

# The influence of wildfire smoke on ambient chemical species concentrations in the contiguous US

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## 22 **Abstract**

23 Wildfires are a significant contributor to ambient air pollution and pose a growing public health  
24 threat in many parts of the world. Increased wildfire activity over the past few decades has ex-  
25 acerbated smoke exposure across the US, yet our understanding of how wildfire influences spe-  
26 cific chemicals and their resulting concentration in smoke remains incomplete. We combine 15  
27 years of daily measures of species-specific  $PM_{2.5}$  concentrations from 700 air pollution monitors  
28 with satellite-derived estimates of ambient wildfire smoke  $PM_{2.5}$ , and use panel regression to es-  
29 timate the contribution of wildfire smoke to the concentrations of 27 different chemical species in  
30  $PM_{2.5}$ . We find that wildfire smoke drives detectable increases in the concentration of 25 of the  
31 27 species, with the largest increases observed for chemicals previously associated with biomass  
32 burning: organic carbon, elemental carbon, and potassium. We find that smoke originating from  
33 wildfires that burned structures had higher concentrations of copper, lead, zinc and nickel relative  
34 to smoke from fires that did not burn structures. Wildfire smoke is responsible for an increasing  
35 share of ambient species concentrations for multiple species, especially in the Western US. Using  
36 existing estimated relationships between ambient chemical exposure and cancer risk, we find that  
37 wildfire enhancement of carcinogenic species concentrations could be enough to cause small in-  
38 creases in cancer risk, but these increases are very small relative to other risk factors. Our results  
39 demonstrate that fixed ground monitors in combination with satellite-derived data can be used to  
40 understand how wildfire smoke influences chemical concentrations at large scales and measure  
41 population-level exposures.

## 42 **Significance Statement**

43 Wildfire risk is growing in the US and many other parts of the world, with demonstrable im-  
44 pact on surface air quality. We use daily measurements from 700 monitors around the US over  
45 15 years to characterize the chemical species present in wildfire smoke, quantify how chang-  
46 ing wildfire smoke concentrations are affecting observed species concentrations, and estimate  
47 health risks from exposure. We find that increases in wildfire smoke lead to measurable increases  
48 in over two dozen chemical species' concentrations and that this influence has grown over time  
49 for many species. We show that the burning of structures significantly elevates specific chemi-  
50 cal concentrations in smoke and that wildfire enhancement of carcinogenic species could lead to  
51 small increases in population cancer risk.

## 52 **Introduction**

53 Wildfires are a major source of ambient air pollution and represent a growing environmental  
54 health hazard in the United States (US) and globally. Over the last few decades, increased wildfire  
55 activity has exacerbated smoke exposure across the US and begun to reverse decades of hard-won  
56 progress in improving air quality<sup>1</sup>. Studies suggest that, in recent years, wildfire smoke has con-  
57 tributed 25% of all ambient fine particulate matter (PM<sub>2.5</sub>) in the US<sup>2,3</sup>, and substantially more  
58 in some areas of the Western US. This contribution is projected to grow in the coming years due  
59 to higher wildfire risks under a warming climate<sup>4</sup>, the expansion of at-risk areas from a grow-  
60 ing wildland-urban interface (WUI)<sup>5,6</sup>, and increasing policy stringency around reducing other  
61 sources of ambient PM<sub>2.5</sub> pollution<sup>7</sup>.

62 A wide body of epidemiological research has demonstrated that exposure to ambient PM<sub>2.5</sub> -  
63 from smoke or otherwise - is associated with a number of negative health impacts that vary in  
64 severity and timing of onset from exposure<sup>8-12</sup>. Documented impacts of exposure to wildfire  
65 smoke in particular include adverse birth outcomes<sup>13,14</sup>, respiratory exacerbations<sup>12,15</sup>, cognitive  
66 impairment<sup>16</sup>, emergency department visits and hospital admissions<sup>12,17-19</sup>, and cardiovascular  
67 mortality and morbidity among others<sup>20,21</sup>. While most previous epidemiological research has  
68 focused on exposure to total PM<sub>2.5</sub> pollution, toxicological evidence indicates that exposures to  
69 specific chemical constituents of PM<sub>2.5</sub> may shape how specific health outcomes manifest and  
70 evolve<sup>22</sup>. Limited epidemiological evidence also supports species-specific impacts<sup>23-26</sup> and  
71 explicitly argues for the consideration of species composition in studies of the health effects of  
72 PM<sub>2.5</sub><sup>27,28</sup>.

73 While existing evidence documents the effects of wildfire smoke on total PM<sub>2.5</sub> concentrations, a  
74 comprehensive understanding of the types and amount of different chemical species in smoke and  
75 the corresponding population exposure remains elusive. Most previous research on the chemical  
76 composition of wildfire smoke PM<sub>2.5</sub> comes from the collection of air samples during controlled  
77 burns in laboratory settings or at field locations near specific active wildfires<sup>29-34</sup>. These studies  
78 find that wildfire smoke is a complex mixture of gases (e.g., carbon monoxide, carbon dioxide,  
79 nitrogen oxides) and particles (e.g., trace metals, elemental carbon, organic compounds), and that  
80 the mixture of gases and particles emitted by wildfires depends on the materials burned, size of  
81 the fire, combustion efficiency, distance the smoke has traveled, and environmental conditions  
82 like wind speed, temperature, and humidity<sup>35</sup>. Some recent observational studies use air quality  
83 monitoring station data across broader spatial and temporal scales and find that wildfire smoke  
84 increases the fractions of elemental and organic carbon during “smoke waves” in the western  
85 US<sup>36</sup> and elevates concentrations of trace metals (copper and lead) on “smoke-impacted” days

86 from specific fires in California<sup>37</sup>. While these studies offer important insights into the composi-  
87 tion and potential toxicity of wildfire smoke in specific settings, they are often limited in spatial  
88 scale (e.g., one sampling location, one region) and thus may not be representative of the types  
89 of exposures that a population experiences. In turn, they can neither be used to characterize how  
90 overall temporal and spatial variation in wildfire smoke exposure is altering species concentra-  
91 tions, nor to assess the potential population health burden from these alterations.

92 Here, we characterize the chemical composition of wildfire smoke  $PM_{2.5}$  at a broad temporal  
93 and spatial scale by combining 15 years of daily measurements of species-specific  $PM_{2.5}$  con-  
94 centrations from 700 air pollution monitors across the continental US, with satellite-derived con-  
95 centrations of ambient wildfire smoke  $PM_{2.5}$ <sup>2</sup> from 2006 to 2020 (**Figure 1**). We estimate the  
96 impact of variation in wildfire smoke on the concentrations of 27 different species of  $PM_{2.5}$  us-  
97 ing a panel regression approach that measures how a one unit change in smoke  $PM_{2.5}$  affects the  
98 mass concentrations of different chemical species (**Materials & Methods**). Our approach relates  
99 daily measurements of species concentrations at each monitor to daily variation in  $PM_{2.5}$  from  
100 wildfire smoke, using high-dimensional fixed effects (dummy variables) to isolate the impact  
101 of wildfire  $PM_{2.5}$  from a range of other time-invariant, seasonal, and time-varying factors that  
102 could also contribute to variation in species concentrations. We estimate both pooled national  
103 models as well as region-specific models to quantify how regional factors might influence smoke-  
104 species relationships. Building on earlier work that links observed smoke  $PM_{2.5}$  concentrations  
105 to specific source fires<sup>38</sup>, we study whether characteristics of those source fires – in particular,  
106 whether structures were burned in the fire – influence measured species concentrations in the re-  
107 sulting smoke. We then combine our statistical estimates of the smoke-species relationships for  
108 each species with estimates of nationwide trends in smoke  $PM_{2.5}$  and calculate the overall change  
109 of species concentrations due to increasing wildfire  $PM_{2.5}$  around the country. Finally, we inte-  
110 grate our predicted changes in species concentrations due to wildfire smoke with established rela-  
111 tionships between ambient chemical exposure and cancer risks to estimate the increase in cancer  
112 cases resulting from wildfire-induced elevations in ambient species concentrations<sup>39,40</sup>.

## 113 **Results**

### 114 **Chemical composition of wildfire smoke $PM_{2.5}$**

115 We find that wildfire smoke significantly increases the concentration of the majority (25 of 27)  
116 of chemical species measured by US EPA’s CSN and IMPROVE monitoring networks in our  
117 sample (**Supplementary Table 1**). The largest increases in concentration relative to an average

118 non-smoke day concentration in our sample (i.e., background species-specific ambient pollu-  
119 tion) are for organic carbon (OC), elemental carbon (EC), potassium (K), and phosphorus (P)  
120 (**Figure 2**). An additional  $1 \mu\text{g}/\text{m}^3$  of wildfire smoke  $\text{PM}_{2.5}$  significantly increases the concen-  
121 trations of OC by  $0.39 \mu\text{g}/\text{m}^3$  (39% increase relative to  $0.99 \mu\text{g}/\text{m}^3$  background, 95% confidence  
122 interval (CI):  $0.36$  to  $0.41 \mu\text{g}/\text{m}^3$ ); EC by  $0.06 \mu\text{g}/\text{m}^3$  (25% increase relative to  $0.25 \mu\text{g}/\text{m}^3$  back-  
123 ground, 95% CI:  $0.05$  to  $0.06 \mu\text{g}/\text{m}^3$ ); K by  $0.007 \mu\text{g}/\text{m}^3$  (14.3% increase relative to  $0.05 \mu\text{g}/\text{m}^3$   
124 background, 95% CI:  $0.006$  to  $0.007 \mu\text{g}/\text{m}^3$ ); and P by  $0.00005 \mu\text{g}/\text{m}^3$  (7.3% increase relative to  
125  $0.0007 \mu\text{g}/\text{m}^3$  background, 95% CI:  $0.00004$  to  $0.00006 \mu\text{g}/\text{m}^3$ ). Findings are robust to alterna-  
126 tive model specifications, covariate inclusion/exclusion, and an alternative measure of monitor-  
127 captured total  $\text{PM}_{2.5}$  concentration (**Supplementary Figure 1**).

128 While the composition and concentration of chemicals can significantly vary based on what is  
129 burning, the temperature of the burn, and surrounding atmospheric conditions<sup>41,42</sup>, studies have  
130 consistently shown that certain species are disproportionately likely to be present in wildfire  
131 smoke. For instance, K is emitted from biomass burning even with varying fuel type and burn  
132 conditions<sup>43,44</sup>, and elevated OC and EC have been associated with increasing wildfire activity  
133 as well as biomass combustion from prescribed fires, wildfires, or residential wood combus-  
134 tion<sup>10,45,46</sup>. Our large observed enhancements in EC and OC during smoke events are consistent  
135 with this earlier work<sup>47,48</sup>. We further confirm this consistency by calculating the potassium-to-  
136 elemental carbon and the potassium-to-organic carbon ratios – common measures of biomass  
137 burning. We estimate that these are  $0.11 \pm 0.14$  for K/EC and  $0.017 \pm 0.027$  for K/OC in our  
138 data, which is similar to published estimates<sup>49</sup>.

139 For the majority of the other species in our sample, we find that variation in wildfire smoke drives  
140 smaller (in terms of mass) but statistically significant increases in concentration relative to the av-  
141 erage ambient background species concentrations (**Figure 2; Supplementary Table 2**). While  
142 small, these mass concentration increases could have meaningful impacts on ecosystems and  
143 human health, particularly as some are essential for the functioning of different organisms (e.g.,  
144 zinc)<sup>50</sup>. Other factors, such as the age of smoke and reactivity of the chemical compound in  
145 smoke, are also important considerations that may matter more for downstream impacts beyond  
146 changes in mass concentration but are not considered here<sup>51</sup>.

147 We detect only small or negligible effects of wildfire-smoke-driven concentration changes for  
148 chromium (Cr) and sodium (Na). The null result is expected for Na because it is highly reactive  
149 and unlikely to be captured in monitoring readings, which aligns with previous work that does not  
150 consistently find elevated Na on smoke days<sup>37</sup>. Previous research shows that Cr production in soil  
151 is highly dependent on fire severity and burn temperature, where low severity fires lead to min-

152 imal differences in the amount of reactive and toxic forms of Cr between burned and unburned  
153 soils, which could in turn impact how much is mobilized in the air<sup>52</sup>; our finding of a small aver-  
154 age effect of smoke on Cr concentrations could mask important heterogeneity by fire type.

### 155 **Variation in the chemical composition of wildfire smoke PM<sub>2.5</sub>**

156 Our main results – using a nationwide model that accounts for regional differences – show that  
157 wildfire smoke increases concentrations of several chemicals of interest. However, this approach  
158 averages over potentially important regional differences in smoke composition; we therefore also  
159 estimate models that allow the effect of wildfire smoke on species concentrations to vary by re-  
160 gion. While this approach does not directly link monitor measurements to specific fires (and  
161 therefore information about what burned), it does allow the relative compositional relationships  
162 to vary across the country. We find regional heterogeneity in the composition of wildfire smoke  
163 that is consistent with existing evidence (**Supplementary Figure 2**)<sup>36</sup>. There are large and sig-  
164 nificant differences in the concentration of EC and OC in smoke PM<sub>2.5</sub> between the western re-  
165 gions (Rocky Mountain, Pacific, and Southwest) compared to the eastern regions (Northeast and  
166 Southeast), whereas for K, the biggest differences are between the Midwest and Southeast regions  
167 (**Supplementary Table 3**). These regional differences could be because each region’s smoke is  
168 a combination of local burned material and transported smoke from burns elsewhere. Given the  
169 prevailing burn and circulation patterns, atmospheric mixing is stronger in the eastern regions,  
170 with likely strong influence from Canadian fires across our study area and period.

171 Beyond regional variation in the chemical composition of smoke PM<sub>2.5</sub>, we explore how overall  
172 species concentrations are affected by growing wildfire activity by predicting the concentration  
173 of each species that is attributable to wildfire over time (**Figure 3**). Given the increased wild-  
174 fire activity during this period, we would expect the concentration of chemicals associated with  
175 wildfire activity (e.g., from the burning of biomass) to also increase over time. We predict the  
176 concentration of each chemical species attributable to wildfire smoke in total PM<sub>2.5</sub> (see **Ma-**  
177 **terials and Methods**) and find that this quantity is significantly trending upward for the major-  
178 ity of species even as the background species-specific concentrations are decreasing for some  
179 chemicals (**Supplementary Figure 3**). By dividing the predicted daily wildfire attributable con-  
180 centrations by the observed species-specific concentrations in total PM<sub>2.5</sub> measured at ground  
181 monitors, we calculate the share (in %) of each species concentration attributable to wildfire and  
182 determine if this share is significantly increasing over time. We find the largest rate of increase in  
183 the concentration attributable to wildfire for species linked to biomass burning, with EC and OC  
184 both increasing at 0.5% per year and P increasing at 2.7% per year on average. For OC, the 2020  
185 annual average wildfire attributable concentration is 5.2x that of the 2006 concentration (from  
186 0.10  $\mu\text{g}/\text{m}^3$  in 2006 to 0.52  $\mu\text{g}/\text{m}^3$  in 2020). This jump is even larger for EC: from an annual av-

187 erage concentration attributable to wildfire in 2006 of  $0.01 \mu\text{g}/\text{m}^3$  to  $0.07 \mu\text{g}/\text{m}^3$  in 2020 (6.1x  
188 higher). For P, the 2020 annual average concentration of wildfire-attributable particles is 5.4x  
189 that of the 2006 concentration, increasing from  $0.00001 \mu\text{g}/\text{m}^3$  in 2006 to  $0.00007 \mu\text{g}/\text{m}^3$  in 2020  
190 (**Supplementary Table 4**).

