

Evaluating estimation methods for wildfire smoke and their implications for assessing health effects

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1 Abstract

2 Growing wildfire smoke represents a substantial threat to air quality and human health in the
3 US and across much of the globe. However, the impact of wildfire smoke on human health re-
4 mains imprecisely understood, due to uncertainties in both the measurement of population wildfire
5 smoke exposure and dose-response functions linking exposure to health. Here, we compare daily
6 wildfire smoke-related surface fine particulate matter ($PM_{2.5}$) concentrations estimated using three
7 approaches, including two chemical transport models (CTMs): GEOS-Chem and Community Mul-
8 tiscale Air Quality (CMAQ), and one machine learning (ML) model over the contiguous US in 2020,
9 a historically active fire year. We study the consequences of these different approaches for estimat-
10 ing smoke $PM_{2.5}$ concentrations and the effects of smoke $PM_{2.5}$ on mortality. In the western US,
11 compared against surface $PM_{2.5}$ measurements from US Environmental Protection Agency (EPA)
12 and PurpleAir sensors, we find that CTMs overestimate $PM_{2.5}$ concentrations during extreme smoke
13 episodes by up to 3-5 fold, while ML estimates are largely consistent with surface measurements.
14 However, in the eastern US, where smoke levels were much lower in 2020, CTMs show modestly
15 better agreement with surface measurements. We develop a calibration framework that integrates
16 CTM- and ML-based approaches and yields estimates of smoke $PM_{2.5}$ concentrations that outper-
17 form each individual approach. When combining the estimated smoke $PM_{2.5}$ concentrations with
18 county-level mortality rates, we find consistent effects of low-level smoke on mortality but large
19 discrepancies on the effects of high-level smoke exposure across different methods. Our research
20 highlights the benefits and costs of different estimation methods for understanding the health im-
21 pacts of wildfire smoke, and demonstrates the importance of bench-marking estimates with available
22 surface measurements.

23 Introduction

24 Wildfires and the smoke they generate pose a substantial threat to the environment and public
25 health globally. In the United States, wildfire burned area has more than quadrupled over the last
26 three decades (1), largely driven by human-induced climate change (2–4), historical fire suppression
27 (5), and the expansion of human activities into forested areas (6). Increased wildfire activity and
28 associated smoke emissions have also contributed to significant increases in ambient air pollution
29 (specifically fine particulate matter, $\text{PM}_{2.5}$) (7–10). In many parts of the western US, recent
30 estimates suggest that wildfire smoke $\text{PM}_{2.5}$ has accounted for over 50% of the annual concentration
31 of $\text{PM}_{2.5}$ in extreme smoke years (11, 12), and has led to stagnation or even reversal of the otherwise
32 declining trend in ambient $\text{PM}_{2.5}$ over the last two decades (13). As a result of increased wildfire
33 risks under future climate change, wildfire smoke pollution and the associated health burdens are
34 projected to increase substantially in the US in the coming decades (14–17).

35 While accumulating evidence suggests that exposure to wildfire smoke $\text{PM}_{2.5}$ can negatively
36 impact physical and mental health outcomes, large uncertainties remain in estimating the mortality
37 and disease burden attributable to wildfire smoke (18, 19). Such uncertainty is a result of 1) the
38 difficulty in estimating pollutant concentrations associated with wildfire smoke (and thus population
39 exposures), and 2) the uncertainty of derived dose-response functions that relate wildfire smoke to
40 various health outcomes. For instance, existing work has shown that estimates of $\text{PM}_{2.5}$ enhance-
41 ment due to the same wildfire events can differ dramatically depending on the data and models
42 used in the process (20, 21), which ultimately lead to widely different estimates of health burdens.
43 The broader literature on the health impacts of wildfire smoke exposure remains similarly mixed
44 (18, 19, 22), perhaps in part related to methodological challenges in modelling wildfire-specific
45 $\text{PM}_{2.5}$ exposure (23). For example, previous studies have shown both positive, negative, and no
46 associations between smoke $\text{PM}_{2.5}$ and cardiovascular outcomes (19, 24), in contrast to robustly
47 identified effects of total $\text{PM}_{2.5}$ on cardiovascular mortality and morbidity (25). Compared to total
48 all-source $\text{PM}_{2.5}$, modelling wildfire smoke $\text{PM}_{2.5}$ concentrations is more challenging due to the lack
49 of benchmark measurements, as surface monitors only measure $\text{PM}_{2.5}$ concentrations in the ambient
50 atmosphere which include contributions from fire and non-fire sources, and because wildfire smoke
51 emissions and concentrations change more dynamically across space and time.

52 Broadly speaking, researchers have used two approaches to estimate wildfire smoke impacts on
53 surface $\text{PM}_{2.5}$. One widely-used approach is mechanistic atmospheric chemical transport models
54 (CTM) that simulate the effects of wildfires on surface $\text{PM}_{2.5}$. Studies often use CTMs paired
55 with wildfire emissions inventories to simulate two scenarios: one including wildfire emissions and
56 one excluding wildfire emissions. They then attribute the differences between the two scenarios as
57 the estimated wildfire smoke $\text{PM}_{2.5}$ concentrations. Wildfire smoke $\text{PM}_{2.5}$ simulated by CTMs are
58 widely used in epidemiological studies to estimate dose-response functions (26–28) and to quantify

59 health burdens due to wildfire smoke (17, 26, 29, 30). However, the estimated wildfire smoke $PM_{2.5}$
60 from CTMs is subject to uncertainty in emission inventories (31, 32), plume rise (20, 33), and fire-
61 weather interactions (34), which results in modeled smoke concentrations potentially differing by an
62 order of magnitude when compared to surface observations (35, 36). Some studies have calibrated
63 CTM outputs against available surface measurements to correct for their potential biases before
64 using them in downstream health impacts analysis (37), but such practices have not been widely
65 adopted.

66 Another increasingly popular approach is to use statistical and machine learning (ML) methods
67 to characterize the relationship between input variables (such as remotely sensed atmospheric vari-
68 ables, meteorology, and fire information) and $PM_{2.5}$ measured at surface monitors during wildfire
69 episodes. CTM outputs are sometimes also used as input features for predicting surface $PM_{2.5}$
70 concentrations (38). By applying such a relationship to locations without surface monitors, these
71 studies can generate wildfire smoke estimates continuously in space and time (similar to CTMs).
72 Prior studies have used various statistical algorithms of different complexity (23, 39, 40). More
73 recently, ML methods have been used to estimate wildfire smoke concentrations (9, 10, 41). Similar
74 to CTM outputs, wildfire smoke $PM_{2.5}$ generated by these statistical methods are widely used for
75 establishing dose-response functions and assessing the overall health effects of wildfire smoke (14,
76 42–45).

77 Despite the popularity of CTM and ML methods for quantifying the health impacts of wildfire
78 smoke, little are known about the implications and influence of the different wildfire smoke esti-
79 mation approaches for downstream health impact assessment. Previous studies have quantified the
80 performance and uncertainty of each approach (often by comparing against surface observations)
81 or compared CTMs to other non-ML approaches (23, 38, 46–49). However, to our knowledge,
82 there is no inter-comparison between CTM and the increasingly popular ML approaches in terms
83 of their ability to predict wildfire smoke $PM_{2.5}$. More importantly, for downstream users of these
84 datasets, very little is known about the differences in the established dose-response functions and
85 health burdens across the different smoke estimation methods. Differences across wildfire smoke
86 $PM_{2.5}$ datasets can affect estimated health burdens in two ways. First, when applied to an existing
87 dose-response function, disparity in smoke exposure can lead to widely different attributed health
88 impacts. Second, the use of smoke $PM_{2.5}$ data to estimate novel dose-response relationships can
89 yield biased estimates of these relationships if the smoke $PM_{2.5}$ are themselves estimated with error.
90 In a case study that evaluated the effects on hospitalizations in Washington State over a four-month
91 period, Gan et al. demonstrated that the choice of wildfire estimation methods can generate dif-
92 ferent health impact estimates across CTM, spatial interpolation, and regression methods (23).
93 Better understanding the potential measurement error associated with leading smoke estimation
94 approaches to estimating smoke $PM_{2.5}$, and the implications for estimating dose-response func-
95 tions and health burdens, is essential for quantifying the health impacts of wildfire smoke, and for

96 informing researchers on how to best use the wildfire smoke datasets for health impact analysis.

97 Here, we compare the estimated daily wildfire smoke $PM_{2.5}$ in 2020 over the contiguous US
98 using two CTMs (GEOS-Chem and Community Multiscale Air Quality Model (CMAQ)) and one
99 ML model. Figure 1 shows an overview of our research methods. For the GEOS-Chem CTM, we
100 simulate two scenarios, a baseline scenario that includes wildfire emissions derived from the fourth
101 version of the Global Fire Emissions Database with small fires (GFED4s) (50), and a no-fire scenario
102 that excludes wildfire emissions. For the CMAQ CTM, we use daily wildfire smoke $PM_{2.5}$ archived
103 in (11). For ML estimates, we use the daily smoke $PM_{2.5}$ from (9). All three estimates are compared
104 to surface $PM_{2.5}$ concentrations measured by EPA reference-grade monitors and PurpleAir sensors.
105 We design two methods for evaluating the performance of the three estimation approaches: 1)
106 directly comparing to anomalous increases of surface $PM_{2.5}$ at monitors when smoke plumes are
107 overhead, and 2) comparing to total $PM_{2.5}$ measurements after adding estimated smoke $PM_{2.5}$ to the
108 same non-smoke $PM_{2.5}$ estimates. To ensure fair comparisons between CTM and ML approaches,
109 we use ML model predictions obtained from held-out monitor locations that were not used in model
110 training. We then develop a calibration framework that integrates the three individual estimates to
111 generate improved smoke $PM_{2.5}$ estimates.

112 Finally, we empirically estimate the effects of annual smoke $PM_{2.5}$ concentration, derived from
113 the calibrated model and three individual methods, on annual mortality rates using county-level
114 data from 2006 to 2020 on all recorded deaths in the US. Due to data availability, we combine the
115 ML estimates from 2006-2019 and the estimates from each method in 2020. Thus, our approach only
116 quantifies the differences in the dose-response functions due to one year of wildfire smoke data coming
117 from different estimation methods. Finally, we calculate the smoke-related excess deaths across the
118 calibrated model and three individual methods, using both the dose-response function derived as
119 described above for each method as well as smoke-specific dose-response functions documented in
120 prior literature.

121 **Methods**

122 **Chemical transport models**

123 We use GEOS-Chem version 14.0.2 (<https://doi.org/10.5281/zenodo.1343546>) driven by assimilated
124 meteorological data from the NASA Global Modeling and Assimilation Office (GMAO). GEOS-
125 Chem computes the evolution of atmospheric composition by a successive application over model
126 time steps of the operators simulating emissions, transport, chemistry, and deposition (51). Here, we
127 conduct nested simulations at $0.25^\circ \times 0.3125^\circ$ horizontal resolution over the North America domain
128 ($140^\circ W-40^\circ W$, $10^\circ N-70^\circ N$) using the GEOS forward processing (GEOS-FP) meteorological data
129 set. Chemical boundary conditions at the edges of the nested domain are updated every 3-h from

136 run two scenarios to estimate wildfire effects on surface $\text{PM}_{2.5}$: 1) with GFED4s emissions turned
137 on simulating fire smoke in North America, and 2) with GFED4s turned off which produces a “no
138 smoke” control.

139 By default, GEOS-Chem distributes biomass burning emissions uniformly within the boundary
140 layer. To test the influence of injection heights and alternative emission inventory on the simulated
141 $\text{PM}_{2.5}$ concentrations in GEOS-Chem, we further perform three sensitivity simulations over Cali-
142 fornia (latitude: 27°N - 47°N , longitude: 110°W - 130°W) that: (1) distributes 65% of the biomass
143 burning emissions within the boundary layer and the other 35% of emissions in the first ten sigma
144 layers above the boundary layer following (54); (2) distributes 5% of the biomass burning emissions
145 within the boundary layer and 95% biomass burning emissions in the first ten sigma layers above the
146 boundary layer; (3) use CAMS Global Fire Assimilation System (GFAS) based on satellite observed
147 fire radiative power (55) with dynamic injection heights (56).

148 We also use the Community Multiscale Air Quality Model (CMAQ), another widely-used CTM
149 to quantify $\text{PM}_{2.5}$ enhancements due to wildfire smoke in 2020. We use the model output archived
150 from Li et al., 2021 (11). Li et al. used the CMAQ model version v5.3.1 to simulate daily surface
151 $\text{PM}_{2.5}$ at a spatial resolution of 12km for one scenario including wildfire emissions and one scenario
152 excluding them. The CMAQ simulation used the biomass burning emission inventory from GBBEPx
153 v3 system (57) and the injection height scheme from (58) based on a previous evaluation of 8
154 combinations of biomass burning emission inventory and injection height schemes (59). GBBEPx
155 v3 system estimates daily global biomass burning emissions at 0.1 degree or 3km resolution using
156 fire radiative power from a suite of satellite products, using the same algorithm as Quick Fire
157 Emissions Data set (QFED). More details about the CMAQ simulation can be found in (11). All
158 CTM outputs are regridded at 10km resolution to be comparable to the ML output (see below).

