 station weights depending on distance and topographic similarity to the target point. The
 240 leave-one-out cross-validation returned a mean absolute error (MAE) of 1.1 mm, as averaged
 241 across all the meteorological stations and months. Each precipitation field within the dataset
 242 represents the total precipitation accumulated over a 24-hour period, specifically from 08:00
 243 UTC of the day before the observation to 08:00 UTC of the observation day ([Crespi et al.,](#)
 244 [2021](#)). In addition to precipitation, the day of the year (*DoY*) was included as a dynamic
 245 predictor. *DoY* is a sequential number representing each day during a year, starting from 1
 246 on January 1st and ending with 365-366 on December 31st. This value is derived from the
 247 assigned date of each SU, and it mainly serves as a proxy for capturing seasonal effects, e.g.,
 248 vegetation or temperature changes ([Steger et al., 2023](#)).

Table 1: Static and dynamic predictor summary table. Continuous properties were aggregated at the SU level by calculating the average value, while categorical properties were aggregated by determining the proportion of each class and the predominant class.

Predictor	Unit	Value	Reference
Slope steepness	degrees	average	(Zevenbergen and Thorne, 1987)
Relief	m	n.a.	(Mark, 1975)
Concavity	%	average	(Iwahashi and Pike, 2007)
Lithology	1	majority	
Land cover	%	proportion	
Catchment ID	1	majority	
Effectively surveyed area	1	average	(Bornaetxea et al., 2018 ; Steger et al., 2021a)
Slope Unit area	ha	sum	(Moreno and Steger, 2023)
Day of the year	1	n.a.	(Steger et al., 2023)
Year	1	n.a.	(Steger et al., 2023)
Daily precipitation	mm	average	(Crespi et al., 2021)

249 2.3 Methods

250 The methodical framework depicted in Figure 2 is divided into three main stages. The first
 251 stage encompasses the analysis of static predisposing factors, which resulted in the estimation
 252 of landslide susceptibility. The dynamic component was then analyzed in a second stage,
 253 yielding the best two precipitation time windows in terms of preparatory and triggering

254 precipitation. Ultimately, the outcomes of the first two stages are integrated to develop
 255 a dynamic landslide model accounting for static conditions, the dynamics of precipitation,
 256 and seasonal effects. In the following subsection, the three stages are described, along with
 257 background information on the GAMMs we implemented.

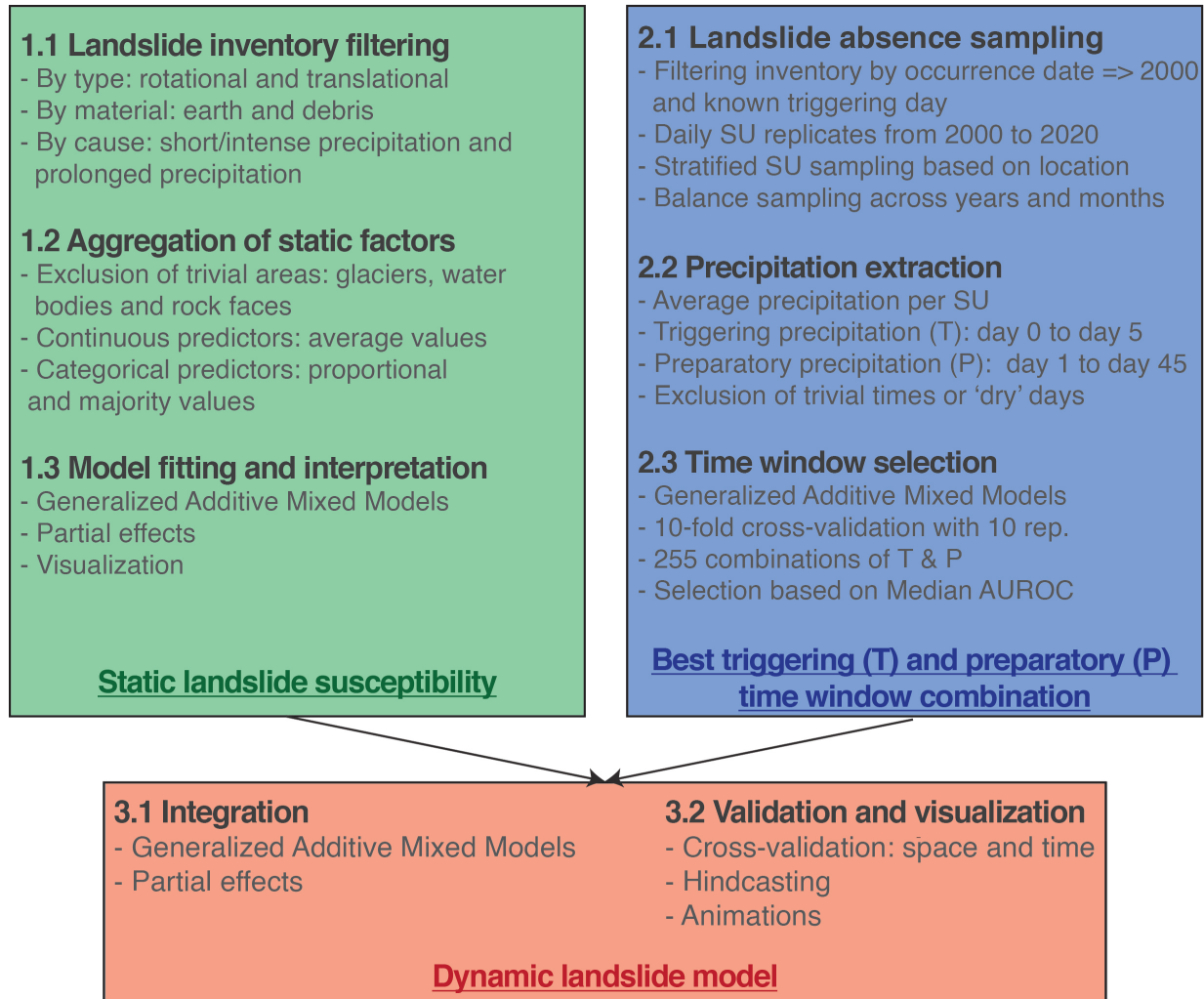


Figure 2: Methodical approach divided into the three main stages we explored in this work: (i) the static component, (ii) the precipitation component, and (iii) the dynamic landslide model.

258 2.3.1 Modeling landslides using Generalized Additive Mixed Models

259 GAMMs are a flexible extension of the well-known generalized linear model (GLM) frame-
 260 work. The latter offers the ability to fit a number of exponential probability distributions
 261 (e.g., Poisson, Gamma, Gaussian, among others) as a function of a predictor set (Zuur et al.,
 262 2009). However, binomial GLMs based on a logit function (i.e., logistic regression) restrict

263 the relation between the response and the independent predictors to the linear case. Al-
264 though this assumption might be reasonable in some cases, it may not hold for many natural
265 processes. For instance, shallow landslides are known to be nonlinearly dependent on slope
266 steepness because terrains with low and very high slope angles (i.e., with no soil cover) may
267 not host failures, while terrains with middle steepness are naturally prone. In the context of
268 multivariable statistical analyses, GAMMs demonstrate their strength, particularly in these
269 situations, enabling nonlinear effects to be incorporated into the modeling procedure (Wood,
270 2006; Bolker et al., 2009). In the case of susceptibility studies, this is achieved by assuming
271 that landslide presences and absences are distributed over space and time according to a
272 binomial probability distribution.