191 Notably, we find that the wildfire-attributable shares of chemicals known to have no safe level  
192 of exposure — As, Ni, and Pb<sup>53</sup> — are also increasing over time. On average, the concentration  
193 attributable to wildfire is increasing 0.16% per year for As (2020 concentration is 1.9x that of  
194 2006), 0.15% per year for Ni (2.6x), and 0.14% per year for Pb (3x) (**Figure 3**). In other words,  
195 wildfire is responsible for an increasing amount of total exposure to many chemicals in our sam-  
196 ple. For some of these chemicals (e.g., As, V, SO<sub>4</sub> among others), exposure to non-wildfire sources  
197 of a chemical is decreasing even as wildfire driven exposures are increasing (**Supplementary**  
198 **Figure 3**).

### 199 **Comparison of species concentrations in smoke from wildfires that burn structures**

200 More man-made structures are burning in wildfires in recent years (**Supplementary Figure 4**),  
201 which existing work suggests could play a role in the concentration of chemicals in smoke<sup>37</sup>.  
202 We leverage our larger temporal and spatial sample to test this hypothesis at scale (i.e., more  
203 chemicals across multiple fires), building on earlier work<sup>38</sup> to link smoke exposure at monitors  
204 to source fires and then using administrative data to link these fires to the number of structures  
205 burned (**Materials & Methods**).

206 We find that an additional structure burned per fire per day significantly increases the concen-  
207 tration of Ni, Cu, Zn, and Pb in smoke compared to smoke from fires that did not burn struc-  
208 tures in our sample (assuming significance at the  $p < 0.01$  level). (**Figure 4A, Supplementary**  
209 **Table 5**). Multiplying our estimates by 1,000 for interpretability, an additional 1,000 structures  
210 burned per fire day increases the concentration of Cu by  $1.7 \times 10^{-4} \mu\text{g}/\text{m}^3$  (0.009% increase rel-  
211 ative to a  $0.002 \mu\text{g}/\text{m}^3$  average concentration in smoke PM<sub>2.5</sub> when no structures burned, 95%  
212 CI:  $6.8 \times 10^{-5}$  to  $2.7 \times 10^{-4} \mu\text{g}/\text{m}^3$ ) and Pb by  $2.5 \times 10^{-4} \mu\text{g}/\text{m}^3$  (0.02% increase relative to  $0.0014$   
213  $\mu\text{g}/\text{m}^3$ , 95% CI:  $1.8 \times 10^{-5}$  to  $3.2 \times 10^{-4} \mu\text{g}/\text{m}^3$ ). We find that the largest increase in concentra-  
214 tion due to an additional structure burning is for Zn, with an average increase of  $1.2 \times 10^{-3} \mu\text{g}/\text{m}^3$   
215 (0.02% increase relative to  $0.006 \mu\text{g}/\text{m}^3$ , 95% CI:  $6.5 \times 10^{-4}$  to  $1.8 \times 10^{-3} \mu\text{g}/\text{m}^3$ ). For some ele-  
216 ments that are more common in vegetation than structure building materials (OC, Se, P, and Mg),  
217 we find a decreasing association between structures burned and species concentrations, consis-  
218 tent with the fact that fires that burn structures are often not those that burn the most vegetation  
219 (**Supplementary Figure 5**). Our results are robust to a randomization inference test in which  
220 structures are randomly re-assigned to alternate (incorrect) fires (**Figure 4B, Supplementary Ta-**

221 **ble 6**), which helps rule out that our results are being driven by region- or time-period-specific  
222 characteristics of nearby fires rather than by burned structures themselves.

223 When applied to specific fires that burned many structures, our estimates suggest that these fires  
224 can result in substantial downwind increases in key species. For example, we estimate the ex-  
225 cess concentrations in smoke PM<sub>2.5</sub> from the Camp Fire, which was the most destructive fire  
226 in California's history, burned over 18,000 structures, and lasted for about 18 days<sup>54</sup>. We find  
227 that smoke from the Camp Fire increased species concentrations in smoke PM<sub>2.5</sub> relative to if no  
228 structures had burned, with increases in Pb concentration by 17.0% per fire day or 0.00025 μg/m<sup>3</sup>  
229 (95% CI: 0.00018 to 0.00032 μg/m<sup>3</sup>) and Ni by 7.2% or 0.00003 μg/m<sup>3</sup> (95% CI: 0.0000031 to  
230 0.00005 μg/m<sup>3</sup>).

### 231 **Cancer risk assessment from exposure to wildfire attributable concentrations**

232 To understand potential risk to public health from wildfire-driven changes in exposure to specific  
233 chemical species, we combine model-estimated changes in country-wide ambient exposure to  
234 three carcinogenic species – As, Ni, and Pb – with existing estimates of changes in cancer risk  
235 that result from changes in ambient exposure to these chemicals. We estimate that increasing  
236 wildfire activity has generated substantial increases in ambient concentrations of these species  
237 between the first five years of our sample (2006-2010) to the last five years (2016-2020), with par-  
238 ticularly large regional increases in areas where changes in wildfire smoke exposure have been  
239 the largest (**Figure 5A**). We then estimate how many additional cancer cases might occur as a re-  
240 sult of this exposure, using existing estimates of changes in lifetime cancer risk due to changes in  
241 chronic exposure, and assuming that our estimated species concentrations represent average long-  
242 term (lifetime) residential exposures (**Materials and Methods**). Total excess cases are then the  
243 product of the estimated change in exposure in a given location, the change in cancer risk for that  
244 exposure, and the number of people who were exposed in that location.

245 We find that increases in wildfire activity have likely increased lifetime cancer risk over our study  
246 period, perhaps substantially in percentage terms, but that the overall increase in attributable  
247 cases is likely very small (**Figure 5B, Supplementary Table 7**). For both Pb and Ni, we estimate  
248 that chronic exposure to wildfire-driven species concentrations in the recent 2016-2020 period  
249 would lead to zero to four additional cancer cases across the US population. For As, chronic ex-  
250 posure to recent concentrations would lead to roughly 45 additional cases. The largest total num-  
251 ber of attributable cases are in populated centers in Midwest and Eastern parts of the US, while  
252 the largest increase in attributable cases is in the Western US. These increases are very small rel-  
253 ative to other risk factors (e.g. exposure to outdoor air pollution causes over half a million deaths  
254 due to lung cancer each year, in comparison<sup>55</sup>), suggesting that wildfire-driven changes in ambi-



255 ent concentration of these particular species is unlikely to be a meaningful source of cancer risk.  
256 However, this does not mean that other chemicals not measured here or that other exposure sce-  
257 narios not considered (e.g., acute exposures) would not lead to meaningful increases in risk, nor  
258 that exposure to these chemicals could not result in other non-cancer health burdens (e.g., autoim-  
259 mune disease).

## 260 **Discussion**

261 Our study characterizes the chemical composition of wildfire smoke using daily concentration  
262 of 27 chemicals measured at EPA ground monitors and statistical techniques that can isolate  
263 the contribution of wildfire smoke to species concentrations from a broad range of other sea-  
264 sonal, trending, or time-invariant factors that could also affect concentrations. Our research com-  
265 plements previous research that has used a variety of techniques (e.g., chemical transport mod-  
266 els<sup>29,56,57</sup>, field collected air quality samples<sup>29-33</sup>, and laboratory techniques<sup>29-34</sup>) to understand  
267 the impact of wildfire on ambient air quality. While these alternate methods offer important in-  
268 sights and detail into the composition of smoke and the mechanisms that drive compositional  
269 changes, they are often limited spatially or temporally and/or are highly computationally or re-  
270 source intensive. Our work enables characterization of the contribution of wildfire smoke to  
271 species concentrations at a broad scale and very low computational and resource cost, comple-  
272 menting these other approaches that can provide mechanistic and fire-specific detail.

273 We detect the signature of wildfire smoke across a broad array of chemical species concentra-  
274 tions and show that this signature has strengthened over time. Our work provides comprehensive  
275 new insight into the composition and toxicity of wildfire smoke, with implications for popula-  
276 tion exposure and corresponding health impacts. We find that wildfire drives significant increases  
277 in 25 of the 27 chemical species in our sample, with the largest increases in the concentration of  
278 OC, EC, K, and P, consistent with the abundance of these elements in biomass. Given the up-  
279 ward trend in wildfire activity and that the majority of historical wildfires burned in vegetated  
280 areas, our results align with previous research that has identified these chemicals as markers of  
281 biomass combustion<sup>10,45-47,49</sup>. We find that wildfire also influences many other chemical species'  
282 concentrations in our sample, albeit sometimes to a lesser degree, yet these changes may still be  
283 meaningful to ecosystem function and human health outcomes given that some of the chemicals  
284 are considered essential nutrients (e.g., zinc) or are harmful toxins (e.g., arsenic).

285 For many of the chemicals in our sample, the *share* of each species' concentration attributable to  
286 wildfire has significantly increased over time, with wildfire-attributed concentrations of multiple

287 carcinogenic metals significantly higher by the end of our sample. This trend is consistent with  
288 the observed increase in wildfire activity over our sample period and will likely continue to grow  
289 given the projections of future wildfire under climate change scenarios. When assessing the risk  
290 that these concentration increases pose to human health, we estimate that exposure to wildfire  
291 attributable concentrations of three of these metals – As, Ni, and Pb – could result in a combined  
292 ~50 excess cancer cases. Though these risks are quite small compared to other environmental  
293 risks, there are likely many other chemicals in smoke not measured in our study that are cancer-  
294 and health-relevant. Our approach could be readily extended to other settings where consistent  
295 monitoring of these other chemicals is available. Our approach could also be extended to study  
296 the chemical composition of smoke from prescribed burning, a key fire risk management tool  
297 where information on the quantity and toxicity of generated smoke is a key open scientific and  
298 public health question.

299 Our work offers insight into one specific channel through which wildfires can impact the chemi-  
300 cal components of ambient air quality: the burning of infrastructure. In line with previous work,  
301 we find that wildfires that burn structures result in higher concentrations of certain metals, in-  
302 cluding Cu, Pb, Zn, and Ni, compared to fires that do not burn structures, and that fires that burn  
303 many structures (such as California’s 2018 Camp Fire) can lead to large compositional changes  
304 in air quality, including substantial enhancement of known carcinogens. Some of these species  
305 are considered trace elements in soil (Zn and Cu) and are present naturally but at relatively low  
306 concentrations compared to other soil components (like Al, Si, and Fe). Thus, enhancements in  
307 the concentrations of Zn and Cu points to smoke from burned structures as an important con-  
308 tributor to ambient concentrations for these species<sup>37,58</sup>. As houses are increasingly constructed  
309 in the wildland-urban interface, risk of structure-driven changes in air quality and toxicity will  
310 also likely increase, potentially posing significant health risks for nearby populations. However,  
311 burned structures are only one among many fire-specific features that could affect the toxicity  
312 and health impacts of smoke. Differences in soil type or vegetation as well as the distance the  
313 smoke traveled are a few among many factors that could affect how wildfire activity translates  
314 into downwind air quality. Future work could again build on existing efforts to link smoke back  
315 to its source fire<sup>38</sup>, and hone in on the drivers of observed air quality differences and their impor-  
316 tance for exposure mitigation and health outcomes.

317 Our approach has additional limitations. We focus on PM<sub>2.5</sub> pollution due to its known link to  
318 harmful health outcomes, but PM of other size bins and ultrafine particles could also be harm-  
319 ful and are an important avenue to explore in future work. Furthermore, our analysis does not  
320 determine the effect that wildfire smoke has on the concentration of sub-species variants (e.g.,  
321 trivalent and hexavalent chromium), as this information is not captured at monitoring stations.

322 Better measurement of sub-species concentrations will likely be important for health risk analy-  
323 ses. Similarly, we did not explicitly quantify compound chemical exposures here (simultaneous  
324 increases in exposure to multiple chemicals) or try to anticipate their health impacts. Finally, over  
325 the course of our sample of air quality measurements, there were changes to air quality monitor-  
326 ing<sup>59</sup>, such as changes to detection limits, management, or monitoring methods that were not al-  
327 ways clearly documented. We identified a few of these changes (**Figure 1C**), and while our fixed-  
328 effects approach will account for many differences in management or measurement error between  
329 monitors and any common changes to monitoring across monitors, it is possible we did not ac-  
330 count for others.

331 We contribute to a growing body of literature that demonstrates the impact that wildfires have on  
332 a range of chemical concentrations in ambient air. Our findings underscore the growing signif-  
333 icance of wildfire as a multidimensional threat to air quality and health. This threat is projected  
334 to grow under a warming world. However, wildfires' contribution to air quality is currently not  
335 considered in conventional regulatory frameworks at the US federal level. Currently, national air  
336 quality standards are focused on six principal pollutants (“criteria” air pollutants), which are reg-  
337 ulated by the National Ambient Air Quality Standards in the Clean Air Act, and are intended to  
338 provide public health protection<sup>60</sup>. Notably, wildfires and their emissions are excluded from these  
339 standards as they are considered “exceptional events,” which are defined as natural or unusual  
340 events that can overwhelm existing strategies to control anthropogenic pollution. Yet as we show  
341 here, wildfires are a significant contributor to the concentration of chemicals in ambient air and  
342 may lead to meaningful health impacts, only some of which we can currently document. We take  
343 an important step towards demonstrating that not all PM<sub>2.5</sub> is created equal in terms of compo-  
344 sition and toxicity. Given the measurable, detectable, and significant impact that wildfire has on  
345 ambient chemical concentrations, our findings could contribute to future policy and regulatory  
346 efforts to protect human health from this growing exposure risk.

## 347 **Materials and Methods**

### 348 **Data**

349 All of the data used in our study are from publicly available sources.

350 ***Total PM<sub>2.5</sub> chemical speciation data.*** We retrieve chemical speciation data (concentration mea-  
351 surements) from the Federal Land Manager Environmental Database (FED), which is an online  
352 repository of air quality data and metadata sponsored by the National Park Service and the U.S.  
353 Forest Service.<sup>61</sup> The FED contains and regularly updates ground-based measurements with

354 standardized metadata from dozens of monitoring networks as well as other air quality outputs.  
355 We select annual data from the “Pre-Generated Data” webpage (the same data are also available  
356 via query for more customized downloads). Here, we retrieve daily monitor-level chemical con-  
357 centration measurements from 2006–2020 from two national air monitoring networks in the US:  
358 the Environmental Protection Agency’s (EPA) PM<sub>2.5</sub> Chemical Speciation Network (CSN) (con-  
359 sisting of Speciation Trends Network (STN) sites and supplemental speciation sites) and the In-  
360 teragency Monitoring of Protected Visual Environments (IMPROVE) program. We select only  
361 the chemical species that are measured at both monitoring networks and are in both datasets,  
362 which include: aluminum (Al), arsenic (As), bromine (Br), calcium (Ca), organic carbon (OC),  
363 elemental carbon (EC), chlorine (Cl), chromium (Cr), copper (Cu), iron (Fe), lead (Pb), magne-  
364 sium (Mg), manganese (mg), nickel (Ni), nitrate (NO<sub>3</sub>), phosphorus (P), potassium (K), rubidium  
365 (Rb), selenium (Se), silicon (Si), sodium (Na), strontium (Sr), sulfate (SO<sub>4</sub>), sulfur (S), titanium  
366 (Ti), vanadium (V), zinc (Zn) (**Supplementary Table 1**).