159 **Machine learning estimates of smoke $\text{PM}_{2.5}$**

160 We use gridded daily wildfire smoke $\text{PM}_{2.5}$ predictions for the contiguous US at 10 km resolution
161 from January 1, 2006 to November 30, 2020 derived from (9). Childs et al. constructed a ML
162 model that uses satellite-derived smoke plume data, remotely-sensed atmospheric variables, and
163 meteorological variables to predict the anomalous increases in surface $\text{PM}_{2.5}$ measured by surface
164 air quality monitors during wildfire. Their model achieved a R^2 of 0.67 when evaluated against
165 held-out samples at the daily monitor level. The dataset has been widely used for establishing
166 dose-response functions and assessing the overall health effects of wildfire smoke (14, 42, 43).

167 The ML model in (9) was trained to predict surface $\text{PM}_{2.5}$ data measured at EPA sensors,
168 and thus direct comparisons between the ML model outputs and surface measurements could lead
169 to inflated performance. To address this issue, we use the out-of-sample predictions from their
170 machine learning algorithm for grid cells that contain EPA sensors. In other words, predicted

171 smoke concentrations in those grid cells are derived from a model trained on a sample that excludes
172 monitor measurements in that grid cell. For grid cells with only PurpleAir sensors and no EPA
173 sensors, we directly use the output from Childs et al. as PurpleAir data was not used in the original
174 model training.

175 **Surface measurements of PM_{2.5}**

176 We use surface PM_{2.5} measurements derived from the reference-grade sensors administered by the US
177 EPA as well as measurements from low-cost PurpleAir sensors. Our final data includes 373,203 daily
178 measurements from 1,276 EPA sensors, and 1.3 million daily measurements from 6,553 PurpleAir
179 sensors. Surface PM_{2.5} concentrations are derived from the US Air Quality Systems administered
180 by the US EPA (60), and all publicly reporting outdoor PurpleAir monitors in the US. We first
181 regrid surface measurements at the 10km grid cell level to be consistent with the wildfire smoke
182 estimates, and calculate the daily mean concentration for each grid cell over the EPA and PurpleAir
183 sensors separately. If one grid cell has both EPA and PurpleAir sensors, we then take their average
184 as the daily mean concentration for that grid cell. We drop all observations with negative daily
185 PM_{2.5} concentrations (<0.5% of our full data). One thing to note is that samplers used in certain
186 surface monitors can get clogged due to overload of smoke (61). We use all measurements that
187 are available on the Air Quality Systems website, but the malfunctioning monitors during extreme
188 wildfire smoke could influence our evaluations.

189 For PurpleAir sensors, we only use measurements from outdoor sensors in 2020. The raw
190 temporal resolution of measurements is 10 minutes, and we temporally aggregate them to the daily
191 level after removing unrealistic 10-minute observations (62). Prior studies have found that PurpleAir
192 sensors can generally characterise enhancements of surface PM_{2.5} due to wildfire smoke, yet the
193 quantitative magnitude needs to be calibrated to match measurements obtained from reference-
194 grade air quality monitors (63). We use the method from Barkjohn et al. to calibrate the raw daily
195 concentrations from PurpleAir sensors for wildfire conditions (64). Following PurpleAir guidelines,
196 we then drop all measurements with a daily mean concentration above 1,000 $\mu\text{g}/\text{m}^3$ and top-code
197 all concentrations at 500 $\mu\text{g}/\text{m}^3$ if the raw concentrations are in the range of 500-1,000 $\mu\text{g}/\text{m}^3$. Only
198 96 records from 25 sensors are either dropped or top-coded at 500 $\mu\text{g}/\text{m}^3$ out of 1.3 million records
199 in our sample.

200 We combine surface measurements of PM_{2.5} with satellite-derived smoke plume data to calculate
201 the anomalous increases in surface PM_{2.5} due to wildfire smoke, following similar approaches from (8,
202 9). The smoke plume data is derived from the National Oceanic and Atmospheric Administration
203 (NOAA) Hazard Mapping System (HMS), which includes analyst-identified plume boundaries based
204 on visible bands of satellite imagery (65). For any given monitor, we determine a day to be “smoke
205 day” if identified plume was overhead the monitor location. For any smoke day, we calculate PM_{2.5}

206 anomalies as deviations from recent month- and location-specific non-smoke baselines. For each
207 monitor location, the non-smoke baseline is calculated as the median of all measurements from all
208 non-smoke days that fall in the same month in 2018–2020. The calculated anomalies in surface
209 $PM_{2.5}$ due to wildfire smoke thus consist of zero estimates on all non-smoke days and non-zero
210 deviation on smoke days.

211 **Evaluating wildfire smoke $PM_{2.5}$ estimates using surface measurements**

212 We design two approaches to evaluate the three wildfire smoke $PM_{2.5}$ datasets using surface $PM_{2.5}$
213 measurements from EPA and PurpleAir sensors:

214 First, we compare the estimates of smoke $PM_{2.5}$ from three approaches with the inferred smoke
215 $PM_{2.5}$ anomalies for each monitor (Evaluation 1 in Figure 1B). We calculate the root-mean square
216 error (RMSE) between the estimated smoke concentrations and monitor-level anomalies and use
217 it as the evaluation metric. We further evaluate the models’ performance under three conditions
218 with different levels of wildfire smoke. “No smoke” includes all monitor days with no smoke plume
219 overhead and low estimated smoke concentrations from both CTMs ($<0.5 \mu\text{g}/\text{m}^3$). “Medium and
220 high smoke” includes monitor-days with high estimated smoke $PM_{2.5}$ from both CTMs ($>5 \mu\text{g}/\text{m}^3$).
221 “Low smoke” category includes all the other conditions. As we noted above, we use the out-of-sample
222 predictions from the ML model to ensure a fair comparison between surface measurements and the
223 ML model.

224 Second, we design an evaluation method that uses *total* $PM_{2.5}$ measurements from surface air
225 quality sensors due to the potential uncertainty in the inferred smoke $PM_{2.5}$ anomalies (Evaluation 2
226 in Figure 1B). As the ML model did not estimate non-smoke $PM_{2.5}$ concentrations, we first create a
227 common non-smoke $PM_{2.5}$ baseline using non-smoke $PM_{2.5}$ simulated by GEOS-Chem and CMAQ:

$$PM^{monitor} = f(PM_{nonsmoke}^{GC}, PM_{nonsmoke}^{CMAQ}, \dots) \text{ — only trained on non-smoke days} \quad (1)$$

228 To do this, we construct a XGBoost ML model (function f in equation 1) to predict total $PM_{2.5}$
229 measurements during all “non-smoke days” using non-smoke $PM_{2.5}$ simulated by GEOS-Chem and
230 CMAQ as model features. The XGBoost model also uses temperature, precipitation, wind speed,
231 humidity, latitude, longitude, and month-of-year as model features to account for potential seasonal
232 and spatial model biases. Using this model, we then estimate non-smoke $PM_{2.5}$ for all monitor
233 days in our sample (including smoke days). Finally, we calculate three total $PM_{2.5}$ estimates over
234 all monitor locations by adding the different smoke $PM_{2.5}$ estimates to the same non-smoke $PM_{2.5}$
235 estimated above (Figure 1A). Therefore, the only differences in the three constructed $PM_{2.5}$ series
236 are due to their different estimates of smoke $PM_{2.5}$ as they share the same non-smoke $PM_{2.5}$ estimate.
237 We compare the three constructed total $PM_{2.5}$ series against surface measurements of total $PM_{2.5}$
238 by calculating RMSE for each monitor location.

239 **Constructing improved wildfire smoke $PM_{2.5}$ estimates by integrating all three** 240 **sources**

241 Previous work has shown that integrating multiple exposure estimation methods can improve wild-
242 fire smoke $PM_{2.5}$ modeling performance (39, 48, 66). Therefore, to generate improved wildfire
243 smoke estimates, we construct an additional XGBoost-based ML model to predict wildfire smoke
244 $PM_{2.5}$ estimates at all monitor locations that uses outputs from all three modeling approaches
245 (Figure 1C):

$$PM_{smoke}^{monitor} = g(PM_{smoke}^{GC}, PM_{smoke}^{CMAQ}, PM_{smoke}^{ML}, \dots) \text{ — trained on all days} \quad (2)$$

246 For each monitor-day, we first calculate wildfire smoke $PM_{2.5}$ ($PM_{smoke}^{monitor}$) as the difference between
247 total $PM_{2.5}$ measurements and the constructed non-smoke $PM_{2.5}$ (estimated using equation 1).
248 Our constructed non-smoke estimates better characterize the variability in daily non-smoke $PM_{2.5}$
249 relative to the constant non-smoke $PM_{2.5}$ baselines used in prior research (9). Furthermore, our
250 estimates of smoke $PM_{2.5}$ at the monitor level do not depend on the HMS smoke plume boundaries,
251 which are found to have large uncertainty under low and medium smoke conditions (67).

252 Separate calibration models are trained for the western US and the eastern US. The input
253 features include smoke $PM_{2.5}$ estimated by GEOS-Chem, CMAQ, and ML, aerosol optical thickness
254 from MERRA-2, meteorological variables including temperature, precipitation, wind speed, dew
255 point temperature, planet boundary layer height, surface pressure, latitude, longitude, and day-of-
256 year from ERA5. As our main purpose is to estimate wildfire smoke $PM_{2.5}$ for locations that are not
257 covered by surface monitors, we evaluate the model performance using 5-fold spatial cross-validation.
258 The spatial folds are defined considering the coarsest resolution of input features (in our case, the
259 MERRA-2 inputs at 0.5° latitude \times 0.625° longitude). Splitting train and test data sets by monitor
260 locations rather than the more conventional method of random splitting by observation (in which
261 a given monitor can contribute data to both train and test) is a more realistic evaluation of model
262 performance as it avoids leakage of information between training and test sets. To measure variable
263 importance in the XGBoost model, we use the contribution to model performance improvements
264 from “tree splits” made on each feature.

265 **Estimating dose-response function between mortality and wildfire smoke**

266 We empirically estimate a dose-response function between smoke $PM_{2.5}$ and all-cause mortality
267 rates using 2006-2020 county-level data (Figure 1D). As we only have CTM outputs in the year
268 2020, we construct four panels that include wildfire smoke from different estimation methods in
269 2020 (calibrated, GEOS-Chem, CMAQ, and ML), but the same wildfire smoke $PM_{2.5}$ estimates in
270 2006-2019 from the ML method (9). Therefore, our approach tests how a one-year difference in the
271 wildfire smoke estimates influences the derived dose-response function.

272 Following a similar method in (14), we combine county-level population-weighted annual smoke
 273 $PM_{2.5}$, with county-level all-cause mortality rates by different age groups. We obtain individual-
 274 level multiple cause of death mortality data from the National Center for Health Statistics to
 275 calculate age-standardized mortality rates for all ages (68). County-level mortality rates were age-
 276 standardized using the direct method and 5-year bins (0-4, 5-9, ..., 85 and over) based on the 2000
 277 US Census Standard Population. Monthly mortality rates were standardized per 100,000 people.
 278 To fully capture damages from ambient wildfire smoke concentrations, our preferred outcome is
 279 age-standardized, all-cause, all-age mortality rates at the county-year level.

280 In our main analysis, we estimate a Poisson model in which we allow non-linear impacts of
 281 annual smoke $PM_{2.5}$ on mortality rates at the county-year level following method from (14):

$$D_{csy} = \exp \left(\sum_i \beta_i smokeBIN_{csy}^i + \gamma W_{csy} + \eta_{sy} + \theta_c + \epsilon_{csy} \right) \quad (3)$$

282 where D_{csy} denotes the age-adjusted all-cause mortality rates in county c , state s , and year y .
 283 $smokeBIN_{csy}^i$ is a dummy variable for whether annual population-weighted smoke $PM_{2.5}$ in county
 284 c , state s , and year y falls into the range of bin i (0-0.1, 0.1-0.25, 0.25-0.5, 0.5-0.75, 0.75-1, 1-2, 2-3,
 285 3-4, 4-5, 5-6, $>6 \mu g/m^3$; 0-0.1 is the reference category). The main coefficients of interest are the
 286 β_i 's, which estimate the effects of a year with annual smoke concentration of bin i on mortality rates,
 287 relative to a year with annual mean smoke $PM_{2.5}$ concentration below $0.1 \mu g/m^3$. The reference
 288 category included <0.1 because only 4 county-year observations had exactly zero ambient wildfire
 289 smoke. W_{csy} denotes a flexible control of temperature (the number of days that fall in different
 290 temperature bins) and linear and quadratic terms of annual population-weighted precipitation. η_{sy}
 291 denotes a vector of state-year fixed effects (i.e. separate intercepts for each year in each state)
 292 that accounts for all factors that differ across states in a given year (e.g. California 2018 versus
 293 Oregon 2018) as well as all factors that differ within states across years (e.g. California 2017 versus
 294 California 2018). θ_c denotes a set of county-level fixed effects that accounts for any county-specific
 295 time-invariant factors that could be correlated with both smoke exposure and mortality. In essence,
 296 we identify the effect of wildfire smoke on mortality using within-county variation over time, after
 297 accounting for any factors that trend over time within that county's state, and for any correlation
 298 between smoke variation and variation in temperature and precipitation.