273 We use this modeling framework in the three stages shown in Figure 2. The first one is
274 equivalent to a traditional landslide susceptibility, with only static information appearing in
275 the binomial GAMM to explain the distribution of landslides purely in space as presented
276 in Section 2.3.2. The second stage also corresponds to a binomial GAMM, but this time,
277 only dynamic predictors are featured, aiming at explaining landslide occurrence in time
278 purely from the meteorological perspective and accounting for seasonality (e.g., including
279 a day of the year predictor). This operation is closely linked to Steger et al. (2023) and
280 is intended to estimate the best combination of triggering and preparatory precipitation
281 time windows to explain landslide occurrence. More details on this stage are presented in
282 Section 2.3.3. Ultimately, the outcomes of the two previous stages are integrated into a
283 third stage to produce a unified dynamic landslide prediction model as explained in Section
284 2.3.4. This model jointly presents temporally invariant ground conditions and precipitation
285 characteristics that are changing dynamically as a function of time.

286 Notably, binomial models are common in the literature, and assessing their classification
287 performance has been explored in depth. The Receiver Operating Characteristic (ROC) be-
288 longs to the class of cutoff-independent metrics and is the most commonly used to distinguish
289 how well a binary classifier performs (Hosmer et al., 2013). We consider the area under the
290 ROC curve (AUROC, hereafter) as a metric to assess the model performance (Faraggi and
291 Reiser, 2002).

292 2.3.2 Static component

293 Landslide inventory filtering

294 We narrowed down the initial landslide data by applying three criteria: (i) movement type,
295 (ii) material type, and (iii) cause type. As a result, we obtained a subset of 1,821 landslides
296 characterized by *translational* and *rotational* movement types, involving *earth* and *debris*
297 materials, and triggered by *short-intense precipitation* or *prolonged precipitation*.

298 Static factor aggregation

299 In this study, we utilized the SUs to aggregate both the target variable, which comprises land-
300 slide observations, and the predisposing factors previously described in Section 2.2.3. The

301 continuous predictors were aggregated by calculating the average values per SU. Regarding
302 the categorical predictors, we adopted a distinct aggregation method for the geological and
303 land use information. To capture the categorical information of geology within each SU,
304 we implemented the majority rule approach. This method involved identifying the predom-
305 inant lithotype that covered the majority of the SU polygon. This decision was motivated
306 by the spatial coarseness of geological units, where calculating lithology percentages per SU
307 usually resulted in binary outcomes of either 0% or 100%. Conversely, the land cover layer
308 displayed higher spatial variability, with frequent transitions between different classes. To
309 account for this, we treated the land cover information as a continuous variable, expressed
310 as a percentage of a given class intersecting the SU layer.

311 Furthermore, the land cover data provided valuable insights for identifying and excluding
312 other trivial terrains, such as rocky faces, glaciers, and water bodies. These areas were
313 excluded from the aggregation process by masking out pixels not classified as trivial terrains.
314 In other words, we performed the aggregation considering only the pixels outside the trivial
315 terrains. By excluding these areas, we aimed to render the classification problem more
316 topic-specific by a priori excluding terrain that does not induce shallow landslides.

317 Moreover, for the calculation of areal properties (i.e., SU area and proportion of land
318 cover classes), we implemented a correction procedure. This correction aimed to account
319 for potential underestimation caused by the use of conventional planar projection in steep
320 terrains. By applying this correction, we obtained more accurate surface area measurements
321 and mitigated any distortions introduced by the projection method (Steger [et al.](#), 2021a;
322 [Moreno et al.](#), 2023).

323 **Model fitting and interpretation**

324 We initially fitted a binomial GAMM to the filtered landslide data (i.e., 1,821 landslide
325 observations) using the static predictor set. The layer representing the *effectively surveyed*
326 *area* was utilized to account for potential biases in the final susceptibility map. A detailed
327 description of the predictor effects can be found in Section 3.1.1. The bias removal procedure
328 we employed can be summarized as follows: Considering that the mapping of landslides in
329 the province is systematic only close to infrastructure, we incorporated a predictor that
330 describes the potential areas that the geological office systematically surveys in South Tyrol
331 on the basis of previously modeled landslide data-collection effects (Steger [et al.](#), 2021a).
332 This predictor, referred to as the *effectively surveyed area* (Bornaetxea [et al.](#), 2018), was
333 included to capture the spatial variability in the data caused by the spatial mapping bias.
334 By including the *effectively surveyed area*, we ensured that this specific bias was accounted
335 for, enabling the determination of other predictor effects without its confounding influence.
336 To generate an unbiased susceptibility map, we opted to remove the effect of the *effectively*
337 *surveyed area* from the predictive function. This step constituted one of the two elements
338 involved in the third stage of the analysis detailed in Section 2.3.4).

339 2.3.3 Precipitation component

340 **Landslide absence sampling**

341 We further filtered the landslide inventory by considering only landslide events with a known
342 date of occurrence, resulting in a final sample size of 676 landslide records. Retrieving this
343 information is a straightforward process. However, sampling the landslide absence informa-
344 tion poses a much more complex challenge. Opting for a temporal unit of a single day leads
345 to a substantial increase in the potential number of SUs in both space and time (i.e., approx-
346 imately forty million = $5,379 \text{ SU} \times 21 \text{ years} \times 365 \text{ days}$). Therefore, a crucial requirement is
347 to extract a suitable number of representative stable SU from this extensive spatiotemporal
348 dataset. To fulfill this requirement, we employed a suitability criterion based on the method
349 proposed in [Steger et al. \(2023\)](#), that encompasses the following elements:

- 350 1. Balanced sampling across SUs;
- 351 2. Balanced sampling across years;
- 352 3. Balanced sampling across months.

353 We implemented the mentioned criteria as follows. Initially, we randomly selected 10
354 SU replicates for each location from the possible 21-year dataset. This process resulted in
355 an initial stable dataset consisting of 53,790 SUs. However, we noticed that certain years
356 were represented more frequently than others, leading to an uneven temporal distribution.
357 For this reason, we further constrained the selection process to achieve a balanced yearly
358 absence distribution from the potential 40 million cases. This step ensured that the landslide
359 absence data size remained the same, with 53,790 SUs. Although this operation left a
360 consistent distribution of SUs across the study area and over each year, it did not address
361 the uneven proportion of absences at the monthly level. To address this, we repeated the
362 yearly constraint at a finer monthly resolution. As a result, the final dataset comprised
363 53,790 SUs representing locations where landslides did not occur across the entire province
364 of South Tyrol, spanning each year and each month from 2000 to 2020. The subsequent step
365 required merging the presence and absence data, an operation that returned a spatiotemporal
366 domain made of 54,460 SU (i.e., 670 presences and 53,790 absences).

367 **Precipitation extraction**

368 We initially assigned the corresponding daily precipitation amounts to each of the 54,460
369 SUs mentioned above. For detailed methodological information regarding precipitation pro-
370 cessing and its assignment to a respective mapping unit, please refer to [Steger et al. \(2023\)](#).
371 The precipitation extraction consisted of 46 days of daily precipitation, starting from the
372 given date assigned to the SUs replicates (i.e., day 0) and moving backward 45 days (i.e.,
373 day 45). Then, we calculated the daily cumulative precipitation for each SU replicate up
374 to the 46th day. In addition to this operation, we added a subsequent check to exclude the
375 SUs that did not report precipitation in a period before the observation time. In this step,
376 equivalent to the one applied in Section 2.3.2 where we excluded trivial terrains, we opted

377 for excluding trivial or “dry” periods to ensure problem-specific results. Whenever the total
378 cumulative precipitation was estimated below 1.1 mm (i.e., below the MAE resulting from
379 the precipitation dataset cross-validation) in the first two days (i.e., day 0 and day 1), we
380 removed that SU from the analysis. As a result, the overall space-time domain was reduced
381 to 24,466 SUs, out of which 588 corresponded to unstable slopes (i.e., presences) and 23,878
382 were stable ones (i.e., absences). In summary, the final sample represents the spatiotemporal
383 distribution of rainy observations with and without landslides.

