

# Wildfire smoke exposure and mortality burden in the US under future climate change

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# 1 Abstract

2 Wildfire activity has increased in the US and is projected to accelerate under future climate change.  
3 However, our understanding of the impacts of climate change on wildfire smoke and health remains  
4 highly uncertain. Here we quantify the mortality burden in the US due to wildfire smoke fine  
5 particulate matter (PM<sub>2.5</sub>) under future climate change. We construct an ensemble of statistical  
6 and machine learning models that link climate to wildfire smoke PM<sub>2.5</sub>, and empirically estimate  
7 smoke PM<sub>2.5</sub>-mortality relationships using georeferenced data on all recorded deaths in the US  
8 from 2006 to 2019. We project that climate-driven increases in future smoke PM<sub>2.5</sub> could result  
9 in 27,800 excess deaths (95% confidence interval: 13,100 - 43,400) per year by 2050 under a high  
10 warming scenario (SSP3-7.0) – a 76% increase relative to estimated 2011-2020 averages. Cumulative  
11 excess deaths from wildfire smoke PM<sub>2.5</sub> could exceed 700,000 between 2025-2055. When monetized,  
12 climate-induced smoke deaths result in annual damages of \$244 billion, comparable to prior *aggregate*  
13 estimates of all other economic damage due to climate change. Our research suggests that the health  
14 cost of climate-driven wildfire smoke could be among the most important and costly consequences  
15 of a warming climate in the US, and an urgent adaptation priority.

## 16 Introduction

17 Wildfire activity has increased substantially over the US in the last two decades, with the largest  
18 increases observed in the western US (1–5). As a result, air pollution that is associated with  
19 wildfire smoke (specifically fine particulate matter, PM<sub>2.5</sub>) has significantly increased (6–9). Given  
20 established relationships between ambient smoke PM<sub>2.5</sub> exposure and poor health (10–13), these  
21 increases have likely worsened several health outcomes. In many parts of the western US, smoke  
22 PM<sub>2.5</sub> accounted for over 50% of the annual concentration of PM<sub>2.5</sub> in extreme smoke years (14,  
23 15), and has led to stagnation or even reversal of the otherwise declining trend in ambient PM<sub>2.5</sub>  
24 over the last two decades due to the Clean Air Act (16).

25 Mounting evidence has suggested that human-induced climate change is a leading cause for the  
26 increased wildfire activity, especially in forested areas in the western US (2–4, 17–19), alongside  
27 other important causes that include historical fire suppression and the expansion of human activities  
28 into forested areas (20). A warming climate can influence wildfire activities by altering the aridity  
29 of the fuel (2, 21), conditions for fire spread (22, 23), as well as lightning ignitions (24). For the  
30 western US, many studies have projected increasing wildfire risks under a warming climate primarily  
31 due to increasing fuel aridity under higher ambient temperature (25–27).

32 However, the relationship between a warming climate and the resulting increase in wildfire smoke  
33 and health impacts is not fully understood and highly uncertain. Several studies use regression  
34 models or land-vegetation-fire models to first project the wildfire activities under future climate  
35 and then utilize chemical transport models to estimate changes in smoke PM<sub>2.5</sub> concentrations (28–  
36 32) and associated health outcomes (33–36). However, prior projections of future mortality due to  
37 climate-driven fire smoke span a wide range of uncertainties (37) – reflecting an important knowledge  
38 gap given the large potential impacts. Uncertainties in the prior projections come from three key  
39 sources. First, large uncertainties exist in how wildfire emissions respond to climate change (38).  
40 Second, modeling fire impacts on surface PM<sub>2.5</sub> often faces large uncertainty in emission inventories  
41 (39, 40), the vertical distribution of emission profiles (41), and fire-weather interactions (42), which  
42 results in modeled smoke concentration sometimes differ by an order of magnitude when compared  
43 to surface observations (43). Third, most prior studies quantify the health impacts of smoke PM<sub>2.5</sub>  
44 by applying existing concentration-response functions derived from total PM<sub>2.5</sub> exposures, which  
45 could fail to capture unique health impacts of smoke PM<sub>2.5</sub> exposure, such as from smoke-specific  
46 chemical composition and toxicity (44) or behavioral responses unique to smoke events (13).

47 Because of these challenges, very few studies to date have projected future smoke PM<sub>2.5</sub> con-  
48 centrations using empirically grounded relationships between climate, wildfire, and PM<sub>2.5</sub> (35, 45).  
49 To our knowledge, no studies have estimated the future smoke mortality burden accounting for  
50 the unique health impacts of smoke PM<sub>2.5</sub> using dose-response functions that are specific to smoke  
51 pollution exposure. Absent this quantification, leading estimates of the societal impact of climate

52 change – many of which are directly used to guide policy – do not incorporate potential mortality  
53 impacts due to wildfire smoke  $\text{PM}_{2.5}$  (46–48). Detailed projections of future smoke  $\text{PM}_{2.5}$  exposure  
54 and health burden are crucial to inform policies to mitigate and adapt to the negative impacts of  
55 smoke  $\text{PM}_{2.5}$  on humans.

56 In this paper, we develop a comprehensive, data-driven approach that directly address all three  
57 of the above challenges. First, to improve understanding of the climate-fire emissions relationship,  
58 we construct an ensemble of statistical and machine learning models that predict fire emissions as  
59 a function of climate and land-use variables over North America (including Mexico and Canada),  
60 using observational data from 2001-2021. By using historical data that includes recent years with  
61 extreme weather conditions (e.g., drought in the western US in 2020), which is projected to increase  
62 under future climate change, our ensemble of models can better characterize how climate influences  
63 wildfire emissions in future scenarios. We use dry matter emissions from the fourth version of  
64 the Global Fire Emissions Database with small fires (GFED4s) (49), which include fire emissions  
65 from agriculture fires and land-use change. However, as wildfire emissions dominate in most of  
66 our studied regions (especially in the western US and Canada where we see the largest effects, see  
67 Table S1), we refer to our estimates as “wildfire emissions” and “wildfire smoke  $\text{PM}_{2.5}$ ” for simplicity  
68 and consistency. By modeling changes in wildfire emissions in Canada and Mexico, our approach  
69 can also capture important transboundary influences on US smoke  $\text{PM}_{2.5}$  and health effects, such  
70 as those that occurred in the summer of 2023 (50). Second, we use surface wildfire smoke  $\text{PM}_{2.5}$   
71 estimates from (8) to establish an empirical relationship between wildfire emissions and smoke  
72  $\text{PM}_{2.5}$  concentration across the contiguous US at 10 km resolution, accounting for variation in wind  
73 directions and spatial transport. Our approach fits the observed surface  $\text{PM}_{2.5}$  data well (Figure  
74 S2), and allows us to efficiently predict smoke concentration in one location from changes in wildfire  
75 emissions in another (Methods). Third, to address the challenge of accurately estimating the health  
76 impacts of ambient smoke exposure, we empirically estimate the effects of annual smoke  $\text{PM}_{2.5}$   
77 concentration on annual mortality rates using county-level data from 2006 to 2019 on all recorded  
78 deaths in the US. We estimate dose-response functions using a Poisson model in which we allow non-  
79 linear impacts of smoke  $\text{PM}_{2.5}$  on mortality rates, consistent with prior papers that examine smoke  
80 impacts on mortality and other health outcomes (13, 51), while flexibly controlling for temperature,  
81 precipitation, and a broad range of possible spatial and temporal confounds (Methods).

82 Finally, we combine the empirical relationships between climate, wildfire emissions, smoke  $\text{PM}_{2.5}$ ,  
83 and mortality rates derived above with projected climate variables derived from CMIP6 global  
84 climate model ensembles to generate future projections of smoke  $\text{PM}_{2.5}$  and mortality burden. We  
85 project the annual average smoke  $\text{PM}_{2.5}$  concentration in each 10 km location across the contiguous  
86 US (48 states and the District of Columbia) between 2046 and 2055 under different climate scenarios.  
87 We then quantify changes in mortality rates in each county in the contiguous US between 2050 and  
88 the historical period, and the difference across three climate scenarios (SSP1-2.6, SSP2-4.5, and

89 SSP3-7.0) to quantify the potential health benefits from climate mitigation and adaptation. We  
90 quantify the uncertainty in the final projected mortality burden across the different components of  
91 our modeling framework and compare our mortality estimates with estimates of direct temperature-  
92 related mortality burden and aggregate climate costs from prior work (46, 52, 53) to contextualize  
93 the importance of climate-smoke channels relative to other known climate impacts.

## 94 Results

### 95 Empirical relationship between climate and smoke $PM_{2.5}$

96 We considered three different statistical and machine learning frameworks for modeling the climate-  
97 fire relationship (Methods). To account for geographical heterogeneity, we estimated each of our  
98 frameworks separately by region, resulting in five ensembles of climate-fire models. Our models  
99 can capture the variability of wildfire dry matter emissions at 10-year intervals (to account for fire  
100 stochasticity at the annual level, see Methods), highlighting their ability to quantify changes in  
101 wildfire emissions under different climate conditions (Figure 1A). When evaluating through cross-  
102 validation of temporal blocks (i.e. randomly splitting a time series of observations into disjoint sets of  
103 training and testing years), our models achieve high prediction performance, especially in the western  
104 US, Canada, and Mexico, with correlation coefficients of 0.87-0.95 in the out-of-sample evaluations  
105 (Table S2). Under these evaluation criteria, our model achieves higher performance relative to other  
106 commonly-used regression methods such as a log-linear model to model climate impacts on burned  
107 area (2), as well as more flexible machine learning methods (38) (Figure S1). However, the model  
108 performance indicates that climate conditions are not the only factors influencing the variability  
109 of wildfire emissions over time. For example, we find that the model performs less well in the  
110 southeastern US and northeastern US, where many fires are agricultural or prescribed fires, which  
111 are less directly influenced by climate factors (54). Furthermore, while our models can predict  
112 spatially- and temporally-aggregated emissions effectively, the predictive performance deteriorates  
113 when the same model is evaluated at finer temporal and spatial resolutions (Figures S3 and S4).  
114 Such evaluation results are consistent with prior literature on global fire modeling (55). Our findings  
115 suggest that, although climate conditions such as low soil moisture and high ambient temperatures  
116 are related to enhanced fire activity in aggregate, whether a fire occurs in a specific location depends  
117 on more stochastic factors such as lightning and human ignitions that are very hard to predict (56).

118 Combining our statistical and machine learning models with future climate projections from  
119 CMIP6 global climate models, we project that wildfire emissions will increase by 2050 in all study  
120 regions except for the eastern US (Figure 1B). The largest increases in wildfire emissions are pro-  
121 jected in the western US, where the model estimates that the annual wildfire emissions will increase  
122 by between 248% (SSP1-2.6) and 470% (SSP3-7.0) in the 2050s relative to average emissions during

123 2011-2020. When compared to 2020, the largest wildfire year for the western US in our historical  
124 data, projected annual wildfire emissions during the 2050s will either reach (as in the case of SSP1-  
125 2.6) or exceed (by 34% under SSP2-4.5 or 62% under SSP3-7.0) emissions observed in 2020. This  
126 magnitude of increases is largely consistent with prior estimates of the western US derived from  
127 statistical models and process-based models (26, 27, 31). Consistent with prior literature, we find  
128 that decreased soil moisture and increased ambient temperature, especially in the forest areas in the  
129 western US, are the leading contributors to increased wildfire emissions (Figure S5, Table S3, Table  
130 S4). In the eastern US, we estimate a decrease of wildfire emissions by 15% under SSP1-2.6 and an  
131 increase of wildfire emissions by 10% under SSP3-7.0. These opposing predictions are driven by a  
132 combination of two conflicting factors: projected increases in ambient temperature, which increase  
133 emissions, and projected increases in precipitation, which decrease projected emissions (Figure S5).  
134 Our projected patterns in the eastern US are consistent with a prior study that used a process-based  
135 fire-climate model (31). By the 2050s, we project an increase in emissions of 33-43% in Mexico, and  
136 of 30-49% in Canada, relative to average emissions during 2011-2020, in large part due to projected  
137 increases in Vapor Pressure Deficit (VPD).

138 To link wildfire emissions to smoke  $PM_{2.5}$  concentrations, we develop an empirical relationship  
139 that accounts for wind direction, distance from fire, and geographical region (Figure 2). As shown  
140 in Figure 2A, we find that wildfire emissions increase smoke  $PM_{2.5}$  concentrations near an active  
141 fire, with the effects gradually decaying as the distance from the fire increases. Consistent with  
142 previous evidence of long-range transport of smoke (57, 58), we find a statistically significant effect  
143 ( $p < 0.05$ ) of wildfire emissions on downwind locations up to 1000 km away. We find substantial  
144 regional heterogeneity in the impacts of dry matter emissions on wildfire  $PM_{2.5}$  (Figure 2B). For  
145 example, we find that one ton of dry matter emissions (as estimated in GFED4s fire emissions  
146 database) can generate as much as 3x surface smoke  $PM_{2.5}$  in the Northwest compared to the  
147 Southwest and South. Such regional heterogeneity likely reflects a multitude of factors, such as  
148 vegetation type, vegetation density, and fire intensity (Methods).

## 149 **Projected smoke $PM_{2.5}$ concentration under future climate**

150 As a result of projected rising wildfire emissions, we find increases in annual smoke  $PM_{2.5}$  con-  
151 centrations throughout the US in 2050 under all future climate scenarios (Figure 3A). Under our  
152 highest warming scenario (SSP3-7.0), we estimate that annual average smoke  $PM_{2.5}$  concentration  
153 could reach  $10 \mu\text{g}/\text{m}^3$  in some regions on the west coast, a level that has only been observed in  
154 extreme smoke years such as 2020 (8). While the most substantial changes in smoke  $PM_{2.5}$  happen  
155 across the western US, smoke  $PM_{2.5}$  concentrations are also projected to increase in the northeast  
156 US, largely due to projected increases in wildfire emissions in the western US and Canada and  
157 subsequent increases in cross-boundary transport of wildfire smoke from these fires.

