DeepGEM-EGF: A Bayesian strategy for joint estimates of source-time functions and empirical Green's functions

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11	Key Points:
12	• We propose a new Bayesian framework for estimating source time functions and
13	optimizing one or several prior Empirical Green's Functions (EGFs)
14	• DeepGEM-EGF accounts for epistemic uncertainties and efficiently discriminates
15	source parameters given approximations made in the forward model
16	• DeepGEM-EGF robustly estimates complex source time functions with high-frequency
17	content and exhibits limited sensitivity to prior $EGF(s)$ selection

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Abstract

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An earthquake record is the convolution of source radiation, path propagation and site 19 effects, and instrument response. Isolating the source component requires solving an 20 ill-posed inverse problem. Whether the instability of inferred source parameters arises 21 from varying properties of the source, or from approximations we introduce in solving the 22 problem, remains an open question. Such approximations often derive from limited 23 knowledge of the forward problem. The Empirical Green's function (EGF) approach offers 24 a partial remedy by approximating the forward response of larger events using the records 25 of small events. Indeed, the choice of the « best » small event drastically influences the 26 properties estimated for the larger earthquake. Discriminating variability in source 27 properties from epistemic uncertainties, stemming from the forward problem or other 28 modeling assumptions, requires us to reliably account for, and propagate, any bias or 29 trade-off introduced in the problem. We propose a Bayesian inversion framework that 30 aims at providing reliable and probabilistic estimates of source parameters (here, for the 31 source-time function or STF), and their posterior uncertainty, in the time domain. We 32 jointly solve for the best EGF using one or a few small events as *prior* EGF. Our 33 approach is based on DeepGEM, an unsupervised generalized expectation-maximization 34 framework for blind inversion (Gao et al., 2021). We demonstrate, with toy models as well 35 as an application to an earthquake swarm in California, the potential of DeepGEM-EGF 36 37 to disentangle the variability of the seismic source from biases introduced by modeling assumptions. 38

Plain Language Summary

Our understanding of earthquakes is based on the analysis of earthquakes waveforms 40 recorded at the surface. This non-trivial analysis consists in discriminating properties of 41 the earthquake itself from the footprints of the medium the waves propagate through. In 42 turn, it remains unclear whether the instability of estimated earthquake parameters is 43 inherent to the earthquake nature, or derives from the approximations used in our 44 analysis. The empirical Green's function (EGF) approach offers a partial remedy, in using 45 a small earthquake as a proxy for the propagation term of a larger event. However, the 46 choice of the "best" small event significantly impacts parameters estimated for the larger 47 earthquake. We introduce a probabilistic unsupervised machine learning framework that 48 estimates earthquake source parameters in the time domain (source time functions), and 49 related uncertainties, while discarding the need to select one "best" small event. The 50 approach sequentially updates the properties of an *a priori* selected set of small events 51 towards an optimal solution. We demonstrate the potential of the framework using both 52 simplified models and a case study of an earthquake swarm in California. Our method 53 proves efficient at disentangling earthquake source properties from uncertainties in the 54 modeling process. 55

1 Introduction

Our understanding of how earthquakes start, grow, and stop resides in our ability to build 57 physical models from observed source and rupture properties. Variations in rupture 58 pattern, velocity and directivity, radiated energy and stress drop are examined to provide 59 constraints on fault zone processes (e.g., Kanamori & Brodsky, 2004). The variability of 60 source parameters has practical implications for hazard assessment as it impacts ground 61 shaking predictions (e.g., Anderson & Brune, 1999; Pavic et al., 2000; Boore, 2003; Cotton 62 et al., 2013; A. S. Baltay et al., 2013; Oth et al., 2017; Gerstenberger et al., 2020). Source 63 (and slip) scaling currently dictates how we relate observations, physical processes and hazard (e.g., Aki, 1996; Cotton et al., 2013; Cocco et al., 2016; Lambert et al., 2021). Yet, 65 our insight on whether large and small earthquakes strictly share similar properties (Aki, 66 1967), or if their physics diverge, remains nuanced because of the puzzling instability of 67

the source parameters we infer (e.g., Shearer et al., 2006; Viesca & Garagash, 2015; Lin & Lapusta, 2018; Hardebeck, 2020; Abercrombie, 2021; Bindi et al., 2023).