384 **Time window selection of the cumulative precipitation**

385 In line with the rainfall threshold approach observed in landslide early warning systems
386 (Segoni [et al.](#), 2014; Wang [et al.](#), 2021; Chleborad, 2003), we determined the most ap-
387 propriate cumulative precipitation time windows from the possible 46 to describe critical
388 landslide conditions. To address this, we employed a binomial GAMM, in which we aimed
389 to predict landslide presences and absences using two precipitation time windows as predic-
390 tors. One precipitation time window characterizes the short-term triggering precipitation T
391 (i.e., precipitation shortly before the event), and the other, the medium-term preparatory
392 precipitation P (i.e., precipitation prior to T). The predictor T was built considering the
393 cumulative precipitation windows from day 0 to day 5, whereas P was set by considering
394 the cumulative precipitation time windows prior to T and up to day 45. Additionally, to
395 account for potential seasonal effects and interannual data variability, we included a circular
396 spline effect defined for each DoY and an effect defined for each year ($Year$). Unlike other
397 splines, a year exhibits a cyclic temporal characteristic, necessitating a distinct approach.
398 DoY utilized then a cyclic spline that introduces an additional seasonal constraint. Hence,
399 each day within a year exhibits similar performance across the entire spatiotemporal do-
400 main, and the effect of the last day of the year is reciprocally dependent on the first day
401 of the year. The predictor $Year$ guarantees that any potential discrepancies in reporting
402 landslide occurrences over different years do not influence the modeled relationships between
403 our predictors of interest and landslide occurrence. Subsequently, the binomial GAMM was
404 iterated over the 255 possible combinations of the predictors P and T . This procedure is
405 explained in detail in Steger [et al.](#) (2023). Differently from this previous study, the cur-
406 rent one also incorporated a 10-fold random cross-validation step, conducting the procedure
407 255,000 times. These iterations corresponded to the 255 pairwise combinations, which we
408 bootstrapped 10 times over 10-randomized repetitions for further robustness. We calculated
409 the AUROC at each step throughout the iterations and stored the results. Ultimately, we
410 identified the best pair of triggering and preparatory precipitation time windows to be the
411 highest median AUROC.

412 **2.3.4 Dynamic landslide prediction model**

413 **Integration of static and precipitation components**

414 This final modeling step integrates the two components presented in the previous sections.

415 The temporally invariant susceptibility of the terrain described in Section 2.3.2 is combined
416 with the temporally-varying component pertaining to the best pair of triggering and prepara-
417 tory precipitation time windows reported in Section 2.3.3 in a single space-time model. We
418 emphasize that we included a methodical procedure in the static component to address
419 the spatial bias resulting from the landslide mapping strategies. Consequently, the actual
420 susceptibility utilized in the dynamic landslide prediction model is unbiased. This static
421 susceptibility map served as a nonlinear effect in the landslide dynamic model. In other
422 words, the susceptibility map consistently conveys the same signal at each temporal repli-
423 cate, irrespective of the variations in the spatial distribution of stable and unstable SUs.

424 The dynamic information provided to the model primarily consisted of three additional
425 predictors representing the triggering and preparatory precipitation effects, as well as sea-
426 sonal variations captured using the *DoY* predictor, which were employed in a non-linear
427 manner.

428 **Validation and visualization**

429 The extensive spatiotemporal domain offers numerous possibilities for testing the dynamic
430 model results. We conducted a standard 10-fold random cross-validation (RCV), repeated
431 ten times for a total of 100 iterations. In addition, we included random spatial cross-
432 validation (SCV), utilizing a 10-fold structure repeated ten times (Brenning, 2012). We
433 also featured two temporal validation routines to complement the overall model testing.
434 The first corresponded to a 21-fold temporal validation, where one year is excluded at a time
435 for prediction (i.e., leave-one-year-out), and this process is repeated until every year in the
436 dataset has been predicted. Similarly, we performed the same operation at a monthly level,
437 employing a 12-fold temporal validation. In this case, one month is removed at a time to
438 serve as the prediction target, and the routine is repeated until every month within a year
439 has been predicted.

440 Ultimately, as a demonstration of its capabilities, we also used our dynamic model for
441 simulation purposes. Ideally, one could employ the model to simulate any day of the year
442 and for any desired duration, provided that the precipitation data is available. However, this
443 operation is computationally intensive in practice. In fact, not only does the precipitation
444 data need to be preprocessed but also aggregated at the SU level before this information is
445 incorporated into the dynamic landslide predictive model. Therefore, to showcase a practical
446 application of our model, we generated two specific scenarios, one for a hypothetical case
447 and one for a specific historical precipitation event. In the first case, the scenario includes
448 temporally varying but spatially homogeneous hypothetical precipitation (i.e., with no vari-
449 ations across the landscape) and the other predictors. In the second case, we conducted
450 a hindcast of the event recorded on August 5th, 2016, by incorporating the spatiotemporal
451 precipitation distribution from July 15th to August 15th, 2016.

3 Results

3.1 Static component

This section shows an overview of the model components and the estimated landslide susceptibility.

3.1.1 Model relationships and landslide susceptibility

Figure 3 offers a comprehensive view of the predictor effects we considered in our model. Below, we provide a concise overview of these predictors by describing their partial effects (i.e., assuming all the other predictor contributions to be fixed).

For instance, the predictor *slope steepness* exhibits a positive nonlinear contribution to the estimated landslide probability. At first, it shows a positive increasing effect that escalates until slope inclinations of $\sim 30^\circ$. However, within steeper SUs surpassing $\sim 30^\circ$, this contribution gradually diminishes, ultimately tapering off to zero for slope angles exceeding $\sim 50^\circ$. The characteristic of *concavity* follows a moderately nonlinear trend, suggesting that SUs displaying a moderate degree of concavity (i.e., $\sim 40\% - 50\%$) hold the highest contribution to the estimated shallow landslide susceptibility. Similarly to the *slope steepness*, the effect of *concavity* also shows a gradual reduction, eventually converging to zero when concavity values approach to $\sim 65\%$.