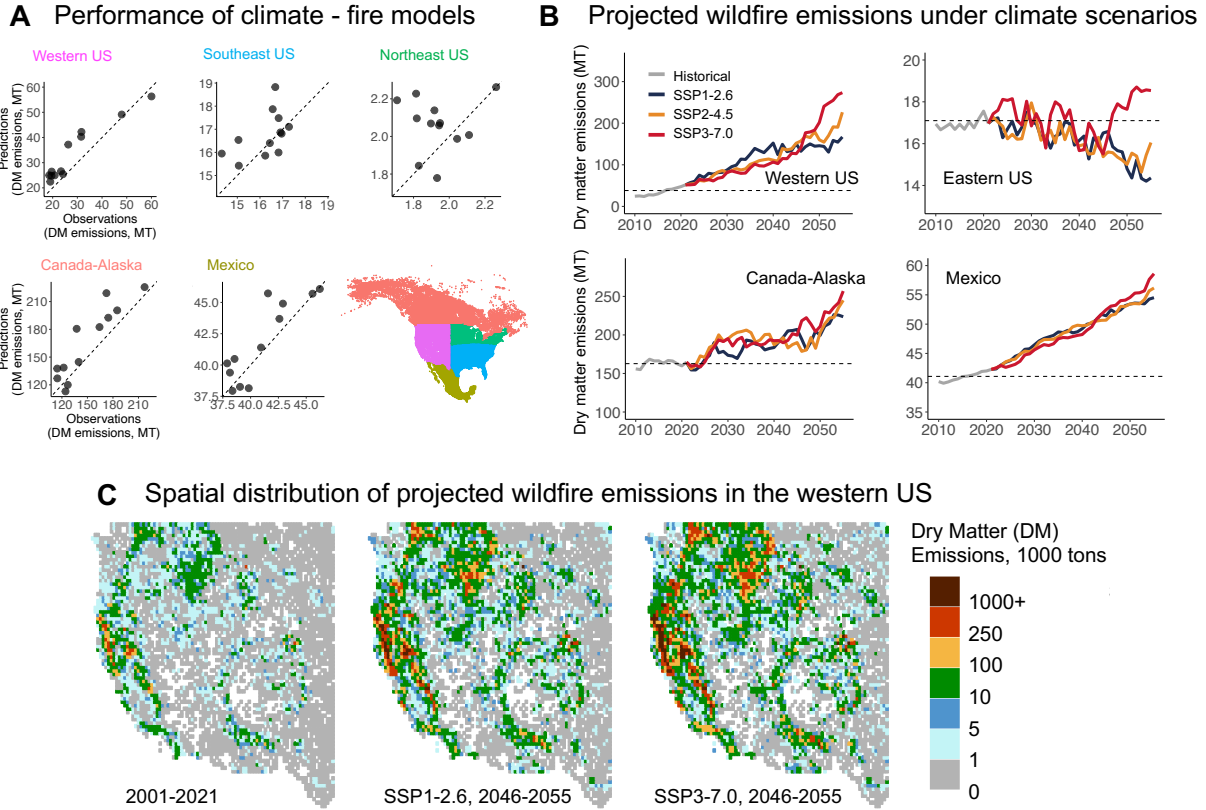


Figure 1: **Projected wildfire emissions under future climate change scenarios.** Panel A: Performance of the statistical and machine learning ensemble models. We build separate models to predict wildfire Dry Matter (DM) emissions for five regions respectively: Western US, Southeast US, Northeast US, Canada-Alaska, and Mexico. The plot shows the 10-year moving average of predicted emissions (y-axis) against the observed emissions (x-axis), aggregated at the regional level. Panel B: Projected wildfire emissions (unit: Million Tons, MT) under the historical scenarios and three future climate scenarios (SSP1-2.6, SSP2-4.5, and SSP3-7.0). The plot shows the 10-year moving average of the wildfire emission projections. The dashed line represents the average observed emissions over 2001-2021 for each region. For presentation purpose, we aggregate predictions from northeast US and southeast US to calculate the total for eastern US. Panel C: Observed DM emissions at the native resolution (0.25 degree) in 2001-2021 from GFED4s, and projected annual emissions averaged between 2046-2055 under SSP1-2.6 and SSP3-7.0 scenarios (down-scaled from aggregated projections).

158 We find that the relative contribution of wildfire smoke to total population-weighted  $PM_{2.5}$   
 159 increases by 240-320% in 2050. This finding holds even if non-smoke  $PM_{2.5}$  remains constant – a  
 160 conservative assumption given recent and ongoing declines in non-smoke  $PM_{2.5}$  concentrations (16).  
 161 We estimate that smoke  $PM_{2.5}$  will account for 13-17% of total population-weighted  $PM_{2.5}$  in the  
 162 US in 2050, which is 2-3x its contribution of 5.4% during 2011-2020. Wildfire smoke will account

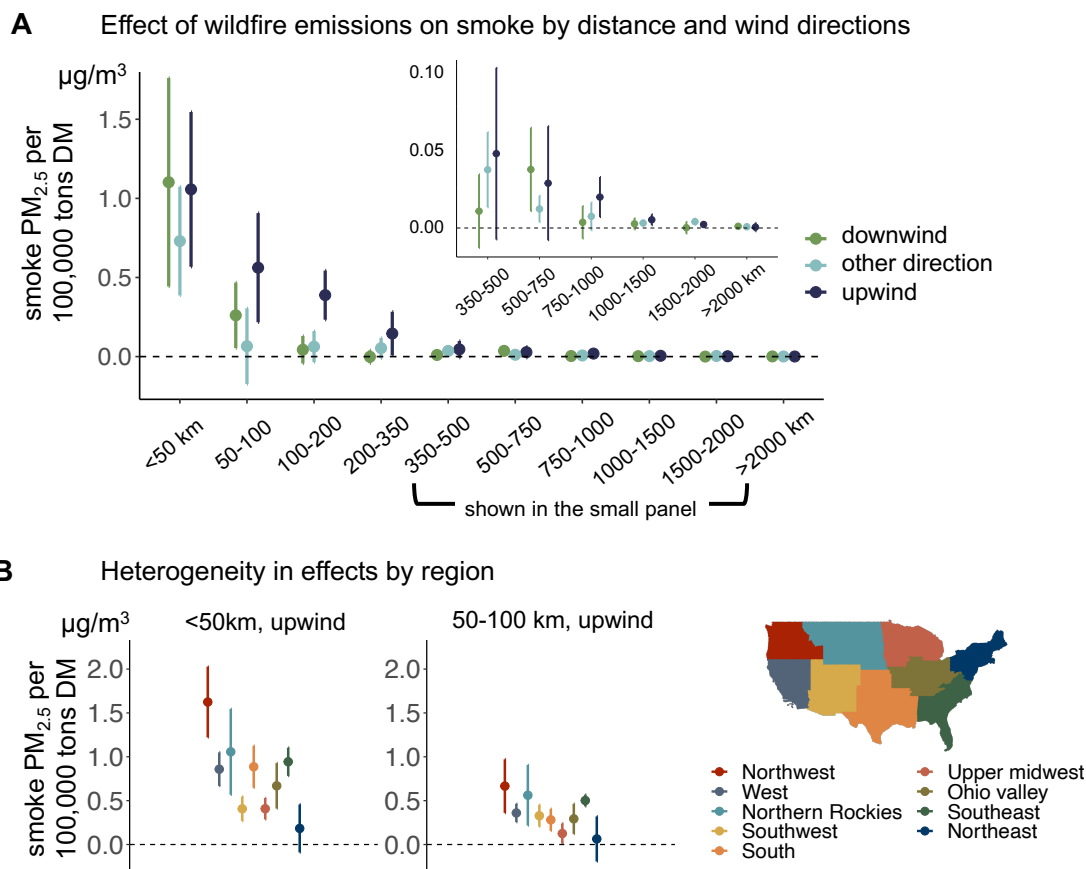


Figure 2: **Wildfire emissions increase the observed smoke  $PM_{2.5}$  concentration in the neighboring and downwind areas.** Panel A: The empirically estimated effects of wildfire emissions on smoke  $PM_{2.5}$  by distance from emissions and wind directions. “Upwind” means the fire is upwind of the location at which  $PM_{2.5}$  is measured. Wildfire emissions are estimated to have larger impacts on smoke  $PM_{2.5}$  when smoke location is closer to fire (distance to emissions is shown on the x-axis), and when wildfire emissions happen upwind of the smoke locations (wind patterns shown in colors). Separate models are estimated for the 9 climatic regions in the US determined by National Centers for Environmental Information (as shown in Panel B). Panel A shows the results in the *Northern Rockies* region. Panel B: Regional heterogeneity in emission impacts on smoke  $PM_{2.5}$ . Panel B shows the estimated effects of *upwind* emissions in the <50 km and 50-100 km bins, across the nine regions in the US.

163 for at least 15% of total population-weighted  $PM_{2.5}$  in 17 states, including states both in the West  
 164 such as Oregon (with 61% smoke contribution), Washington (56%), and California (30%), as well as  
 165 states in the South and Midwest such as Oklahoma (19%) and Minnesota (16%). Figure 3B shows  
 166 the smoke contribution in the top 10 states (see Table S5 for more states).

167 Under the SSP3-7.0 scenario, average population-weighted smoke  $PM_{2.5}$  exposure is projected to  
 168 reach  $1.47 \mu\text{g}/\text{m}^3$ , an increase of over 200% relative to the average level between 2011-2020 (Figure



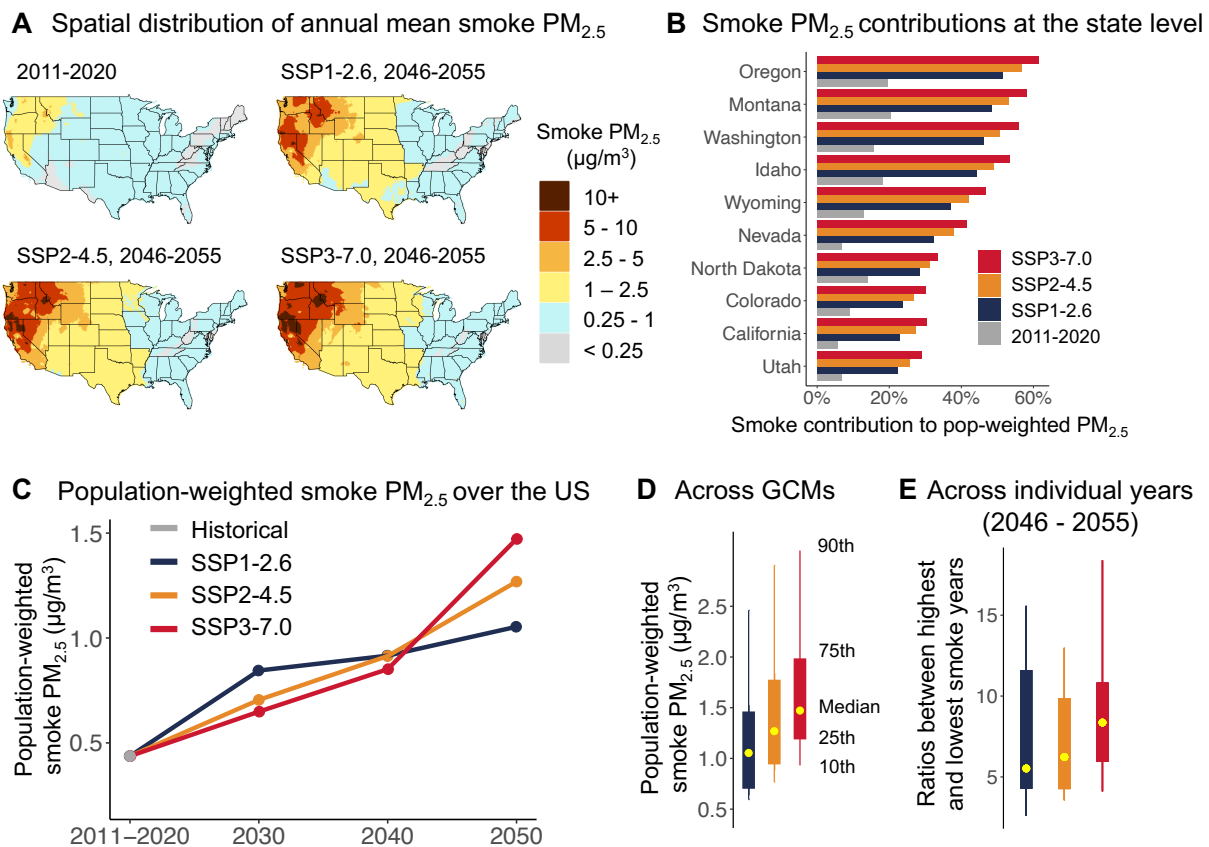


Figure 3: **Population exposure to wildfire smoke  $PM_{2.5}$  increases by 2- to 3-fold under future climate change scenarios.** Panel A: The annual mean smoke  $PM_{2.5}$  concentration in the historical data (2011-2020), and projected annual mean smoke  $PM_{2.5}$  concentration under the three climate scenarios in 2046-2055. Panel B: the contribution of smoke  $PM_{2.5}$  to total population-weighted  $PM_{2.5}$  at the state level. Non-smoke  $PM_{2.5}$  is calculated as the difference between total  $PM_{2.5}$  (derived from (59)) and smoke  $PM_{2.5}$  in 2016-2020, and is assumed to be constant in future. The panel only lists the top ten states with the highest smoke contribution under SSP3-7.0 scenario in 2050. Panel C: population-weighted smoke  $PM_{2.5}$  over the US in different decades. Panel D: uncertainty in the population-weighted smoke  $PM_{2.5}$  across the 28 GCMs used in the projection. Panel E: for each GCM, we calculate the ratio between the highest and lowest projected population-weighted smoke  $PM_{2.5}$  during 2046-2055. The panel shows the quantiles of these ratios across the 28 GCMs.

169 3C), and 1.6x the population-weighted smoke  $PM_{2.5}$  concentration in the historically extreme year  
 170 of 2020 ( $0.90 \mu g/m^3$ ). The differences across the three climate scenarios are negligible in 2030 and  
 171 2040 due to little difference in projections of the climate variables (Figure S5). However, by the  
 172 2050s, population-weighted smoke  $PM_{2.5}$  is meaningfully smaller in the low warming scenarios, at  
 173  $1.05 \mu g/m^3$  under SSP1-2.6 or  $1.27 \mu g/m^3$  under SSP2-4.5, averaged across GCMs. Some individual

174 GCMs project much larger or smaller increases (Figure 3D). Also, these estimates represent decadal  
175 averages of annual smoke PM<sub>2.5</sub> concentrations, in this case averaged 2046 to 2055. Given interan-  
176 nual climate variability, projections suggest that average smoke PM<sub>2.5</sub> concentrations in individual  
177 years could differ substantially, with the highest projected smoke year having roughly 5-10x the  
178 concentration of the lowest year (Figure 3E). Our method likely underestimates the interannual  
179 variability as it does not capture variability in non-climate factors.

## 180 Mortality burden due to smoke PM<sub>2.5</sub> exposure

181 We find that exposure to annual smoke PM<sub>2.5</sub> increases all-age mortality rates (Figure 4A), even  
182 at low smoke concentrations ( $<1 \mu\text{g}/\text{m}^3$ ), consistent with recent evidence from studies of low levels  
183 of all-source PM<sub>2.5</sub> (60). Compared to a year of zero or minimal smoke PM<sub>2.5</sub> (annual mean  
184 concentration  $<0.1 \mu\text{g}/\text{m}^3$ ), we find that a year with annual average smoke PM<sub>2.5</sub> of 0.75-1  $\mu\text{g}/\text{m}^3$   
185 increases county-level mortality rate by 1.3% (95%CI: 0.6%, 2.0%). Years with extreme ambient  
186 wildfire smoke concentrations ( $>6 \mu\text{g}/\text{m}^3$ ) increase annual mortality rates by 5.8% (95%CI: 2.2%,  
187 8.9%). Wildfire smoke increases mortality rates among both the elderly and the general population  
188 (Figure S6). Our estimated smoke-mortality relationship is similar in shape to the results estimated  
189 by (51) at the county-month level. For a given increase in PM<sub>2.5</sub> concentration by 1  $\mu\text{g}/\text{m}^3$ , our  
190 observed effects for smoke PM<sub>2.5</sub> exceed a recent meta-analysis estimate for all-source PM<sub>2.5</sub> (0.8%  
191 increase in mortality rates per 1  $\mu\text{g}/\text{m}^3$  (61)), although our confidence interval contains this lower  
192 estimate.

193 Combining our empirically-derived dose-response function and historical smoke PM<sub>2.5</sub> concen-  
194 trations, we estimate that smoke PM<sub>2.5</sub> caused 15,800 excess deaths (95% CI: 6900, 25300) per  
195 year during 2011-2020 (Figure 4B), relative to a counterfactual of no smoke PM<sub>2.5</sub>. This number  
196 of smoke-related deaths would account for 9.2% of total estimated deaths due to total (smoke and  
197 non-smoke) PM<sub>2.5</sub> exposure during the same period (estimated using the response function from  
198 (61) and total PM<sub>2.5</sub> estimates from (59)). As shown in Figures 4B and S7, roughly 90% of esti-  
199 mated excess deaths from wildfire smoke exposure come from relatively low but frequent exposures  
200 to annual concentrations below 1  $\mu\text{g}/\text{m}^3$ .