An earthquake record is the convolution of source radiation, path propagation and site 70 effects, and instrument response. Isolating the source component requires solving an 71 ill-posed inverse problem; however there are questions regarding whether such instabilities 72 arise from the source, or from approximations we introduce in the problem. Since the 73 pioneering work of Hartzell (1978), the Empirical Green's function (EGF) approach is a 74 widely used assumption that spares us the need to model a costly forward problem (e.g., 75 76 Frankel & Kanamori, 1983; Mueller, 1985; Frankel et al., 1986; Hutchings & Wu, 1990; Dreger, 1994; Ihmlé, 1996; Courboulex et al., 1996; Hough, 1997a). Assuming that the 77 only difference in the recordings of two similar and co-located earthquakes is due to their 78 respective rupture, then an event small enough to have an impulsive source can be used as 79 a Green's function for a larger one. However, the choice of the "best" EGF remains a 80 large source of epistemic uncertainties, because of discrepancies in focal mechanism, 81 non-impulsivity of the small event, relative hypocentral distance, finite bandwidth, and 82 noise (e.g., Viegas et al., 2010; Kane et al., 2011, 2013; Abercrombie, 2015, and references 83 therein). Using several potential events (e.g., Hough, 1997b; Abercrombie et al., 2017) or 84 averaging over a dense cluster of events and stations (generalized spectral decomposition, 85 e.g., Andrews, 1986; Prieto et al., 2006; Shearer et al., 2006; Bindi et al., 2009; Trugman 86 & Shearer, 2017) are good palliative approaches, but they fail at accounting for the actual 87 impact of approximations in the forward problem on our understanding of source physics. 88 Approximations we introduce in the blind source deconvolution problem also derive from 89 data errors, non-physically justified priors and any modeling assumptions. Discriminating 90

variability in source properties from such epistemic uncertainties requires us to reliably 91 account for, and propagate, any uncertainty and trade-off (as already suggested by many 92 authors, e.g., Abercrombie & Rice, 2005; Prieto et al., 2006; Shearer et al., 2019; 93 Trugman, 2022). To do so, spectral ratio analyses have seen recent progress towards 94 Bayesian inversion frameworks, facilitated by the few number of parameters involved (e.g., 95 Godano et al., 2015; Garcia-Aristizabal et al., 2016; Van Houtte & Denolle, 2018; Supino 96 et al., 2019; Törnman et al., 2021; Trugman, 2022). Frequency-domain analyses are 97 usually more popular, for computational speed and the reduced number of parameters. 98 But, whether they be generalized decompositions or spectral divisions, those analyses 99 often require non physically justified regularization (Bertero et al., 1997) or assumptions 100 (e.g., Andrews, 1986; Trugman, 2022), and can be biased by artifacts produced in the 101 time domain (Zollo et al., 1995). By contrast, time-domain analyses could reduce the 102 posterior variability of inferred source parameters (Courboulex et al., 2016) while 103 improving constraints on a few properties such as directivity (Boatwright, 1984; McGuire, 104 2004; Trugman, 2022). Yet, time-domain analyses are, so far, limited to deterministic 105 optimization approaches (e.g., Vallée, 2004; Plourde & Bostock, 2017; Gallegos & Xie, 106 2020) that hinder our ability to monitor and assess epistemic variability of the source. 107

Here, we propose a time-domain Bayesian inversion framework that aims to provide
reliable and probabilistic estimates of source parameters and their uncertainty. To do so,
we consider the EGF assumption as a potential cause of epistemic uncertainties and
discard any oversimplification of the problem. We apply our method on toy models and
an application to a swarm in California to demonstrate the potential of this unsupervised
approach to deepen our understanding of the source of earthquakes, and in particular of
its variability.