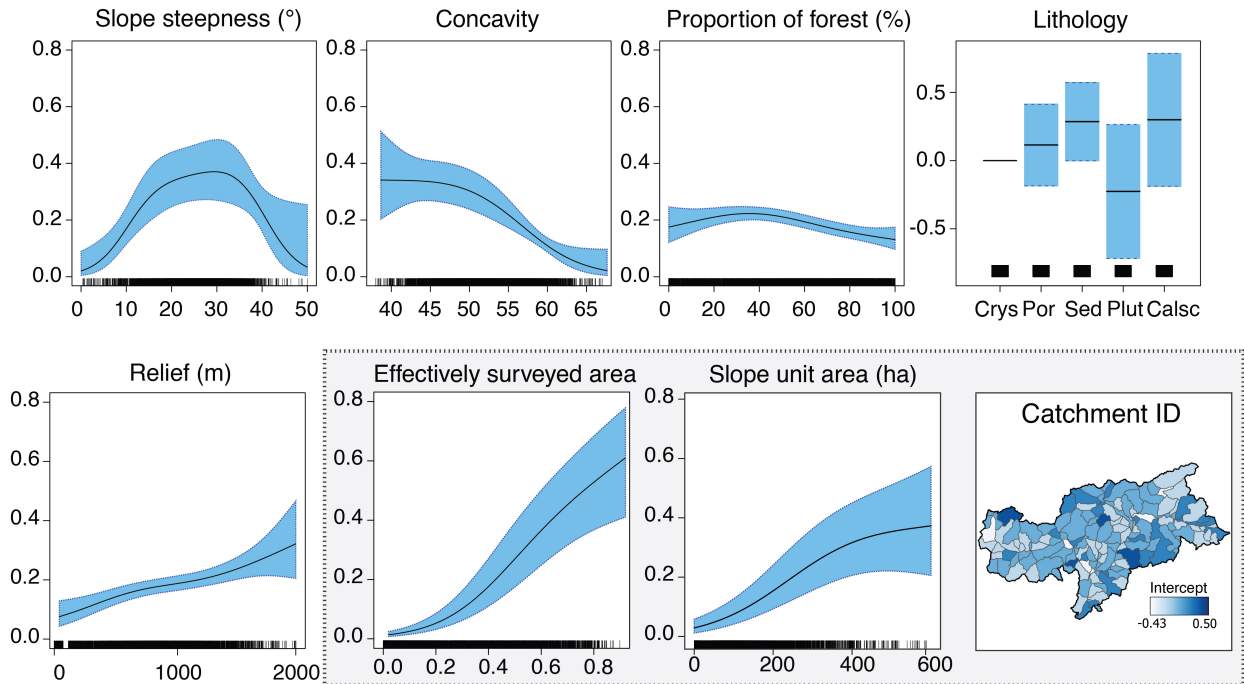


Figure 3: Partial effects of the selected static predictors for the landslide susceptibility estimation. The y-axes are expressed at the response scale probabilities.

469 *Proportion of forest* was the only land cover class that showed statistical significance
 470 in our model. However, its partial effect seems to remain relatively constant, posing a
 471 challenge in describing its specific contribution to landslide susceptibility. Among the scope
 472 of the five lithological classes evaluated, a singular distinction emerged in relation to our
 473 reference class (i.e., crystalline). Specifically, the “sedimentary rocks” category emerged
 474 as statistically significant, positively influencing the estimated landslide susceptibility. The
 475 local *relief* exhibited a predominantly linear impact on the model, with its most substantial
 476 contributions observed in SUs with relief differences of ~ 2000 m.

477 In addition, three more predictors (i.e., *effectively surveyed area*, *slope unit area*, and
 478 *catchment ID*) were included in the model fit but zeroed (i.e., averaged out) from the model
 479 predictions. This step ensures that the model isolates their contributions without directly
 480 or indirectly reproducing their effects in the predictions. Among these predictors, *effectively*
 481 *surveyed area* had the overall highest contribution to the estimated landslide susceptibility.

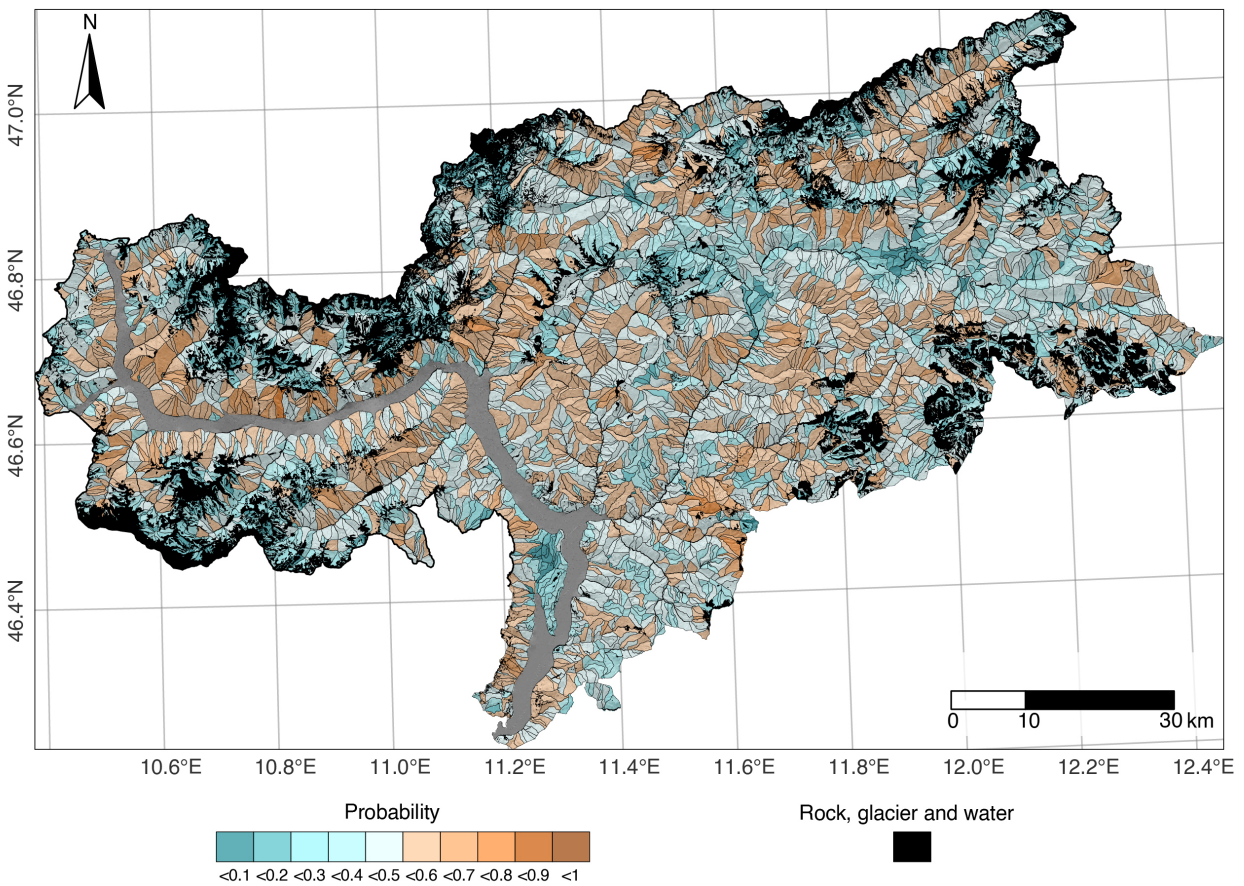


Figure 4: Static landslide susceptibility map. The areas in black consist of the trivial terrains (i.e., rock, glacier, and water bodies). Note that the grey area corresponds to the flood plain of the Etsch-Adige River removed during the SU delineation procedure.

482 A robust linear trend was evident, illustrating that slopes with high effectively surveyed
 483 area scores exerted a considerably greater influence compared to slopes characterized by lower

effectively surveyed area scores. For this reason, we removed its effect from the predictions to ensure that our model predictions do not reproduce this effect. Likewise, *slope unit area* displayed a relatively linear trend with positive influences on the estimated susceptibility.

Notably, the SUs characterized by larger areal extents appear to contribute much more to the estimated landslide susceptibility compared to their counterparts with smaller areal extents. By excluding this effect from the predictions, we avoided potentially misleading interpretations that arise from the mapping unit choice.

The map shown in Figure 4 visually presents the spatial distribution of the modeled relationships described in Subsection 3.1.1. This resulting landslide susceptibility model was subsequently integrated with the best precipitation time windows derived in Section 3.2 to construct a dynamic landslide predictive model.

3.2 Precipitation component

The text below presents the results of the best precipitation time windows to describe landslide occurrence.