299 The coefficients are estimated using weighted Poisson regression models, with the function "fe-
 300 pois" from R package "fixest". The estimations are weighted by county-level population counts to
 301 enable estimates of population-averaged effects, as well as to reduce statistical uncertainty. The
 302 uncertainty of the coefficients is estimated using bootstrap of 500 runs. ϵ_{csy} represents the error
 303 terms.

304 **Quantifying smoke-related health burdens**

305 We quantify the excess mortality attributable to wildfire smoke in 2020, using the four different
306 wildfire smoke estimates (from the three individual approaches and the calibrated model) and dif-
307 ferent dose-response functions. We first calculate the excess mortality using four empirically derived
308 dose-response functions that are associated with each estimation method (as developed above). To
309 understand the difference in health burdens associated with smoke $\text{PM}_{2.5}$ estimations alone, we fur-
310 ther calculate the mortality burdens using the same dose-response functions applied to the wildfire
311 smoke derived from different estimation methods. We use two dose-response functions, one at the
312 annual level, and one at the daily level. At the annual level, we use the dose-response function
313 derived using the wildfire smoke data from the ML method. At the daily level, we use the dose-
314 response function from a recent meta-analysis (19). Meta-analyzing eight prior published estimates,
315 Gould et al. estimated an increase in daily mortality rate by 0.15% (95% CI: 0.01%, 0.28%) per 1
316 $\mu\text{g}/\text{m}^3$ of smoke $\text{PM}_{2.5}$ on the same day (without considering lagged effects).

Results

Figure 2 shows the annual average wildfire smoke $\text{PM}_{2.5}$ concentration over the contiguous US (averaged from February-October in 2020) across the three estimation methods. We find that estimated wildfire smoke $\text{PM}_{2.5}$ concentrations vary substantially among the three approaches. In the western US, GEOS-Chem and CMAQ estimate much higher wildfire smoke concentrations compared to the ML estimates. For example, in Oregon and northern California, GEOS-Chem estimates an annual average smoke $\text{PM}_{2.5}$ concentration of $>50 \mu\text{g}/\text{m}^3$, while the ML estimates are $15\text{-}25 \mu\text{g}/\text{m}^3$. Population-weighted wildfire smoke $\text{PM}_{2.5}$ exposure also differs dramatically. While all three methods estimate that the population-weighted average smoke $\text{PM}_{2.5}$ reaches the highest of the year on September 10-13, the max daily population-weighted mean smoke $\text{PM}_{2.5}$ estimated by GEOS-Chem ($54.5 \mu\text{g}/\text{m}^3$) is 3x the ML estimates ($17.5 \mu\text{g}/\text{m}^3$) and 2x the CMAQ estimates ($25.4 \mu\text{g}/\text{m}^3$). However, regional differences exist in the intercomparison of the three methods. In the southeastern US, CMAQ estimates are the highest among the three methods, and GEOS-Chem estimates are generally the lowest (Figure 2D). This could be due to the biased global Aerosol Optical Depth-based scaling factor used in the QFED emission inventory which results in higher estimated emissions in the southeast US (31, 69).

The largest difference across three methods comes from their estimations of the smoke events in September 2020 in Oregon and California. When comparing model estimates over this region during a two-week period (September 6-20), we find even larger differences across methods (Figure 3). While the spatial distributions of wildfire smoke are similar across the three approaches, the estimated magnitude can differ by as much as 40x. The maximum daily smoke concentration for any grid cell in this region is estimated to be $32,700 \mu\text{g}/\text{m}^3$ in GEOS-Chem, $17,300 \mu\text{g}/\text{m}^3$ in CMAQ and $774 \mu\text{g}/\text{m}^3$ in ML model. When comparing against surface measurements from EPA and PurpleAir sensors in this area, we find that the ML approach shows a much better agreement with surface measurements, while GEOS-Chem and CMAQ substantially overestimate surface $\text{PM}_{2.5}$ concentration (Figure 3D and 3E). The maximum daily smoke $\text{PM}_{2.5}$ concentration measured by any EPA or PurpleAir sensors is $821 \mu\text{g}/\text{m}^3$ over this two-week period. We further investigate the significant model bias of the CTMs over this period, by conducting sensitivity simulations with varying injection heights and an alternative emission inventory in GEOS-Chem. We find the upward model bias in our original GEOS-Chem simulation (with default injection height and GFED emissions) is likely due to both the unrealistic injection heights and overestimated emissions in GFED, with the emissions being the more important factor (Figure S1).

Given the underlying uncertainty in the inferred monitor-level smoke $\text{PM}_{2.5}$, we further evaluate the performance of GEOS-Chem and CMAQ against surface measurements of total $\text{PM}_{2.5}$ (Figure 4 and S2). When there is no or low wildfire smoke, we find that CTMs can capture the overall spatial and temporal variability of $\text{PM}_{2.5}$ (Figure 4A). CTMs are generally able to predict the

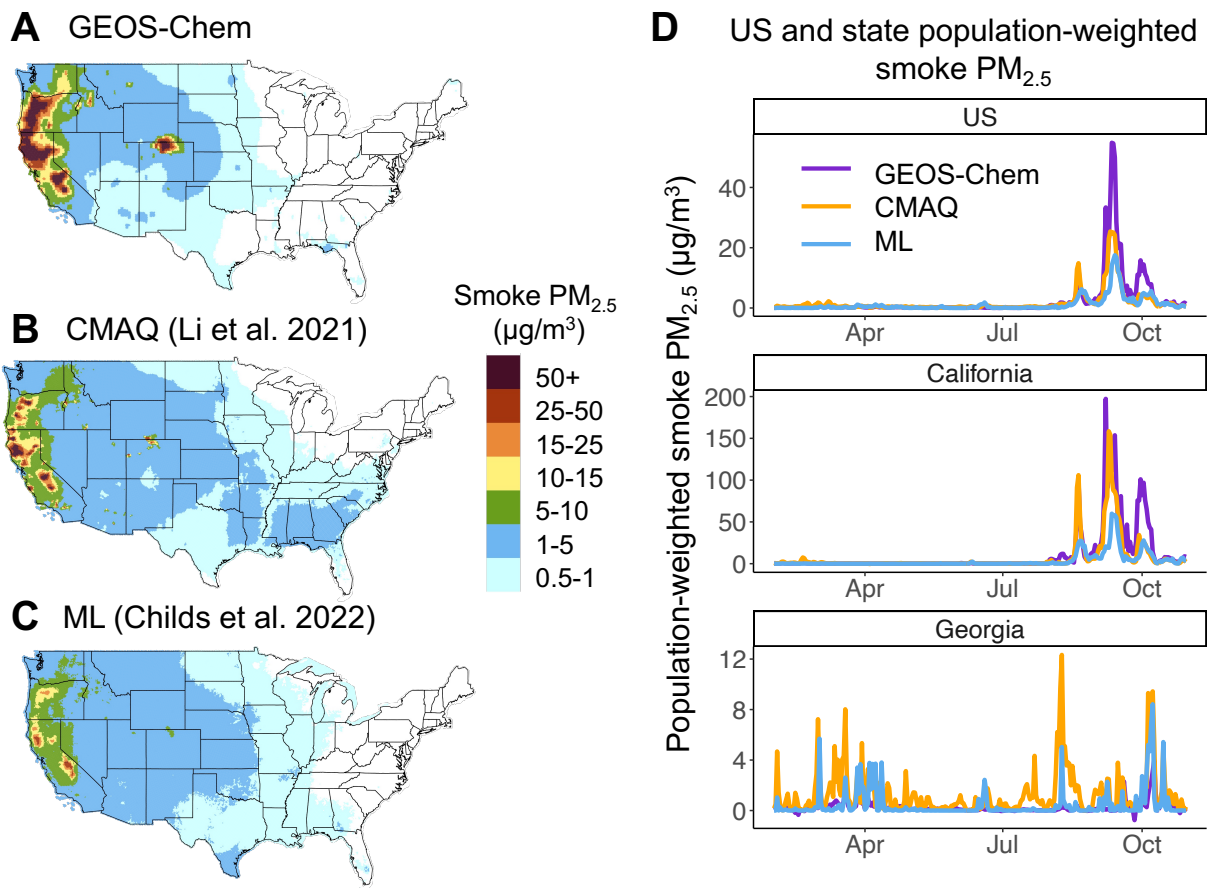


Figure 2: **Average smoke $PM_{2.5}$ concentration estimated by the three methods.** Panels A-C show the average smoke $PM_{2.5}$ concentration during February to October in 2020 estimated by GEOS-Chem (A), CMAQ (B, derived from (11)), and a ML method (C, derived from (9)). Panel D shows the population-weighted smoke $PM_{2.5}$ concentration in the contiguous US, California, and Georgia in 2020.

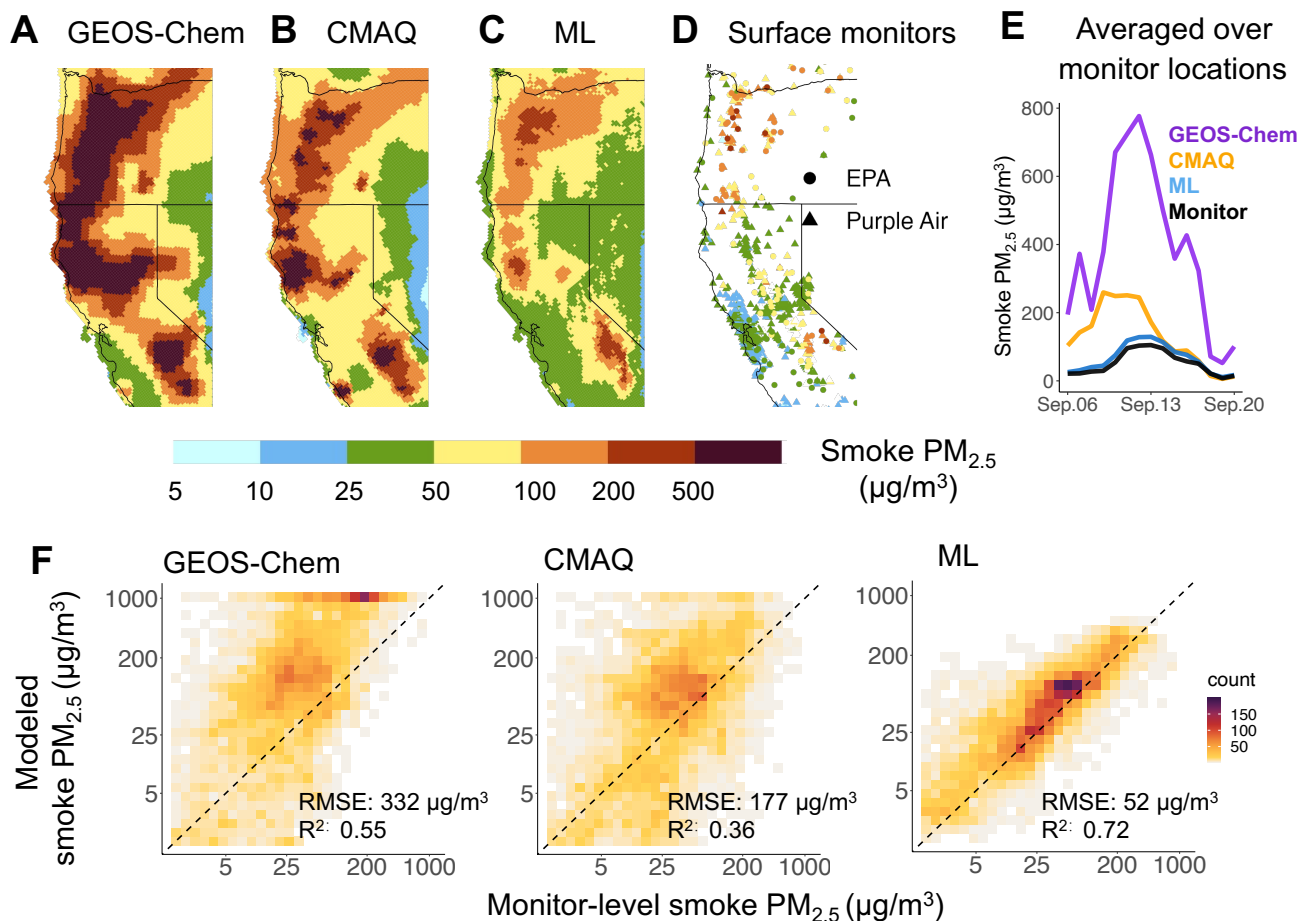


Figure 3: **Smoke $PM_{2.5}$ concentrations in extreme wildfire episodes in Oregon and northern California in September 2020.** Panels A-C show the average smoke $PM_{2.5}$ concentrations during September 6–20 estimated by GEOS-Chem, CMAQ, and ML, respectively. Panel D shows the average smoke $PM_{2.5}$ estimates derived from surface measurements of surface reference-grade monitors administered by US EPA and PurpleAir sensors. Panel E shows the average smoke $PM_{2.5}$ concentration estimated by different methods over all monitor locations. Panel F shows the estimated wildfire smoke $PM_{2.5}$ (y-axis) against the inferred smoke $PM_{2.5}$ at the monitor level. Inferred wildfire smoke $PM_{2.5}$ is estimated as the anomalous increases in total $PM_{2.5}$ when there is wildfire smoke overhead relative to baselines of $PM_{2.5}$ on non-smoke days, using a similar method as in Childs et al., 2022.