367 Measurements from ground monitors, such as those utilized here from CSN and IMPROVE air  
368 monitoring networks, are the most accurate available source of information on ambient surface  
369 pollutant concentration. However, the monitors also have limitations particularly with respect to  
370 measuring wildfire smoke. Monitors are not uniformly spread across the US and monitor den-  
371 sity varies dramatically across space. Additionally, monitors have their own set of measurement  
372 biases and those biases may vary over space and time. Both CSN and IMPROVE monitoring net-  
373 works underwent significant changes in the methods and operations of measuring air quality over  
374 our sample duration, including: management shifts of who operates specific monitors, technologi-  
375 cal improvements that enhance the ability of monitors to detect lower concentrations or measure  
376 new chemicals, and methodological changes that correct previous biases in measuring concen-  
377 tration of different species<sup>62–64</sup>. In our data cleaning process, we retain any measured negative  
378 concentrations values as these are reported in both monitoring networks data. Both the changes  
379 in methods as well as the inclusions of negative concentration values are unlikely to impact our  
380 main results because our regression approach should capture any baseline changes between moni-  
381 tors in reporting or any changes in the methods or detection limits in a given year. Lastly, wildfire  
382 smoke PM<sub>2.5</sub> is only one of many source that contribute to ambient total PM<sub>2.5</sub> and these moni-  
383 tors do not distinguish between PM<sub>2.5</sub> from wildfires and that from other sources of ambient  
384 PM<sub>2.5</sub>. Additional steps must therefore be taken to estimate the part of monitor measured PM<sub>2.5</sub>  
385 that comes from wildfire smoke (see **Wildfire Smoke PM<sub>2.5</sub> Data**).

386 **Total PM<sub>2.5</sub> data.** In addition to chemical concentration measurements, we retrieve total PM<sub>2.5</sub>  
387 concentration data for both monitoring networks from the annual-level pre-generated chemical  
388 speciation data files as well as additional data located on the EPA’s AQS site called “EPA PM<sub>2.5</sub>

389 Mass FRM (88101) - Daily” under the “EPA Air Quality System (AQS).” For our analysis, we se-  
390 lect a measure for total  $PM_{2.5}$  called “ $PM_{2.5}$  - Local Conditions” (MF, AQS code: 88101). There  
391 can be many measures of total  $PM_{2.5}$  within the same monitoring networks as well as between  
392 networks, and previous studies have shown that there can be large differences in the reported con-  
393 centration of total  $PM_{2.5}$  at the same monitor. We explore how using the two different measures  
394 of total  $PM_{2.5}$  may impact our results, and we show that our results are very similar regardless of  
395 the total  $PM_{2.5}$  measure (**Supplementary Figure 1**).

396 In the chemical speciation data, neither monitoring network records the coordinate reference sys-  
397 tem (CRS) for the coordinate pairs (latitude and longitude) that indicate the location of a given  
398 monitoring site. To ensure that the CRS is consistent across all our data, we retrieve spatial in-  
399 formation about each site from the Air Quality System (AQS) website<sup>65</sup> and standardize the co-  
400 ordinate reference system across monitoring networks. The location of the monitors used in our  
401 analysis along with the duration that each monitoring site is online is shown in **Figure 1A**. The  
402 majority of monitors are online for 10-15 years (our full sample period is 15 years).

403 Our dataset includes chemical speciation data for total  $PM_{2.5}$  from both IMPROVE and CSN  
404 monitors from 2006-2020 at the daily-monitor level in the contiguous US. To produce a clean  
405 dataset, we adjust values that are “Inf”, “NaN”, “-999”, or “-499” to “NA.” We leave negative  
406 concentration values in our dataset as they could indicate that a measurement was below the  
407 method detection limit, except for those mentioned above, which likely indicate a data input er-  
408 ror or missing data<sup>66</sup>. We drop any monitor-dates that are missing data for all of our selected  
409 species to ensure the composition of our monitoring stations is not changing dramatically over  
410 our sample. All chemical concentration measurements as well as total  $PM_{2.5}$  are reported in units  
411 of  $\mu g/m^3$ . To explore seasonality (**Figure 1C**) and regional variation (**Supplementary Figure**  
412 **2**) in the chemical composition of wildfire smoke  $PM_{2.5}$ , we classify each monitor-date in one of  
413 six regions (Midwest, Northeast, Pacific, Rocky Mountain, Southeast, and Southwest) and one of  
414 four seasons (autumn [September-November], winter [December-February], spring [March-May],  
415 summer [June-August]).

416 **Wildfire smoke  $PM_{2.5}$  data.** For daily smoke  $PM_{2.5}$  concentrations, we use a dataset of gridded  
417 daily wildfire smoke  $PM_{2.5}$  predictions on smoke days at 10 km resolution from January 1, 2006  
418 to December 31, 2020 for the contiguous US from Childs et al., 2022<sup>2</sup>. Predictions of wildfire  
419 smoke  $PM_{2.5}$  are non-zero numbers, such that predictions on non-smoke days are by construc-  
420 tion zero. When there are predictions of zero for wildfire smoke  $PM_{2.5}$  on smoke days, this in-  
421 dicates that a given grid cell-day was determined to be a “smoke-day” but did not have elevated  
422 total  $PM_{2.5}$ . For further details on the daily wildfire smoke  $PM_{2.5}$  dataset, see Childs et al., 2022.

423 For every monitor day, we construct a measure of total  $PM_{2.5}$ , smoke  $PM_{2.5}$ , and non-smoke  
424  $PM_{2.5}$ . We match the gridded daily wildfire smoke  $PM_{2.5}$  predictions with the daily chemical  
425 species concentration data from the air quality monitors to construct a measure of smoke  $PM_{2.5}$   
426 at monitoring stations. We assume that predicted smoke  $PM_{2.5}$  in the same grid cell as monitor-  
427 ing station(s) is representative of what measured smoke  $PM_{2.5}$  would be at that station(s). For  
428 monitor-days that do not have a wildfire smoke  $PM_{2.5}$  prediction (i.e., missing), we assume these  
429 are non-smoke days, and assign values of zero for smoke  $PM_{2.5}$  on these days. For days without  
430 wildfire smoke (non-smoke days), we set total  $PM_{2.5}$  equal to non-smoke  $PM_{2.5}$ . For days with  
431 a predicted concentration of wildfire smoke  $PM_{2.5}$ , we calculate non-smoke  $PM_{2.5}$  by subtract-  
432 ing smoke  $PM_{2.5}$  from the measure of total  $PM_{2.5}$  from the air quality monitors. We expect that  
433 the concentration of total  $PM_{2.5}$  should be at least equal to or greater than the measure of wild-  
434 fire smoke  $PM_{2.5}$ , given that total  $PM_{2.5}$  should be the total of both smoke and non-smoke  $PM_{2.5}$   
435 components. For any total  $PM_{2.5}$  concentrations that are either missing or lower than our measure  
436 of wildfire smoke  $PM_{2.5}$  at a given monitor and day, we adjust the total  $PM_{2.5}$  values to be equal  
437 to the wildfire smoke  $PM_{2.5}$  at a given monitor-date. This is a reasonable assumption given that  
438 wildfire smoke  $PM_{2.5}$  may comprise the majority of total  $PM_{2.5}$  on smoke-days.

439 ***Estimating the effect of wildfire smoke  $PM_{2.5}$  on chemical species concentration.*** To assess  
440 wildfire smoke's influence on different chemicals in ambient air pollution, we estimate changes  
441 in species' concentrations that come from changes in the level of wildfire smoke  $PM_{2.5}$ , while  
442 accounting for any other co-varying emission sources that could also change chemical concen-  
443 trations in ambient air. A simple comparison of the concentration of chemical species on days  
444 with wildfire smoke as compared to days without wildfire smoke would be biased because, in  
445 our sample, total  $PM_{2.5}$  pollution is about  $5 \mu\text{g}/\text{m}^3$  higher on wildfire smoke days, conditional  
446 on controlling for unobserved variation at a given monitor-month and year. Therefore, if wildfire  
447 smoke  $PM_{2.5}$  and non-smoke  $PM_{2.5}$  have the same proportion of a given species, total concen-  
448 tration will be higher on a wildfire smoke day. Here, we define (and calculate) non-smoke  $PM_{2.5}$   
449 as the ambient pollution represented by the difference of total  $PM_{2.5}$  and wildfire smoke  $PM_{2.5}$ .  
450 Thus, a comparison of the relative species concentrations on wildfire smoke versus non-smoke  
451 days will not be informative for per-unit toxicity. Similarly, most of our selected chemical species  
452 have background seasonal variation (Fig 1C), since the majority of wildfire smoke days occur  
453 between June–September in the US. Accordingly, a comparison of species' concentrations may  
454 mischaracterize true seasonal differences as differences in smoke and non-smoke  $PM_{2.5}$ , though  
455 the direction of this seasonal bias is unknown *ex ante* and could be species-specific.

456 To do this, we combine observed ambient species concentration at ground monitors with gridded  
457 smoke  $PM_{2.5}$  predictions at a monitor-day level and estimate a regression framework that com-

458 pares within-monitor changes over time after flexibly accounting for seasonality and time trends.  
 459 We do not compare across monitoring stations because monitors in different regions experience  
 460 different levels of wildfire smoke  $PM_{2.5}$  and so observed differences in species concentration  
 461 could be due to baseline differences in wildfire smoke exposure. We also do not make compar-  
 462 isons for a given monitor across years, because wildfire smoke and some chemical species are  
 463 trending over time (**Figure 1B and C**). Instead, our regression framework compares monitors to  
 464 themselves within the same month-of-year after controlling flexibly for time trends. Namely, we  
 465 relate monitor-specific deviations in average species concentration to monitor-specific deviations  
 466 in the average level of wildfire smoke, after accounting for baseline differences across monitors  
 467 and any year-specific shocks experienced by all monitoring stations. Using a fixed effects panel  
 468 regression, we estimate:

$$spec_{idmt} = \beta_s PM_{idmt}^{smoke} + \beta_{ns} PM_{idmt}^{non-smoke} + \alpha_{im} + \delta_t + \varepsilon_{idmt} \quad (1)$$

469 via ordinary least squares (OLS) where the outcome is the concentration of a given species at  
 470 monitor  $i$  in month  $m$  on day  $d$  in year  $t$ . Our estimate of interest is  $\beta_s$ , which can be interpreted  
 471 as the average effect of a one  $\mu g/m^3$  increase in daily wildfire smoke  $PM_{2.5}$  on the same-day con-  
 472 centration of the chemical species of interest in ambient air. We dived our estimates by baseline  
 473 non-smoke species-specific concentrations to get the change in concentration relative to the aver-  
 474 age concentration of that species on non-smoke days. We estimate these impacts both in percent-  
 475 age (**Figure 2A**) and in absolute terms (**Figure 2B**).

476 Our preferred model regresses daily wildfire smoke  $PM_{2.5}$  on non-smoke  $PM_{2.5}$  with fixed ef-  
 477 fects for monitor by month-of-year  $\alpha_{im}$  (i.e., one intercept for each monitoring station’s location  
 478 in each of the 12 months, e.g., a monitoring station in July in CA), and fixed effects for each year.  
 479 The monitor-by-month-of-year fixed effects account for any baseline differences between moni-  
 480 toring stations and local seasonality. Year fixed effects account for year-specific shocks common  
 481 to all monitoring stations (e.g., a particularly bad wildfire year, changes in monitoring methods  
 482 or equipment, etc). Our approach assumes that day-to-day within monitor variation in wildfire  
 483 smoke after accounting for seasonality and time trends is plausibly exogenous with respect to  
 484 chemical species concentrations.

485 We show that our estimated coefficients are robust to alternative model specifications (**Supplementary**  
 486 **Figure 1**). To check whether sources of emissions other than those from wildfire smoke  $PM_{2.5}$   
 487 change species’ concentrations differentially we estimate versions of our models with and without  
 488 non-smoke  $PM_{2.5}$  concentration included as a covariate. We find that our estimates are consistent

489 regardless of whether our models include non-smoke  $PM_{2.5}$  and we include it in our preferred  
490 model because it is both highly predictive of our outcome and independent of our treatment, and  
491 thus reduces our error variance and corresponding standard errors (**Supplementary Figure 1**).  
492 We cluster our standard errors at the monitoring station-level to account for potential serial auto-  
493 correlation. Main model results can be seen in **Supplementary Table 2**.

494 ***Estimating the contribution of wildfire smoke  $PM_{2.5}$  to a chemical species' concentrations over***  
495 ***time***. We predict the contribution of smoke  $PM_{2.5}$  to daily chemical species' concentrations over  
496 our sample period of 2006–2020. Using the coefficients from our regional model, we multiply the  
497  $\beta_s$  by the predicted smoke  $PM_{2.5}$ , which provides an estimated concentration of a given species  
498 at a monitoring station on a given date. We compare these predictions to the observed species  
499 concentration measured at monitoring stations by dividing our predictions by the observed con-  
500 centration to get the fraction (%) of a given species concentration attributable to wildfire smoke  
501  $PM_{2.5}$ . We test if the trend in the wildfire attributable fraction of each species is significantly in-  
502 creasing over our sample period by estimating:

$$AttFrac_{imt} = \beta_y year_{imt} + \alpha_{im} + \varepsilon_{idmt} \quad (2)$$

503 via ordinary least squares (OLS) where the outcome is the wildfire attributable fraction of a given  
504 species concentration at monitor  $i$  in month  $m$  on day  $d$  in year  $t$ . The estimate of interest is  $\beta_y$ ,  
505 which can be interpreted as the average annual change (%) in the attributable fraction of the chem-  
506 ical species of interest. We include the monitor-by-month-of-year fixed effects to account for  
507 monitor-specific changes and local seasonality. Standard errors are clustered at the monitoring  
508 station-level.

509 We then average the daily concentration predictions by region, species, year, and month to get the  
510 average concentration at a given monitor, month, and year, and then we average again by species,  
511 year, and month, to get the average concentration of a given chemical species across all moni-  
512 tors in a given year and month. We calculate the average observed concentration in total  $PM_{2.5}$   
513 measured at monitoring stations for each species, month, and year. We plot this as percent of con-  
514 centration due to wildfire smoke  $PM_{2.5}$  (**Figure 3**).

515 **Supplementary Figure 3** shows our predicted attributable concentration in comparison to ob-  
516 served species. **Figure 3** condenses this information only showing the attributable fraction of  
517 concentration due to wildfire smoke  $PM_{2.5}$  for each species, and species that demonstrate signifi-  
518 cant annual increases in the concentration attributable to wildfire smoke  $PM_{2.5}$  have their percent



519 increase year-over-year included in the plot in black. Chemical species in which no detectable  
520 significant trend was detected are shown, but the percent increase is greyed out.

521 ***Estimating regional chemical species concentration in wildfire smoke  $PM_{2.5}$ .*** Using our main  
522 model specification (Equation 1), we run a regression for each region (Midwest, Northeast, Pa-  
523 cific, Rocky Mountain, Southeast, and Southwest) to estimate the  $\beta_s$ , or the average effect of a  
524 one  $\mu\text{g}/\text{m}^3$  increase on the concentration of each chemical species within smoke within that re-  
525 gion. Similar to our full sample model, we normalize the regional coefficients by dividing the  $\beta_s$   
526 for each chemical species within a given region by the regional concentration average on non-  
527 smoke days as a way to understand how much concentration is changing relative to a baseline  
528 ambient species-specific concentration in that region. Our results can be interpreted as the aver-  
529 age percent change of a given chemical species' concentration due to a one unit increase in smoke  
530  $PM_{2.5}$  within a region relative to a non-smoke day (**Supplementary Figure 2, Supplementary**  
531 **Table 3**).