3.2.1 Time window selection

We utilized the final modeling sample in a 10-fold RCV framework with 10 repetitions to determine the optimal time windows for representing triggering precipitation T and preparatory precipitation P . The pairwise comparison of model performance is depicted in Figure 5a. The combination of T_1 and P_{15} yielded the best performance, with a median AUROC of 0.85. Figure 5b illustrates the practical interpretation of the chosen time windows. T_1 represents a 2-day cumulative precipitation window (i.e., precipitation on day 0 plus day 1), while P_{15} represents a 14-day cumulative precipitation window prior T (i.e., P_{15} minus T_1).

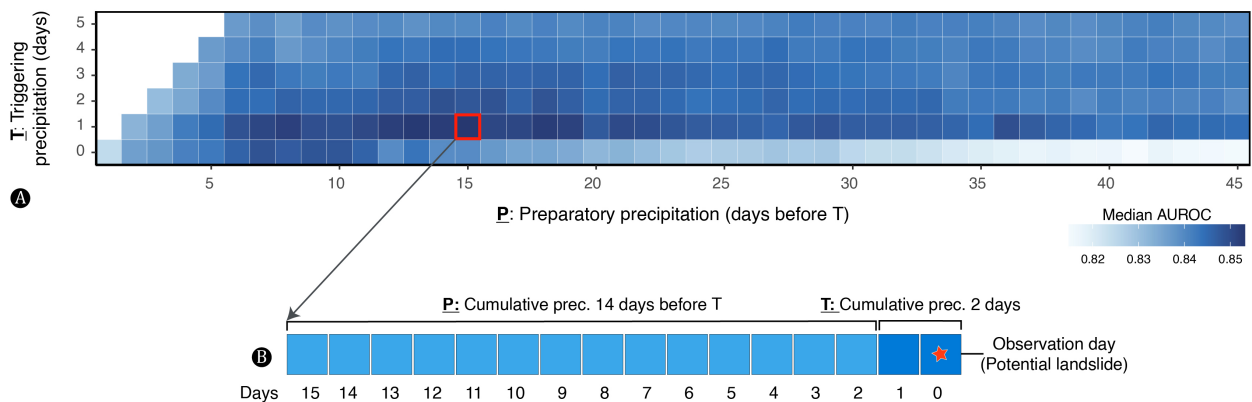


Figure 5: Panel A shows the best combination of precipitation time windows using 10-fold cross-validation with 10 repetitions. The highest AUROC was obtained when combining a triggering precipitation of T_1 and a preparatory precipitation of P_{15} . Panel B shows an interpretation of the best precipitation time windows.

506 Generally, high AUROCs were achieved when combining the 2-day triggering cumulative
 507 precipitation T_1 with preparatory cumulative precipitation ranging from 5 to 20 days (i.e.,
 508 P_6 to P_{21}). Conversely, lower AUROCs were observed when pairing longer triggering precip-
 509 itation windows (e.g., T_3 to T_5) with either very long or very short preparatory precipitation
 510 windows (e.g., P_0 to P_3 or $> P_{30}$). Additionally, lower performances were observed for almost
 511 any combination involving T_0 , as this time window only covers 16 hours of the observation
 512 day and excludes 8 hours of precipitation considered in subsequent time windows (for more
 513 details, see [Steger et al., 2023](#); [Crespi et al., 2021](#)).

514 The resulting best time windows T_1 and P_{15} were then integrated with the static ground
 515 conditions modeled in Section 3.1 to construct a dynamic landslide model.

516 3.3 Dynamic landslide model

517 In this section, we present an interpretation of the dynamic model components, showcase
 518 examples of the estimated probability maps, and discuss the validation of the model.

519 3.3.1 Model relationships

520 Figure 6 offers an overview of the components included in the dynamic model. Similarly to
 521 the static component, we provide a summary of the predictors by describing their partial
 522 effects. The *static susceptibility* exhibits a nonlinear trend, indicating that as the static
 523 landslide susceptibility increases, the dynamic landslide probabilities also increase. In terms
 524 of the preparatory precipitation P_{15} , its contribution shows a relatively linear behavior. This
 525 suggests that as the amount of antecedent 14-day cumulative precipitation increases, the
 526 chance of landslide occurrence also increases. Likewise, the triggering precipitation T_1 can
 527 be described in a similar manner to the preparatory precipitation. However, it shows notably
 528 higher relative probabilities in the domain of high precipitation values. Lastly, it is observed
 529 that the predictor *DoY* indicates slightly lower probabilities around DoY_{200} , corresponding
 530 to the summer season.

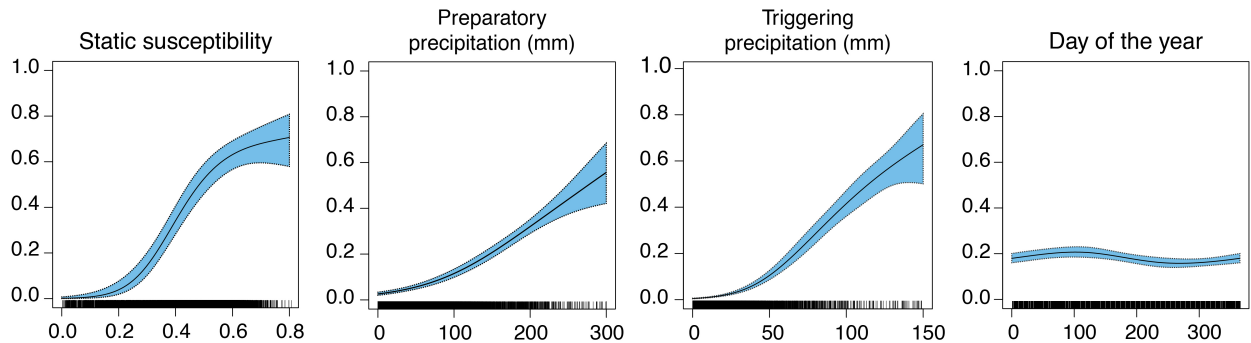


Figure 6: Partial effects of the selected predictors for the dynamic landslide probability estimation. The y-axes are expressed at the response scale probabilities.

531 **3.3.2 Validation and visualization**

532 The validation results are summarized in Figure 7. Overall, the adopted cross-validation
 533 approaches depicted a relatively high model generalization and transferability level, with
 534 AUROCs consistently above 0.85. This level of performance is considered “excellent dis-
 535 crimination” according to Hosmer et al. (2013).

536 Regarding the two cross-validation approaches used to assess the spatial component (i.e.,
 537 RCV and SCV), RCV slightly estimates higher performance scores than SCV. This out-
 538 come is expected because SCV eliminates residual dependence from the spatially distributed
 539 dataset, providing fair indications where the model does not rely solely on its inherent spatial
 540 structure.

541 In terms of the two cross-validation approaches employed to evaluate the temporal com-
 542 ponent, we observe nearly identical median AUROCs, but differences arise in terms of uncer-
 543 tainty or the InterQuartile Range (IQR). The larger IQR observed in the “leave-one-year-out”
 544 approach can mainly be attributed to the imbalanced distribution of landslide observations
 545 across years. Instances with a limited number of landslide observations in specific years may
 546 result in poorer predictions for those particular years.