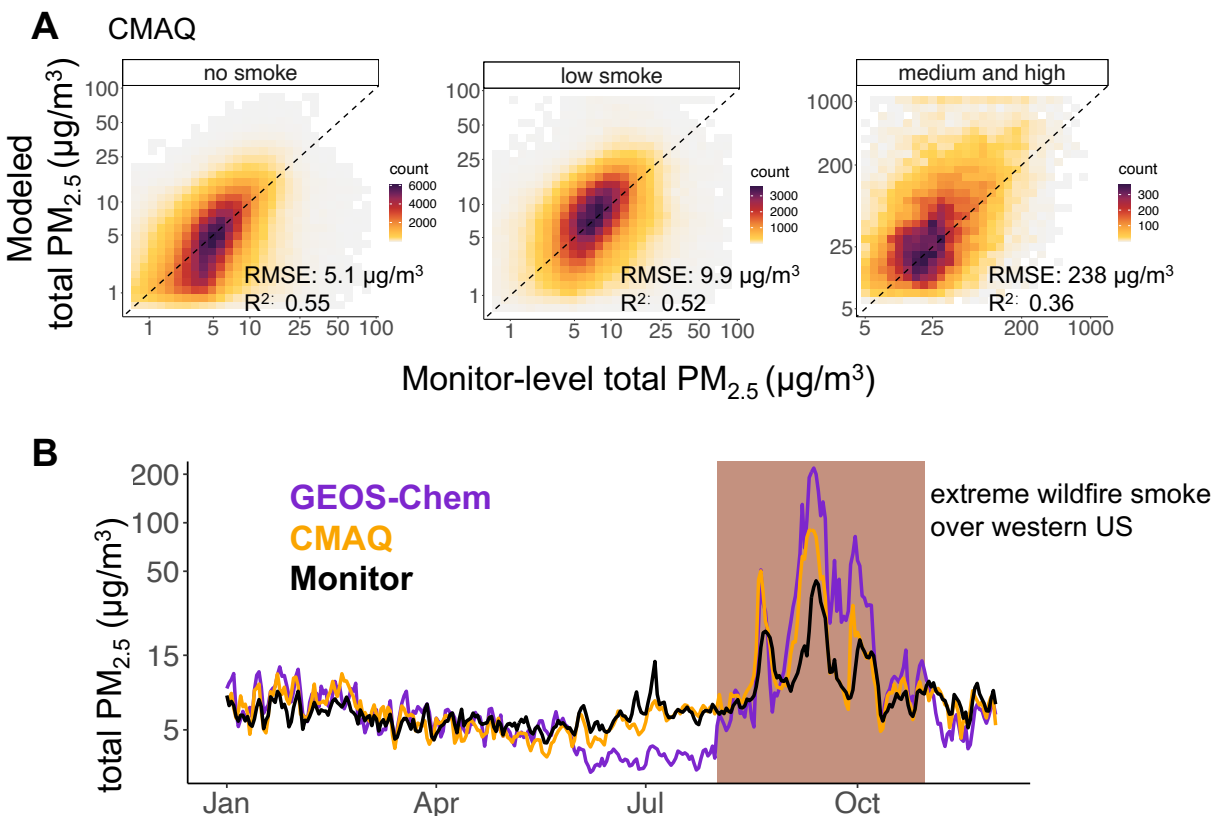
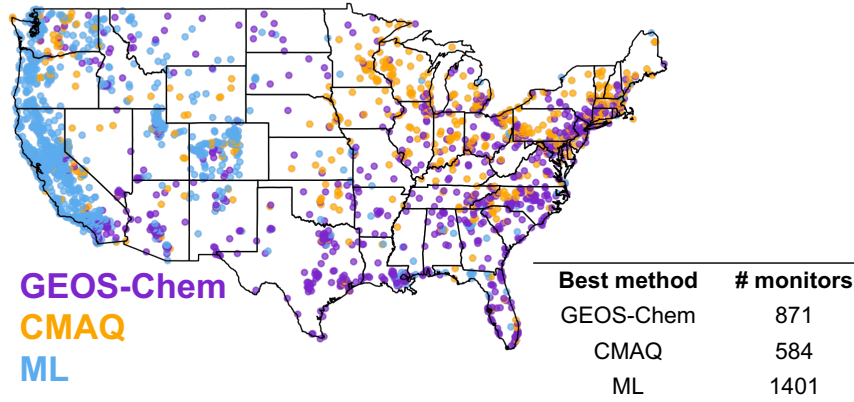


Figure 4: **Comparing CTM simulations of total $\text{PM}_{2.5}$ against surface measurements of total $\text{PM}_{2.5}$.** Panel A shows the total $\text{PM}_{2.5}$ concentration simulated by CMAQ (y-axis) against total $\text{PM}_{2.5}$ measurements from surface monitors under different levels of wildfire smoke. “No smoke” includes all monitor days with no smoke plume overhead and low estimated smoke concentrations from both CTM ($<0.5 \mu\text{g}/\text{m}^3$). “Medium and high smoke” includes monitor-days with high estimated smoke $\text{PM}_{2.5}$ from both CTMs ($>5 \mu\text{g}/\text{m}^3$). “Low smoke” category includes all the other conditions. Panel B shows the average total $\text{PM}_{2.5}$ concentration across all monitor locations in the US. The shade denotes a period when the western US experienced extreme wildfire smoke. Note the non-linear scale of y-axis for better visualization.

353 overall variability of $\text{PM}_{2.5}$ (Figure 4B), with the exception that the GEOS-Chem simulation shows
 354 a substantial downward bias in June and July, possibly due to over-partition of inorganic $\text{PM}_{2.5}$
 355 species in the gas phase as suggested by (70) (see Figure S3). However, when there are medium
 356 and high levels of smoke in the air (defined by location-days that all three approaches agree that
 357 there is some smoke), we observe substantially larger differences between model simulations and
 358 surface measurements of total $\text{PM}_{2.5}$. The average $\text{PM}_{2.5}$ concentration simulated by GEOS-Chem
 359 and CMAQ is 5x and 3x the $\text{PM}_{2.5}$ concentrations measured by surface monitors, over all monitor
 360 locations and high smoke days.

A which method performs best at each location



B performance differences against the best method

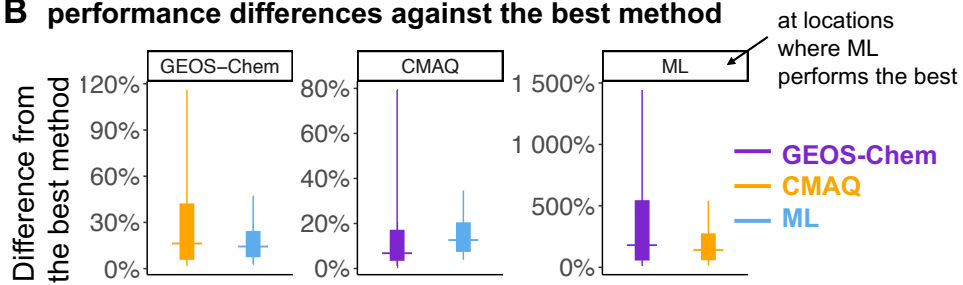


Figure 5: Evaluating the performance of three wildfire smoke estimation methods against total $\text{PM}_{2.5}$ measurements from surface monitors. Panel A shows the optimal estimation approach for each surface monitor location in our study. The color of the dots shows the optimal method at each location with the smallest RMSE against total $\text{PM}_{2.5}$ measurements. Table A shows the number of monitor locations that correspond to each of the estimation method. Panel B shows the difference in model performance from the best approach. For each panel, the bars show the range of performance difference against the best estimation approach (as shown in the panel title). The performance difference is measured as the relative difference between RMSE and RMSE of the best method. The solid bar shows the 25th and 75th percentile, and the line shows the median, 10th, and 90th percentile of the relative differences.

361 Using total $\text{PM}_{2.5}$ measurements from surface monitors as the evaluation benchmark, we identify
 362 the estimation approach with the best performance of characterizing smoke $\text{PM}_{2.5}$ in each monitor
 363 location in the US (Figure 5). We find that the ML approach performs the best in characterizing
 364 $\text{PM}_{2.5}$ concentrations in the western US where CTMs show a large model bias. However, GEOS-
 365 Chem and CMAQ outperform the ML approach in the eastern US, where low and medium levels
 366 of smoke are more prevalent. CTMs outperform the ML approach in these areas, possibly due
 367 to the uncertainty in non-smoke baselines used in the ML algorithm and the uncertainty of HMS
 368 smoke plumes over low and medium smoke conditions outside of the western US (67). As shown

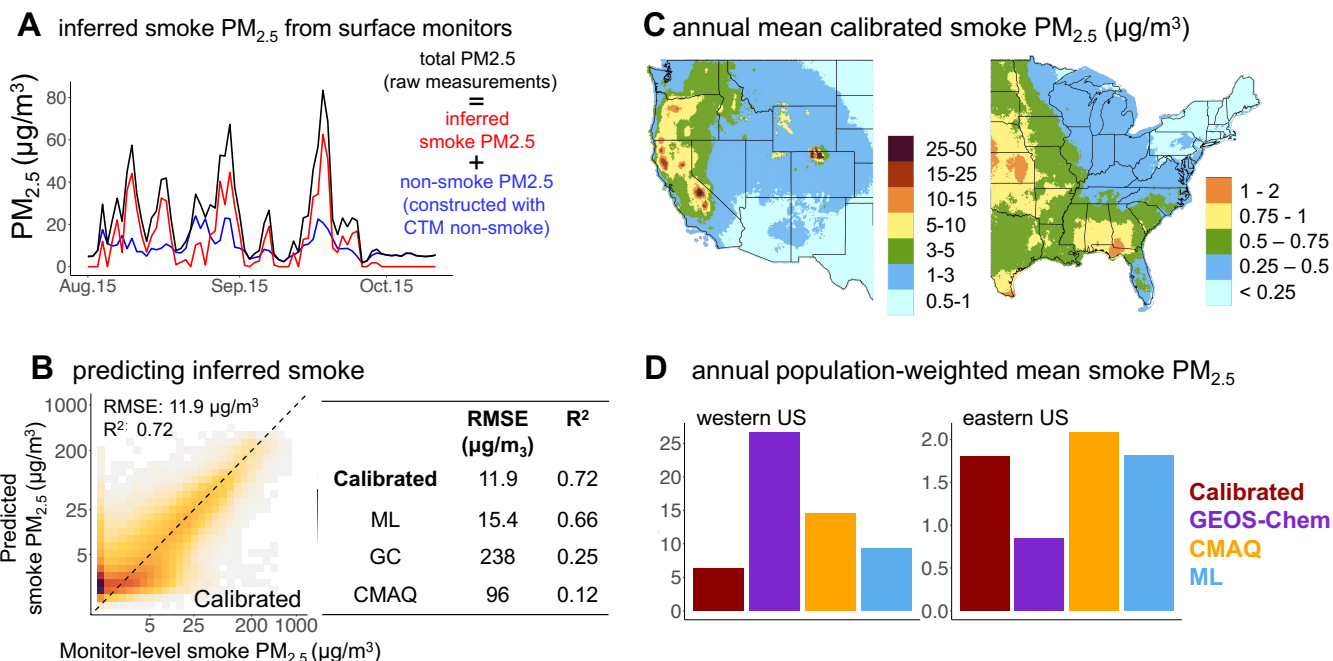


Figure 6: **Creating calibrated smoke $PM_{2.5}$ estimates using three individual smoke estimates.** Panel A shows how we infer smoke $PM_{2.5}$ from surface measurements. The inferred smoke $PM_{2.5}$ (shown in red) is calculated as the difference between the total $PM_{2.5}$ measurements (shown in black) and the non-smoke $PM_{2.5}$ constructed with non-smoke $PM_{2.5}$ estimates from CTMs and other co-variates (shown in blue). Panel B shows the performance of our calibration model. We use an XGBoost model to predict the inferred smoke $PM_{2.5}$ at the monitor level using the three wildfire smoke estimates and other features. The performance of the calibrated model is evaluated using a spatial 5-fold CV. Panel C shows the annual mean calibrated smoke $PM_{2.5}$ concentrations (from Feb-Oct). Note the different scales in the western and eastern US. Panel D shows the annual population-weighted mean smoke $PM_{2.5}$ estimates using the calibrated model and three individual estimates.