532 ***Estimating the effect of fires that burn structures on species' concentrations in smoke  $PM_{2.5}$ .***  
533 We first match wildfires to the smoke  $PM_{2.5}$  that they emit. We recover the incident name and  
534 number of structures burned for fires in our sample period by processing and combining two  
535 datasets: Globfire and monitoring trends in burn severity (MTBS) burned area perimeter data.  
536 We combine Globfire's shapefiles of the final area of fires from 2006-2020 with the MTBS burned  
537 area perimeter data for the same overlapping period of 2006-04-19 – 2020-12-31. We join Glob-  
538 fire and MTBS data by area, calculate area of intersection, and only keep if area of intersection  
539 is  $>75\%$  of the burned area. We remove fires that have MTBS ignition date more than one week  
540 before Globfire start or end dates.

541 We then incorporate the National Interagency Fire Center (NIFC) data to address some spatial in-  
542 consistencies in the burned structures data. Some fires in the burned structure data have integer  
543 lat/lon values. To resolve this, we merge the NIFC data with the burned structures data by filter-  
544 ing to the relevant year and correct coordinates. In cases where the lat/lon from the burned struc-  
545 tures dataset is an integer, we use the NIFC lat/lon. In the NIFC data, there were 515 duplicated  
546 unique identifiers, where some were actual duplicates and others appeared to have slightly dif-  
547 ferent county, FIPS, etc. This is likely because this dataset aggregates reports of fires to different  
548 agencies so there could be duplicates. We solve for the duplicate issues by grouping by unique  
549 ID, year, incident number, state, county, FIPS, and averaging the lat/lon coordinates. After this,  
550 there were still about 35 fires with duplicates IDs and some structures with integer lat/lon, so we  
551 remove these from the dataset.

552 We combine the burned structures dataset with the Globfire data and only keep incidents in which  
 553 the start date from the structures burned dataset is within two weeks of the Globfire ignition date.  
 554 To account for the fact that fires burn multiple days and the data provides only the final total num-  
 555 ber of structures destroyed, we divide the number of structures destroyed by the duration of the  
 556 fire, using the end-start date of a given fire from Globfire to get the number of structures a partic-  
 557 ular fire burned in a given day. We aggregate all data to the 10x10km grid cell to match the res-  
 558 olution of our smoke dataset. However, one grid cell may experience multiple fires on the same  
 559 date, we leverage a method from Wen et al., 2023, where we use a share value to distribute the  
 560 number of structures destroyed within a grid cell to a specific fire’s smoke. We sum up the total  
 561 structures destroyed from each contributing fire for the given date and grid cell.

562 To understand if chemicals species’ concentrations are changing due to fires that that burn struc-  
 563 tures, we construct a dataset that matches the concentration of each chemical species (from mon-  
 564 itoring stations in that grid cell) in wildfire smoke  $PM_{2.5}$  to the number of burned structures per  
 565 day per fire in that smoke. We filter to fires that have burned at least one structure. We are essen-  
 566 tially asking if smoke with structures burned in a fire has different levels of species concentra-  
 567 tions than smoke that did not burn structures. We estimate if the effect that an increase in wild-  
 568 fire smoke  $PM_{2.5}$  has on species’ concentrations differs as a function of the number of structures  
 569 burned in the source fire. Using a panel regression with fixed effects, we estimate:

$$spec_{idmt} = \beta_s PM_{idmt}^{smoke} + \beta_{bs} PM_{idmt}^{smoke} * BurnStruc_{idmt} + \beta_b BurnStruc_{idmt} + \alpha_{im} + \delta_t + \varepsilon_{idmt} \quad (3)$$

570 via ordinary least squares (OLS) where the outcome is the concentration of a given species at  
 571 monitor  $i$  on day  $d$  in month  $m$  in year  $t$  in smoke that has burned at least one structure. The esti-  
 572 mate of interest is the interaction term  $\beta_{bs}$ , which is a weighted measure of the number of struc-  
 573 tures burned in the attributed smoke, and can be interpreted as the effect that one additional struc-  
 574 ture burning in a fire has on the same-day concentration of a given chemical species in wild-  
 575 fire smoke  $PM_{2.5}$  (**Figure 4, Supplementary Table 5, Supplementary Figure 4**). We cluster  
 576 our standard errors at the monitor-level, and we include fixed effects for monitor by month-of-  
 577 year and year to account for any baseline differences between monitoring stations, seasonality of  
 578 smoke, and any year-specific shocks common to all monitoring stations. Our approach holds PM  
 579 constant but does not directly control for fire-specific differences (e.g., low smoke and high struc-  
 580 tures burned versus high smoke no structures burned), nor does it differentiate timing of struc-  
 581 tures burned within a fire event.

582 **Permutation/randomization-style inference.** We allocate a share of burned structures per day to  
583 smoke associated with specific fires, but there may be cases of mis-allocation of our treatment.  
584 Mis-allocation could impact our estimated effect of structures burning on species' concentra-  
585 tion. Therefore, randomization style inference allows us to test whether the effect we observe is  
586 unlikely by chance. Using a sharp null hypothesis that our estimated effect of an additional struc-  
587 ture burning is zero, we permute the allocation of structures to specific fire's smoke by explic-  
588 itly breaking the potential temporal autocorrelation of a given fire burning multiple days (i.e.,  
589 same fire is overhead multiple days) and spatial autocorrelation (e.g., two or more fires burning at  
590 same time and one destroys structures at a monitor in the middle). For example, if we use state-  
591 level shuffling of structures burned, the attributed smoke would be the same even if we swap the  
592 destroyed structures between the fires. To address this, we explicitly sample from observations  
593 within state, but ensure that the samples are from different years, thereby explicitly breaking tem-  
594 poral autocorrelation. We run this simulation 1,000 times for each species and compare our true  
595 coefficient to the distribution of the simulated coefficients. We calculate the percentiles of the  
596 observed regression coefficient estimates (**Supplementary Table 6**).

597 **Estimating wildfire attributable species concentrations across the US.** We estimate species con-  
598 centration attributable to wildfire smoke in grid cells that have monitoring stations using the loca-  
599 tion of each monitor (latitude and longitude) via:

$$\begin{aligned}
spec_{idmt} = & \beta_s PM_{idmt}^{smoke} + \beta_{ns} PM_{idmt}^{non-smoke} + \beta_s PM_{idmt}^{smoke} * \beta_y lat_{idmt} + \\
& \beta_s PM_{idmt}^{smoke} * \beta_x long_{idmt} + \beta_s PM_{idmt}^{smoke} * \beta_x long_{idmt} * \beta_y lat_{idmt} + \alpha_{im} + \delta_t + \varepsilon_{idmt}
\end{aligned}$$

600 where the outcome is the concentration of a given species at monitor  $i$  located at longitude  $x$  and  
601 latitude  $y$  on day  $d$  in month  $m$  and year  $t$ . The coefficients of interest are the  $\beta_s$  and the inter-  
602 acted terms, which allow us to flexibly predict the concentration of a given chemical species in  
603 grid cells without monitors. To estimate population-level exposure to chemical concentrations at-  
604 tributable to wildfire smoke  $PM_{2.5}$ , we predict daily concentrations of each chemical species in  
605 10x10km grid-cells that match those from the smoke  $PM_{2.5}$  dataset.

606 Using our estimated coefficients, we predict across the US:

$$spec_{gmt} = PM_{gmt}^{smoke} (\beta_s + \beta_s * X_g + \beta_s * Y_g + \beta_s * X_g * Y_g) \quad (4)$$

607 where the outcome is the predicted attributable concentration of a given species in grid cell  $g$   
608 located at longitude  $x$  and latitude  $y$  on day  $d$  in month  $m$  and year  $y$ .

609 To see how the wildfire-attributable concentration of three health-relevant chemicals, As, Ni, and  
610 Pb, have changed spatially over our sample period, we calculate the average wildfire attributable  
611 concentration in each grid cell across the first five years (2006-2010) as well as the last five years  
612 (2016-2020) of our sample and plot this in **Figure 5A**.

613 ***Assessing the human health risk from chronic exposures to historical wildfire attributable con-***  
614 ***centrations***. Following a similar approach to Dickinson et al. (2022)<sup>40</sup> and building on guidance  
615 from the US EPA and California’s Office of Environmental Health Hazard Assessment<sup>67</sup>, we as-  
616 sess the potential cancer risks of chronic smoke exposure by exploring different wildfire exposure  
617 scenarios. We consider one standard chronic exposure scenario used to assess cancer risk<sup>68</sup>. The  
618 “lifetime” exposure scenario is defined by the projection of chronic and repeated exposure over  
619 the average human’s life expectancy, which the EPA’s designates as 70 years<sup>40,68</sup>. We do not as-  
620 sess acute exposure risk to chemicals in wildfire smoke in this study.

621 We focus on three chemicals that are measured in our sample and have no known safe level of ex-  
622 posure: As, Ni, and Pb. As and Ni are known carcinogens, while lead has been associated with a  
623 number of harmful health outcomes that will likely occur before a person develops cancer; still,  
624 Pb is considered a carcinogen<sup>67</sup>. To conduct a risk assessment of exposure to these three chemi-  
625 cals, we first calculate an exposure concentration (EC) using a standard equation developed by the  
626 US EPA<sup>68</sup>:

$$EC = \frac{\sum(CA) \times ET \times EF \times ED}{AT} \quad (5)$$

627 where CA is the chemical concentration in air in  $\mu\text{g}/\text{m}^3$  (here the CA is the summed wildfire at-  
628 tributable concentration over the five year period of interest); ET is the exposure time in hours per  
629 day (24 hr/day); EF is the exposure frequency in days per year (here the EF is one since each day  
630 has a unique value); ED is the exposure duration in years (70 for lifetime); and AT is the averag-  
631 ing time in hours (or 24 hr/day  $\times$  365 days/yr  $\times$  70 years).

632 Previous work has set a specific window of exposure to wildfire and assumed that exposure con-  
633 centrations (ECs) remain similar for the entire window<sup>40</sup>. We modify these assumptions in two  
634 key ways: (1) we use our predicted daily wildfire attributable concentrations of As, Ni, and Pb  
635 in each 10x10km grid cell over a five year period. As a result, we allow species’ concentrations  
636 to vary over the exposure window rather than assuming a constant concentration over that win-

637 dow, which is standard practice often due to limited air quality concentration sampling data; (2)  
638 we choose a much longer exposure window (5 years versus Dickinson et al. (2022) which uses  
639 30 days) to evaluate the risk from exposure to wildfire attributable concentrations. This longer  
640 exposure window has the advantage of being more representative of what a population may have  
641 been exposed to over that window and allows us to compare how exposure on average in the first  
642 five years (2006-2010) of our sample compared to the last five years (2016-2020) has changed  
643 population-level cancer risk.

644 We use our EC to calculate the cancer risk associated with each chronic exposure scenario for  
645 As, Ni, and Pb. We apply the standard equation used by the US EPA and other risk assessment  
646 agencies<sup>69-71</sup>:

$$\text{CancerRisk} = IUR * EC \tag{6}$$

647 where IUR ( $ug/m^3$ ) is the inhalation unit risk, which represents the incremental risk posed by  
648 a specific concentration unit in air. We use IUR values from the California's Office of Environ-  
649 mental Health Hazard Assessment (OEHHA), which are widely used in risk assessment practice  
650 because they are considered more protective than the current federal standards. For Pb, the IUR is  
651  $1.2E-5 m^3/ug$ , As IUR is  $3.3E-3 m^3/ug$ , and Ni is  $2.6E-4 m^3/ug$ <sup>67,71</sup>.

652 The calculated outcome of cancer risk represents the number of excess cases above the base-  
653 line cancer rate from exposure to wildfire attributable concentrations of individual carcinogenic  
654 species. When calculated at the population-scale (rather than the individual), this is known as the  
655 excess cancer burden. In practice, if cancer risk is calculated to be 0.0006, this can be interpreted  
656 as for every 10,000 people similarly exposed, we would expect 6 excess cases of cancer. Multi-  
657 plying this estimate of risk by the population will provide the excess cancer burden (or number of  
658 cancer cases) from exposure to a specific chemical<sup>67,71,72</sup>.

659 We calculate the population and regional cancer burden by multiplying the cancer risk in each  
660 grid cell by the population in that grid cell (gathered from Landscan data<sup>73</sup>. We re-weight the  
661 gridded population data using US state-wide census data as Landscan overestimates the popu-  
662 lation of the US by roughly 40,000,000 people and the US census data is not available at a grid-  
663 cell scale. We create weights by dividing the population in each year from the US census data  
664 by the Landscan data, and then multiply the weight by the population in each grid cell so that to-  
665 tal population from Landscan will match the total population in the US census data<sup>74</sup>. Once we  
666 multiply cancer risk by population (we assume the entire population in a grid cell is exposed to

667 wildfire smoke), we get a number that represents the excess cancer burden due to exposure. In  
668 other words, this is the expected number of excess cases above background in that grid cell for  
669 that exposure level and population. We then we sum across all grid cells in a region as well as  
670 across the US to produce a measure of cancer burden from exposure to each wildfire attributable  
671 carcinogenic concentration **Figure 5B**.

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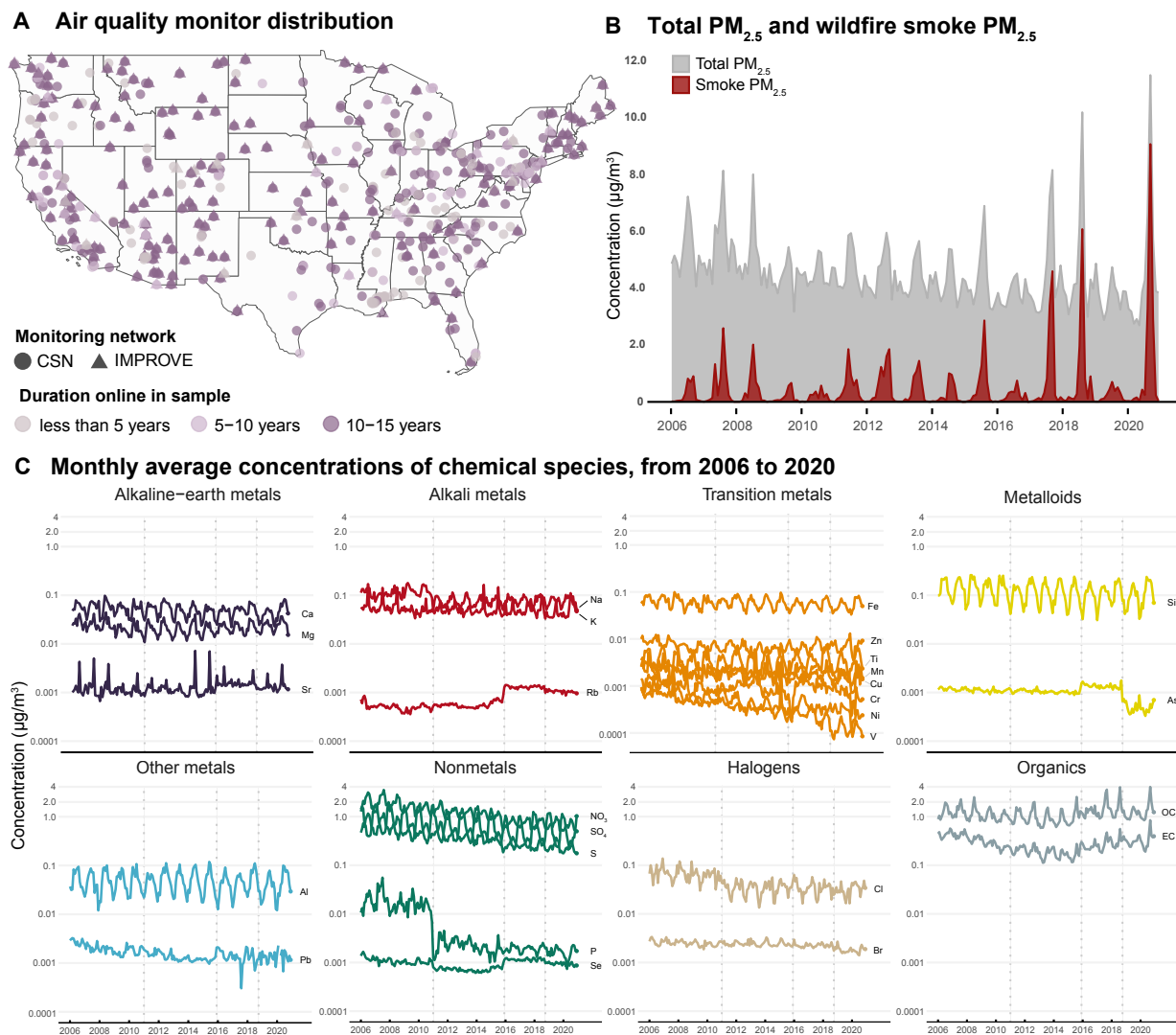
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**Figure 1: Temporal trends of total PM<sub>2.5</sub>, wildfire smoke PM<sub>2.5</sub>, and chemical species in the United States.** (A) Chemical Speciation Network (CSN) and IMPROVE air quality monitor locations across the United States, colored by duration each monitor is online over our sample period (2006 - 2020). Most monitors sample air quality for the full 15 years. (B) Monthly average concentrations ( $\mu\text{g}/\text{m}^3$ ) of total PM<sub>2.5</sub> and wildfire smoke PM<sub>2.5</sub> from 2006 to 2020. (C) Monthly average concentrations ( $\mu\text{g}/\text{m}^3$ ) of individual chemical species measured by CSN and IMPROVE monitors from 2006 to 2020. Dates at which monitoring networks had documented changes in their monitor methods or other operations are indicated in dashed lines (some changes may not be represented). There is clear seasonality in all of the chemical species.