201 We estimate that smoke PM<sub>2.5</sub> will cause 23,800 to 27,800 annual excess deaths by mid-century  
202 across the three climate scenarios – an increase of 51-76% in mortality burden from smoke relative  
203 to 2011-2020. Even under the low warming scenario (SSP1-2.6), we estimate that smoke PM<sub>2.5</sub> will  
204 lead to 8,000 more annual excess deaths in the 2050s relative to today. Over the period of 2025-2055,  
205 we estimate that wildfire smoke PM<sub>2.5</sub> could lead to cumulative excess deaths of 690,000 (SSP1-2.6)  
206 to 720,000 (SSP3-7.0). Although in the historical period, annual mean wildfire smoke concentrations  
207 above 5  $\mu\text{g}/\text{m}^3$  were rare and represented only 3% of the total mortality burden (Figure 4A), we  
208 estimate that these more extreme years will account for between 20-26% of the total excess deaths

209 from smoke in the 2050s (Figure S7). The climate-induced smoke deaths are distributed across  
210 populous counties in the western US as well as in the Midwest, Northeast, and South (Figure 4C).  
211 The top five states that are predicted to experience the largest increases in annual smoke  $PM_{2.5}$   
212 deaths in 2050s under SSP3-7.0 are California (3300 excess deaths per year), Washington (900),  
213 Texas (680), Oregon (610), and Florida (380). While projected smoke concentrations are highest  
214 in the western US, almost half of the smoke mortality come from eastern states (east of  $95^\circ$  W)  
215 due to higher population densities and damages from low wildfire smoke concentrations (Figure S8  
216 and Table S7). Estimated mortality effects are largely robust across alternative specifications of  
217 the smoke-mortality models including alternative functional forms, temporal aggregations, and bin  
218 definitions (Figure S9 and S10).

219 We contextualize the magnitude of these mortality impacts in two ways. First, we compare  
220 our estimates of excess deaths from climate-driven smoke  $PM_{2.5}$  to the direct effects of extreme  
221 temperatures on mortality – an impact which has been the primary focus of climate change impacts  
222 on mortality and is projected to be one of the leading economic costs of global climate change  
223 (47, 48, 52, 62). Recent studies find that, by mid-century in the US, increasing mortality from  
224 more frequent extreme heat is likely to be more than offset by declining mortality due to cold  
225 weather with a projected decrease in annual excess deaths of 15,800 by mid-century (under the  
226 SSP2-4.5 scenario) compared to 2001-2010 (52). Our projected increase in smoke mortality over  
227 the same period represents 62% of this reduction in direct temperature-related deaths (Figure 4D),  
228 significantly offsetting a potential benefit of future warming in the US. However, as shown in Figure  
229 4E, the size of this offset differs across the US, with certain states likely to suffer compounded  
230 consequences from increases in both smoke-related and heat-related deaths (e.g., CA, TX, FL), and  
231 other states likely to see minimal smoke-related mortality and a substantial decline in heat-related  
232 deaths (e.g., IL).

233 As a second comparison, we compare our estimates of climate-induced smoke damages with two  
234 prior estimates of aggregated monetized damage due to climate change. Using a Value of Statistical  
235 Life (VSL) of \$10.95 million dollars (year 2019 dollars, as suggested by EPA (63)), we find that the  
236 projected 12k increase in annual excess deaths due to climate-driven wildfire smoke would result in  
237 annual damages of \$244 billion in 2050 (not discounted, in year 2019 dollars, see Methods). Under  
238 a similar projected warming level of SSP3-7.0 scenario, Hsiang et al. (46) estimated annual damage  
239 of 0.4%-0.8% of US GDP or \$166-332 billion (in year 2019 dollars, using annual projected GDP of  
240 \$38.5 trillion from (53)), which included damages from temperature-related mortality, agriculture,  
241 crime, coastal storms, energy, and labor channels. The Framework for Evaluating Damages and  
242 Impacts (FrEDI), developed by US EPA (53), considered more sectors (including estimated wildfire  
243 damages from the western US (35)) and estimated annual damage of \$292 billion in 2050s. Our  
244 estimates suggest that damages from increase smoke-related mortality could roughly equal damages  
245 from all other estimated causes by mid-century in the US.

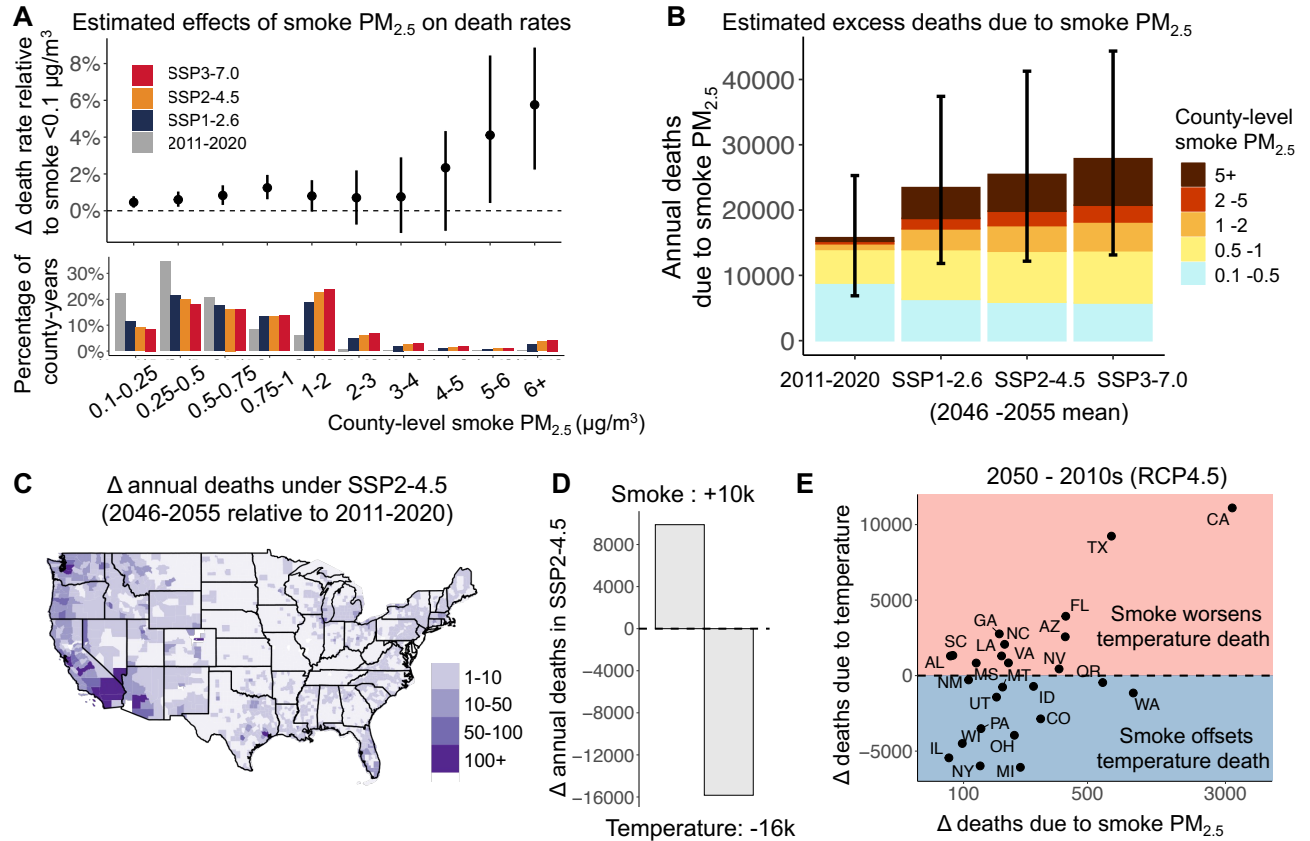


Figure 4: **Mortality impacts of wildfire smoke  $PM_{2.5}$  and estimated mortality due to smoke  $PM_{2.5}$  under future climate scenarios.** Panel A: empirically estimated effects of annual smoke  $PM_{2.5}$  concentration on county-level all-age annual mortality rates. The figure shows the effects of exposure to different annual mean concentration of smoke  $PM_{2.5}$  (shown in the x-axis) relative to a year with smoke concentration  $<0.1 \mu\text{g}/\text{m}^3$ , estimated using a Poisson model at the county and annual level and data from 2006-2019. The error bars show the 95% confidence interval estimated using bootstrap. The bottom part of panel A shows the percentage of county-years in each smoke concentration bin over the historical period (2011-2020) as well as future climate scenarios (2046-2055). Panel B: estimated annual excess deaths due to smoke  $PM_{2.5}$ , and contribution to total smoke excess deaths from different smoke concentration bins. The error bars show the 95% bootstrapped confidence intervals. Panel C: county-level projected increases in annual excess deaths due to smoke  $PM_{2.5}$  in 2050; increases are calculated as the differences between the average deaths under SSP2-4.5 scenario over 2046-2055 and the 2011-2020 average. Panel D shows US-wide total estimated annual smoke deaths and direct temperature-related deaths in 2050, with increasing smoke deaths offsetting 62% of the reduction in temperature deaths. Panel E: projected increase in rate deaths offsets projected reductions in direct temperature-related deaths by 2050s, the latter as estimated in a recent study (52). The x-axis shows the changes in deaths due to smoke  $PM_{2.5}$  in 2050s (note the log-scale), and the y-axis shows the changes in deaths due to temperature change, where only the 25 states with  $> 75$  smoke related deaths per year are visualized.

## 246 Discussion

247 While the effects of climate change on wildfire smoke and human health have become an emerging  
248 research topic, these effects are rarely incorporated into estimates of climate impacts. In this study,  
249 we estimate that climate-induced smoke  $\text{PM}_{2.5}$  could lead to 12k additional excess deaths per year  
250 under the SSP3-7.0 scenario in the US, substantially offsetting the reduction in direct temperature-  
251 related deaths expected due to climate change. These estimated deaths lead to an amount of  
252 monetized damage on par with quantified damages from all other sectors combined. Our results  
253 suggest that increasing wildfire smoke pollution due to climate change could be one of the most  
254 important and costly consequences of a warming climate in the US.

255 We find that aggressive mitigation of global greenhouse gas emissions would limit increases in  
256 smoke-related deaths, but that such deaths are likely to increase substantially even under low-  
257 emission scenarios. This finding points to the urgent need for future adaptation if damages are  
258 to be avoided. Adaptation could occur at many points along the wildfire-smoke-mortality chain.  
259 Increased fuel management, such as prescribed burning, could reduce the likelihood of extreme  
260 wildfire activity during adverse climate conditions, but will create smoke of its own; while the  
261 reduction in smoke from high-intensity fire is likely to substantially outweigh the increase from  
262 purposeful low-intensity fire, quantifying such tradeoffs is another critical area for work (64–66).  
263 Adaptation could also target the relationship between smoke and adverse health outcomes. This  
264 could include better informing individuals of, and protecting them from, smoke that does occur as  
265 current reliance on individuals to self-protect appears highly inadequate and inequitable (67, 68).  
266 Improved indoor filtration, including low-cost portable filters, appears a particularly promising and  
267 scalable solution, and ensuring that such filtration is affordable, accessible, and used is a potential  
268 policy priority (69).

269 Using georeferenced data on deaths and ambient wildfire smoke concentrations, we show that  
270 increasing annual exposures to smoke  $\text{PM}_{2.5}$  are associated with higher county-level annual mortality  
271 rates across the contiguous US. Our work contributes to a large literature documenting the impacts  
272 of annual exposures to total  $\text{PM}_{2.5}$  on mortality, which has shaped decades of policy to improve  
273 ambient air quality in the US. Due to our annual level projections of wildfire smoke, impacts of  
274 wildfire smoke on mortality were necessarily conducted at the annual level. However, wildfires are  
275 episodic and typically generate short-term spikes in ambient air pollution, which our measure of  
276 exposure may partly obscure (70). As such, our results are a complement to other studies on the  
277 health effects of short-term (e.g., daily) wildfire smoke exposures (12).

278 We find that elevated long-term average smoke  $\text{PM}_{2.5}$  concentrations increase mortality rates  
279 at both low and high concentrations. These increases lead to two important implications. First,  
280 we project large mortality burden not only in regions where large fires occur but also in populous  
281 regions with low smoke concentrations (e.g., the eastern US) that have historically received less

282 focus in wildfire studies. We find that 67% of the estimated historical smoke mortality and 42% of  
283 the projected future mortality come from the eastern US, as a result of increases in low-level smoke  
284 concentrations, consistent with previous historical estimates from (57). Second, despite larger  
285 differences in projected smoke  $PM_{2.5}$  concentration across the three climate scenarios, we estimate  
286 substantial mortality increases even in the low warming scenario (SSP1-2.6), again because this  
287 scenario generates low-level annual concentration increases that we estimate can have substantial  
288 mortality impacts. Our projected mortality impacts are in the uncertainty band of one prior study  
289 that applied a range of dose-response functions of total  $PM_{2.5}$  exposure (34), while substantially  
290 higher than the other estimate which only focuses on the western US (35), in part due to the  
291 mortality impacts we find at low exposure levels.

292 Our approach can isolate the “direct” impacts of climate change on wildfire air pollution, but does  
293 not account for potential “indirect” effects of climate on wildfire through channels such as climate’s  
294 influence on vegetation growth or lightning-related ignitions. Existing evidence has suggested that  
295 vegetation overall would increase under higher warming levels, which could lead to higher wildfire  
296 emissions and smoke (27). Furthermore, we did not attempt to model the many non-climate factors  
297 that contribute to wildfire activity, including the location of energy infrastructure, distance to road,  
298 housing development, and fire suppression efforts. Instead, we sought a model that could isolate the  
299 influence of climate while holding these other factors fixed. If these factors change dramatically in  
300 the future, then our estimates could understate or overstate future emissions, smoke, and mortality.  
301 For example, if expansions of houses near wildland vegetation continue (20), the effects of a warming  
302 climate on wildfire emissions could be larger given more human ignitions, particularly as population  
303 growth in the wildland-urban interface has been most rapid in areas where the vegetation is most  
304 vulnerable to wildfire (71). Alternatively, large increases in wildfire activity could be self-limiting  
305 as fires regulate the amount and availability of fuel load for future combustion. Existing studies  
306 suggest that this feedback is likely modest (26), but constraining this feedback empirically is a  
307 critical area for future work.

308 Our projection analysis quantifies the key uncertainties in climate-wildfire-smoke-mortality es-  
309 timations (Figure S11). Addressing these uncertainties could further improve understanding of the  
310 climate influences on wildfire pollution and health, and thus inform relevant policies. One of the  
311 largest uncertainties is how climate change will influence wildfire emissions and smoke  $PM_{2.5}$ . The  
312 statistical models we train can predict the emissions well given observational data, but we know  
313 little about their ability to predict wildfire levels under unprecedented climate conditions. Also,  
314 we could only robustly establish the climate-wildfire relationship when evaluated at aggregated  
315 spatial and temporal scales; predicting wildfire ignitions and growth at local scales remains very  
316 challenging. In the future, combining statistical models that can leverage the observational con-  
317 straints with process-based climate-vegetation-fire models could likely generate a useful framework  
318 for understanding climate impacts on wildfire pollution. Another critical uncertainty is the health

319 effects of smoke  $PM_{2.5}$  exposure. Quantifying health impacts of smoke  $PM_{2.5}$  at both low and high  
320 concentrations in the context of the unique chemical composition of smoke  $PM_{2.5}$  and fire influence  
321 on human behaviors remains an important area of future research. Furthermore, our estimated  
322 health cost is likely only a subset of the overall health burden due to possible morbidity effects of  
323 smoke, or health costs from other wildfire-driven pollutants.