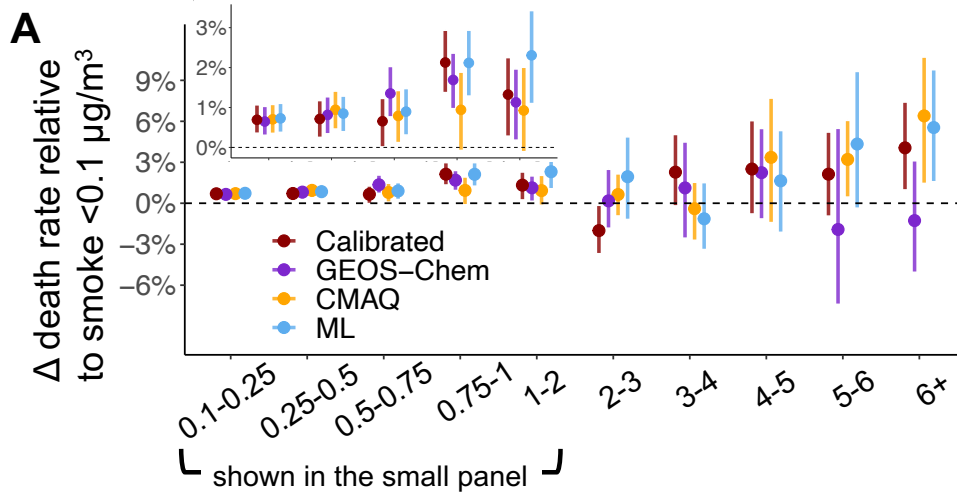
369 in Figure 5B, we find that in the locations where the ML method performs the best, ML methods
 370 significantly outperform CTM approaches (RMSE of CTM approach is 200-500% higher than the
 371 ML method). In the locations where CTM methods perform better, the relative gain in RMSE is
 372 smaller yet meaningful due to overall lower levels of smoke concentrations.

373 Integrating the three wildfire smoke estimates, we construct improved estimates of wildfire smoke
 374 $PM_{2.5}$ (referred to as “calibrated smoke $PM_{2.5}$ ” thereafter) constrained by surface measurements of
 375 total $PM_{2.5}$ and constructed non-smoke $PM_{2.5}$ concentrations (see Figure 6). The calibrated smoke
 376 $PM_{2.5}$ shows a better agreement with the inferred smoke $PM_{2.5}$ from surface monitors compared to
 377 the three individual approaches both over the entire study period (Figure 6B and S4) and during
 378 extreme smoke episodes (Table S1 and Figure S5). Consistent with the spatial pattern shown in

379 Figure 5A, we find that “ML smoke estimate” is the most important feature in the calibration model
380 in the western US while all three estimates contribute to the calibration model in the eastern US
381 (see Figure S6 for feature importance metrics). After calibration, we calculate that the annual
382 average population-weighted smoke $\text{PM}_{2.5}$ concentration is $6.4 \mu\text{g}/\text{m}^3$ in the western US in 2020,
383 lower than all individual estimates. The calibrated estimates are even lower than ML estimates
384 ($7.7 \mu\text{g}/\text{m}^3$) possibly because the calibrated model is also trained to predict measurements from
385 PurpleAir sensors which are often placed in less polluted areas. In the eastern US, the calibrated
386 model estimates an annual average population-weighted smoke $\text{PM}_{2.5}$ concentration of $1.8 \mu\text{g}/\text{m}^3$,
387 which is in the range of the three individual estimates.

388 The different smoke $\text{PM}_{2.5}$ estimates (the three individual approaches and the calibrated) have
389 important implications for empirically estimating dose-response function between smoke and all-
390 cause mortality (Figure 7A), as well as the estimated mortality burden (Figure 7B). Even replacing
391 one year of smoke data from different sources, we find important differences in the empirically
392 estimated dose-response functions. While all dose-response functions show similar effects of smoke
393 on mortality rates at lower annual smoke concentrations (e.g., below $0.5 \mu\text{g}/\text{m}^3$), the effects can be
394 quite different for extreme smoke exposure. This is because 2020 contributes to more observations in
395 the high smoke bin compared to the low smoke bins. We estimate that years with extreme ambient
396 wildfire smoke concentrations ($>6 \mu\text{g}/\text{m}^3$) increase annual mortality rates by 4.0% (95%CI: 1.0%,
397 7.4%) when using the calibrated smoke concentrations. The point estimate is smaller than point
398 estimates derived from using ML (5.4%) or CMAQ smoke estimates (6.6%). However, if we use
399 GEOS-Chem smoke estimates in 2020, we would find a negative (but statistically not significant)
400 association between high annual smoke $\text{PM}_{2.5}$ concentration and mortality rates.

401 The different dose-response functions and estimated wildfire smoke exposures generate different
402 estimates of excess deaths attributable to smoke $\text{PM}_{2.5}$ exposure in 2020, though the estimates are at
403 comparable magnitudes (ranging from 18,200 to 43,957). When using different dose-response func-
404 tions, the differences in estimated mortality can be as much as 2x across methods. The differences
405 are smaller when using the same dose-response function, likely due to the fact that dose-response
406 functions are estimated with binned smoke exposures and thus we treat all years with annual mean
407 smoke above $6 \mu\text{g}/\text{m}^3$ as the same. When using a linear dose-response function (which is used to
408 estimate same-day mortalities), the differences across methods can be as much as 3x driven by the
409 differences in estimated smoke concentrations.



B

Outcome	Model	Mortality estimates
Annual deaths	CMAQ	35,588 (13,166 56,439)
Different DRFs	GEOS-Chem	18,200 (2,456 33,300)
	ML	32,292 (17,498 46,865)
	Calibrated	23,224 (10,316 36,897)
Annual deaths	CMAQ	43,957 (25,551 61,991)
Same DRF (from ML)	GEOS-Chem	31,628 (16,107 47,672)
	ML	32,292 (17,498 46,865)
	Calibrated	30,330 (16,724 44,075)
Daily short-term deaths	CMAQ	5,296 (353 9,886)
DRF from Gould et al., 2023	GEOS-Chem	7,375 (492 13,767)
	ML	3,283 (219 6,128)
	Calibrated	2,818 (188 5,261)

Figure 7: **Empirically estimated impacts of wildfire smoke $\text{PM}_{2.5}$ on all-cause mortality rates and attributed mortality burden across different smoke estimates.** Panel A shows the effects of exposure to different annual mean concentration of smoke $\text{PM}_{2.5}$ (shown in the x-axis) relative to a year with smoke concentration $<0.1 \mu\text{g}/\text{m}^3$, estimated using a Poisson model at the county and annual level using data from 2006–2020. The wildfire smoke data is the same across estimations in 2006–2019 (derived from the ML method), but is different in 2020 depending on the estimation approach. Panel B shows the estimated excess deaths due to smoke $\text{PM}_{2.5}$ in 2020 using different dose-response functions (DRF) and smoke $\text{PM}_{2.5}$ estimated by different methods. “annual deaths, different DRF” is estimated using the DRFs shown in panel A. “annual deaths, same CRF” is estimated using the same DRF (derived from ML estimates) but different exposure estimates from the four methods. Thus, the difference is only due to different estimates of wildfire smoke $\text{PM}_{2.5}$ concentration. “daily short-term deaths” is estimated using the same DRF across four estimations methods by using the DRF from (19), which estimates the effects of daily wildfire smoke $\text{PM}_{2.5}$ on same-day mortality rates. The error bars and values in the parenthesis show the 95% confidence interval estimated using bootstrapping.

410 Discussion

411 We find substantial discrepancies in estimated wildfire smoke $\text{PM}_{2.5}$ concentrations across three
412 widely-used approaches in the US during 2020, the most extreme smoke year in the recent history
413 of the US but likely to be more prevalent under future climate change. Therefore, our results
414 have implications for estimating future fire smoke exposure and health impacts. Among the three
415 approaches we evaluate, we find that the ML approach performs better in the western US with
416 high-level of smoke, while CTM approaches exhibit large biases due to biased emissions inventory
417 and incorrectly-modeled injection heights. On the other hand, CTM approaches outperform the
418 ML approach in the eastern US, due to the uncertainty of the underlying smoke plume data and
419 the calculated non-smoke baselines used in the ML approach. To address the discrepancies across
420 methods, we develop a calibration approach that uses all three wildfire smoke estimates to generate
421 improved estimates of wildfire smoke that better match surface measurements of air quality. Given
422 the improved performance against surface measurement throughout the US, our improved estimates
423 of smoke $\text{PM}_{2.5}$ are appropriate for downstream health impacts analysis.

424 Consistent with the large literature on environmental health and epidemiology, our research
425 demonstrates that the measurement of environmental exposures matters for estimating downstream
426 health impacts. Perhaps surprisingly, we find that only one year of different exposures in our 15
427 years of data can generate large differences in the estimated dose-response function. In our case,
428 incorrect exposures can lead to suspicious associations between wildfire smoke and mortality (e.g.,
429 GEOS-Chem model results show negative associations between extreme wildfire smoke and mortality
430 rates). Therefore, one important message our work has for health researchers is to use non-validated
431 exposure metrics with extreme caution, and to validate/calibrate them against in-situ measurements
432 before using them for estimating health impacts. We also show that the differences in exposure can
433 lead to larger uncertainties in dose-response functions compared to estimates of health burdens.
434 This occurs because noisy yet unbiased exposure estimates can substantially attenuate estimates of
435 the dose-response function, effectively biasing down the relationship. Conversely, applying a roughly
436 linear dose-response function to various exposures may result in similar levels of health burdens,
437 as noise in the exposure data cancels out. Given the widespread use of empirically-estimated dose-
438 response functions, our research demonstrates the importance of correctly modelling smoke pollution
439 exposure in the first place, and how calibrations of the model estimates can improve the downstream
440 health impact analysis.

441 Despite the uncertainty in wildfire smoke estimates, we find that increasing annual exposures to
442 smoke $\text{PM}_{2.5}$ are associated with higher county-level annual mortality rates across the contiguous
443 US. Using our calibrated smoke estimates, we estimate more than 23,000 deaths attributable to
444 exposure to wildfire smoke in 2020. Our work contributes to a large and growing body of literature
445 documenting the impacts of annual exposures to total $\text{PM}_{2.5}$ on mortality, which has shaped decades

446 of policy to improve ambient air quality in the US. However, wildfires are episodic and typically
447 generate short-term spikes in ambient air pollution (71). In a sensitivity analysis, we show that using
448 short-term daily extremes (i.e., the number of extreme smoke days in a year) to estimate mortality
449 yields similar (but more noisy) estimates to our main estimates using annual average smoke exposure
450 (see Figure S7), likely because annual changes in exposure are driven substantially by increases in
451 short-run extremes. We observe similar discrepancies in the derived dose-response functions that
452 link the number of extreme smoke days to annual mortality across the smoke estimation methods.

453 Our research points to several potential areas for future research to improve our understanding
454 of the health effects of wildfire smoke. A more in-depth comparison based on multiple years of
455 wildfire smoke estimates will help understand the generalizability of our findings and evaluate the
456 influences of smoke estimation methods on health effects beyond 2020. While our research selects
457 three widely-used wildfire smoke estimation methods, we do not fully explore the whole suite of
458 wildfire smoke estimation approaches, such as other CTMs using different emission inventories and
459 injection heights and alternate ML algorithms. Nevertheless, our framework of evaluation and
460 calibration of wildfire smoke estimates can be extended to other wildfire estimates generated by
461 other CTMs or data-driven techniques. Due to the limited temporal resolution of the mortality
462 data, we demonstrate the influence of wildfire smoke estimates on the annual smoke-mortality
463 relationship. Given the established short-term effects of wildfire smoke on health outcomes (19),
464 future research can evaluate how different wildfire smoke estimates can influence short-term health
465 dose-response functions.