### Effect of wildfire smoke on species' concentrations

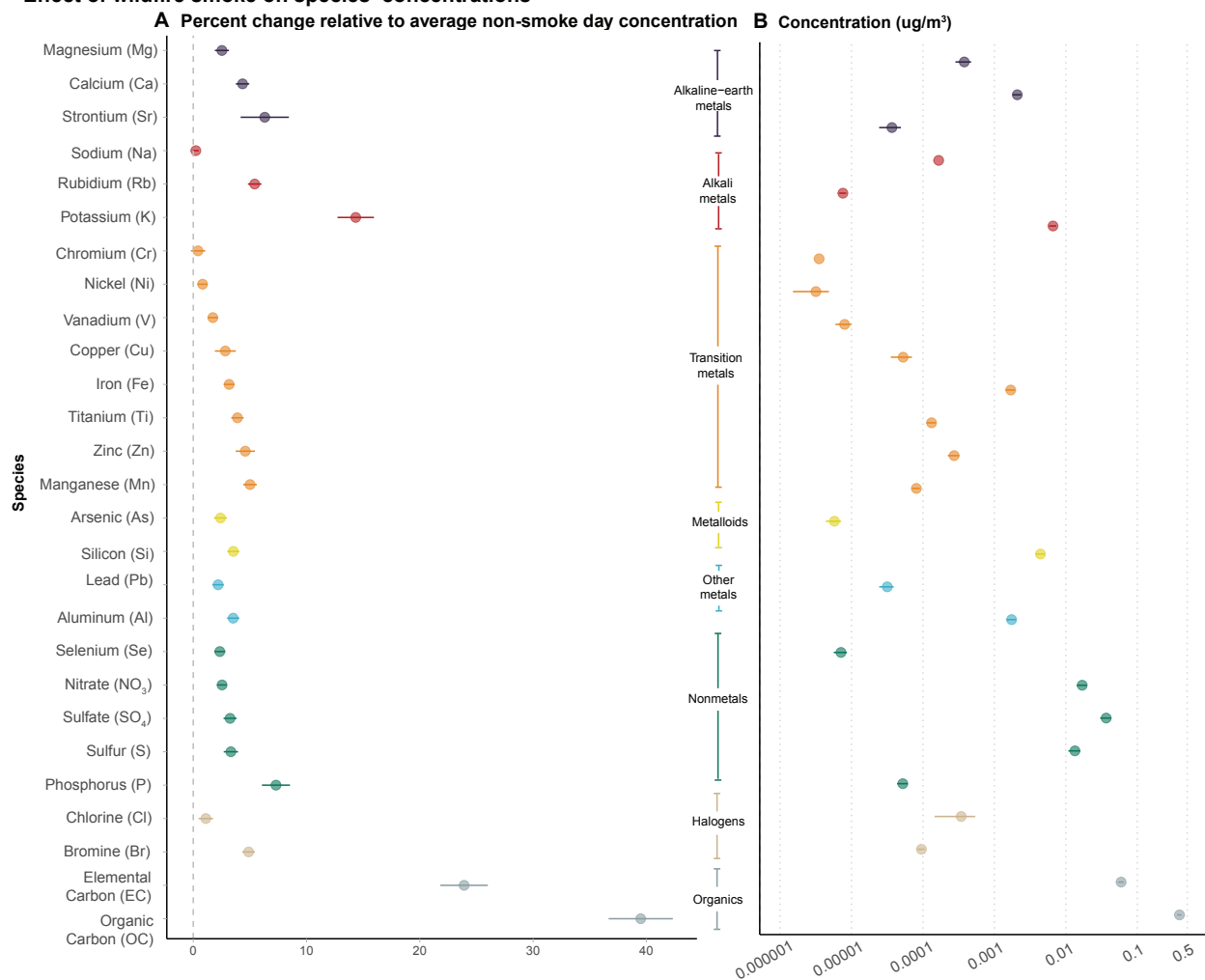
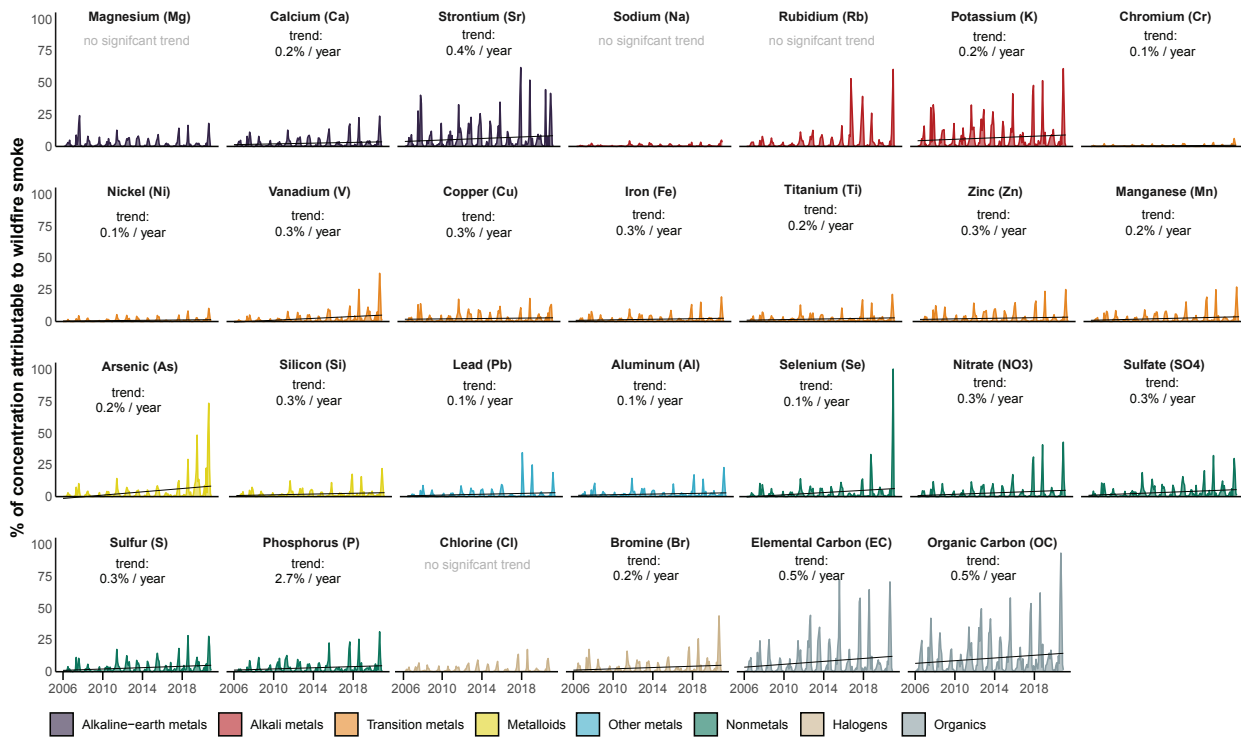


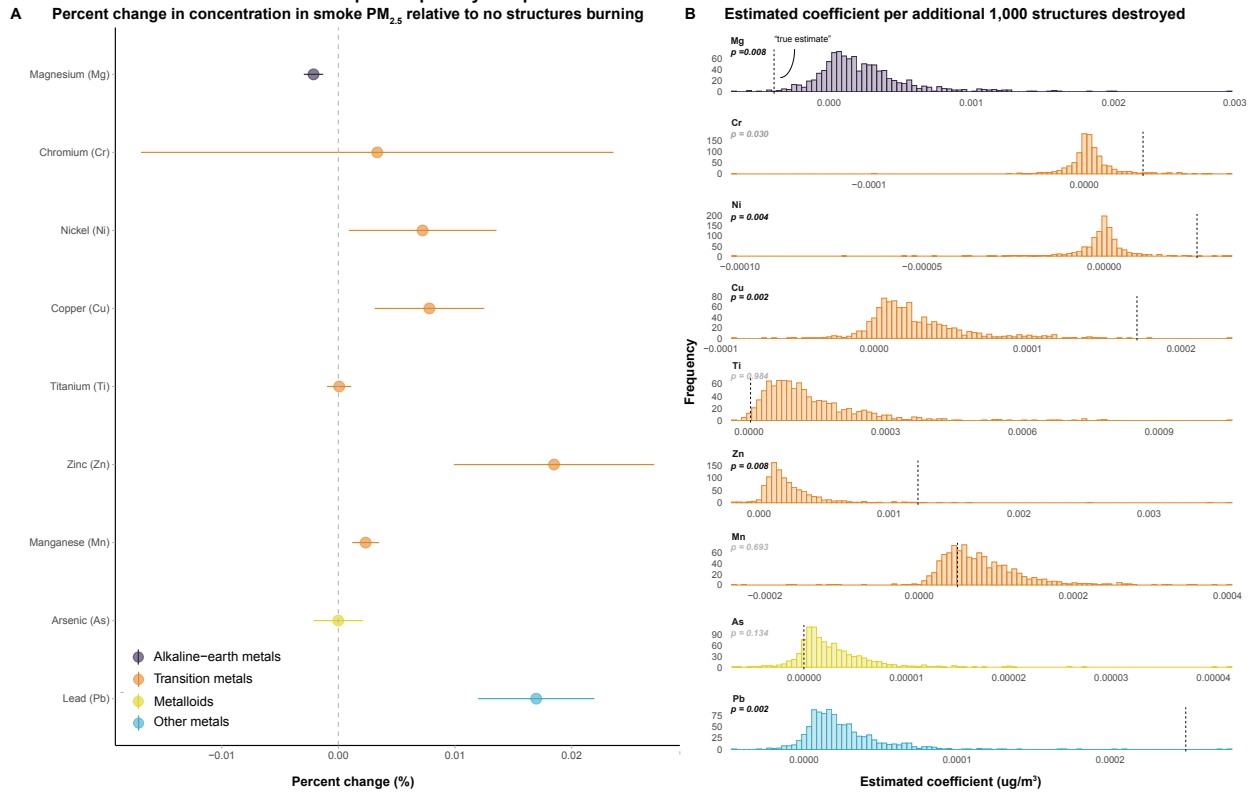
Figure 2: **Effect of wildfire smoke  $\text{PM}_{2.5}$  on chemical species' concentrations measured at monitoring stations across the contiguous US.** Species are categorized and colored by their chemical properties. (A) The percent change in a chemical species concentration relative to its non-smoke day sample average (i.e., relative to a day when there is no smoke in ambient air) given a one  $\mu\text{g}/\text{m}^3$  increase in wildfire smoke  $\text{PM}_{2.5}$ . The largest increases relative to baseline are in species associated with biomass burning, including OC, EC, and K (B) The concentration change due to an additional  $\mu\text{g}/\text{m}^3$  of wildfire smoke  $\text{PM}_{2.5}$  (log scale).

**Percent of concentration attributable to wildfire smoke  $PM_{2.5}$  is increasing for many species**



**Figure 3: The percentage of each species' concentration attributable to wildfire smoke  $PM_{2.5}$  is increasing for many species.** Predicted concentration of each species due to wildfire smoke over time as proportion of total observed species concentration. Monitor- and species-specific estimates are averaged to the monitor-month and then averaged across monitors to generate the time series shown in each panel. Annual rate of increase for each species due to wildfire smoke is estimated using linear regression, with percentage annual increase annotated in each panel. When wildfire is not driving an increasing fraction of a chemical species, there is no trend line or rate indicated, such as for Mg, Na, Cl.

**Effect of an additional structured burned per fire per day on species' concentrations**



**Figure 4: Smoke  $PM_{2.5}$  originating from wildfires that destroyed structures has higher concentrations of Ni, Cu, Zn, Mn, and Pb than smoke from wildfires that did not burn structures.** **A** Estimated effect of one additional structure burned in source fire on species concentration at affected monitor, for specific species previously associated with infrastructure burning. **B** Randomization inference test in which we estimate placebo “treatments” by randomly re-assigning structures burned across fires within each state and then re-estimating the effect of structures burned on species concentration (1,000 replications). Distribution of placebo treatment effects centered at zero, and the estimated treatment effect under true assignment (vertical dashed line) is unlikely to happen by chance (Ni, Cu, Zn, Pb, and Mg are significant at the 1% level).