324 Our projections of smoke  $PM_{2.5}$  and mortality effects can support climate science, health, and  
325 policy research to better understand drivers and consequences of smoke  $PM_{2.5}$  under climate change,  
326 and help inform policy priorities to address their negative impacts. Our estimates suggest that  
327 health costs due to climate-induced smoke  $PM_{2.5}$  could be among the most damaging consequences  
328 of climate change in the US. Based on our results, designing and implementing policies to reduce  
329 wildfire smoke and protect vulnerable communities has the potential to deliver substantial health  
330 benefits now and in the coming decades.

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## 341 **Author contributions**

342 MQ and MB designed the study. MQ led the smoke projection modeling with inputs from all  
343 authors. MQ, JL, RJ led the statistical and machine learning modeling of fire emissions. MQ,  
344 SHN, CFG, MLC led the health impacts analysis. MQ and MB led the writing of the manuscript  
345 with inputs from all authors. All authors contributed to interpretation of results and reviewed the  
346 manuscripts.

## 347 **Competing interests**

348 The authors declare no competing interests.

## 349 **Data and materials availability**

350 Data and code to replicate all results in the main text and supplementary materials will be made  
351 available at a public repository, except for county-level mortality data for low-population counties,  
352 which are not publicly available and were obtained through application to the National Center for  
353 Health Statistics.



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## 519 **Materials and Methods**

### 520 **Wildfire and smoke PM<sub>2.5</sub> datasets**

521 We use annual fire emissions from the fourth version of the Global Fire Emissions Database with  
522 small fires (GFED4s) from 2001-2021 (1). The native spatial resolution of GFED4s is 0.25×0.25  
523 degrees. We use the estimated dry matter (DM) emissions as our primary variable for the emissions.  
524 DM emissions capture the amount of biomass being consumed in the burning process. We choose  
525 DM emissions as the proxy for overall fire emissions (rather than individual emissions species such  
526 as black carbon or NO<sub>x</sub>) due to uncertainty in the emission factors used in GFED4s. GFED4s  
527 include fire emissions from agriculture fires and land-use change as well. However, as wildland fire  
528 emissions dominate in most study regions (especially in western US and Canada where we see the  
529 largest effects), we refer to our estimates as “wildfire emissions” and “wildfire smoke” for simplicity  
530 and consistency (Table S1).

531 For smoke PM<sub>2.5</sub>, we use gridded daily wildfire smoke PM<sub>2.5</sub> predictions for the contiguous US at  
532 10 km resolution from January 1, 2006 to December 31, 2020 derived from (2). This dataset specif-  
533 ically estimates the ambient PM<sub>2.5</sub> concentration due to wildfire smoke influence by constructing a  
534 machine learning model that uses smoke plume data, remotely-sensed variables, and meteorologi-  
535 cal variables to predict the anomalous increases in surface PM<sub>2.5</sub> measured by surface air quality  
536 monitors during wildfire. To estimate contributions of smoke PM<sub>2.5</sub> to total PM<sub>2.5</sub>, we use the  
537 total PM<sub>2.5</sub> estimates from (3), which combines satellite retrievals of aerosol optical depth, chemi-  
538 cal transport modeling, and ground-based measurements to estimate monthly total ambient PM<sub>2.5</sub>  
539 concentrations.

### 540 **Climate and meteorological datasets**

541 We use climate and land use variables to predict wildfire DM emissions. The climate variables  
542 include 2m air temperature, precipitation, relative humidity, soil moisture (of the top soil layer),  
543 vapor pressure deficit (VPD), wind speed (at 10m level), and runoff (sum of surface and subsur-  
544 face). We include these climate variables because they are available in both the historical data and  
545 the climate projections from CMIP6 climate model ensembles. Our models do not include other  
546 potentially important variables such as fire weather index and fuel moisture (as used in (4)) be-  
547 cause they are unavailable in future projections. These climate variables are derived from the North  
548 American Regional Reanalysis (NARR) (5), with the exception of soil moisture. Soil moisture is  
549 derived from the VIC land-surface model of phase 2 of the North American Land Data Assimilation  
550 System (NLDAS-2) (6) and only available in the contiguous US. The native spatial resolution is 32  
551 km for NARR variables and 0.125 degree for NLDAS-2 variables. Land use variables are derived  
552 from the North American Land Change Monitoring System (NALCMS) for the year 2015 (7). More

553 specifically, we use three land use variables which each represents the percentage of area in three  
554 categories: cropland, forest, and grassland. The native resolution of land use variables is 30m. Be-  
555 cause high-resolution projections of future land use change are not available, the land use variables  
556 are held constant across time in both the historical and future periods.

557 For future climate change scenarios, we use the projected climate variables from the Coupled  
558 Model Intercomparison Project Phase 6 (CMIP6). We examine three primary climate-forcing sce-  
559 narios featured by the IPCC, which are constructed as pairs between the Shared Socio-economic  
560 Pathways (SSPs) and the Representative Concentration Pathways (RCPs) (8). We use SSP1-2.6  
561 (which the IPCC refers to as the “Low” scenario), SSP2-4.5 (which the IPCC refers to as the “In-  
562 termediate” scenario), and SSP3-7.0 (which the IPCC refers to as the “High” scenario). We use  
563 projections from 28 global climate models that include the selected variables that cover the study  
564 region (Table S6). Following practice of IPCC, we select only one ensemble realization for each  
565 model – we use the first ensemble variant of each model (“r1i1p1f1”) when possible.

566 When modeling the relationship between wildfire emissions and smoke  $PM_{2.5}$ , we also include  
567 meteorological variables in the regression model. The daily gridded meteorological variables are  
568 derived from gridMET (9). In our main specification, we aggregate the meteorological variable  
569 to the monthly and smoke grid cell level. We include the splines of daily surface temperature,  
570 precipitation, dewpoint temperature, boundary layer height, air pressure, 10m wind direction (U  
571 and V components) and wind speed.

## 572 **Predicting wildfire emissions**

573 We construct an ensemble of statistical and machine learning models to predict wildfire emissions  
574 using climate and land use variables. Our models predict the annual dry matter (DM) emissions de-  
575 rived from GFED4s emission inventory using climate and land-use variables from 2001 to 2021. We  
576 build separate models for each of the five regions (western US, southeastern US, northeastern US,  
577 Canada-Alaska, and Mexico) to capture the regionally heterogeneous relationships between climate,  
578 land type and wildfire emissions. For each region, we construct six different models as potential  
579 model candidates: linear regression model, linear regression model with log outcomes, Least Abso-  
580 lute Shrinkage and Selection Operator (LASSO) models, LASSO models with log outcomes, 2-layer  
581 Neural Network (NN) model, and NN models with log outcomes. These six algorithms are selected  
582 to cover a possible range of model candidates with varying desired characteristics – including simple  
583 models that are commonly used in prior studies (e.g., the linear and log-linear regression models),  
584 models that are easy to interpret (e.g., the linear regression and LASSO models), and more flexible  
585 machine learning models that are used in prior studies (e.g., the NN model).

586 One key challenge for this prediction problem is that the fire occurrence, spread, and resulting  
587 emissions at local scales are often fairly stochastic due to varying and hard-to-predict non-climate

588 factors, including where and when human and natural ignitions occur and how much suppression  
589 effort is applied. Therefore, to better capture the predictable components of the climate-wildfire  
590 relationship, we create models to predict annual emissions aggregated at different spatial scales  
591 for each of the six model types mentioned above. We aggregate the outcome variables and model  
592 features at four spatial scales: the grid scale (0.25 deg, 26956 cells in total), the North America Level-  
593 3 Ecoregion scale (177 regions in total), the North America Level-2 Ecoregion scale (51 regions in  
594 total), and the regional scale (5 regions in total). We then select the spatial resolution that optimizes  
595 model performance for each model type (as described below), allowing the optimal spatial resolution  
596 to differ across different model types and regions (see Figure S3 for model performances across spatial  
597 scales).

598 To evaluate the model performance, we use nested leave-one-out cross-validations (LOOCV) at  
599 the temporal scale. We divide our data into 21 temporal folds, each including one year of data. For  
600 each holdout fold, we train the model using the remaining 20 folds of data with hyper-parameters  
601 selected using an inner-loop 5-fold CV within the training data. We then obtain out-of-sample  
602 predictions for the holdout fold and repeat this process to obtain out-of-sample predictions for the  
603 entire time period. As we focus on projecting the future wildfire emissions over a 10-year period  
604 (i.e. decadal averages) under future climate scenarios, we thus evaluate the performance of our  
605 models on similar 10-year intervals. We compute the moving averages of predicted and observed  
606 emissions over 10-year moving windows. We compute two metrics and use them as the basis for  
607 evaluating the performance of each model: 1) the root mean square error between predictions and  
608 observations, and 2) the prediction biases of the highest-emitting 10-year period. The first metric  
609 allows us to assess the model performance across years with different climate conditions to detect  
610 differences between current and future climate for different climate scenarios. The second metric  
611 allows us to assess the model performance under the extreme smoke conditions which are more likely  
612 to occur under future climate. To obtain the final model that can be used for future projections,  
613 we create an “ensemble model” which combines the predictions from the selected base models with  
614 the corresponding optimal spatial resolution. The selected models and their performances can be  
615 found in Table S2.

## 616 **Quantifying fire impacts on smoke $PM_{2.5}$**

617 To estimate smoke  $PM_{2.5}$  concentrations associated with future wildfire emissions, we design a sta-  
618 tistical approach to establish an empirical relationship between ambient smoke  $PM_{2.5}$  from (2) and  
619 wildfire emissions derived from GFED4s. We estimate the relationship between wildfire emissions  
620 and smoke  $PM_{2.5}$  concentration across the contiguous US (48 states and the District of Columbia)  
621 at 10 km resolution, accounting for variation in wind directions and atmospheric transport. This  
622 approach allows us to efficiently predict smoke concentration in one location from changes in wild-



623 fire emissions in another. Despite using estimated DM emissions from GFED4s as an input, our  
 624 estimates of smoke PM<sub>2.5</sub> concentrations strongly predict the variations in the empirical estimates  
 625 of surface smoke PM<sub>2.5</sub> concentrations, and are thus directly constrained by surface PM<sub>2.5</sub> measure-  
 626 ment during wildfire episodes.

Specifically, we use the following regression equation to empirically quantify the impacts of the  
 wildfire DM emissions on smoke PM<sub>2.5</sub> in the US in our historical data:

$$Smoke_{iym} = \sum_{d,w} \beta_{dw} \Delta Emis_{dw,iym} + \gamma \mathbf{W}_{iym} + \eta y + \psi_m + \theta_i + \epsilon_{iym} \quad (1)$$

627 where  $Smoke_{iym}$  denotes the smoke PM<sub>2.5</sub> at grid cell  $i$  (resolution: 10 km), year  $y$  and month-of-  
 628 year  $m$ .  $Emis_{dw,iym}$  denotes the wildfire DM emissions that in the distance bin  $d$  and wind direction  
 629  $w$  ( $w \in \{upwind, other, downwind\}$ ) of the smoke location  $i$  on month-of-year  $m$  and year  $y$ . In our  
 630 main specification, we estimate the impacts of wildfire DM emissions at different distances from the  
 631 smoke location:  $<50$  km,  $50-100$  km,  $100-200$  km,  $200-350$  km,  $350-500$  km,  $500-750$  km,  $750-1000$   
 632 km,  $1000-1500$  km,  $1500-2000$  km,  $>2000$  km.  $\mathbf{W}_{iym}$  are the meteorological variables at the grid  
 633 cell  $i$  (as described in the dataset section). We include these meteorological variables to capture  
 634 potential meteorological variability that could influence ambient PM<sub>2.5</sub> concentrations. Our main  
 635 specification includes linear year trend ( $\eta y$ ) and month-of-year fixed effects ( $\psi_m$ ) to capture the  
 636 long-term trend and seasonality of smoke PM<sub>2.5</sub> concentration, and grid cell-level fixed effects ( $\theta_i$ )  
 637 to control for the time-invariant unobserved factors at the grid cell location.  $\epsilon_{iym}$  represents the  
 638 error term.

639 To better capture the atmospheric transport of smoke PM<sub>2.5</sub>, we divide the wildfire emissions  
 640 (from a given distance bin) into three categories depending on wind direction and the location of fire.  
 641 Following methods in (10), wildfire emissions are classified into “upwind” or “downwind”, depending  
 642 on whether the wildfire location is at the upwind or downwind direction of the smoke grid cell. We  
 643 combine daily emissions with daily wind direction (10m wind) to calculate the daily emission from  
 644 each wind direction and further aggregate to the monthly level.

645 Many previous studies have demonstrated that wildfire emission factors (e.g., mass of organic  
 646 carbon particles emitted from burning one kg fuel) strongly depend on the combustion conditions  
 647 (e.g., the combustion completeness) and the underlying fuel type among many other factors (11–14).  
 648 As many of these characteristics (e.g., the combustion efficiency of different fires) are not available  
 649 at the national scale, we use a data-driven approach and estimate different models/equations for  
 650 the nine US climate regions determined by National Centers for Environmental Information (see  
 651 Figure 2 for region definitions), which allows the relationship between emissions and surface smoke  
 652 PM<sub>2.5</sub> to differ by region. The resulting regional estimates therefore implicitly account for some  
 653 heterogeneity in the vegetation fuel types, fire intensities (as characterized in historical fires), and  
 654 topographies for different locations. For example, prior studies have shown that smoldering fires

655 often have higher  $\text{PM}_{2.5}$  emission factors compared to flaming fires due to incomplete combustion  
656 (13), which might partly explain the relatively high emissions factors in the Southeast as smoldering  
657 fires are more common there due to high humidity (15).

## 658 **Projecting wildfire emissions and smoke $\text{PM}_{2.5}$ under future climate**

659 We combine our ensemble of statistical and machine learning models with climate projections from  
660 ensembles of global climate models to project the wildfire emissions and smoke  $\text{PM}_{2.5}$  under future  
661 climate scenarios. Consistent with the optimal spatial resolutions selected for each region, we  
662 predict the annual wildfire DM emissions at different spatial resolutions, from 2001-2055. We then  
663 statistically downscale the predicted regional emissions to the native grid cell level (0.25 degree) by  
664 distributing predicted DM emissions using average historical spatial distribution of emissions at the  
665 grid cell level (2001-2021).

666 We combine the downscaled predicted DM emissions at GFED4s grid cell level (0.25 degree)  
667 with the empirical relationship we established between smoke  $\text{PM}_{2.5}$  and GFED4s DM emissions to  
668 calculate predicted smoke  $\text{PM}_{2.5}$  in each smoke grid cell (resolution of 10 km). When calculating the  
669 smoke  $\text{PM}_{2.5}$  in future scenarios, the wind direction and meteorological conditions are held constant  
670 at the average conditions in the historical period. We further calculate the difference between the  
671 estimated smoke in any future year and the average *estimated* smoke between 2011-2020. The delta  
672 difference is then added to the average *observed* smoke  $\text{PM}_{2.5}$  concentration between 2011-2020 to  
673 obtain the final smoke predictions for each grid cell in the future years.