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475 **Author contributions**

476 MQ and MB designed the study. MK and XJ performed the GEOS-Chem model simulations.
477 DT provided the CMAQ simulation. SHN collected and calibrated the PurpleAir data. MQ led
478 the evaluation of the smoke products and health impacts analysis, with input from all co-authors.
479 MQ and MB drafted the manuscript with input from all co-authors. All authors contributed to
480 interpretation of results and reviewed the manuscript.

481 **Competing interests**

482 The authors declare no competing interests.

483 **Data and materials availability**

484 Data and code to replicate all results in the main text and supplementary materials will be made
485 available at a public repository, except for county-level mortality data for low-population counties,
486 which are not publicly available and were obtained through application to the National Center for
487 Health Statistics.

488 References

- 489 1. V. Iglesias, J. K. Balch, W. R. Travis, US fires became larger, more frequent, and more
490 widespread in the 2000s. *Science advances* **8**, eabc0020 (2022).
- 491 2. J. T. Abatzoglou, A. P. Williams, Impact of anthropogenic climate change on wildfire across
492 western US forests. *Proceedings of the National Academy of Sciences* **113**, 11770–11775 (2016).
- 493 3. Y. Zhuang, R. Fu, B. D. Santer, R. E. Dickinson, A. Hall, Quantifying contributions of natural
494 variability and anthropogenic forcings on increased fire weather risk over the western United
495 States. *Proceedings of the National Academy of Sciences* **118**, e2111875118 (2021).
- 496 4. N. S. Diffenbaugh, A. G. Konings, C. B. Field, Atmospheric variability contributes to increasing
497 wildfire weather but not as much as global warming. *Proceedings of the National Academy of
498 Sciences* **118**, e2117876118 (2021).
- 499 5. Z. L. Steel, H. D. Safford, J. H. Viers, The fire frequency-severity relationship and the legacy
500 of fire suppression in California forests. *Ecosphere* **6**, 1–23 (2015).
- 501 6. V. C. Radeloff *et al.*, Rising wildfire risk to houses in the United States, especially in grasslands
502 and shrublands. *Science* **382**, 702–707 (2023).
- 503 7. C. D. McClure, D. A. Jaffe, US particulate matter air quality improves except in wildfire-prone
504 areas. *Proceedings of the National Academy of Sciences* **115**, 7901–7906 (2018).
- 505 8. K. O’Dell, B. Ford, E. V. Fischer, J. R. Pierce, Contribution of wildland-fire smoke to US
506 PM_{2.5} and its influence on recent trends. *Environmental science & technology* **53**, 1797–1804
507 (2019).
- 508 9. M. L. Childs *et al.*, Daily Local-Level Estimates of Ambient Wildfire Smoke PM_{2.5} for the
509 Contiguous US. *Environmental Science & Technology* **56**, 13607–13621 (2022).
- 510 10. D. Zhang *et al.*, Wildland Fires Worsened Population Exposure to PM_{2.5} Pollution in the
511 Contiguous United States. *Environmental Science & Technology* **57**, 19990–19998 (2023).
- 512 11. Y. Li *et al.*, Dominance of wildfires impact on air quality exceedances during the 2020 record-
513 breaking wildfire season in the United States. *Geophysical Research Letters* **48**, e2021GL094908
514 (2021).
- 515 12. M. Burke *et al.*, The changing risk and burden of wildfire in the United States. *Proceedings of
516 the National Academy of Sciences* **118**, e2011048118 (2021).
- 517 13. M. Burke *et al.*, The contribution of wildfire to PM_{2.5} trends in the USA. *Nature* **622**, 761–
518 766 (2023).
- 519 14. M. Qiu *et al.*, Wildfire smoke exposure and mortality burden in the US under future climate
520 change. (2024).

- 521 15. B. Ford *et al.*, Future fire impacts on smoke concentrations, visibility, and health in the con-
522 tiguous United States. *GeoHealth* **2**, 229–247 (2018).
- 523 16. J. C. Liu *et al.*, Future respiratory hospital admissions from wildfire smoke under climate
524 change in the Western US. *Environmental Research Letters* **11**, 124018 (2016).
- 525 17. J. E. Neumann *et al.*, Estimating PM_{2.5}-related premature mortality and morbidity associated
526 with future wildfire emissions in the western US. *Environmental Research Letters* **16**, 035019
527 (2021).
- 528 18. C. E. Reid *et al.*, Critical review of health impacts of wildfire smoke exposure. *Environmental*
529 *health perspectives* **124**, 1334–1343 (2016).
- 530 19. C. F. Gould *et al.*, Health effects of wildfire smoke exposure. *Annual Review of Medicine* **75**
531 (2023).
- 532 20. X. Ye *et al.*, Evaluation and intercomparison of wildfire smoke forecasts from multiple modeling
533 systems for the 2019 Williams Flats fire. *Atmospheric Chemistry and Physics* **21**, 14427–14469
534 (2021).
- 535 21. P. Makkaroon *et al.*, Development and evaluation of a North America ensemble wildfire air
536 quality forecast: Initial application to the 2020 Western United States “Gigafire”. *Journal of*
537 *Geophysical Research: Atmospheres* **128**, e2022JD037298 (2023).
- 538 22. C. Black, Y. Tesfaigzi, J. A. Bassein, L. A. Miller, Wildfire smoke exposure and human health:
539 Significant gaps in research for a growing public health issue. *Environmental toxicology and*
540 *pharmacology* **55**, 186–195 (2017).
- 541 23. R. W. Gan *et al.*, Comparison of wildfire smoke estimation methods and associations with
542 cardiopulmonary-related hospital admissions. *GeoHealth* **1**, 122–136 (2017).
- 543 24. V. A. Williams *et al.*, Impact of wildfires on cardiovascular health. *Circulation research* **134**,
544 1061–1082 (2024).
- 545 25. S. Rajagopalan, S. G. Al-Kindi, R. D. Brook, Air pollution and cardiovascular disease: JACC
546 state-of-the-art review. *Journal of the American College of Cardiology* **72**, 2054–2070 (2018).
- 547 26. G. Chen *et al.*, Mortality risk attributable to wildfire-related PM_{2.5} pollution: a global time
548 series study in 749 locations. *The Lancet Planetary Health* **5**, e579–e587 (2021).
- 549 27. A. Heaney *et al.*, Impacts of fine particulate matter from wildfire smoke on respiratory and
550 cardiovascular health in California. *GeoHealth* **6**, e2021GH000578 (2022).
- 551 28. T. Ye *et al.*, Short-term exposure to wildfire-related PM_{2.5} increases mortality risks and
552 burdens in Brazil. *Nature Communications* **13**, 7651 (2022).
- 553 29. S. Pan *et al.*, Quantifying the premature mortality and economic loss from wildfire-induced
554 PM_{2.5} in the contiguous US. *Science of The Total Environment* **875**, 162614 (2023).

- 555 30. Y.-Y. Meng *et al.*, Health and economic cost estimates of short-term total and wildfire PM_{2.5}
556 exposure on work loss: using the consecutive California Health Interview Survey (CHIS) data
557 2015–2018. *BMJ Public Health* **2** (2024).
- 558 31. X. Pan *et al.*, Six global biomass burning emission datasets: intercomparison and application
559 in one global aerosol model. *Atmospheric Chemistry and Physics* **20**, 969–994 (2020).
- 560 32. T. S. Carter *et al.*, How emissions uncertainty influences the distribution and radiative impacts
561 of smoke from fires in North America. *Atmospheric Chemistry and Physics* **20**, 2073–2097
562 (2020).
- 563 33. Y. Li *et al.*, Impacts of estimated plume rise on PM_{2.5} exceedance prediction during extreme
564 wildfire events: a comparison of three schemes (Briggs, Freitas, and Sofiev). *Atmospheric Chem-
565 istry and Physics* **23**, 3083–3101 (2023).
- 566 34. X. Huang *et al.*, Smoke-weather interaction affects extreme wildfires in diverse coastal regions.
567 *Science* **379**, 457–461 (2023).
- 568 35. Y. Xie, M. Lin, L. W. Horowitz, Summer PM_{2.5} pollution extremes caused by wildfires over
569 the western United States during 2017–2018. *Geophysical Research Letters* **47**, e2020GL089429
570 (2020).
- 571 36. X. Feng *et al.*, Improved estimates of smoke exposure during Australia fire seasons: importance
572 of quantifying plume injection heights. *Atmospheric Chemistry and Physics* **24**, 2985–3007
573 (2024).
- 574 37. J. C. Liu *et al.*, Wildfire-specific fine particulate matter and risk of hospital admissions in
575 urban and rural counties. *Epidemiology* **28**, 77–85 (2017).
- 576 38. C. E. Reid, E. M. Considine, M. M. Maestas, G. Li, Daily PM_{2.5} concentration estimates
577 by county, ZIP code, and census tract in 11 western states 2008–2018. *Scientific data* **8**, 112
578 (2021).
- 579 39. W. Lassman *et al.*, Spatial and temporal estimates of population exposure to wildfire smoke
580 during the Washington state 2012 wildfire season using blended model, satellite, and in situ
581 data. *GeoHealth* **1**, 106–121 (2017).
- 582 40. G. Geng *et al.*, Satellite-based daily PM_{2.5} estimates during fire seasons in Colorado. *Journal
583 of Geophysical Research: Atmospheres* **123**, 8159–8171 (2018).
- 584 41. R. Aguilera *et al.*, A novel ensemble-based statistical approach to estimate daily wildfire-
585 specific PM_{2.5} in California (2006–2020). *Environment international* **171**, 107719 (2023).
- 586 42. Y. Ma *et al.*, Wildfire smoke PM_{2.5} and mortality in the contiguous United States. *medRxiv*
587 (2023).

- 588 43. S. Heft-Neal *et al.*, Emergency department visits respond nonlinearly to wildfire smoke. *Proceedings of the National Academy of Sciences* **120**, e2302409120 (2023).
589
- 590 44. C. Chen, L. Schwarz, N. Rosenthal, M. E. Marlier, T. Benmarhnia, Exploring spatial hetero-
591 geneity in synergistic effects of compound climate hazards: Extreme heat and wildfire smoke
592 on cardiorespiratory hospitalizations in California. *Science Advances* **10**, eadj7264 (2024).
- 593 45. Q. Zhu *et al.*, Wildfires are associated with increased emergency department visits for anxiety
594 disorders in the western United States. *Nature Mental Health*, 1–9 (2024).
- 595 46. A. L. Johnson, M. J. Abramson, M. Dennekamp, G. J. Williamson, Y. Guo, Particulate matter
596 modelling techniques for epidemiological studies of open biomass fire smoke exposure: a review.
597 *Air Quality, Atmosphere & Health* **13**, 35–75 (2020).
- 598 47. X. Jiang, E.-H. Enki Yoo, Modeling wildland fire-specific PM_{2.5} concentrations for uncertainty-
599 aware health impact assessments. *Environmental science & technology* **53**, 11828–11839 (2019).
- 600 48. W. Yuchi *et al.*, Blending forest fire smoke forecasts with observed data can improve their
601 utility for public health applications. *Atmospheric Environment* **145**, 308–317 (2016).
- 602 49. E. M. Considine *et al.*, Evaluation of model-based PM_{2.5} estimates for exposure assessment
603 during wildfire smoke episodes in the western US. *Environmental Science & Technology* **57**,
604 2031–2041 (2023).
- 605 50. G. R. Van Der Werf *et al.*, Global fire emissions estimates during 1997–2016. *Earth System
606 Science Data* **9**, 697–720 (2017).
- 607 51. I. Bey *et al.*, Global modeling of tropospheric chemistry with assimilated meteorology: Model
608 description and evaluation. *Journal of Geophysical Research: Atmospheres* **106**, 23073–23095
609 (2001).
- 610 52. S. Akagi *et al.*, Emission factors for open and domestic biomass burning for use in atmospheric
611 models. *Atmospheric Chemistry and Physics* **11**, 4039–4072 (2011).
- 612 53. M. Mu *et al.*, Daily and 3-hourly variability in global fire emissions and consequences for atmo-
613 spheric model predictions of carbon monoxide. *Journal of Geophysical Research: Atmospheres*
614 **116** (2011).
- 615 54. E. V. Fischer *et al.*, Atmospheric peroxyacetyl nitrate (PAN): a global budget and source
616 attribution. English, *Atmospheric Chemistry and Physics* **14**, 2679–2698 (2014).
- 617 55. J. W. Kaiser *et al.*, Biomass burning emissions estimated with a global fire assimilation system
618 based on observed fire radiative power. English, *Biogeosciences* **9**, 527–554 (2012).
- 619 56. S. Rémy *et al.*, Two global data sets of daily fire emission injection heights since 2003. *Atmo-
620 spheric Chemistry and Physics* **17**, 2921–2942 (May 2017).