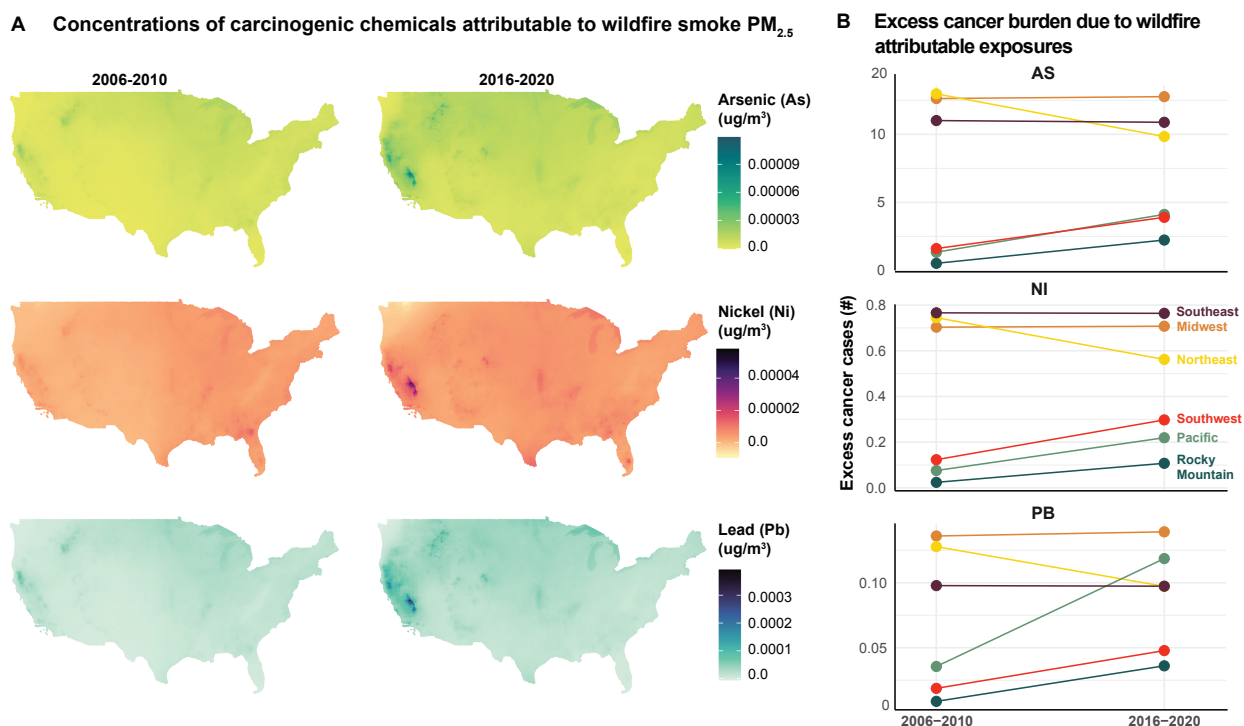
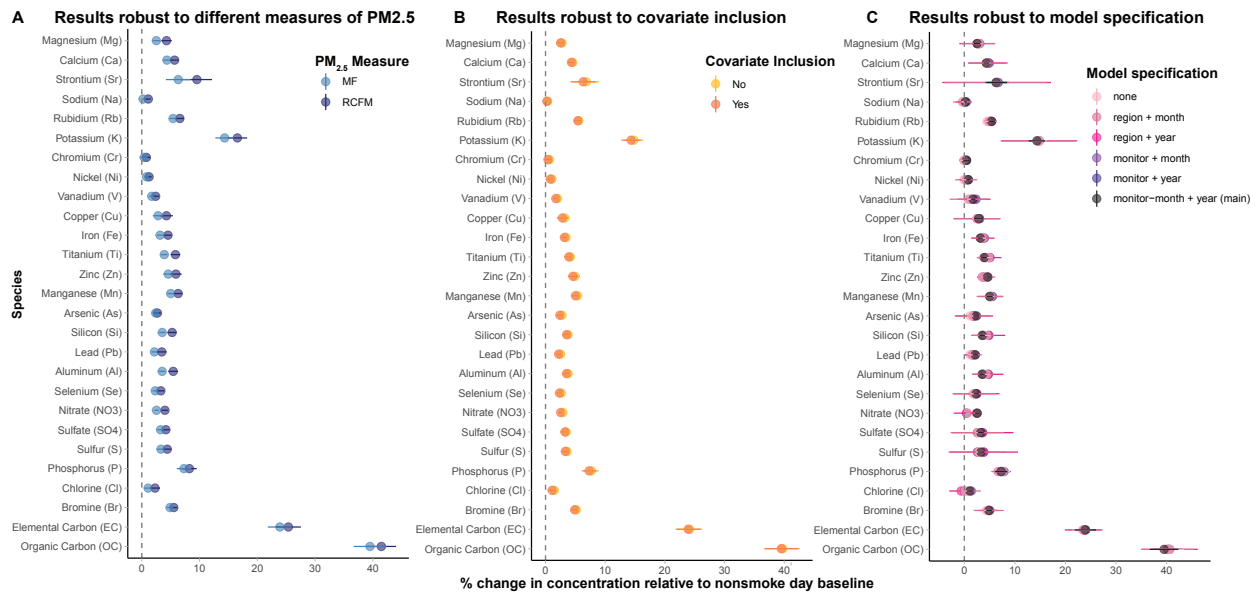


Figure 5: **Health risks from exposure due to wildfire attributable concentrations of carcinogenic species in smoke  $PM_{2.5}$ .** **A** Average concentrations of carcinogenic species (As, Ni, and Pb) attributable to wildfire smoke  $PM_{2.5}$  over two five-year periods (early versus late in our sample) across the US, with noticeable increases in the West and Rocky Mountain Regions. **B** Estimated excess cancer burden due to lifetime exposure to wildfire attributable concentrations observed early versus late in our sample by region. The cancer burden represents the excess number of cancer cases that would be expected above the background rate due to ambient exposure to specific chemicals. Regional differences show the direction and magnitude of difference in estimated cancer risk between exposure scenarios for early and late periods.

939 **Supplementary Information**

**Supplementary Table 1.** List of chemical species in our sample measured at IMPROVE and CSN monitoring sites from 2006 - 2020.

<b>Species type</b>	<b>Species</b>
Alkaline-earth metals	Magnesium (Mg)
	Calcium (Ca)
	Strontium (Sr)
Alkali metals	Sodium (Na)
	Rubidium (Rb)
	Potassium (K)
Transition metals	Chromium (Cr)
	Nickel (Ni)
	Vanadium (V)
	Copper (Cu)
	Iron (Fe)
	Zinc (Zn)
	Manganese (Mn)
Metalloids	Arsenic (As)
	Silicon (Si)
Other metals	Lead (Pb)
	Aluminum (Al)
	Titanium (Ti)
Nonmetals	Selenium (Se)
	Nitrate (NO <sub>3</sub> )
	Sulfate (SO <sub>4</sub> )
	Sulfur (S)
	Phosphorus (P)
Halogens	Chlorine (Cl)
	Bromine (Br)
Organics	Elemental Carbon (EC)
	Organic Carbon (OC)



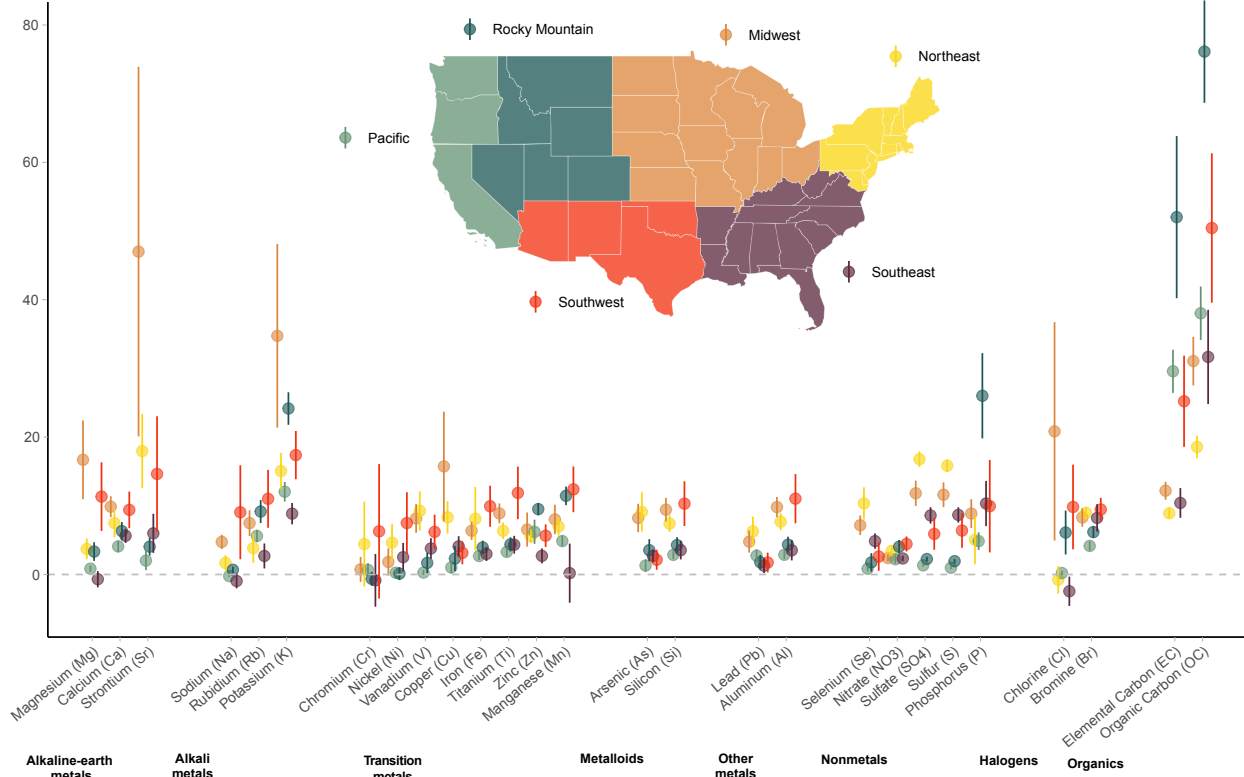
**Supplementary Figure 1. Results are robust to various sensitivity analyses. A** Robustness checks for using different measures of Total PM<sub>2.5</sub>. Using our main model specification, we estimate coefficients for each species using reconstructed fine mass (RCFM) and compare this to our results using measured fine mass (MF) in the main text. We find there our results are quantitatively and qualitatively similar. **B** Robustness checks for including non-smoke PM<sub>2.5</sub> as a covariate in our model. We find that there are no significant differences in our estimates when directly controlling for non-smoke PM<sub>2.5</sub>. **C** Robustness checks for model specification. We find that there are no significant differences in our estimates when trying alternative model specifications.

**Supplementary Table 2. Main model results.** Raw estimates correspond to Figure 2B, which is plotted in concentration ( $\mu\text{g}/\text{m}^3$ ) levels. Normalized estimates are our main model coefficients divided by the baseline non-smoke day concentration and multiplied by 100 to provide a measure of the percent change in each species concentration due to smoke  $\text{PM}_{2.5}$ , and correspond to Figure 2A.

Species	Estimate (raw)	Std. Error (clustered)	t-stat	p-value	Baseline non-smoke conc ( $\mu\text{g}/\text{m}^3$ )	Estimate (normalized)	2.5% CI	97.5% CI
Al	0.0017	0.0001	12.8858	0.0000	0.0491	3.5563	3.0144	4.0982
As	0.0000	0.0000	8.0861	0.0000	0.0002	2.3897	1.8094	2.9699
Br	0.0001	0.0000	17.1362	0.0000	0.0020	4.8568	4.3003	5.4133
Ca	0.0021	0.0001	14.3860	0.0000	0.0481	4.3366	3.7447	4.9285
Cl	0.0003	0.0001	3.4290	0.0006	0.0311	1.1194	0.4784	1.7604
Cr	0.0000	0.0000	1.2843	0.1995	0.0009	0.4079	-0.2157	1.0314
Cu	0.0001	0.0000	5.9792	0.0000	0.0019	2.8076	1.8856	3.7296
EC	0.0595	0.0027	21.6658	0.0000	0.2509	23.6965	21.5488	25.8442
Fe	0.0017	0.0001	13.0489	0.0000	0.0537	3.1617	2.6860	3.6375
K	0.0066	0.0004	17.3209	0.0000	0.0462	14.3469	12.7206	15.9732
Mg	0.0004	0.0000	7.9256	0.0000	0.0149	2.5444	1.9140	3.1748
Mn	0.0001	0.0000	16.2403	0.0000	0.0016	4.9821	4.3798	5.5844
Na	0.0002	0.0001	1.6966	0.0902	0.0726	0.2188	-0.0344	0.4720
Ni	0.0000	0.0000	3.7044	0.0002	0.0004	0.8091	0.3803	1.2380
NO <sub>3</sub>	0.0170	0.0014	12.2009	0.0000	0.6726	2.5324	2.1248	2.9400
OC	0.3903	0.0148	26.3438	0.0000	0.9895	39.4472	36.5054	42.3890
P	0.0001	0.0000	11.3754	0.0000	0.0007	7.1348	5.9033	8.3664
Pb	0.0000	0.0000	8.5022	0.0000	0.0015	2.1866	1.6817	2.6916
Rb	0.0000	0.0000	17.6088	0.0000	0.0001	5.3946	4.7931	5.9962
S	0.0132	0.0013	10.2264	0.0000	0.4013	3.2959	2.6631	3.9287
Se	0.0000	0.0000	9.3185	0.0000	0.0003	2.3113	1.8243	2.7983
Si	0.0044	0.0003	12.9684	0.0000	0.1238	3.5671	3.0270	4.1072
SO <sub>4</sub>	0.0363	0.0033	11.0922	0.0000	1.1283	3.2184	2.6487	3.7881
Sr	0.0000	0.0000	5.7872	0.0000	0.0006	6.3287	4.1815	8.4759
Ti	0.0001	0.0000	13.9269	0.0000	0.0034	3.9198	3.3672	4.4725
V	0.0000	0.0000	7.5670	0.0000	0.0005	1.7224	1.2755	2.1694
Zn	0.0003	0.0000	10.5446	0.0000	0.0060	4.5619	3.7125	5.4114



**Regional variation in the impact of wildfire smoke PM<sub>2.5</sub> on species' concentrations**



**Supplementary Figure 2: Regional comparison of selected species concentration in wildfire smoke measured at CSN and IMPROVE sites in the contiguous US. For a one unit increase in smoke PM<sub>2.5</sub>, the corresponding percent change ( $\mu\text{g}/\text{m}^3$ ) relative to the regional sample average for a given chemical species on non-smoke days.**

**Supplementary Table 3.** Regional estimates of the impact of wildfire smoke on species concentrations

Species	Estimate ( $\mu\text{g}/\text{m}^3$ )	SE	p-value ( $\mu\text{g}/\text{m}^3$ )	Baseline conc. (%)	Norm est (%)	Norm CI25 (%)	Norm CI975
<b>MIDWEST</b>							
Al	0.0031	0.0002	0.0000	0.0311	9.8791	8.4335	11.3248
As	0.0000	0.0000	0.0000	0.0003	8.2685	6.1965	10.3404
Br	0.0002	0.0000	0.0000	0.0023	8.3283	6.9744	9.6822
Ca	0.0056	0.0004	0.0000	0.0560	9.9399	8.4076	11.4722
Cl	0.0025	0.0010	0.0112	0.0120	20.9204	4.8238	37.0170
Cr	0.0000	0.0000	0.4322	0.0011	0.7305	-1.1032	2.5641
Cu	0.0003	0.0001	0.0001	0.0020	15.8906	7.8587	23.9224
EC	0.0339	0.0019	0.0000	0.2838	11.9385	10.5904	13.2867
Fe	0.0039	0.0004	0.0000	0.0622	6.3348	5.0006	7.6690
K	0.0174	0.0034	0.0000	0.0492	35.3996	21.6950	49.1041
Mg	0.0018	0.0003	0.0000	0.0103	17.1306	11.2179	23.0433
Mn	0.0002	0.0000	0.0000	0.0021	8.0250	5.8572	10.1927
Na	0.0017	0.0002	0.0000	0.0355	4.7668	3.7325	5.8011
Ni	0.0000	0.0000	0.0846	0.0004	1.7780	-0.2457	3.8017
NO <sub>3</sub>	0.0319	0.0036	0.0000	1.3267	2.4051	1.8662	2.9440
OC	0.3250	0.0191	0.0000	1.0728	30.2965	26.7577	33.8353
P	0.0000	0.0000	0.0000	0.0005	8.5165	6.4916	10.5415
Pb	0.0001	0.0000	0.0000	0.0022	4.8072	3.1755	6.4390
Rb	0.0000	0.0000	0.0000	0.0002	7.3953	5.4835	9.3071
S	0.0627	0.0048	0.0000	0.5365	11.6928	9.9314	13.4542
Se	0.0000	0.0000	0.0000	0.0005	7.1741	5.7518	8.5964
Si	0.0084	0.0008	0.0000	0.0883	9.5021	7.7384	11.2659
SO <sub>4</sub>	0.1800	0.0138	0.0000	1.5176	11.8599	10.0559	13.6638
Sr	0.0002	0.0001	0.0007	0.0005	48.2687	20.6644	75.8730
Ti	0.0002	0.0000	0.0000	0.0022	8.8641	7.4726	10.2556
V	0.0000	0.0000	0.0000	0.0004	8.1921	6.0849	10.2993
Zn	0.0007	0.0001	0.0000	0.0105	6.5610	4.0452	9.0767