## 674 **Impacts of smoke $\text{PM}_{2.5}$ on mortality**

675 We calculate all-cause mortality associated with wildfire smoke exposure historically and under  
676 future climate scenarios using a dose-response function empirically derived from 2006-2019 county-  
677 level data. We combine county-level population-weighted annual smoke  $\text{PM}_{2.5}$ , derived from (2),  
678 with county-level all-cause mortality rates by different age groups. We obtain individual-level mul-  
679 tiple cause of death mortality data from the National Center for Health Statistics to calculate  
680 age-standardized mortality rates for all ages, those under 65 years of age, and those 65 years and  
681 older (16). County-level mortality rates were age-standardized using the direct method and 5-year  
682 bins (0-4, 5-9, ..., 85 and over) based on the 2000 US Census Standard Population. Monthly mortal-  
683 ity rates were standardized per 100,000 population. To fully capture damages from ambient wildfire  
684 smoke concentrations, our preferred outcome is age-standardized, all-cause, all-age mortality rates  
685 at the county-year level. We also separately estimate impacts among those 65 years and older and  
686 those under 65 years of age (Figure S6).

In our main analysis, we estimate a Poisson model in which we allow non-linear impacts of

annual smoke PM<sub>2.5</sub> on mortality rates at the county-year level:

$$D_{csy} = \exp \left( \sum_i \beta_i \text{smokeBIN}_{csy}^i + \gamma W_{csy} + \eta_{sy} + \theta_c + \varepsilon_{csy} \right) \quad (2)$$

687 where  $D_{csy}$  denotes the age-adjusted all-cause mortality rates in county  $c$ , state  $s$ , and year  $y$ .  
688  $\text{smokeBIN}_{csy}^i$  is a dummy variable for whether annual population-weighted smoke PM<sub>2.5</sub> in county  
689  $c$ , state  $s$ , and year  $y$  falls into the range of bin  $i$  (0-0.1, 0.1-0.25, 0.25-0.5, 0.5-0.75, 0.75-1, 1-2, 2-3,  
690 3-4, 4-5, 5-6, >6  $\mu\text{g}/\text{m}^3$ ; 0-0.1 is the reference category). The main coefficients of interest are the  
691  $\beta_i$ 's, which estimate the effects of a year with annual smoke concentration of bin  $i$  on mortality rates,  
692 relative to a year with annual mean smoke PM<sub>2.5</sub> concentration below 0.1  $\mu\text{g}/\text{m}^3$ . The reference  
693 category included <0.1 because only 4 county-year observations had exactly zero ambient wildfire  
694 smoke.  $W_{csy}$  denotes a flexible control of temperature (the number of days that fall in different  
695 temperature bins) and linear and quadratic terms of annual population-weighted precipitation.  $\eta_{sy}$   
696 denotes a vector of state-year fixed effects (i.e. separate intercepts for each year in each state)  
697 that accounts for all factors that differ across states in a given year (e.g. California 2018 versus  
698 Oregon 2018) as well as all factors that differ within states across years (e.g. California 2017 versus  
699 California 2018).  $\theta_c$  denotes a set of county-level fixed effects that accounts for any county-specific  
700 time-invariant factors that could be correlated with both smoke exposure and mortality (e.g., high  
701 income communities in the mountainous areas on the west coast could have higher smoke exposure  
702 but lower mortality rates due to non-smoke reasons). In essence, we identify the effect of wildfire  
703 smoke on mortality using within-county variation over time, after accounting for any factors that  
704 trend over time within that county's state, and for any correlation between smoke variation and  
705 variation in temperature and precipitation. Because temporal variation in wildfire smoke exposure  
706 is largely a function of idiosyncratic factors such as where a given fire starts and which way the wind  
707 blows, our estimates have a plausibly causal interpretation. The coefficients are estimated using  
708 weighted Poisson regression models, with function "fepois" from R package "fixest". The estimations  
709 are weighted by county-level population counts to enable estimates of population-averaged effects,  
710 as well as to reduce statistical uncertainty. The uncertainty of the coefficients are estimated using  
711 bootstrap of 500 runs.  $\varepsilon_{csy}$  represents the error terms.

712 While we observe historical data on daily smoke PM<sub>2.5</sub> concentrations and monthly cause-specific  
713 mortality rates, we estimate the dose-response functions at the annual level to be consistent with  
714 our smoke concentration projections, which are only feasible at the annual level. This approach  
715 deviates from previous studies estimating health impacts from wildfire smoke which focus primarily  
716 on sub-annual exposures, but it allows for a direct application of the estimated response functions  
717 to annual smoke projections. It also has the advantage of allowing us to capture the net effect  
718 of either behavioral dynamics in response to short-term variation, as has been observed in related  
719 settings (17), or "displacement" of mortality that would otherwise occurred but was hastened as

720 a result of short-term exposure – a common concern in climate impact studies (18).

721 To evaluate the influence of functional forms of the dose-response function, we estimate alter-  
722 native response functions using a Poisson model, a least-squares linear regression, and a quadratic  
723 model where wildfire smoke concentrations were treated as a continuous exposure, and calculate  
724 how different functional forms influence the estimates of projected annual excess deaths (Figure  
725 S9). We find that non-binned models generally fail to capture meaningful impacts of both low-level  
726 and high-level smoke exposure (Figure S12).

727 Further, to assess the sensitivity of our results to multiple assumptions, we estimate several  
728 alternative specifications of the Poisson model. Specifically, we estimate a model which uses al-  
729 ternative bin definitions, a model which includes year 2020, a model which calculates the number  
730 of months or the number of days in a year that fall in different smoke bins to represent different  
731 temporal aggregations, and a model which is estimated at county-month level. While we cannot  
732 calculate the impact on projected mortality under scenarios using these sub-annual measures of  
733 wildfire smoke  $PM_{2.5}$  given the resolution of the wildfire smoke projections, we instead compare  
734 between estimated historical excess deaths during 2011-2020, calculated as the difference between  
735 predicted deaths at observed smoke levels relative to what would have occurred absent any smoke.  
736 We find that the largest differences occur when using monthly bins, likely due to the lagged effects  
737 of smoke on mortality at the monthly level (Figure S10).

738 To calculate smoke attributable deaths in the historical scenario, we use the county-level pop-  
739 ulation data for the year 2019. We use the county-level average death rate between 2006 to 2019  
740 as the baseline mortality rate for calculations with the Poisson model. For projections of future  
741 mortality burden, we scale the population according to the future population projections from the  
742 US Census (19).

### 743 **Monetizing health impacts**

744 The mortality impacts are monetized using a value of statistical life (VSL) of \$10.95 million (year  
745 2019 dollars), as recommended by the US EPA (20) and used in previous studies (21). For future  
746 scenarios, we adjust VSL values using the projected economic growth of 2% and income elasticity  
747 of one, following a similar method from Carleton et al. (21). We compare the monetized health  
748 impacts from climate-induced smoke with two prior estimates of aggregate monetized/economic  
749 damage due to climate change. Hsiang et al. estimated an annual damage of 0.4%-0.8% of US GDP  
750 or \$166-332 billion (in year 2019 dollars, using annual projected GDP of \$38.5 trillion from (22)).  
751 Their approach empirically calculated the effects of climate change on a variety of economic damages  
752 from temperature-related mortality, agriculture, crime, coastal storms, energy, and labor channels  
753 (23). The Framework for Evaluating Damages and Impacts (FrEDI), developed by US EPA (22),  
754 estimated an annual damage of \$292 billion in the 2050s. FrEDI considered 21 sectors (including

755 estimated wildfire damages from western US (24)). The wildfire health damages considered in  
756 FrEDI only accounted for effects of wildfire in the western US and used an empirical climate-fire  
757 relationship derived from historical data before 2013 which did not include recent extreme wildfire  
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821 **Supplementary tables**

Table S1: Estimated dry matter (DM) emissions by land-use type in historical period and future scenarios. For the historical period, the table shows the annual mean DM emissions from each land-use type in each region from 2001-2021, directly derived from GFED4s. For the future scenario, the table shows the annual mean DM emissions from each land-use type in each region under SSP3-7.0 from 2046-2055. Landuse types are derived from GFED4s inventory. “Forest” includes emissions from both temperate forests and boreal forests.

Region	Type	2001 - 2021		2050 SSP3-7.0 emissions (MT)
		emissions (MT)	percent	
Western US	forest	25.8	68%	184.7
	savanna	10.7	28%	76.5
	agriculture	1.7	4%	12.3
	landuse change	0.0	0%	0.0
	peatland	0.0	0%	0.0
Southeastern US	forest	4.2	28%	4.4
	savanna	5.3	35%	5.8
	agriculture	5.6	37%	5.6
	landuse change	0.0	0%	0.0
	peatland	0.0	0%	0.0
Northeastern US	forest	0.6	29%	0.6
	savanna	0.2	11%	0.2
	agriculture	1.2	60%	1.1
	landuse change	0.0	0%	0.0
	peatland	0.0	0%	0.0
Canada-Alaska	forest	152.6	94%	240.8
	savanna	0.2	0%	0.4
	agriculture	1.7	1%	2.7
	landuse change	0.0	0%	0.0
	peatland	8.2	5%	13.0
Mexico	forest	1.2	3%	1.7
	savanna	19.4	47%	29.0
	agriculture	6.4	16%	9.5
	landuse change	14.1	34%	17.1
	peatland	0.0	0%	0.0

Table S2: Performance of the individual statistical and machine learning models. For each region, we train six algorithms  $\{\text{Linear, LASSO, Neural Net}\} \times \{\text{level, log of the outcome}\}$ . The table shows the optimal spatial resolution and three evaluation metrics for each algorithm. The three evaluation metrics are correlation coefficient (R), bias in predicting the highest-emitting 10-year (Bias), and Root Mean Square Error over the mean of the outcome (RMSE/Mean). Bias is calculated as  $(\text{Prediction} - \text{Observation}) / \text{Observation}$  for the 10-year period with the highest emissions. Models selected in the final model ensembles are bolded and labeled “Y” in the “Selected” column. The selection is based on  $\text{RMSE} + |\text{Bias}|$  to consider both metrics. In our main analysis, for each region, only the algorithms with “ $\text{RMSE} + |\text{Bias}|$ ” within 5% of the best algorithm are selected.

Region	Algorithm	Optimal resolution	R	Bias	RMSE/ Mean	RMSE +  Bias	Diff	Selected
<b>Western US</b>	<b>Linear, level</b>	<b>regional</b>	<b>0.98</b>	<b>-10%</b>	<b>20%</b>	<b>29%</b>	<b>0%</b>	<b>Y</b>
Western US	Linear, log	eco2	0.91	-16%	22%	37%	8%	N
Western US	LASSO, level	regional	0.99	-14%	28%	43%	13%	N
<b>Western US</b>	<b>LASSO, log</b>	<b>regional</b>	<b>0.89</b>	<b>-3%</b>	<b>31%</b>	<b>33%</b>	<b>4%</b>	<b>Y</b>
Western US	Neural Net, level	eco2	0.73	0%	90%	91%	61%	N
Western US	Neural Net, log	eco3	0.98	-20%	19%	39%	9%	N
<b>Southeastern US</b>	<b>Linear, level</b>	<b>eco3</b>	<b>0.51</b>	<b>-1%</b>	<b>6%</b>	<b>7%</b>	<b>0%</b>	<b>Y</b>
Southeastern US	Linear, log	eco2	0.36	-18%	14%	32%	25%	N
Southeastern US	LASSO, level	eco2	0.58	-12%	9%	21%	14%	N
Southeastern US	LASSO, log	eco2	0.02	-16%	14%	30%	23%	N
Southeastern US	Neural Net, level	grid	0.11	5%	11%	16%	9%	N
Southeastern US	Neural Net, log	eco2	0.30	-12%	12%	24%	17%	N
<b>Northeastern US</b>	<b>Linear, level</b>	<b>grid</b>	<b>0.05</b>	<b>-2%</b>	<b>11%</b>	<b>13%</b>	<b>0%</b>	<b>Y</b>
Northeastern US	Linear, log	regional	0.06	-11%	9%	20%	7%	N
Northeastern US	LASSO, log	eco2	0.19	-26%	19%	45%	32%	N
<b>Northeastern US</b>	<b>Neural Net, level</b>	<b>eco2</b>	<b>0.07</b>	<b>2%</b>	<b>14%</b>	<b>15%</b>	<b>2%</b>	<b>Y</b>
Northeastern US	Neural Net, log	eco3	0.29	-20%	12%	32%	20%	N
<b>Canada-Alaska</b>	<b>Linear, level</b>	<b>regional</b>	<b>0.91</b>	<b>4%</b>	<b>15%</b>	<b>19%</b>	<b>0%</b>	<b>Y</b>
Canada-Alaska	Linear, log	eco2	0.70	43%	35%	78%	59%	N
Canada-Alaska	LASSO, level	eco2	0.94	-15%	19%	34%	15%	N
Canada-Alaska	LASSO, log	regional	0.73	-13%	16%	29%	10%	N
Canada-Alaska	Neural Net, level	eco3	0.20	-9%	27%	36%	17%	N
Canada-Alaska	Neural Net, log	regional	0.71	-30%	15%	45%	26%	N
<b>Mexico</b>	<b>Linear, level</b>	<b>eco2</b>	<b>0.88</b>	<b>0%</b>	<b>4%</b>	<b>4%</b>	<b>0%</b>	<b>Y</b>
Mexico	Linear, log	eco2	0.85	-2%	14%	16%	12%	N
<b>Mexico</b>	<b>LASSO, level</b>	<b>eco3</b>	<b>0.86</b>	<b>0%</b>	<b>5%</b>	<b>5%</b>	<b>1%</b>	<b>Y</b>
Mexico	LASSO, log	eco2	0.71	-13%	14%	27%	23%	N
Mexico	Neural Net, level	eco2	0.72	0%	10%	10%	6%	N
Mexico	Neural Net, log	regional	0.82	-7%	7%	14%	9%	N



Table S3: Estimated coefficients from the selected linear regression models that use climate features to predict wildfire emissions. The table only shows the coefficients from the final selected models in each region with the corresponding optimal spatial resolution. Statistically significant coefficients ( $p < 0.1$ ) are bolded.

	Western US		Southeastern US		Northeastern US		Canada-Alaska		Mexico	
	coef	p-value	coef	p-value	coef	p-value	coef	p-value	coef	p-value
temperature	2.0E-02	0.25	-9.9E-04	0.15	2.1E-04	0.30	-3.0E-02	0.08	<b>-1.2E-02</b>	<b>0.00</b>
precipitation	-3.3E-02	0.27	<b>-4.9E-03</b>	<b>0.00</b>	-2.5E-04	0.63	-3.6E-02	0.43	<b>8.8E-03</b>	<b>0.00</b>
RH	5.6E-03	0.26	2.4E-04	0.55	-1.2E-04	0.36	2.3E-02	0.23	<b>4.8E-03</b>	<b>0.00</b>
wind speed	7.3E-02	0.12	<b>9.6E-03</b>	<b>0.00</b>	<b>-1.6E-03</b>	<b>0.09</b>	4.6E-02	0.79	-1.1E-03	0.92
VPD	-2.3E-02	0.94	3.3E-03	0.74	-2.5E-03	0.55	<b>2.4E+00</b>	<b>0.02</b>	<b>2.6E-01</b>	<b>0.00</b>
runoff	1.2E-02	0.24	5.5E-04	0.83	5.1E-04	0.49	1.2E-01	0.13	-9.1E-03	0.26
soil moisture	<b>-2.6E-02</b>	<b>0.01</b>	<b>-3.8E-03</b>	<b>0.00</b>	-2.8E-04	0.51				

Table S4: Estimated coefficients from the selected LASSO models that use climate features to predict wildfire emissions. As LASSO models are only selected in the western US and Mexico, the table shows the coefficients from these two final selected models with the corresponding optimal spatial resolution.