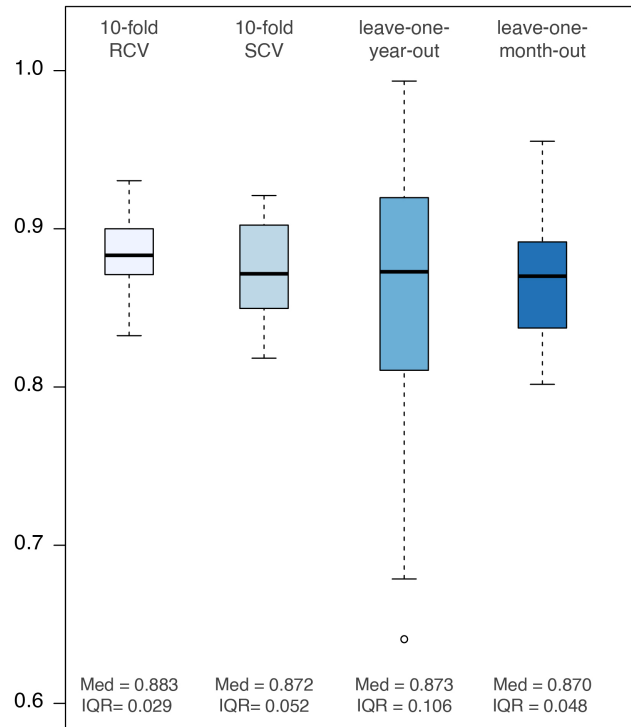


Figure 7: Summary of the model performance featuring the space and time components. The space component is assessed by 10-fold RCV and 10-fold SCV, while the time component is assessed by leave-one-year out and leave-one-month-out validation.

547 To showcase the model’s predictive capabilities, we generated various hypothetical non-
 548 spatially explicit triggering T_1 and preparatory P_{15} precipitation scenarios for the study area,

549 as shown in Figure 8. We considered scenarios with 2-day cumulative precipitation values
550 equal to 0, 50, 100, and 150 mm and with P_{15} equal to 15 (i.e., the equivalent to approxi-
551 mately 1 mm of daily precipitation over the past 14 days) and 63 mm. Furthermore, we use
552 the DoY_{200} to 200, corresponding to July 20th, an arbitrary day during the summer season.
553 Through these hypothetical scenarios, the model demonstrates the dynamic changes in shal-
554 low landslide probabilities based on (i) the static ground conditions, (ii) both preparatory
555 and triggering precipitation components and (iii) seasonality.

556 To further demonstrate the versatility of our model, we conducted a hindcast of a heavy
557 localized precipitation event that occurred on the 5th of August 2016. Using the recorded
558 spatiotemporal distribution of precipitation, we generated predictions from the 15th of July
559 to the 15th of August 2016. The resulting predictions are presented as a GIF file, which can
560 be accessed in the supplementary material of this paper. Figure 9 displays two frames from
561 the animation, specifically covering the 4th and 5th of August. These two frames from the
562 animation clearly depict the relatively high probabilities of landslide occurrence predicted
563 by the model in the area where the event occurred, specifically in the Passeier Valley, in the
564 center north of the study area. This highlights the model’s capability to hindcast landslides
565 in specific regions, even for localized heavy precipitation events.

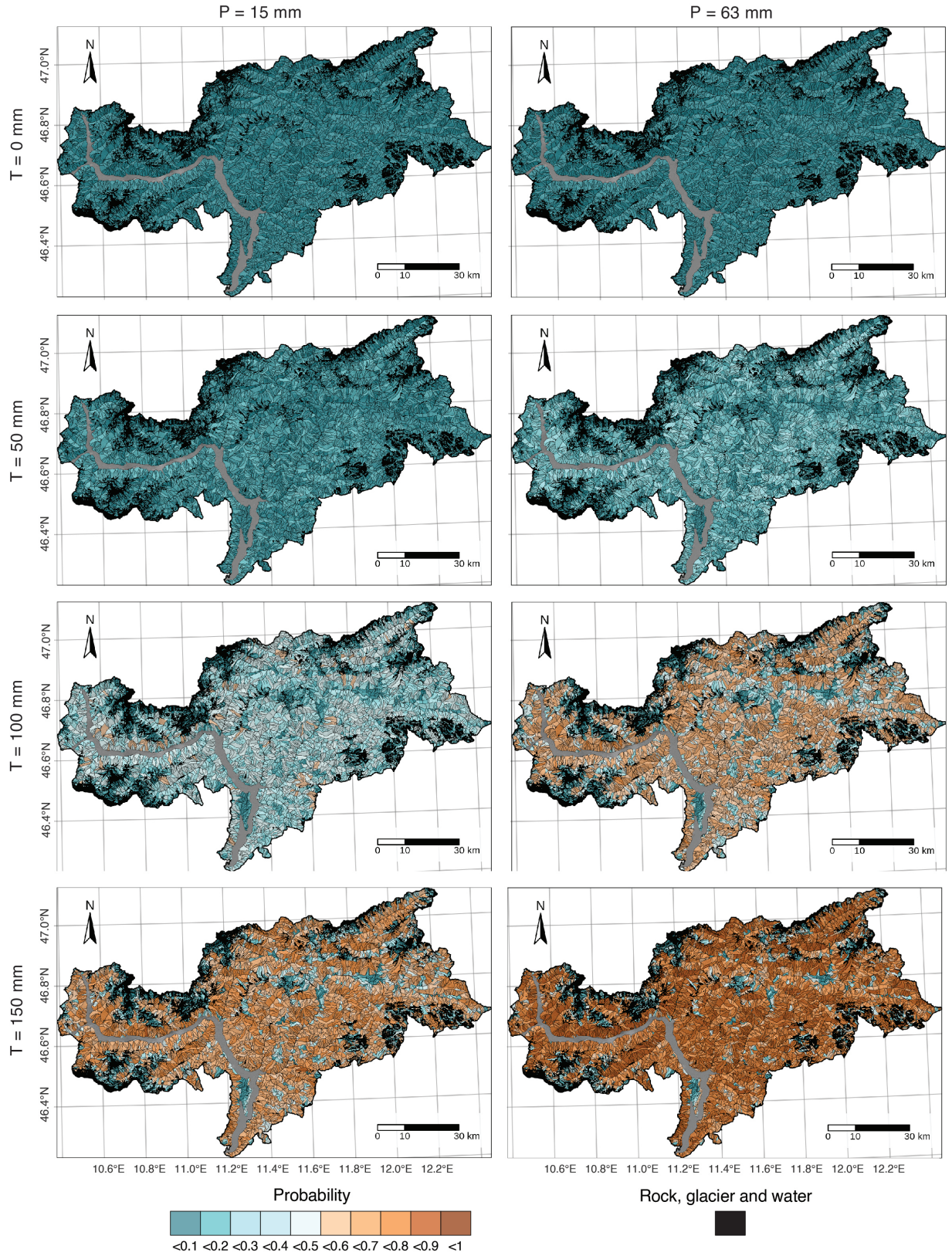


Figure 8: Dynamic landslide predictions with hypothetical non-spatial precipitation scenarios. Preparatory precipitation P_{15} is set to 15 mm and 63 mm, while the triggering precipitation T_1 is set to 0 mm, 50 mm, 100 mm, and 150 mm.