- 621 57. X Zhang *et al.*, *The blended global biomass burning emissions product from MODIS and VIIRS*
622 *observations (GBBEPx) version 3.1*, 2019.
- 623 58. M Sofiev, T Ermakova, R Vankevich, Evaluation of the smoke-injection height from wild-land
624 fires using remote-sensing data. *Atmospheric Chemistry and Physics* **12**, 1995–2006 (2012).
- 625 59. Y Li *et al.*, Ensemble PM_{2.5} forecasting during the 2018 camp fire event using the HYS-
626 PLIT transport and dispersion model. *Journal of Geophysical Research: Atmospheres* **125**,
627 e2020JD032768 (2020).
- 628 60. U.S. Environmental Protection Agency, *Air Data: Air Quality Data Collected at Outdoor Mon-*
629 *itors Across the US*, 2023.
- 630 61. Jenny L. Hand, *IMPROVE DATA USER GUIDE 2023 (VERSION 2)*, 2023.
- 631 62. M. Burke *et al.*, Exposures and behavioural responses to wildfire smoke. *Nature human be-*
632 *haviour* **6**, 1351–1361 (2022).
- 633 63. D. A. Jaffe *et al.*, Wildfire and prescribed burning impacts on air quality in the United States.
634 *Journal of the Air & Waste Management Association* **70**, 583–615 (2020).
- 635 64. K. K. Barkjohn, A. L. Holder, S. G. Frederick, A. L. Clements, Correction and accuracy of
636 PurpleAir PM_{2.5} measurements for extreme wildfire smoke. *Sensors* **22**, 9669 (2022).
- 637 65. W Schroeder *et al.*, Validation analyses of an operational fire monitoring product: The Hazard
638 Mapping System. *International Journal of Remote Sensing* **29**, 6059–6066 (2008).
- 639 66. S. E. Cleland *et al.*, Estimating wildfire smoke concentrations during the October 2017 Cali-
640 fornia fires through BME space/time data fusion of observed, modeled, and satellite-derived
641 PM_{2.5}. *Environmental science & technology* **54**, 13439–13447 (2020).
- 642 67. T. Liu *et al.*, Is the smoke aloft? Caveats regarding the use of the Hazard Mapping System
643 (HMS) smoke product as a proxy for surface smoke presence across the United States. (2023).
- 644 68. National Center for Health Statistics Division of Vital Statistics, *NVSS Restricted-use Mor-*
645 *tality Files, 1999-2021. Hyattsville, Maryland*.
- 646 69. T. Liu *et al.*, Diagnosing spatial biases and uncertainties in global fire emissions inventories:
647 Indonesia as regional case study. *Remote Sensing of Environment* **237**, 111557 (2020).
- 648 70. P. S. Kim *et al.*, Sources, seasonality, and trends of southeast US aerosol: an integrated analysis
649 of surface, aircraft, and satellite observations with the GEOS-Chem chemical transport model.
650 *Atmospheric Chemistry and Physics* **15**, 10411–10433 (2015).
- 651 71. J. A. Casey *et al.*, Measuring long-term exposure to wildfire PM_{2.5} in California: Time-
652 varying inequities in environmental burden. *Proceedings of the National Academy of Sciences*
653 **121**, e2306729121 (2024).

654 **Supplementary tables**

Table S1: RMSE between the inferred smoke $PM_{2.5}$ at the monitor level and different fire smoke $PM_{2.5}$ estimates. The table compares the inferred smoke $PM_{2.5}$ concentrations with smoke estimated by two calibration models and three individual methods. “Calibration” is our main calibration model that uses all surface measurements, and “Calibration-extreme” up-weights the extreme smoke days for a better representation of the extreme conditions in the calibration process. The table shows the results for the western US and eastern US, respectively, and for extreme and non-extreme smoke conditions. Extreme smoke day are defined as days with smoke $PM_{2.5}$ over $50 \mu\text{g}/\text{m}^3$ in the western US or $15 \mu\text{g}/\text{m}^3$ in the eastern US. The RMSE is calculated with spatial 5-fold CV for the calibration models.

	Western US		Eastern US	
	(longitude < -100)		(longitude > -100)	
	non-extreme	extreme	non-extreme	extreme
Calibration	7.3	64	2.2	37
Calibration-extreme	11	61	5.3	35
GEOS-Chem	140	1313	2.7	38
CMAQ	81	445	3.3	37
ML	12	73	3.0	37

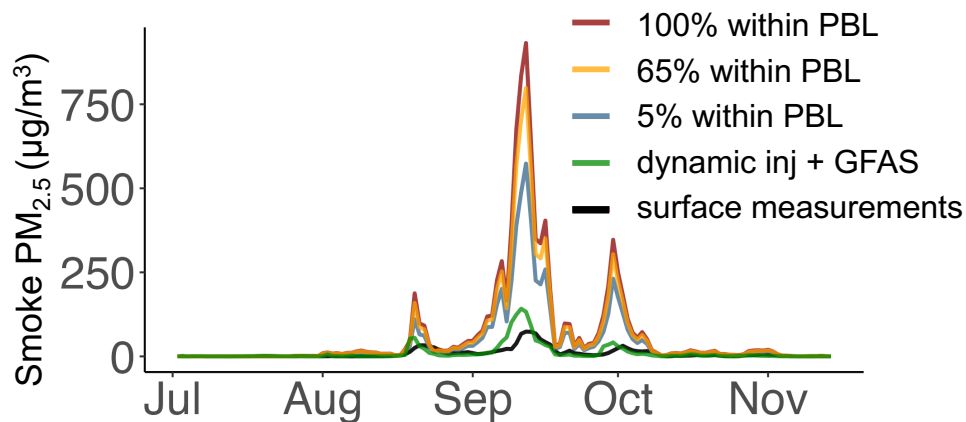


Figure S1: Estimated wildfire smoke PM_{2.5} in Oregon and California by GEOS-Chem, under alternate injection heights and emission inventories. The plot shows the daily smoke PM_{2.5} concentrations averaged over all monitor locations in Oregon and California. The black line shows the smoke PM_{2.5} derived from surface measurements. In the main analysis, our GEOS-Chem simulation uses GFED4s emission inventory and assumes 100% of the emission occurred within Planetary Boundary Layer. The plot shows the simulations from three sensitivity scenarios: GFED4s emissions with alternative injection heights (65% or 5% emissions occurred within PBL), and a scenario using dynamic injection height and GFAS emission inventory (see Method).

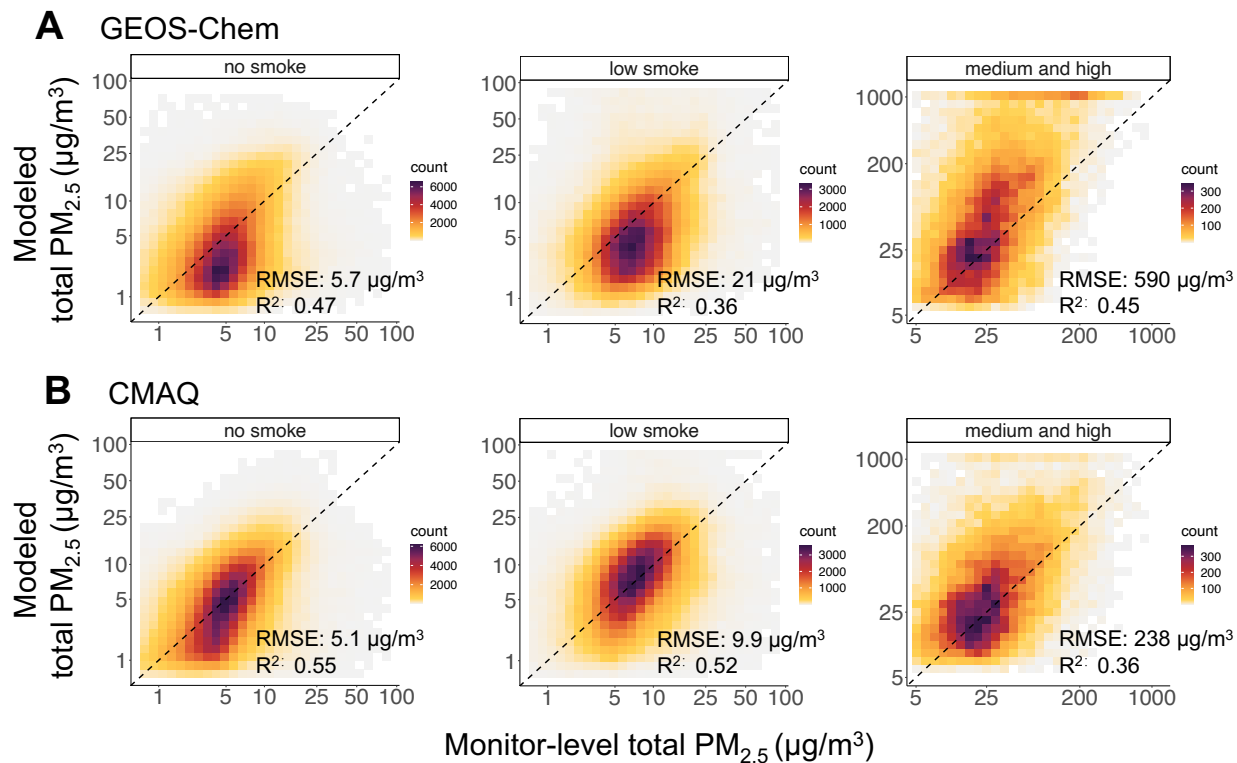


Figure S2: Simulated total $\text{PM}_{2.5}$ concentrations compared against surface measurements across different smoke conditions. “No smoke” includes all monitor days with no smoke plume overhead and low estimated smoke concentrations from both CTM ($<0.5 \mu\text{g}/\text{m}^3$). “Medium and high smoke” includes monitor days with high estimated smoke $\text{PM}_{2.5}$ from both CTMs ($>5 \mu\text{g}/\text{m}^3$). “Low smoke” category includes all the other conditions.

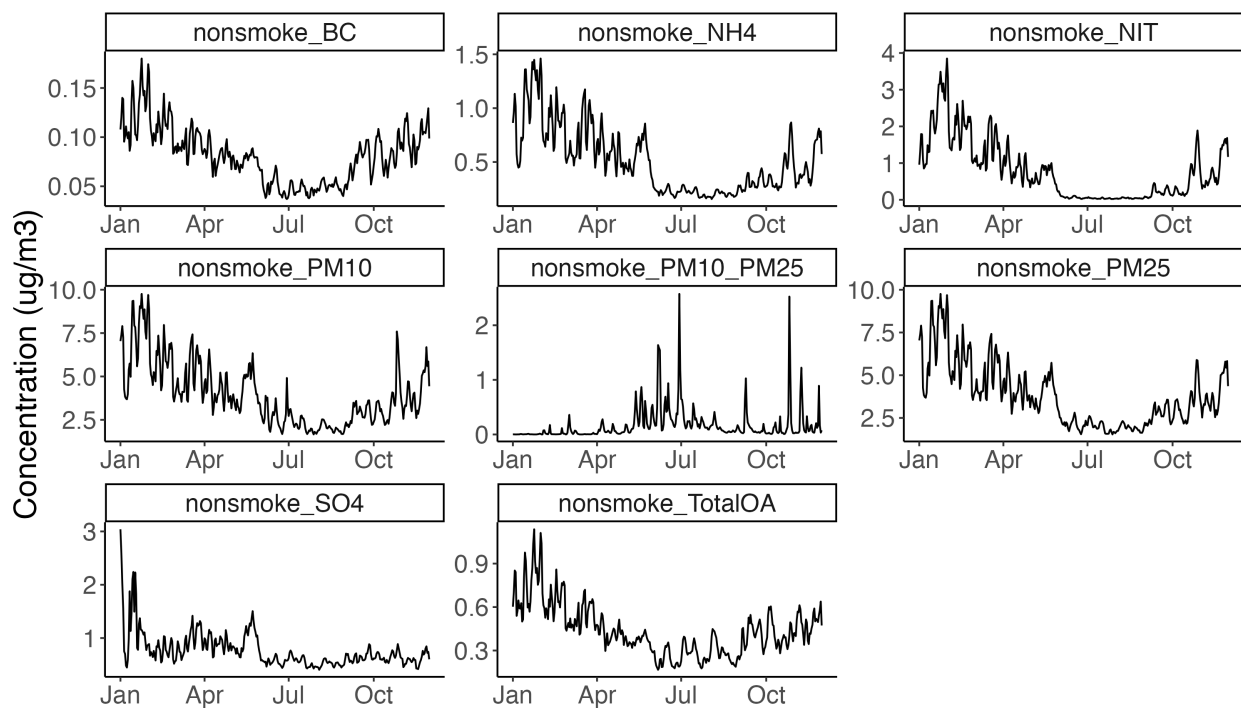


Figure S3: Concentrations of $PM_{2.5}$ components, PM_{10} , and coarse PM ($PM_{10} - PM_{2.5}$) simulated by GEOS-Chem under the no-wildfire scenario. The plot shows the concentration averaged over all US monitor location for black carbon (BC), ammonium (NH_4), nitrate (NIT), PM_{10} , coarse PM ($PM_{10} - PM_{2.5}$), $PM_{2.5}$, sulfate (SO_4), and total organic aerosol (TotalOA).