Western US		Mexico	
Selected variables	coef	Selected variables	coef
soil moisture	-1.2E+00	VPD*grass	4.1E-01
temperature	1.1E+00	VPD*precipitation	9.0E-03
VPD*precipitation	-2.8E+00	VPD*RH	1.3E-03
VPD*runoff	2.5E+00		
RH <sup>2</sup>	-1.5E-04		
runoff <sup>2</sup>	-1.6E-01		
runoff*wind speed	1.3E-02		
temperature <sup>2</sup>	4.6E-05		
wind speed <sup>2</sup>	4.1E-01		

Table S5: Estimated population-weighted average smoke PM<sub>2.5</sub>, total PM<sub>2.5</sub>, and smoke PM<sub>2.5</sub> contribution at the state level. Total PM<sub>2.5</sub> are calculated as the sum of smoke and non-smoke PM<sub>2.5</sub> concentrations. Non-smoke PM<sub>2.5</sub> are assumed to be the same as the average non-smoke PM<sub>2.5</sub> between 2016-2020, calculated as the difference between total PM<sub>2.5</sub> from (3) and smoke PM<sub>2.5</sub> from (2). Only states with >10% smoke contributions under SSP3-7.0 scenario are listed.

State	Smoke PM <sub>2.5</sub> μg/m <sup>3</sup>	Total PM <sub>2.5</sub> μg/m <sup>3</sup>	Smoke percent	State	Smoke PM <sub>2.5</sub> μg/m <sup>3</sup>	Total PM <sub>2.5</sub> μg/m <sup>3</sup>	Smoke percent	Scenario
Oregon	1.3	6.6	20%	Kansas	0.7	7.3	9%	2011-2020
	5.0	9.7	51%		1.5	7.7	19%	SSP1-2.6
	6.2	11.0	57%		1.7	7.9	21%	SSP2-4.5
	7.5	12.3	61%		1.8	8.1	22%	SSP3-7.0
Montana	1.3	6.4	20%	Nebraska	0.7	7.4	9%	2011-2020
	4.7	9.7	48%		1.2	7.4	16%	SSP1-2.6
	5.7	10.7	53%		1.3	7.5	18%	SSP2-4.5
	6.9	11.9	58%		1.5	7.6	19%	SSP3-7.0
Washington	0.9	6.0	16%	Oklahoma	0.6	7.8	8%	2011-2020
	3.8	8.3	46%		1.2	8.0	15%	SSP1-2.6
	4.6	9.0	51%		1.4	8.2	17%	SSP2-4.5
	5.6	10.0	56%		1.5	8.3	19%	SSP3-7.0
Idaho	1.3	7.1	18%	Minnesota	0.6	6.6	9%	2011-2020
	4.4	10.0	44%		0.9	6.5	14%	SSP1-2.6
	5.4	11.0	49%		1.0	6.7	15%	SSP2-4.5
	6.4	12.0	54%		1.1	6.8	16%	SSP3-7.0
Wyoming	0.7	5.4	13%	Arkansas	0.6	8.2	7%	2011-2020
	2.6	7.1	37%		1.1	8.0	14%	SSP1-2.6
	3.2	7.7	42%		1.2	8.2	15%	SSP2-4.5
	3.9	8.4	47%		1.3	8.2	16%	SSP3-7.0
Nevada	0.5	7.0	7%	Texas	0.5	8.3	6%	2011-2020
	3.1	9.6	32%		1.1	8.5	12%	SSP1-2.6
	3.9	10.4	38%		1.2	8.6	14%	SSP2-4.5
	4.6	11.1	42%		1.3	8.7	15%	SSP3-7.0
North Dakota	0.7	5.3	14%	Arizona	0.2	8.2	2%	2011-2020
	1.7	6.1	28%		0.9	8.7	11%	SSP1-2.6
	2.0	6.3	31%		1.1	8.9	13%	SSP2-4.5
	2.2	6.5	33%		1.3	9.1	14%	SSP3-7.0
California	0.6	10.5	6%	Iowa	0.6	7.8	8%	2011-2020
	2.8	12.2	23%		0.8	7.5	11%	SSP1-2.6
	3.5	12.9	27%		0.9	7.6	12%	SSP2-4.5
	4.1	13.5	30%		1.0	7.7	13%	SSP3-7.0
Colorado	0.5	6.1	9%	Wisconsin	0.5	7.4	7%	2011-2020
	1.7	7.2	24%		0.8	7.1	11%	SSP1-2.6
	2.0	7.5	27%		0.9	7.3	12%	SSP2-4.5
	2.3	7.8	30%		1.0	7.3	13%	SSP3-7.0
Utah	0.5	6.9	7%	Louisiana	0.4	8.5	5%	2011-2020
	1.7	7.7	22%		0.8	8.5	10%	SSP1-2.6
	2.0	8.0	26%		1.0	8.6	11%	SSP2-4.5
	2.4	8.4	29%		1.0	8.6	12%	SSP3-7.0
South Dakota	0.7	6.1	11%	Mississippi	0.4	8.3	5%	2011-2020
	1.4	6.3	22%		0.7	8.0	9%	SSP1-2.6
	1.5	6.5	24%		0.9	8.1	11%	SSP2-4.5
	1.7	6.7	26%		0.9	8.2	11%	SSP3-7.0
New Mexico	0.3	5.4	6%	Michigan	0.4	8.0	5%	2011-2020
	1.2	6.0	20%		0.6	7.6	8%	SSP1-2.6
	1.4	6.2	23%		0.7	7.7	9%	SSP2-4.5
	1.6	6.4	25%		0.8	7.7	10%	SSP3-7.0

Table S6: Climate models used in this study for future projections. We use projections from 28 global climate models with available output under the historical and three climate scenarios from the CMIP6 model ensembles. The spatial resolution of each model is shown in latitude  $\times$  longitude (unit: degree). Resolutions are approximated for models with varying latitudes. Data is downloaded in February, 2023.

<b>Model</b>	<b>Ensemble variant</b>	<b>Resolution</b>
ACCESS-CM2	r1ilp1f1	1.25 x 1.88
ACCESS-ESM1-5	r1ilp1f1	1.25 x 1.88
BCC-CSM2-MR	r1ilp1f1	1.12 x 1.12
CanESM5	r1ilp1f1	2.79 x 2.81
CAS-ESM2-0	r1ilp1f1	1.42 x 1.41
CESM2-WACCM	r1ilp1f1	0.94 x 1.25
CMCC-CM2-SR5	r1ilp1f1	0.94 x 1.25
CMCC-ESM2	r1ilp1f1	0.94 x 1.25
CNRM-CM6-1	r1ilp1f2	1.4 x 1.41
CNRM-CM6-1-HR	r1ilp1f2	0.5 x 0.5
CNRM-ESM2-1	r1ilp1f2	1.4 x 1.41
EC-Earth3	r1ilp1f1	0.7 x 0.7
EC-Earth3-Veg	r1ilp1f1	0.7 x 0.7
EC-Earth3-Veg-LR	r1ilp1f1	1.12 x 1.12
FGOALS-f3-L	r1ilp1f1	0.94 x 1.25
FGOALS-g3	r1ilp1f1	2.03 x 2
GFDL-ESM4	r1ilp1f1	1 x 1.25
GISS-E2-1-G	r1ilp1f2	2 x 2.5
GISS-E2-1-H	r1ilp1f2	2 x 2.5
IPSL-CM6A-LR	r1ilp1f1	1.27 x 2.5
KACE-1-0-G	r1ilp1f1	1.25 x 1.88
MIROC-ES2L	r1ilp1f2	2.79 x 2.81
MIROC6	r1ilp1f1	1.4 x 1.41
MRI-ESM2-0	r1ilp1f1	1.12 x 1.12
NorESM2-LM	r1ilp1f1	1.89 x 2.5
NorESM2-MM	r1ilp1f1	0.94 x 1.25
TaiESM1	r1ilp1f1	0.94 x 1.25
UKESM1-0-LL	r1ilp1f2	1.25 x 1.88

Table S7: Estimated annual excess deaths due to wildfire smoke at the state level. For historical period, the table shows average annual excess deaths due to smoke PM<sub>2.5</sub> exposure during 2011-2020. For future climate scenarios, the table shows average annual excess deaths due to smoke PM<sub>2.5</sub> exposure during 2046-2055 (median across 28 GCMs).

State	Historical	SSP1-2.6	SSP2-4.5	SSP3-7.0	State	Historical	SSP1-2.6	SSP2-4.5	SSP3-7.0
California	1381	4164	4657	5700	South Carolina	266	327	353	380
Texas	1276	1974	1958	1999	Tennessee	360	283	358	373
Washington	360	1108	1266	1530	Massachusetts	283	257	330	359
Florida	821	1119	1198	1295	Montana	87	219	253	318
Oregon	411	858	1020	1245	Mississippi	184	295	302	306
New York	800	749	924	979	Arkansas	244	323	305	302
Michigan	610	807	819	825	Kentucky	256	238	285	291
Ohio	651	701	845	821	Iowa	243	277	283	286
Pennsylvania	633	617	759	820	Kansas	224	265	260	269
Illinois	779	746	862	817	Utah	92	186	245	259
North Carolina	442	575	612	667	Maryland	236	168	213	232
Georgia	447	567	606	643	New Mexico	82	175	188	212
Arizona	184	511	558	574	Connecticut	162	150	193	201
Nevada	117	421	463	560	Nebraska	147	168	167	180
Colorado	225	398	497	540	West Virginia	94	69	101	109
Virginia	288	431	467	497	Wyoming	31	72	81	99
Wisconsin	366	461	464	471	South Dakota	63	81	85	91
Missouri	476	406	453	462	Maine	62	59	79	87
Indiana	392	394	464	452	New Hampshire	54	59	74	79
Louisiana	274	440	438	441	North Dakota	53	63	67	77
New Jersey	394	328	406	437	Rhode Island	49	43	55	61
Idaho	100	296	348	431	Delaware	44	24	33	37
Minnesota	340	399	405	418	Vermont	26	25	32	35
Alabama	306	358	391	409	D.C.	29	25	32	33
Oklahoma	320	415	394	404					

822 **Supplementary figures**

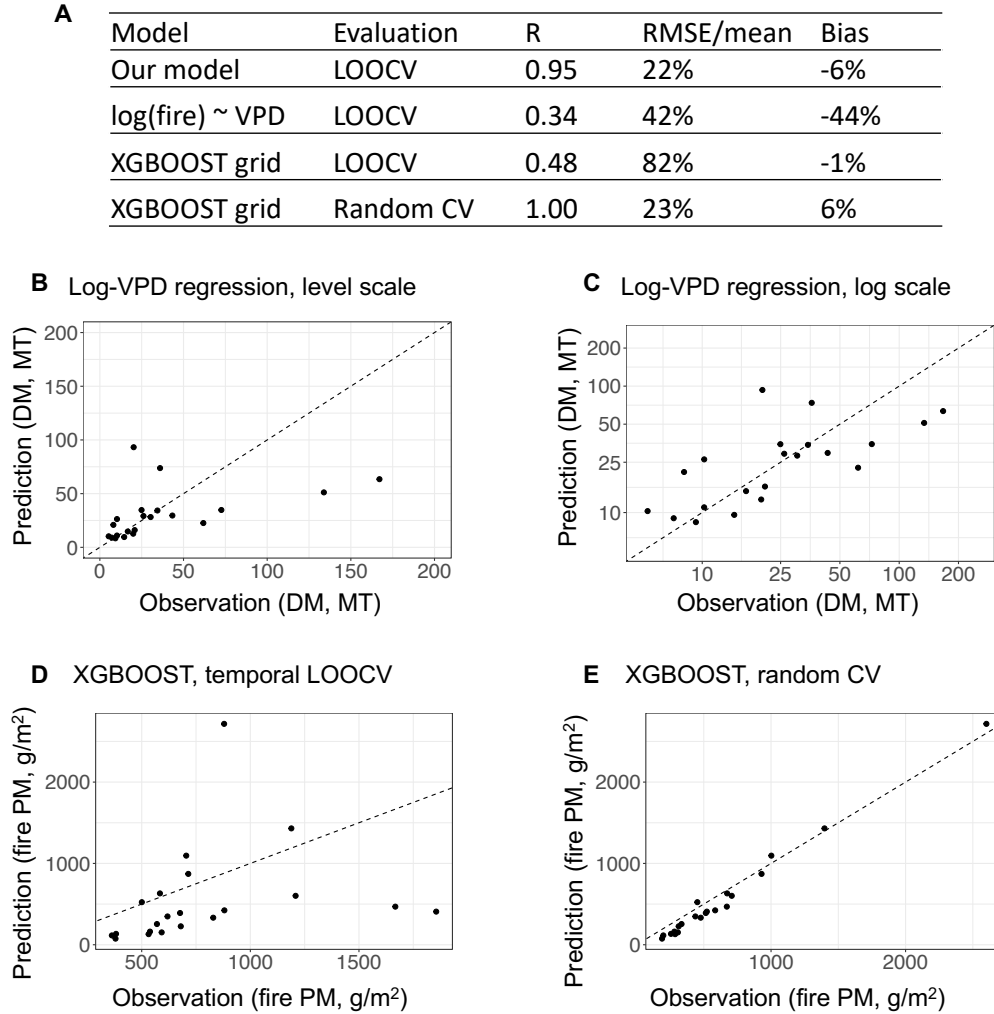


Figure S1: Predictive performance of our model and two other approaches used in previous research to predict wildfire emissions using climate variables. For comparison purposes, this figure only shows the results in western US. Panel A compares the predictive performance between our ensemble statistical and machine learning model (“Our model”), a regression method that uses fire-season VPD to predict the logged fire emissions (“ $\log(\text{fire})\text{-VPD}$ ”) as used in (25), and a XGBOOST model that predicts the fire emissions at the grid cell level as used in (26). The table shows the correlation coefficient (R), RMSE/mean, and bias of the highest-emitting 10-year period. Panels B and C show the out-of-sample prediction from the  $\log(\text{fire})\text{-VPD}$  regressions, with the same underlying data shown in level scale (B) and log scale (C). This demonstrates that while  $\log(\text{fire})\text{-VPD}$  regression achieves reasonable performance in the log scale (as reported by previous papers), its performance is inferior to our models in predicting the absolute levels of fire emissions. Panels D and E show the out-of-sample predictions from XGBOOST model under temporal LOOCV (D) and random CV (E) using the underlying dataset from (26). Random CV randomly partitions data to training and test sets with the same grid cell from different years possibly existing in both training and test sets. Panels D and E suggest the XGBOOST model trained at the grid cell has an inflated performance under random CV which grid cells can contribute data to both training and test sets.

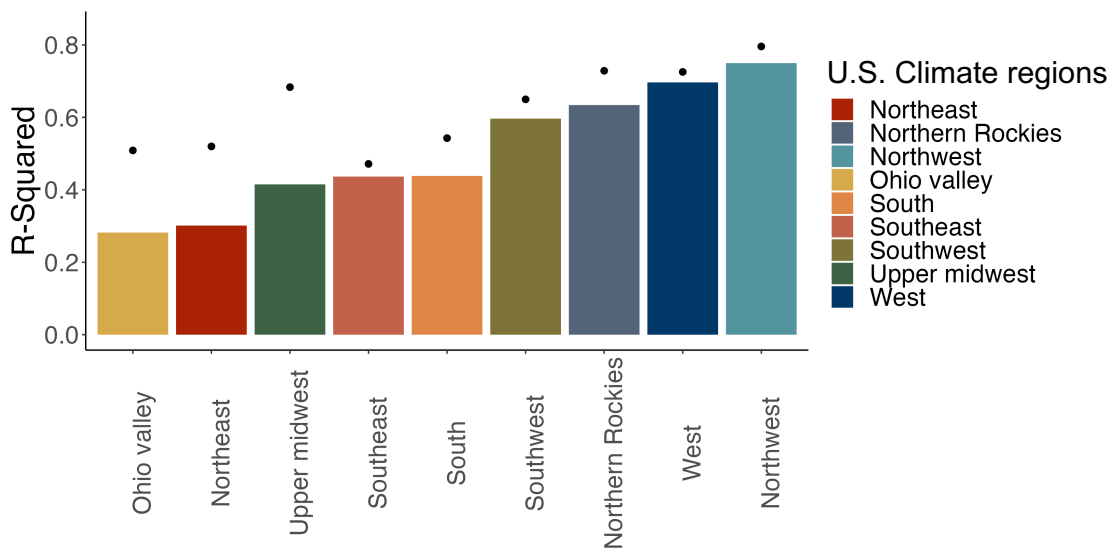


Figure S2: Performance of the fire-smoke regression models. The black dots show the full adjusted  $R^2$  of the regression model. The color bars show the within  $R^2$  after partialing out the month-of-year and grid cell fixed effects. The within  $R^2$  thus quantify the model predictive performance within each grid cell and month-of-year. Each bar shows the performance of a fire-smoke model in one of the nine US climate regions.