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2 DeepGEM-EGF: Generalized Expectation-Maximization for blind deconvolution

The approach we propose, DeepGEM-EGF, is twofold: we explore the solution space for any probable source-time function (STF) using one or more candidate EGFs, while considering these EGFs to be a good-but slightly incorrect-prior that needs updating.

The inverse problem that we are trying to solve can be written as $\mathbf{d} = G_{\theta}(\mathbf{m})$; \mathbf{d} being the observed data, \mathbf{m} the parameters of the inverse problem (the STF), G the forward problem (i.e. convolution with the EGFs), and θ the parameters of G representing the EGFs. Solving for both \mathbf{m} and θ is an ill-posed problem. We want to estimate the

EGFs. Solving for both \mathbf{m} and θ is an ill-posed problem. We want to estimate the posterior uncertainty on the STF, but generally already have good prior knowledge on the forward problem as we have selected one or multiple EGFs. We therefore choose to estimate the posterior probability $p(\mathbf{m}|\mathbf{d},\theta)$ on \mathbf{m} while solving for one best estimate of θ that maximizes the log likelihood $p(\theta|\mathbf{d})$ considering all of the potential solutions for \mathbf{m} .

To do so, we follow an approach derived from DeepGEM (Gao et al., 2021), a variational 128 Bayesian Expectation-Maximization (EM) framework that can be used to solve for the 129 parameters of both an inverse problem and a forward model in an unsupervised manner. 130 DeepGEM was initially designed for blind tomography, but has also proven to be effective 131 for a simple blind image deconvolution problem (Gao et al., 2021). Our EM-like approach 132 iterates between two steps; (1) an E-step that learns an approximation to the posterior 133 distribution of **m** given the fixed forward model parameters $\theta^{(i-1)}$ inferred at the previous 134 iteration i-1; and (2) an M-step that solves for $\theta^{(i)}$ considering the fixed parameters of 135 the inverse problem solved in the prior E-step. We present here the main characteristics, 136 and modifications from the initial framework, of DeepGEM-EGF. 137

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2.1 Expectation step: solving for the Source-Time Function

The expectation step (i.e E-step) aims to approximate the posterior distribution of our target STF **m** knowing the data **d** using a normalizing flow-based generative model based on Deep Probabilistic Imaging (DPI, Sun & Bouman, 2021; Sun et al., 2022). We infer the posterior distribution $q_{\phi}(\mathbf{m})$ using variational inference where the class of normalizing flow-based neural networks defines our variational distribution. Once we have solved for the approximate posterior distribution, $q_{\phi}(\mathbf{m})$, we can sample from it to solve the M-step.

We parameterize $q_{\phi}(\mathbf{m})$ with a normalizing flow F_{ϕ} such that $q_{\phi}(\mathbf{m}) = p(F_{\phi}(z))$. F_{ϕ} allows us to map a complicated distribution $p(F_{\phi}(z))$ as a composition of L invertible

transformations $F_{\phi_L} \circ F_{\phi_{L-1}} \circ \dots \circ F_{\phi_1}$ applied to Gaussian samples z where L is the number of layers for the generative flow. At the *i*-th E-step, we optimize the weights $\phi^{(i)}$ of our $F_{\phi^{(i)}}$ that best approximates our posterior distribution $p(\mathbf{m}|\mathbf{d}, \theta^{(i-1)})$ through the following objective:

$$\phi^{(i)} = \arg\min_{\phi} \operatorname{KL}\left(q_{\phi}(\mathbf{m}) \parallel p(\mathbf{m}|\mathbf{d}, \theta^{(i-1)})\right)$$

$$\approx \arg\min_{\phi} \frac{1}{N} \sum_{n=1}^{N} \left[-\log p(\mathbf{d} \mid \theta^{(i-1)}, \mathbf{m}_{n}) - \log p(\mathbf{m}_{n}) + \beta \log q_{\phi}(\mathbf{m}_{n})\right],$$
(1)