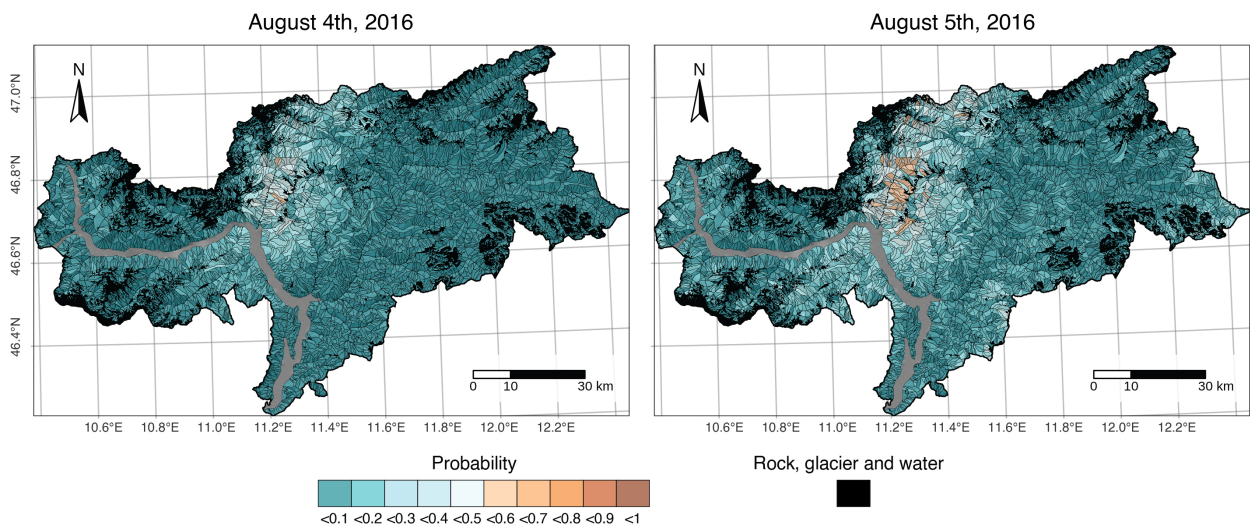


Figure 9: Dynamic landslide predictions for hindcasting landslides associated with precipitation event in the Passeier Valley on the 4th and 5th of August 2016.

4 Discussion

In this research, we developed a data-driven approach to integrate static and dynamic factors for the prediction of precipitation-induced shallow landslides. The resulting model demonstrates a robust performance, consistently achieving AUROCs surpassing 0.85. This reflects the model’s capabilities of accounting for static ground conditions, short-term triggering precipitation, medium-term preparatory precipitation, seasonal effects, and spatial biases. The modeled relationships, the model’s strengths, limitations, and applications are discussed below.

4.1 Understanding the static and dynamic modeled relationships

Fig. 3, we visualized the partial effects of the static predictors used when modeling landslide susceptibility. Below, we provide an interpretation of some of the most relevant ones.

The contribution of *slope steepness* to the landslide susceptibility diminishes for inclinations exceeding $\sim 30^\circ$. This reduction can be explained by the prevalence of thin or absent soil layers on exceedingly steep gradients, caused by, e.g., soil erosion caused by rainfall (Chen et al., 2018). Consequently, given that the initial screening of the landslide inventory was centered around landslides with earth or debris materials, the steep slopes characterized by minimal or absent soil will decrease the probability of landslide incidence.

Only the *proportion of forest* depicted statistical significance among the six distinct land cover classes. Nonetheless, the interpretation of its partial effect encounters some considerations. This ambiguity could potentially be ascribed to the presence of confounding factors (e.g., elevation) that influence the relationship between forest cover and landslide occurrence. Additionally, it is plausible that *proportion of forest* exhibits limited variability across the SUs since, in South Tyrol, forested areas tend to dominate the lower elevation slopes while being less prevalent on higher elevation slopes. Given that the mapping units (i.e., SUs) often span a wide elevation range from the toeslope to the summit, they inherently incorporate an average effect of the *proportion of forest*. This consideration may explain the relatively smoothed effect that *proportion of forest* has in the landslide susceptibility.

The strong influence of *effectively surveyed area* suggests the importance of accounting for it in the analyses. This effect signifies that SUs that have been effectively surveyed, particularly those located in proximity to infrastructure, are estimated as more susceptible to landslides when contrasted with their less-investigated counterparts (i.e., SUs with lower effectively surveyed area scores). From a geomorphological standpoint, this trend does not inherently relate to SUs being more or less susceptible. Instead, it reflects the underlying landslide collection strategy referred to as “data collection effects” (Steger et al., 2021a). The landslide inventory for the study area solely encompasses the landslides that prompted interventions by the provincial offices. Consequently, landslides that did not cause any damage or pose any threat (e.g., landslides far from infrastructure) typically go unrecorded. For these reasons, not accounting for this factor and not excluding it within our model would

604 have resulted in an overrepresentation of high susceptible SUs primarily in close proximity
605 to infrastructure while deeming SUs distant from infrastructure as less susceptible.

606 Similarly, it was found that *slope unit area* substantially influenced the estimated land-
607 slide susceptibility. This may lead to the assumption that as the areal extent of the SU
608 becomes larger, the modeled landslide susceptibility also increases. However, this assump-
609 tion does not hold true because the areal extent of the mapping unit does not necessarily
610 condition whether a terrain is more prone to landsliding. In reality, landslide susceptibility
611 is determined by the interplay of geo-environmental factors and not by the areal extent of
612 the underlying mapping unit. Therefore, neglecting to account for the areal extent of the
613 mapping unit would have resulted in an erroneous estimation, wherein SUs were labeled as
614 more susceptible solely due to their areal extent when compared to their smaller areal extent
615 counterparts. To avoid misleading interpretations caused by the choice of mapping unit, it
616 is important to consider the implications of the areal extent of the SUs.

617 The majority of the literature dedicated to temporal landslide prediction relies on rainfall
618 thresholds, where the presence-absence framework, which considers only the rainfall events
619 leading to landslides while disregarding those that do not trigger landslides, is frequently
620 overlooked. Our optimal time window selection process results reveal that the combination
621 of time windows yielding the best performance includes T 2-day cumulative precipitation and
622 P 14-day cumulative precipitation preceding T , as illustrated in Fig. 5. Notably, our optimal
623 time window combination, particularly the component pertaining to factor P , exhibited a
624 relatively shorter time window compared to the findings in [Steger et al. \(2023\)](#), where the
625 highest-performing combination involved P with 28-day cumulative precipitation prior to T .
626 We attribute this discrepancy mainly to the choice of mapping unit; we employ SUs, whereas
627 pixels were utilized in the previous study. SUs need an aggregation step due to the gridded
628 structure of precipitation datasets. This aggregation process can result in a more averaged
629 representation of precipitation values compared to pixel-scale analyses. Furthermore, it is
630 worth noting that while the combination with the highest median AUROC involves T_1 and
631 P_{15} , the differences among alternative combinations, such as T_1 and P_{29} , are relatively minor.
632 This suggests that the results are relatively stable and not drastically affected by the choice
633 of a slightly different time window.

634 Regarding the partial effects observed in the dynamic model, as depicted in Fig. 6, it
635 is noteworthy that while the variable DoY exhibits a relatively consistent pattern, some
636 fluctuations are evident, particularly along DoY_{200} . Importantly, these fluctuations should
637 not be understood as a reduced likelihood of landsliding across the summer months. Rather,
638 they suggest a more nuanced relationship. Specifically, they hint at the intricate interplay
639 between DoY and the characteristic heavy precipitation patterns typical of Alpine summers,
640 which are, to some extent, captured by the predictors T and P . In practical terms, since
641 our model already accounts for the effects of precipitation, it implies that an equal amount
642 of precipitation (e.g., 100 mm of precipitation) results in a lower probability of landslide
643 occurrence during the summer season. This phenomenon may be attributed to the potentially

644 stabilizing influences of vegetation and temperature on slope stability during this particular
645 season, as also discussed in [Steger et al. \(2023\)](#).