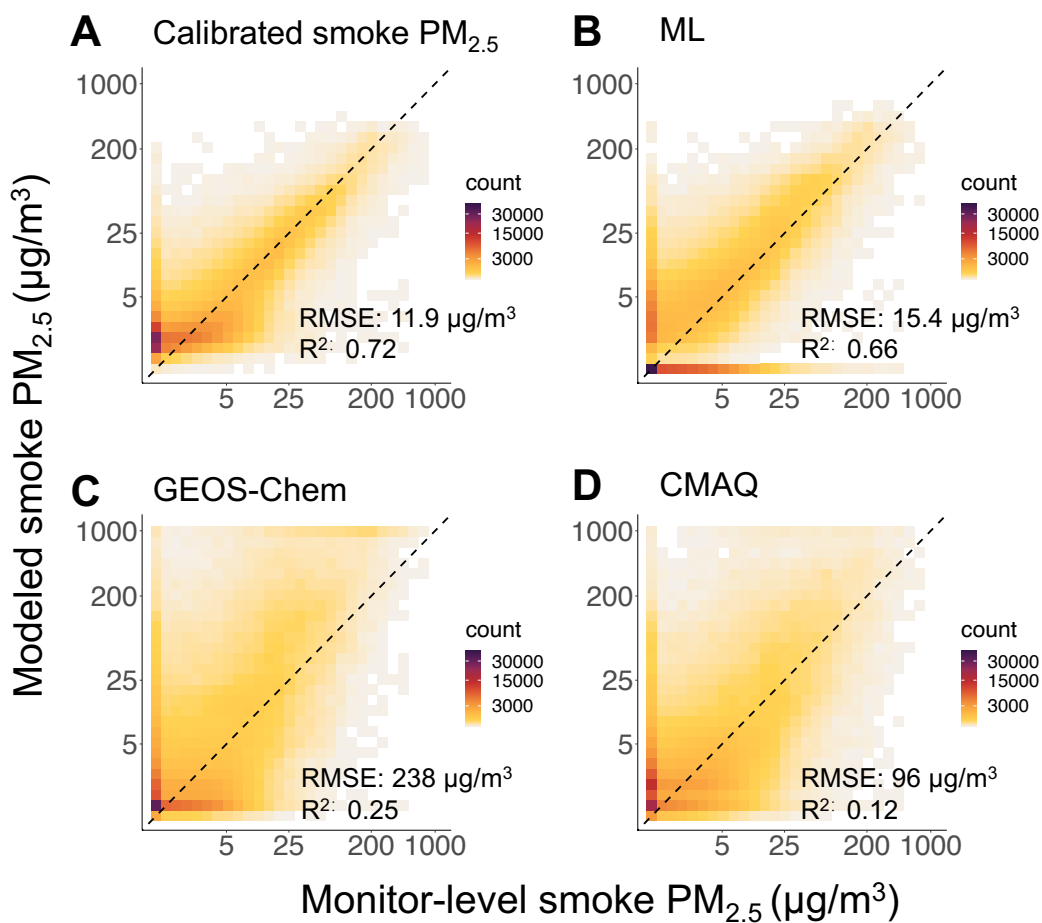


Figure S4: Estimated smoke $PM_{2.5}$ (y-axis) compared against the inferred smoke $PM_{2.5}$ concentrations at the surface monitors (x-axis). Panel A shows the calibrated smoke $PM_{2.5}$ derived from the machine learning model that uses three individual smoke $PM_{2.5}$ simulated by CTMs to predict surface $PM_{2.5}$ under smoke days. Panel B-D show the non-smoke $PM_{2.5}$ concentrations simulated by ML, GEOS-Chem, and CMAQ.

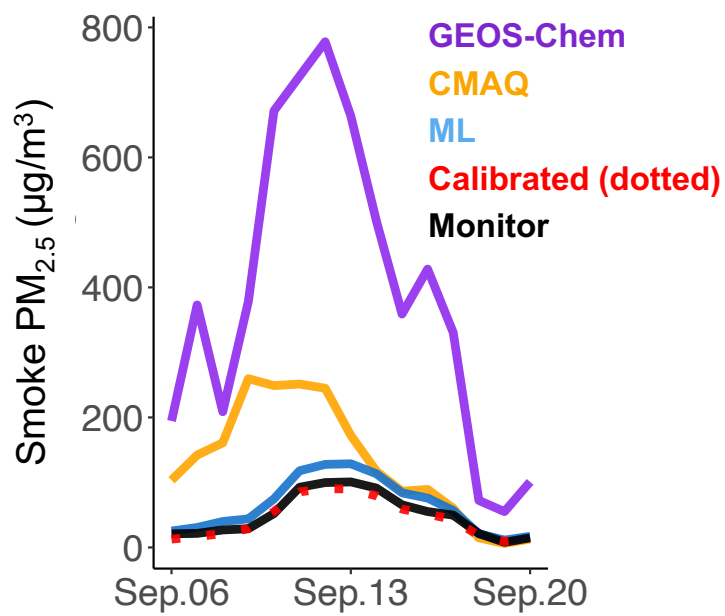


Figure S5: The plot shows the average smoke PM_{2.5} concentrations estimated by different methods over all monitor locations in the western US during September 2020. The spatial range of the monitors is shown in Figure 3.

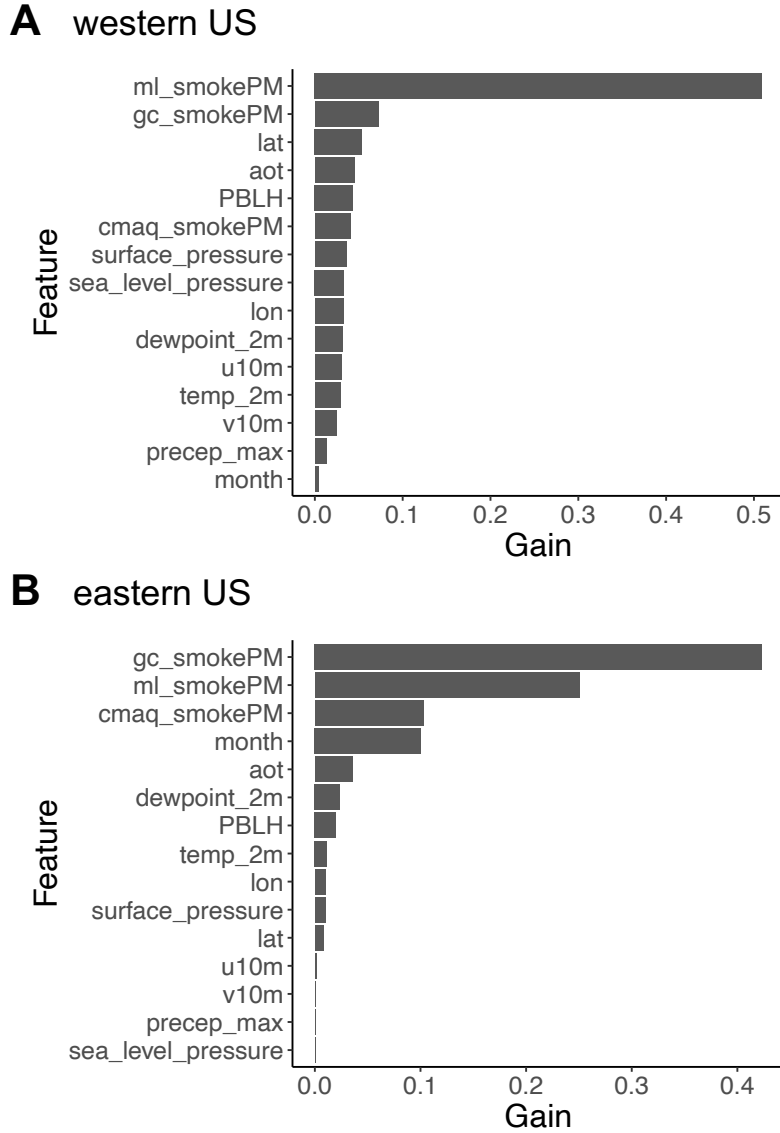
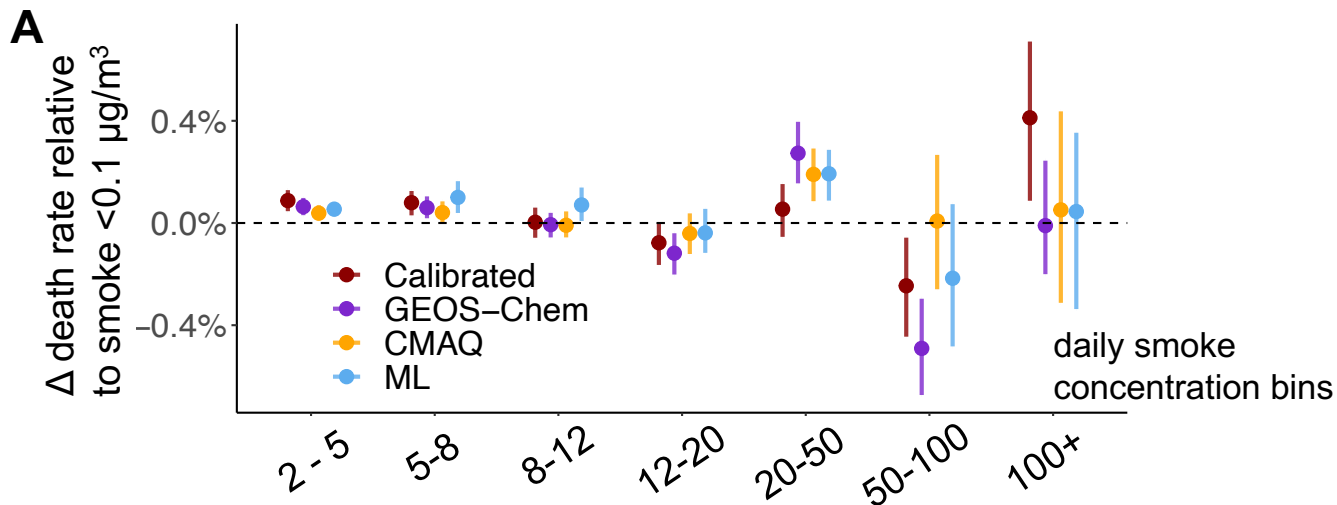


Figure S6: Feature importance of the XGBoost model to calibrate the smoke $PM_{2.5}$ estimates in the western US (Panel A) and the eastern US (Panel B). The y-axis shows the normalized gain of each feature in the model, defined as the average gain in RMSE due to splits on this variable. The sum of gain from all features is one.



B Using dose-response function of # of days in smoke concentration bins

Outcome	Model	Mortality estimates
Annual deaths	CMAQ	37,658
Different DRFs	GEOS-Chem	16,599
	ML	55,213
	Calibrated	69,157
Annual deaths	CMAQ	58,254
Same DRF (from ML)	GEOS-Chem	27,301
	ML	55,213
	Calibrated	50,606

Figure S7: Empirically estimated dose-response functions using the number of days that fall in each smoke concentration bin in each year. Panel A shows the effects of exposure to one day of smoke $\text{PM}_{2.5}$ with different daily smoke concentrations (shown in the x-axis) relative to a day with smoke concentration of $< 2 \mu\text{g}/\text{m}^3$, estimated using a Poisson model at the county and annual level using data from 2006–2020. The wildfire smoke data is the same across estimations in 2006–2019 (derived from the ML method), but is different in 2020 depending on the estimation approach. Panel B shows the estimated excess mortality due to smoke $\text{PM}_{2.5}$ in 2020 using different dose-response functions (DRF) and smoke $\text{PM}_{2.5}$ estimated by different methods. “annual deaths, different DRF” is estimated using the different DRFs shown in panel A. “annual deaths, same CRF” is estimated using the same DRF (derived from ML estimates) but different exposure estimates from four methods. Thus, the difference is only due to different wildfire smoke estimations. The error bars show the 95% confidence interval estimated using bootstrapping.