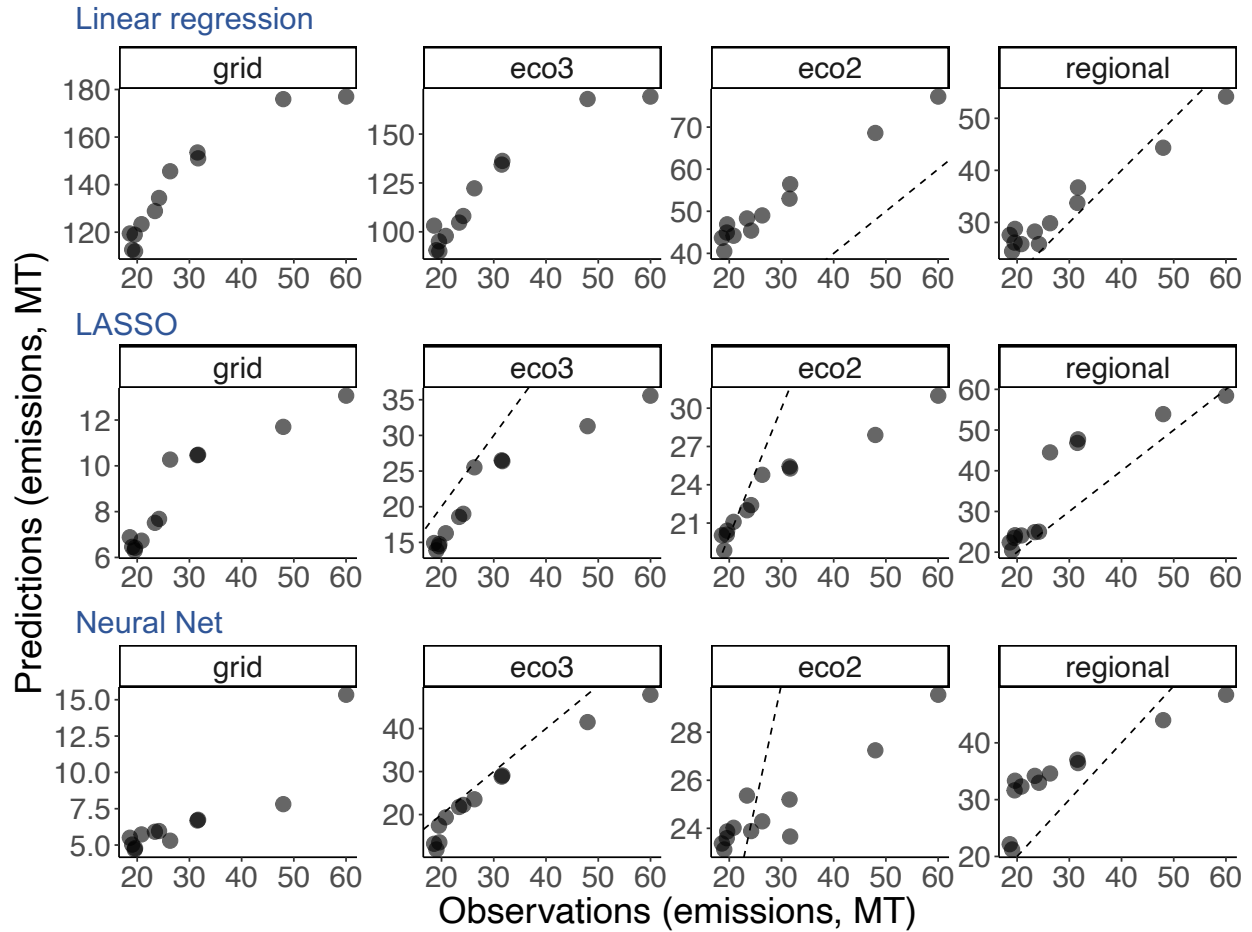


Figure S3: Predictive performance of models trained at different spatial resolutions (Western US). The plot shows the 10-year moving average of predicted emissions (y-axis) against the observed emissions (x-axis) from models trained at different spatial resolutions. For each algorithm (row), results are presented for models trained using grid cell data (“grid”), data aggregated at the level-3 ecoregion (“eco3”), data aggregated at the level-2 ecoregion (“eco2”), and data aggregated at the regional level (“regional”). Despite the different spatial resolutions of training data, the evaluation is at the regional level: we first aggregate the out-of-sample prediction to the regional level and compare the aggregated predictions against the aggregated observations. Dashed lines are 1-1 lines.



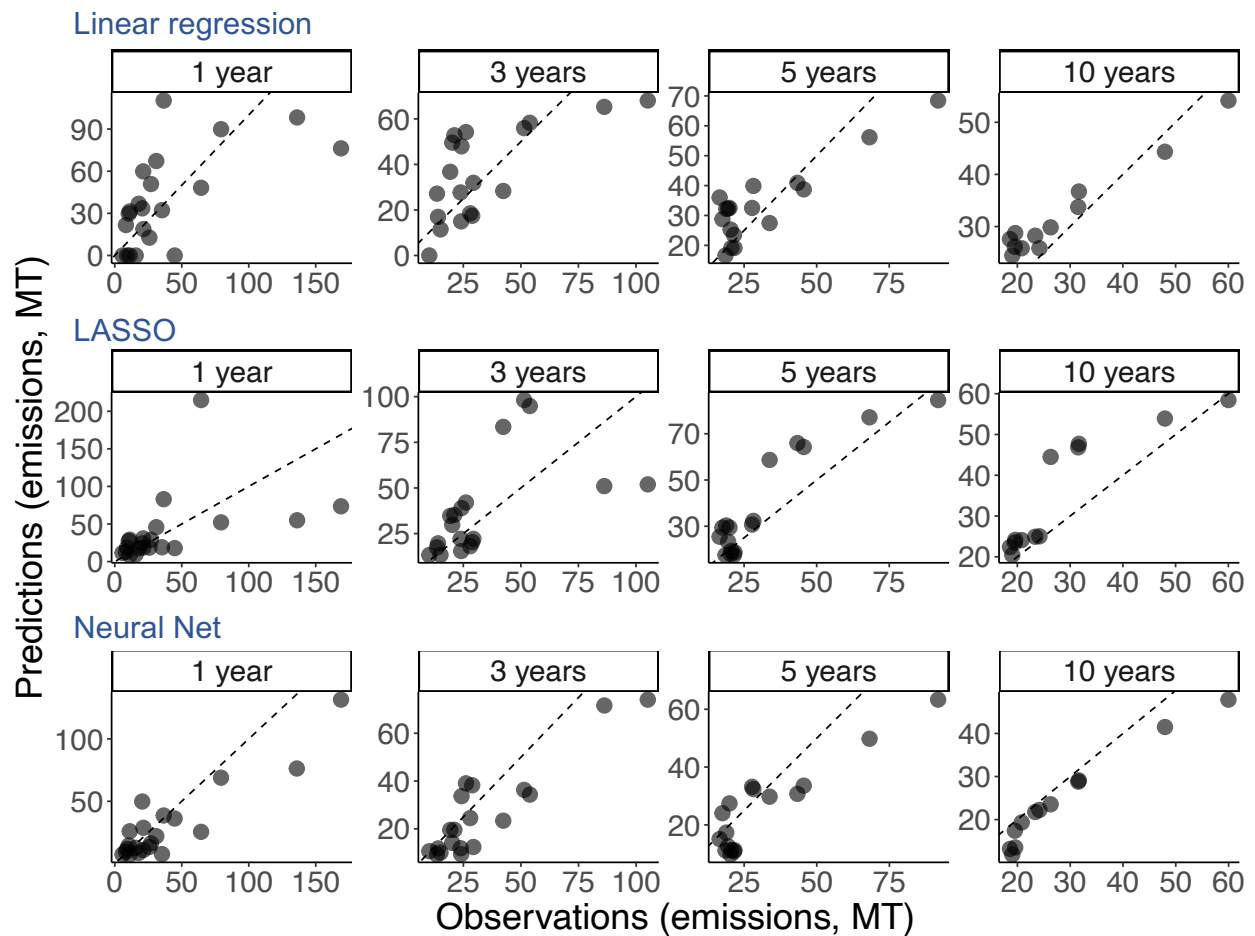


Figure S4: Predictive performance of models evaluated at different temporal scales (Western US). The plot shows the 10-year moving average of predicted emissions (y-axis) against the observed emissions (x-axis) from the same set of model but evaluated at different temporal scales. For each algorithm (row), the results show the out-of-sample prediction aggregated at different temporal scales ranging from no-aggregation (i.e. 1 year), to aggregation at the 10-year intervals. Dashed lines are 1-1 lines.

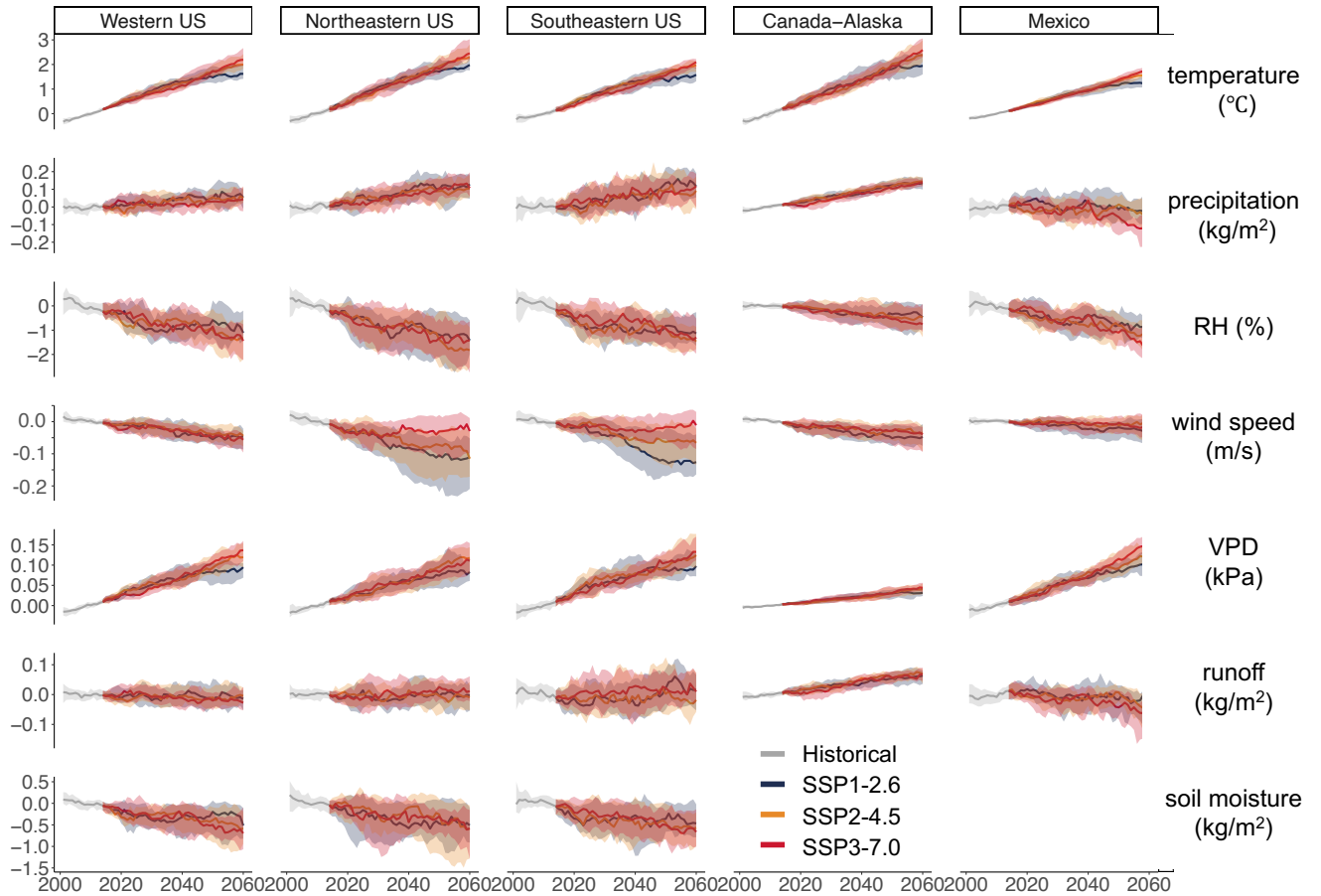


Figure S5: Projections of the climatic variables used in our statistical and machine learning models. Colour line indicates the median across 28 GCMs, and the shade area shows the 25th and 75th percentile across GCMs. The plot shows the 10-year moving average of the anomalies of each variable relative to the average values under historical scenario during 2001-2014. Soil moisture is not shown in Canada-Alaska and Mexico, as historical observations of soil moisture from NLDAS-2 are not available for these two regions.

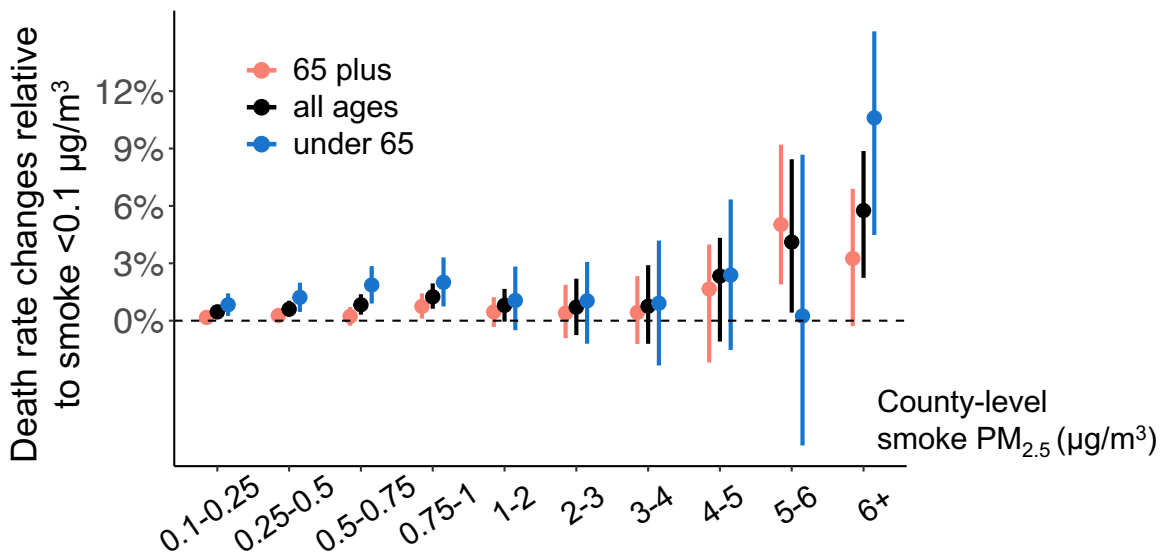


Figure S6: Impacts of smoke  $\text{PM}_{2.5}$  concentration on mortality rates estimated by age group. The figure shows the effects of exposure to different annual mean concentration of smoke  $\text{PM}_{2.5}$  (x-axis) relative to a no-smoke year (defined as a year with smoke  $\text{PM}_{2.5}$  concentration less than  $0.1 \mu\text{g}/\text{m}^3$ ), estimated using a Poisson model at the county and annual level. The error bars show the 95% confidence interval estimated using bootstrap.