646 4.2 Benefits of a space-time approach

647 The landslide hazard definition initially proposed by [Varnes \(1984\)](#) and subsequently mod-
648 ified by [Guzzetti et al. \(1999\)](#) requires the evaluation of three critical components: spatial
649 assessment (i.e., determining where landslides may occur), temporal assessment (i.e., identi-
650 fying when or under which conditions landslides may occur), and intensity assessment (i.e.,
651 estimating the potential destructiveness of landslides in a given area). In the context of our
652 approach, although we do not encompass all three components, we do effectively address
653 the spatial and temporal aspects. Notably, in the scientific literature, there is a prevalence
654 of landslide susceptibility studies that predominantly focus on the analysis of static factors,
655 effectively covering the spatial component, as highlighted in [Reichenbach et al. \(2018\)](#). How-
656 ever, endeavors are scarce to integrate static and dynamic factors or unify the space and time
657 dimensions into a comprehensive model ([Lombardo et al., 2020](#); [Wang et al., 2022](#)). Also, as
658 emphasized in the introduction of this manuscript, landslides are the outcome of a complex
659 interplay among predisposing, preparatory, and triggering environmental factors. In this
660 interplay, the stability of a slope is portrayed as a function of these three categories of fac-
661 tors ([Crozier, 1986](#)). To illustrate, a slope characterized by predisposing factors contributing
662 significantly to its instability may require only minimal preparatory and triggering factors
663 to become unstable. In the context of our model, a slope characterized by a high landslide
664 susceptibility may demand only limited amounts of medium-term (i.e., P) and short-term
665 (i.e., T) precipitation to undergo failure. Conversely, a slope with low landslide suscepti-
666 bility may require substantial quantities of medium-term and short-term precipitation to
667 reach a state of instability. In contrast to traditional regional approaches where reclassified
668 landslide susceptibility is combined with reclassified precipitation levels through heuristic
669 methods, our approach harnesses the power of GAMMs to integrate both elements while
670 accounting for nonlinearities.

671 Another noteworthy aspect to highlight is the high interpretability and high performance
672 that our model affords, largely attributable to the selected GAMM modeling framework. In
673 particular, the partial effect plots offer a great benefit to gaining insights into how static
674 and dynamic environmental factors may, albeit statistically, contribute to the occurrence of
675 landslides in both space and time. The model consistently demonstrates excellent classifi-
676 cation results, as evidenced by the robust performance scores across various cross-validation
677 routines, including RCV, SCV, leave-one-year-out, and leave-one-month-out assessments.

678 The developed approach’s practical utility and implementation possibilities are presented
679 in two instances. The first instance, exemplified in [Fig. 8](#), pertained to hypothetical pre-
680 cipitation scenarios. These scenarios allowed for the estimation of landslide dynamic prob-
681 abilities under varying precipitation conditions, either with or without spatial variation. In
682 our specific illustration, we opted for non-spatially explicit precipitation conditions. In this

683 manner, the outcomes reveal notable differences in the estimated dynamic landslide proba-
684 bilities between scenarios with relatively low preparatory precipitation and those with high
685 preparatory precipitation, with the latter being much more drastic. In the second instance
686 exemplified in 9, we evaluated our model during a localized, heavy precipitation event that
687 occurred between the 4th and 5th of August 2016. The results consistently demonstrate the
688 model’s ability to predict relatively high landslide probabilities within the affected areas
689 during this period.

690 Furthermore, our approach holds promise for other applications in the realm of early
691 warning systems. By extending the approach developed in [Steger et al. \(2023\)](#) to encompass
692 the space-time framework outlined here, it becomes feasible to utilize dynamic precipitation
693 thresholds while also accounting for static terrain factors. This extension entails assessing
694 whether observed or anticipated precipitation amounts are likely to exceed specific thresholds
695 established for particular terrain and seasonal conditions, such as steep slopes, forested
696 terrain, or low elevations in summer or winter.

697 4.3 Limitations and future work

698 As previously mentioned in the discussion of the model’s benefits, it is important to empha-
699 size that a comprehensive landslide hazard assessment needs to consider the evaluation of
700 landslide intensity as one key component. Although our developed approach addresses spa-
701 tial and temporal aspects, it currently lacks the capability to assess intensity. Consequently,
702 it cannot provide an estimation of landslide hazard, which is a crucial part for conducting
703 landslide risk assessment.

704 Additionally, our approach relies on high-resolution gridded precipitation data sourced
705 from [Crespi et al. \(2021\)](#). However, these data may not be readily available or easily obtain-
706 able if one were to extend the approach to different geographic regions. In such cases, the
707 use of satellite-based precipitation products, as implemented in other recent research studies
708 ([Kirschbaum et al., 2015](#); [Kirschbaum and Stanley, 2018](#); [Stanley et al., 2021](#)), could offer a
709 viable alternative.

710 Furthermore, the selection of the mapping unit posed certain challenges, particularly
711 with respect to the aggregation of predictors. This aggregation process proved to be compu-
712 tationally costly and raised issues, particularly when dealing with static predictors. Those
713 inherent issues arise from the fact that a SU encompasses a spatial domain spanning from
714 the toeslope to the summit. Consequently, aggregating these predictors often leads to unde-
715 sirable over-smoothing effects, as was evident in the case of the *proportion of forest*. In light
716 of these challenges, we advocate for further research within the domain of SU-based landslide
717 models, focusing on the development of appropriate aggregation strategies tailored to spe-
718 cific predictor variables or identifying sets of predictors that do not exhibit these excessive
719 smoothing effects. This research direction will be instrumental in improving the accuracy
720 and interpretability of such models in the future.

721 Moreover, concerning the aggregation of precipitation data, this workflow proved to be

722 relatively time-consuming. Specifically, the precipitation extraction process for the 53,790
723 SUs required approximately 16 hours to complete. This demanding task was executed us-
724 ing the ITC Geospatial Data Analysis Platform <https://crib.utwente.nl/> with a computing
725 configuration featuring 72 vCPU Intel x86-64/768 GB RAM/NVIDIA RTX A4000 GPU.

726 Our future research endeavors will center on the integration of additional dynamic pre-
727 dictors, which includes an examination of the intricate relationships between precipitation,
728 temperature, and snow (Camera *et al.*, 2021; Bajni *et al.*, 2023). We plan to expand our
729 investigations by extending our model framework to sub-daily precipitation measurements,
730 which can be derived from meteorological RADAR observations. This finer temporal resolu-
731 tion in precipitation data can provide valuable insights into short-term variations and their
732 influence on landslide occurrences. Another important aspect to explore involves evaluat-
733 ing the predictive capabilities of our model under various projected precipitation scenarios,
734 particularly in the context of climate change. This step may help us understand how our
735 approach performs under altered precipitation patterns, which are anticipated to become
736 increasingly relevant in the future. Furthermore, we aim to enhance our model by integrat-
737 ing an intensity estimation component. In theory, this integration may enable offering a
738 complete assessment of landslide hazard, thereby advancing our ability to assess and reduce
739 landslide risk effectively.

740 5 Conclusion

741 We have presented a novel approach employing GAMMs to effectively integrate static ground
742 conditions and dynamic preparatory and triggering factors for the prediction of shallow land-
743 slides in both space and time within South Tyrol in the northeastern Italian Alps from 2000
744 to 2020. The static predisposing factors, featured by the *static susceptibility* predictor,
745 are integrated with dynamic factors such as daily cumulative precipitation represented by
746 medium-term precipitation, short-term precipitation, and the day of the year to account for
747 seasonal variations. Additionally, our SU-based approach addresses critical issues related to
748 spatial biases in landslide inventory, potential reporting biases across different years, and the
749 areal extent of the chosen mapping unit. Through a comprehensive multi-validation strategy
750 in both space and time, our model consistently exhibited excellent predictive performance,
751 consistently achieving AUROC values above 0.85. Moreover, we have demonstrated the
752 practical applicability of our approach in two specific scenarios. Firstly, we employed our
753 approach to estimate landslide probabilities under hypothetical precipitation scenarios. Sec-
754 ondly, we successfully utilized the model for hindcasting, effectively reproducing landslides
755 triggered by heavy and localized precipitation events. These applications underscore the
756 real-world utility of our approach in enhancing landslide prediction and extending it into
757 space and time.

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