<b>Species</b>	<b>Estimate</b> ( $\mu\text{g}/\text{m}^3$ )	<b>SE</b>	<b>p-value</b> ( $\mu\text{g}/\text{m}^3$ )	<b>Baseline conc.</b> (%)	<b>Norm est</b> (%)	<b>Norm CI25</b> (%)	<b>Norm CI975</b>
<b>NORTHEAST</b>							
Al	0.0015	0.0001	0.0000	0.0194	7.8136	6.5548	9.0724
As	0.0000	0.0000	0.0000	0.0004	9.1260	6.2381	12.0138
Br	0.0002	0.0000	0.0000	0.0022	8.9673	8.0504	9.8842
Ca	0.0019	0.0003	0.0000	0.0250	7.4338	5.3532	9.5145
Cl	-0.0003	0.0004	0.4481	0.0421	-0.7605	-2.7420	1.2209
Cr	0.0001	0.0000	0.1610	0.0013	4.3893	-1.7773	10.5558
Cu	0.0002	0.0000	0.0000	0.0023	8.2895	5.9400	10.6390
EC	0.0317	0.0013	0.0000	0.3613	8.7783	8.0685	9.4881
Fe	0.0036	0.0010	0.0008	0.0444	8.0912	3.4549	12.7276
K	0.0067	0.0006	0.0000	0.0443	15.1059	12.4258	17.7861
Mg	0.0004	0.0001	0.0000	0.0115	3.6899	2.1692	5.2106
Mn	0.0001	0.0000	0.0000	0.0012	6.9762	5.6314	8.3211
Na	0.0014	0.0005	0.0050	0.0866	1.6590	0.5139	2.8041
Ni	0.0000	0.0000	0.0010	0.0009	4.6014	1.9074	7.2955
NO <sub>3</sub>	0.0258	0.0029	0.0000	0.7612	3.3956	2.6496	4.1415
OC	0.2200	0.0103	0.0000	1.1882	18.5181	16.8002	20.2360
P	0.0000	0.0000	0.0105	0.0009	4.6075	1.1049	8.1102
Pb	0.0001	0.0000	0.0000	0.0020	6.3110	4.1456	8.4765
Rb	0.0000	0.0000	0.0006	0.0001	3.8450	1.6979	5.9921
S	0.0878	0.0026	0.0000	0.5522	15.9051	14.9552	16.8551
Se	0.0000	0.0000	0.0000	0.0005	10.4067	8.0097	12.8038
Si	0.0038	0.0003	0.0000	0.0513	7.4335	6.1965	8.6705
SO <sub>4</sub>	0.2630	0.0090	0.0000	1.5694	16.7598	15.6247	17.8948
Sr	0.0001	0.0000	0.0000	0.0004	18.0403	12.5859	23.4946
Ti	0.0001	0.0000	0.0000	0.0019	6.3618	5.1564	7.5672
V	0.0001	0.0000	0.0000	0.0007	9.2215	6.3536	12.0894
Zn	0.0005	0.0000	0.0000	0.0084	5.5566	4.5502	6.5630

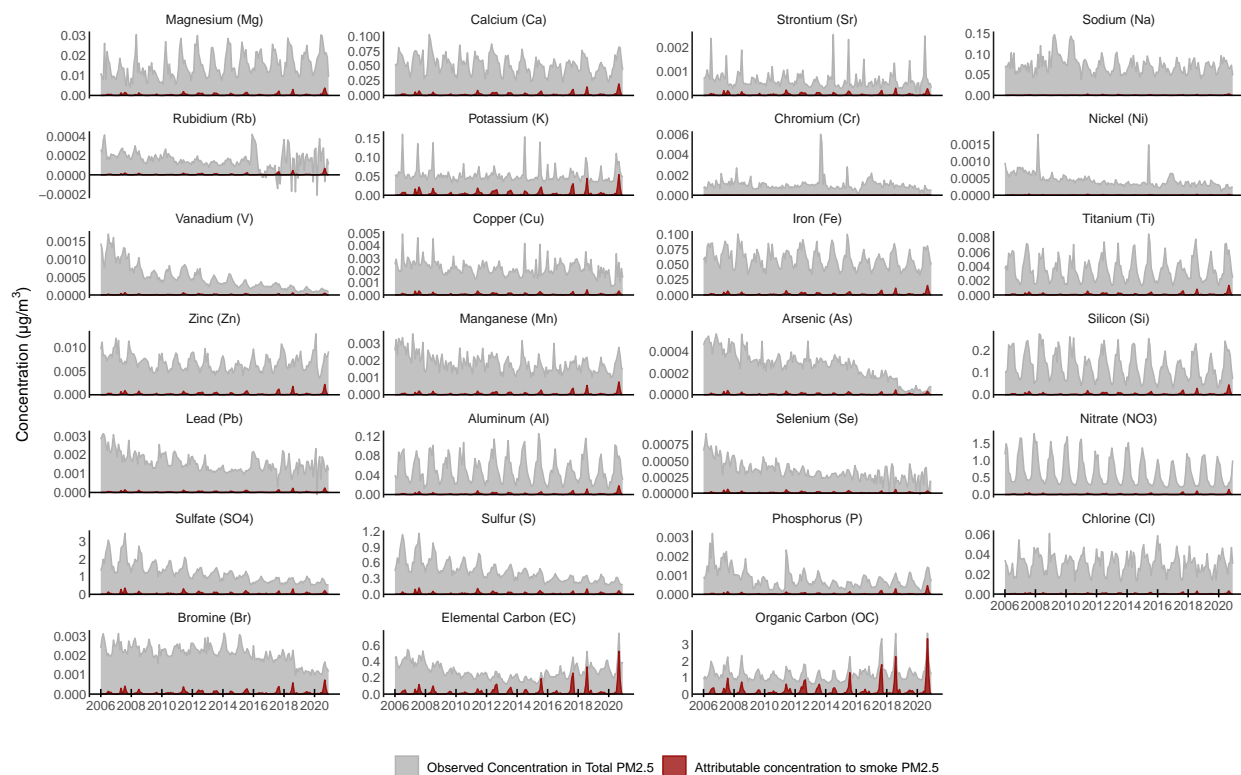
<b>Species</b>	<b>Estimate</b> ( $\mu\text{g}/\text{m}^3$ )	<b>SE</b>	<b>p-value</b> ( $\mu\text{g}/\text{m}^3$ )	<b>Baseline conc.</b> (%)	<b>Norm est</b> (%)	<b>Norm CI25</b> (%)	<b>Norm CI975</b>
<b>PACIFIC</b>							
Al	0.0012	0.0001	0.0000	0.0400	2.9092	2.2639	3.5546
As	0.0000	0.0000	0.0056	0.0002	1.2717	0.3798	2.1636
Br	0.0001	0.0000	0.0000	0.0017	4.1507	3.4855	4.8159
Ca	0.0013	0.0001	0.0000	0.0316	4.1016	3.2549	4.9483
Cl	0.0001	0.0001	0.1898	0.0677	0.2075	-0.1040	0.5191
Cr	0.0000	0.0000	0.1712	0.0005	0.6649	-0.2913	1.6211
Cu	0.0000	0.0000	0.0215	0.0020	1.0159	0.1524	1.8794
EC	0.0642	0.0036	0.0000	0.2191	29.2888	26.0448	32.5329
Fe	0.0012	0.0001	0.0000	0.0455	2.6978	2.1005	3.2950
K	0.0052	0.0003	0.0000	0.0431	11.9852	10.5594	13.4110
Mg	0.0002	0.0000	0.0004	0.0191	0.8490	0.3901	1.3079
Mn	0.0001	0.0000	0.0000	0.0014	4.8449	4.0643	5.6256
Na	-0.0003	0.0001	0.0014	0.1189	-0.2342	-0.3755	-0.0929
Ni	0.0000	0.0000	0.2874	0.0004	0.2604	-0.2219	0.7427
NO <sub>3</sub>	0.0162	0.0025	0.0000	0.7395	2.1897	1.5307	2.8486
OC	0.3746	0.0202	0.0000	0.9886	37.8937	33.8351	41.9524
P	0.0001	0.0000	0.0000	0.0011	4.7655	3.4913	6.0398
Pb	0.0000	0.0000	0.0000	0.0010	2.7287	1.8070	3.6504
Rb	0.0000	0.0000	0.0000	0.0001	5.5915	4.6852	6.4978
S	0.0022	0.0002	0.0000	0.2209	0.9748	0.7811	1.1686
Se	0.0000	0.0000	0.0000	0.0002	0.8331	0.4436	1.2226
Si	0.0029	0.0003	0.0000	0.0997	2.9025	2.2439	3.5610
SO <sub>4</sub>	0.0080	0.0006	0.0000	0.6046	1.3219	1.1415	1.5024
Sr	0.0000	0.0000	0.0048	0.0006	2.0117	0.6236	3.3999
Ti	0.0001	0.0000	0.0000	0.0032	3.3196	2.6248	4.0144
V	0.0000	0.0000	0.1302	0.0005	0.2831	-0.0847	0.6509
Zn	0.0003	0.0000	0.0000	0.0042	6.1736	4.4220	7.9253

Species	Estimate ( $\mu\text{g}/\text{m}^3$ )	SE	p-value ( $\mu\text{g}/\text{m}^3$ )	Baseline conc. (%)	Norm est (%)	Norm CI25 (%)	Norm CI975
<b>ROCKY MOUNTAIN</b>							
Al	0.0020	0.0003	0.0000	0.0468	4.3105	3.1458	5.4751
As	0.0000	0.0000	0.0000	0.0001	3.4709	1.8523	5.0896
Br	0.0001	0.0000	0.0000	0.0012	6.1225	5.2596	6.9855
Ca	0.0025	0.0003	0.0000	0.0406	6.2350	4.9377	7.5322
Cl	0.0003	0.0001	0.0003	0.0052	6.0883	2.8772	9.2994
Cr	-0.0000	0.0000	0.2112	0.0004	-0.6322	-1.6294	0.3650
Cu	0.0000	0.0000	0.0137	0.0006	2.2834	0.4790	4.0878
EC	0.0583	0.0067	0.0000	0.1139	51.2289	39.5167	62.9411
Fe	0.0014	0.0002	0.0000	0.0371	3.8952	2.9128	4.8777
K	0.0070	0.0003	0.0000	0.0293	24.0908	21.7285	26.4530
Mg	0.0004	0.0001	0.0000	0.0113	3.3053	1.9524	4.6583
Mn	0.0001	0.0000	0.0000	0.0009	11.3706	10.0200	12.7212
Na	0.0001	0.0000	0.0038	0.0192	0.6804	0.2248	1.1360
Ni	0.0000	0.0000	0.7787	0.0001	0.1327	-0.8026	1.0681
NO <sub>3</sub>	0.0119	0.0015	0.0000	0.2966	4.0100	3.0119	5.0082
OC	0.4439	0.0220	0.0000	0.5871	75.6151	68.1609	83.0693
P	0.0000	0.0000	0.0000	0.0002	24.9478	18.9425	30.9530
Pb	0.0000	0.0000	0.0015	0.0005	1.7672	0.6959	2.8385
Rb	0.0000	0.0000	0.0000	0.0001	9.1060	7.4208	10.7912
S	0.0028	0.0002	0.0000	0.1473	1.9319	1.6423	2.2216
Se	0.0000	0.0000	0.0158	0.0001	1.6413	0.3152	2.9673
Si	0.0053	0.0007	0.0000	0.1255	4.2554	3.0882	5.4226
SO <sub>4</sub>	0.0091	0.0007	0.0000	0.4020	2.2521	1.9256	2.5787
Sr	0.0000	0.0000	0.0000	0.0004	4.0207	2.7143	5.3271
Ti	0.0001	0.0000	0.0000	0.0028	4.3568	3.3326	5.3811
V	0.0000	0.0000	0.0308	0.0001	1.6720	0.1579	3.1862
Zn	0.0002	0.0000	0.0000	0.0020	9.4327	8.5337	10.3316

<b>Species</b>	<b>Estimate</b> ( $\mu\text{g}/\text{m}^3$ )	<b>SE</b>	<b>p-value</b> ( $\mu\text{g}/\text{m}^3$ )	<b>Baseline conc.</b> (%)	<b>Norm est</b> (%)	<b>Norm CI25</b> (%)	<b>Norm CI975</b>
<b>SOUTHEAST</b>							
Al	0.0019	0.0004	0.0000	0.0498	3.7429	2.2037	5.2822
As	0.0000	0.0000	0.0000	0.0004	2.7697	1.8163	3.7230
Br	0.0002	0.0000	0.0000	0.0025	8.0737	6.0058	10.1417
Ca	0.0021	0.0002	0.0000	0.0366	5.8089	4.7670	6.8508
Cl	-0.0009	0.0004	0.0237	0.0329	-2.6555	-4.9506	-0.3604
Cr	-0.0000	0.0000	0.6323	0.0016	-0.9254	-4.7405	2.8898
Cu	0.0001	0.0000	0.0000	0.0025	4.1221	2.6286	5.6156
EC	0.0386	0.0043	0.0000	0.3789	10.1805	7.9462	12.4148
Fe	0.0017	0.0003	0.0000	0.0552	3.0867	2.0300	4.1433
K	0.0049	0.0004	0.0000	0.0549	9.0026	7.3939	10.6112
Mg	-0.0001	0.0001	0.2925	0.0137	-0.6343	-1.8221	0.5534
Mn	0.0000	0.0001	0.9420	0.0023	0.1595	-4.1638	4.4829
Na	-0.0010	0.0005	0.0544	0.0951	-1.0300	-2.0798	0.0198
Ni	0.0000	0.0000	0.0198	0.0004	2.5830	0.4164	4.7497
NO <sub>3</sub>	0.0121	0.0011	0.0000	0.5127	2.3573	1.9394	2.7752
OC	0.4794	0.0552	0.0000	1.5227	31.4836	24.2496	38.7176
P	0.0001	0.0000	0.0000	0.0006	10.0159	6.6120	13.4198
Pb	0.0000	0.0000	0.0202	0.0020	1.2845	0.2032	2.3659
Rb	0.0000	0.0000	0.0053	0.0002	2.6829	0.8122	4.5536
S	0.0611	0.0040	0.0000	0.7012	8.7198	7.5967	9.8430
Se	0.0000	0.0000	0.0000	0.0005	4.8679	3.7037	6.0320
Si	0.0043	0.0008	0.0000	0.1165	3.7139	2.3137	5.1141
SO <sub>4</sub>	0.1716	0.0125	0.0000	1.9731	8.6992	7.4459	9.9525
Sr	0.0000	0.0000	0.0001	0.0005	6.1818	3.2081	9.1555
Ti	0.0002	0.0000	0.0000	0.0036	4.5132	3.1702	5.8563
V	0.0000	0.0000	0.0000	0.0006	4.0071	2.4073	5.6069
Zn	0.0002	0.0000	0.0000	0.0077	2.7151	1.6073	3.8228