### Percentage of estimated deaths in each smoke bin

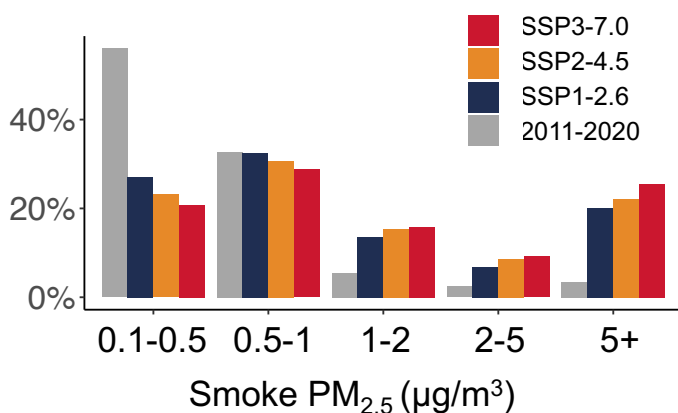


Figure S7: Percentage of estimated death contributions from each smoke concentration bin. The plot shows the contribution to total smoke-related deaths from county-years with annual mean smoke concentrations that fall in different smoke concentration bins under each scenario.

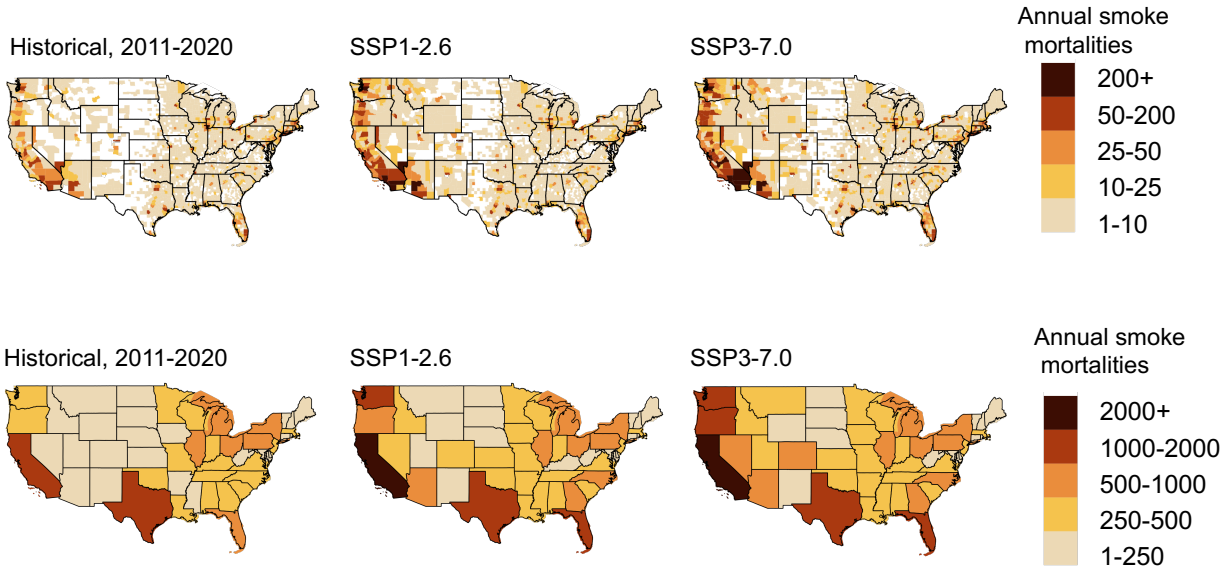


Figure S8: Estimated annual excess deaths due to smoke  $PM_{2.5}$  under the historical, SSP1-2.6, and SSP3-7.0 scenarios. The top panels show estimates at the county level. The bottom panels show estimates at the state level.

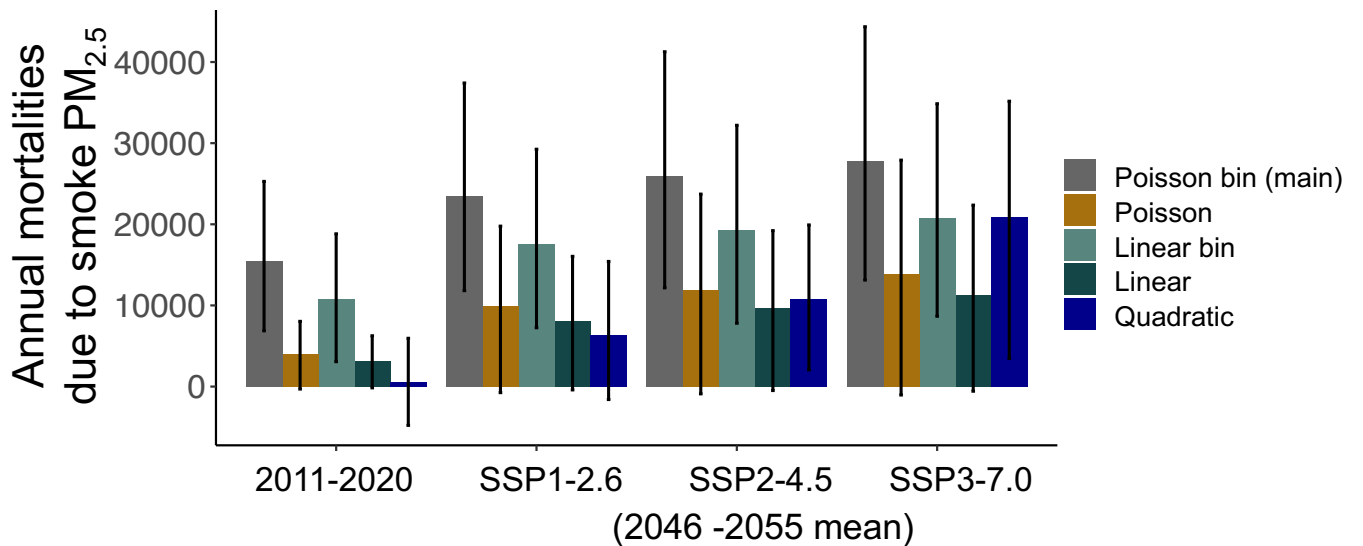


Figure S9: Estimated annual excess deaths due to smoke  $PM_{2.5}$  across alternative dose-response functions. Our main analysis uses the “Poisson bin” specification. The error bars show the 95% confidence interval estimated using bootstrap.

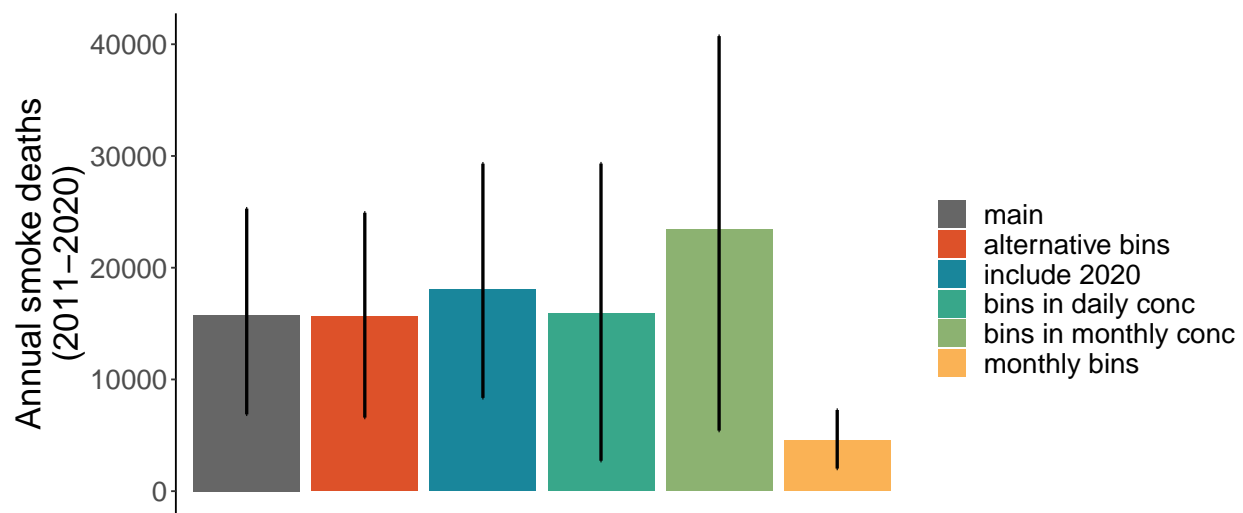


Figure S10: Estimated annual excess deaths due to smoke  $PM_{2.5}$  (2011-2020) across alternative specifications of the Poisson model. In addition to our main model (grey bar), we estimate a model which uses alternative bin definitions, a model which includes year 2020, a model which calculates the number of months or the number of days in a year that fall in different smoke bins to represent different temporal aggregations, and a model which is estimated at the county-month level. The error bars show the 95% confidence interval estimated using bootstrap.

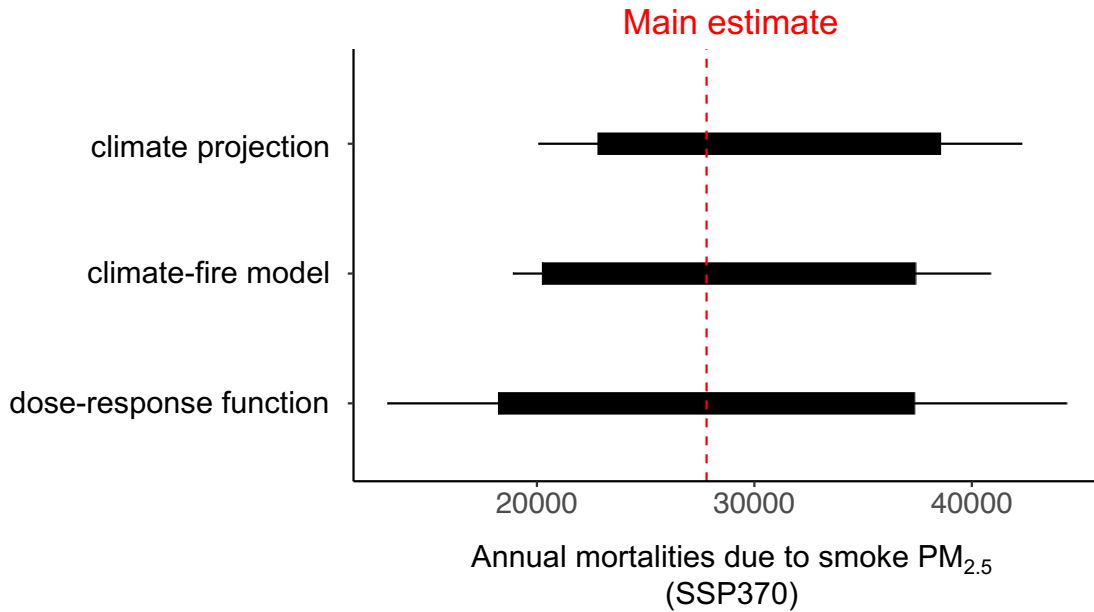


Figure S11: Uncertainty in estimated annual excess deaths due to wildfire smoke PM<sub>2.5</sub> under SSP3-7.0 scenario. The figure shows the uncertainty of the mortality estimates due to climate projections, climate-fire model, and the dose-response function between smoke and mortality. The red dashed line shows the main estimate reported in the paper (i.e. 27,800 excess deaths per year). The solid bar shows the 10th and 90th percentile, and the black line shows the 2.5th and 97.5th percentile. Uncertainty from “climate projection” is calculated using the percentiles of the estimated mortality from the 28 GCMs. Uncertainty from “climate-fire model” is calculated using bootstrap procedures performed on the individual fire-climate models from each region. More specifically, we first construct bootstrapped samples of the fire-climate panel dataset (sample with replacement) and then fit fire-climate model from each bootstrapped sample, and use these models to project smoke deaths. Uncertainty from “dose-response function” is calculated using bootstrap procedures performed on the health response functions. More specifically, we construct bootstrapped samples of the smoke-death dataset and estimate one dose-response function from each sample.

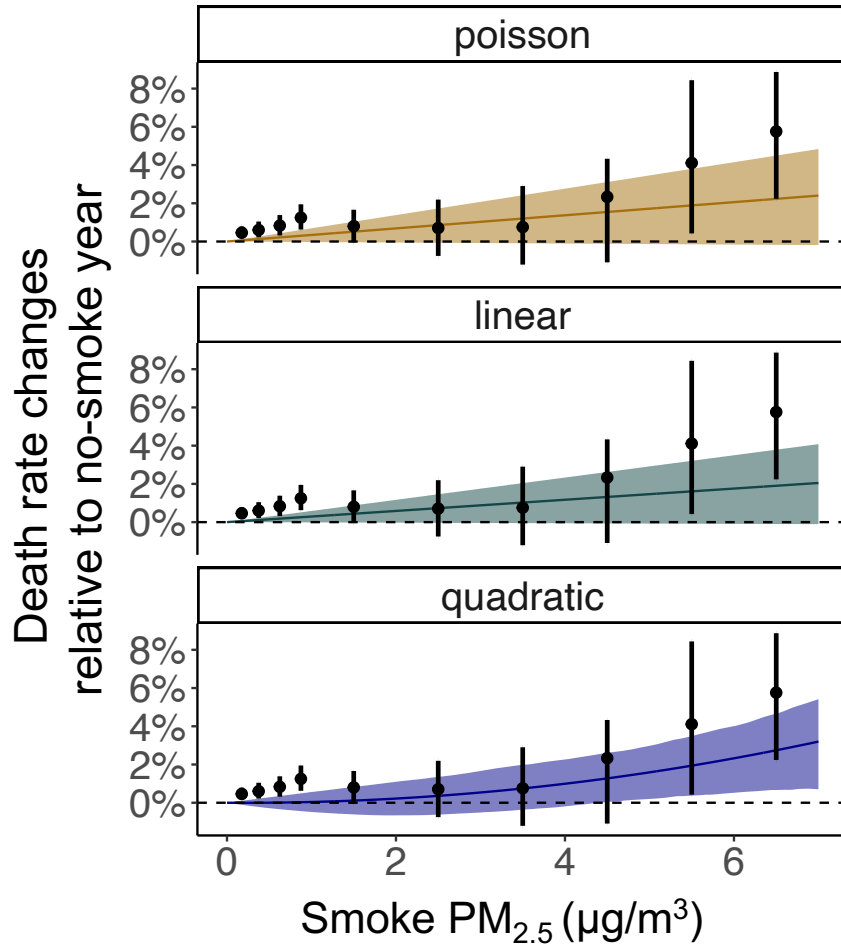


Figure S12: Impacts of smoke PM<sub>2.5</sub> concentration on mortality rates estimated using three alternative dose-response functions. The three colour lines show the estimated results from three non-binned models with poisson, linear, and quadratic specifications. For comparison, the black dots show the estimated coefficients from our main model (Poisson bin model). The shaded areas and the error bars represent the 95% confidence interval estimated using bootstrap procedure.