<b>Species</b>	<b>Estimate</b> ( $\mu\text{g}/\text{m}^3$ )	<b>SE</b>	<b>p-value</b> ( $\mu\text{g}/\text{m}^3$ )	<b>Baseline conc.</b> (%)	<b>Norm est</b> (%)	<b>Norm CI25</b> (%)	<b>Norm CI975</b>
<b>SOUTHWEST</b>							
Al	0.0123	0.0020	0.0000	0.1109	11.1128	7.5098	14.7158
As	0.0000	0.0000	0.0048	0.0002	2.1270	0.6642	3.5899
Br	0.0002	0.0000	0.0000	0.0020	9.4107	7.7182	11.1032
Ca	0.0096	0.0014	0.0000	0.1024	9.4036	6.7272	12.0801
Cl	0.0019	0.0006	0.0021	0.0196	9.7994	3.6575	15.9413
Cr	0.0000	0.0000	0.2063	0.0003	6.2211	-3.4869	15.9291
Cu	0.0001	0.0000	0.0003	0.0020	3.1405	1.4766	4.8045
EC	0.0516	0.0069	0.0000	0.2055	25.0889	18.4435	31.7344
Fe	0.0080	0.0012	0.0000	0.0804	9.9637	6.9654	12.9620
K	0.0101	0.0010	0.0000	0.0582	17.4137	13.8801	20.9474
Mg	0.0027	0.0006	0.0000	0.0234	11.3764	6.3452	16.4076
Mn	0.0002	0.0000	0.0000	0.0020	12.4369	9.0916	15.7823
Na	0.0071	0.0027	0.0104	0.0782	9.0828	2.1915	15.9740
Ni	0.0000	0.0000	0.0014	0.0002	7.4687	2.9692	11.9683
NO <sub>3</sub>	0.0148	0.0018	0.0000	0.3325	4.4508	3.3574	5.5443
OC	0.3974	0.0432	0.0000	0.7896	50.3371	39.4353	61.2390
P	0.0001	0.0000	0.0042	0.0010	9.8600	3.1921	16.5278
Pb	0.0000	0.0000	0.0189	0.0011	1.7285	0.2921	3.1649
Rb	0.0000	0.0000	0.0000	0.0002	11.1188	6.8484	15.3893
S	0.0205	0.0040	0.0000	0.3175	6.4552	3.9224	8.9880
Se	0.0000	0.0000	0.0138	0.0002	2.6388	0.5521	4.7255
Si	0.0279	0.0045	0.0000	0.2689	10.3920	7.1039	13.6801
SO <sub>4</sub>	0.0528	0.0104	0.0000	0.8921	5.9154	3.6103	8.2205
Sr	0.0002	0.0000	0.0009	0.0011	14.6368	6.2049	23.0687
Ti	0.0008	0.0001	0.0000	0.0067	11.9170	8.0799	15.7540
V	0.0000	0.0000	0.0000	0.0005	6.2412	3.6890	8.7935
Zn	0.0002	0.0000	0.0000	0.0037	5.6142	3.9725	7.2560

### Chemical species concentration attributable to wildfire smoke PM<sub>2.5</sub>

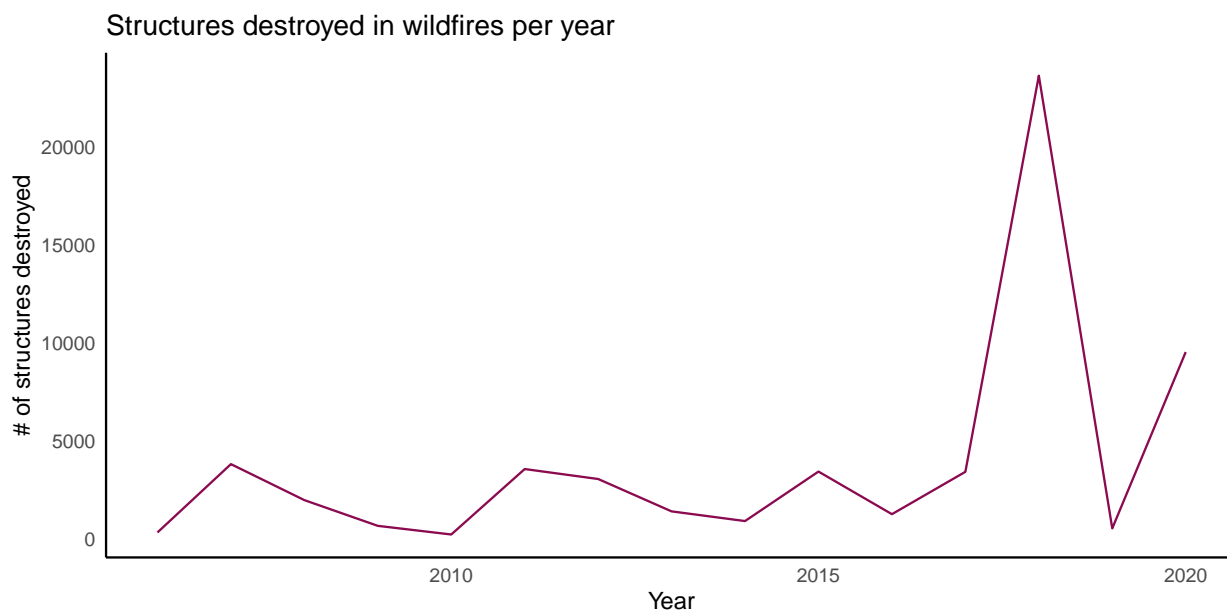


**Supplementary Figure 3: The concentration of each species attributable to wildfire smoke PM<sub>2.5</sub>.** Observed average species concentration across ground monitors is shown in grey, and attributed portion of the concentration due to wildfire smoke is shown in red. Concentrations are averaged to the monitor, year, and month and then averaged across all monitors for a given year and month to depict the monthly time series over our sample. For some chemical species, such as As, EC and OC, the attributable fraction due to wildfire smoke PM<sub>2.5</sub> is increasing over time.



**Supplementary Table 4. Average annual concentration attributable to wildfire in 2006 and 2020.** The multiple is calculated by dividing the average annual wildfire attributable concentration for each species in 2020 by that in 2006.

<b>Species</b>	<b>2006 conc. (<math>\mu\text{g}/\text{m}^3</math>)</b>	<b>2020 conc. (<math>\mu\text{g}/\text{m}^3</math>)</b>	<b>Increase multiple</b>
Al	0.0006158	0.0033442	5.4
As	0.0000031	0.0000062	2.0
Br	0.0000347	0.0001262	3.6
Ca	0.0008030	0.0034977	4.4
Cl	0.0001659	0.0005888	3.5
Cr	0.0000011	0.0000058	5.2
Cu	0.0000305	0.0000621	2.0
EC	0.0126064	0.0774973	6.1
Fe	0.0006225	0.0028002	4.45
K	0.0022521	0.0089506	4.0
Mg	0.0001763	0.0007078	4.0
Mn	0.0000249	0.0001240	5.0
Na	0.0001574	0.0008456	5.4
Ni	0.0000018	0.0000045	2.6
NO <sub>3</sub>	0.0049829	0.0222207	4.5
OC	0.0993863	0.5194907	5.2
P	0.0000134	0.0000724	5.4
Pb	0.0000126	0.0000384	3.1
Rb	0.0000021	0.0000111	5.3
S	0.0085559	0.0152454	1.8
Se	0.0000044	0.0000074	1.7
Si	0.0015560	0.0081340	5.2
SO <sub>4</sub>	0.0249822	0.0451040	1.8
Sr	0.0000213	0.0000558	2.6
Ti	0.0000435	0.0002384	5.5
V	0.0000045	0.0000109	2.4
Zn	0.0000901	0.0003476	3.9



**Supplementary Figure 4.** Number of burned structures per year in our sample.

**Supplementary Table 5.** Estimates for the effect of an additional burned structure on species concentration in smoke.

Species	Baseline conc. ( $\mu\text{g}/\text{m}^3$ )	Estimate ( $\mu\text{g}/\text{m}^3$ )	CI25 ( $\mu\text{g}/\text{m}^3$ )	CI975 ( $\mu\text{g}/\text{m}^3$ )	Est (%)	CI25 (%)	CI975 (%)
As	2.57E-04	-2.71E-11	-5.42E-09	5.37E-09	-1.05E-05	-2.11E-03	2.09E-03
Cr	8.19E-04	2.73E-08	-1.39E-07	1.93E-07	3.33E-03	-1.69E-02	2.36E-02
Cu	2.19E-03	1.71E-07	6.80E-08	2.74E-07	7.80E-03	3.10E-03	1.25E-02
Mg	1.82E-02	-3.87E-07	-5.38E-07	-2.37E-07	-2.13E-03	-2.96E-03	-1.30E-03
Mn	2.13E-03	4.96E-08	2.54E-08	7.38E-08	2.33E-03	1.19E-03	3.47E-03
Ni	3.52E-04	2.54E-08	3.11E-09	4.78E-08	7.22E-03	8.82E-04	1.36E-02
Pb	1.47E-03	2.49E-07	1.76E-07	3.22E-07	1.70E-02	1.20E-02	2.19E-02
Ti	4.74E-03	3.16E-09	-4.61E-08	5.24E-08	6.66E-05	-9.71E-04	1.10E-03
Zn	6.61E-03	1.22E-06	6.54E-07	1.79E-06	1.85E-02	9.90E-03	2.71E-02

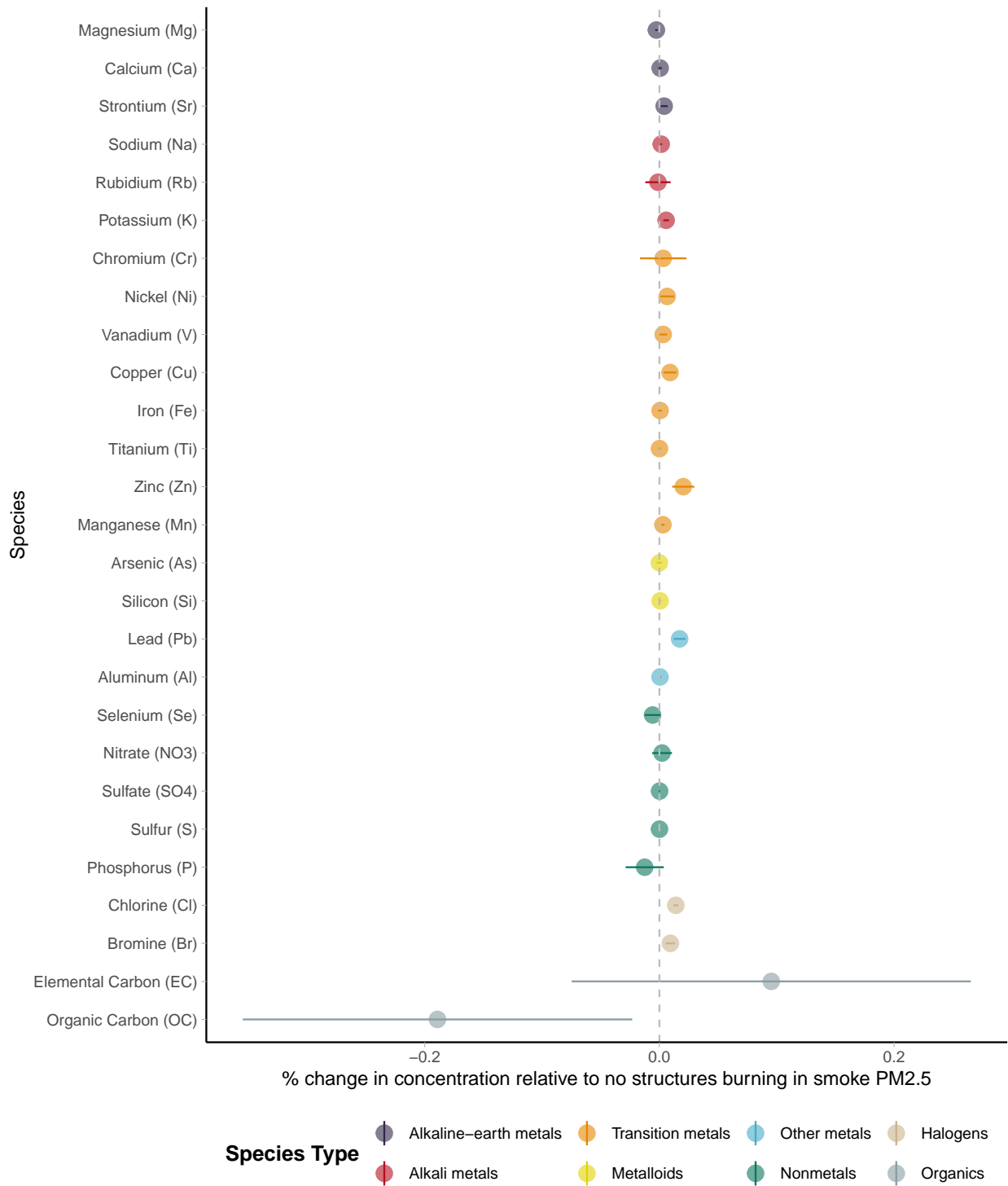
**Supplementary Table 6.** Randomization inference results. Sharp null hypothesis of no effect of a burned structure on species' concentrations ( $H_0$ : Average effect = 0). Comparison of our estimated coefficient (observed estimate) for each species against the distribution of permuted treatment assignment, with the calculated p-value for each comparison.

<b>Species</b>	<b>Observed Estimate</b> ( $\mu\text{g}/\text{m}^3$ )	<b>p-value</b>
Pb	2.49E-07	2.00E-03
Cu	1.71E-07	2.00E-03
Ni	2.54E-08	4.00E-03
Zn	1.22E-06	8.00E-03
Mg	-3.87E-07	8.00E-03
Cr	2.73E-08	3.00E-02
As	-2.71E-11	1.34E-01
Mn	4.96E-08	6.93E-01
Ti	3.16E-09	9.84E-01

**Supplementary Table 7.** Excess cancer burden due to wildfire attributable exposures by region. Expected excess cancer cases due to a lifetime exposure based on the concentrations in two different five-year periods: 2006 - 2010 and 2016 - 2020. Total cases for a given time period and chemical are shown in the last two columns.

<b>Region</b>	<b>2006-2010</b> <b>cases</b>	<b>2016-2020</b> <b>cases</b>	<b>Increase</b> <b>multiple</b>	<b>Total cases</b> <b>(early period)</b>	<b>Total cases</b> <b>(late period)</b>
<b>As</b>					
Midwest	12.6	12.8	1.0	40.1	43.7
Northeast	13.0	9.8	0.8		
Pacific	1.3	4.1	3.1		
Rocky Mountain	0.5	2.2	4.3		
Southeast	11.0	10.9	1.0		
Southwest	1.6	3.9	2.4		
<b>Ni</b>					
Midwest	0.7	0.7	1.0	2.4	2.7
Northeast	0.7	0.6	0.8		
Pacific	0.1	0.2	2.9		
Rocky Mountain	0.0	0.1	4.4		
Southeast	0.8	0.8	1.0		
Southwest	0.1	0.3	2.4		
<b>Pb</b>					
Midwest	0.1	0.1	1.0	0.4	0.5
Northeast	0.1	0.1	0.8		
Pacific	0.0	0.1	3.3		
Rocky Mountain	0.0	0.0	4.1		
Southeast	0.1	0.1	1.0		
Southwest	0.0	0.0	2.5		

### Effect of an additional structure burned per fire per day on species concentrations



**Supplementary Figure 5.** Effect of an additional burned structure on all chemical species concentrations in our sample.