for $\mathbf{m}_n = F_{\phi}(z_n), z_n \sim \mathcal{N}(0, 1)$, a batch size of N, the prior on the STF log $p(\mathbf{m})$, the 145 data likelihood log $p(\mathbf{d}|\mathbf{m}, \theta^{i-1})$, and KL denotes the Kullback-Leibler divergence. Note 146 that the approximation comes from evaluation of an expectation using Monte Carlo 147 sampling. See full derivations in Sun and Bouman (2021) and Gao et al. (2021). An 148 additional hyperparameter β is proposed in DPI to control the entropy of the generative 149 model's posterior samples. We use a Real-NVP network (Dinh et al., 2016) for F_{ϕ} with 150 L=32 affine coupling layers. We use Adam (Kingma & Ba, 2014) as the optimizer with a 151 batch size of 1024. 152

We define a realistic prior over the normalized amplitude of the STF as 153 $p(\mathbf{m}) \sim \mathcal{N}(\overline{\mathbf{m}}, \sigma_m)$. By default, $\overline{\mathbf{m}}$ is a Gaussian-shaped STF and σ_m is of 0.2 sec. The 154 default Gaussian-shaped $\overline{\mathbf{m}}$ is centered on half of the preset STF length, and has a width 155 equal to 10% of the total preset STF duration. We also augment $p(\mathbf{m})$ with a few specific 156 constraints. We want the STF to be close to zero on its boundaries, and we enforce 157 sparsity with a ℓ_1 norm. With sparsity regularization, we avoid overestimating the 158 complexity of the STF due to noise and data over-fitting, while allowing for sharp changes 159 in the solution. We also impose a total variation regularization. Default hyperparameters 160 (weights for the specific constraints) were empirically chosen by inspecting the loss and 161 the fit to synthetic tests on a grid search. 162

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2.2 Maximization step: optimizing the EGFs for the forward model

The maximization step (i.e. M-step) relies on estimates of the approximate posterior $q_{\phi}(\mathbf{m})$ from the preceding E-step to update θ , the parameters of the unknown forward model, from the initially assumed set of EGF(s). We define G_{θ} as the convolution between model parameters \mathbf{m} and a linear kernel k_{θ} of parameters θ . The kernel is a deep network consisting in multiple convolution layers without non-linear activation, as proposed in Bell-Kligler et al. (2019) and successfully applied in Gao et al. (2021):

$$G_{\theta}(\mathbf{m}) = \mathbf{m} * k_{\theta} = \mathbf{m} * \frac{\theta_1 * \theta_2 * \dots * \theta_K}{||\theta_1 * \theta_2 * \dots * \theta_K||_{\infty}},$$
(2)

for a K-layer network. During the *i*-th M-step, we optimize the weights $\theta^{(i)}$ to find the maximum a posteriori (MAP) of the posterior distribution of parameters θ knowing the data **d**:

$$\theta^{(i)} = \arg\max_{\theta} \log p(\theta \mid \mathbf{d})$$

$$\approx \arg\max_{\theta} \left[\frac{1}{N} \sum_{n=1}^{N} \left[\log p\left(\mathbf{d} \mid \theta, \mathbf{m}_{n}\right) \right] + \log p(\theta) \right],$$
(3)

for $\mathbf{m}_n = F_{\phi^{(i)}}(\mathbf{z}_n), \, \mathbf{z}_n \sim \mathcal{N}(0, 1)$. We use K=7 convolution layers to parameterize G_{θ} . We use Adam as the optimizer.

As we already have a good guess for our forward model, we use a prior $\log p(\theta)$ that encourages the forward model to stay close to the initially assumed set of EGF(s). We define $\log p(\theta)$ as a weighted sum of mean absolute error (\mathcal{L}_{ϕ}) , ℓ_2 -norm, and dynamic time warping norm (\mathcal{L}_{DTW}) between the updated and initial EGF(s) k_{θ_0} :

$$\log p(\theta) = \lambda_{\phi} \underbrace{\sum |k_{\theta_0} - k_{\theta^{(i)}}|}_{\ell_0} + \lambda_2 \underbrace{\sum (k_{\theta_0} - k_{\theta^{(i)}})^2}_{\ell_0} + \lambda_{DTW} \underbrace{DTW_{0.1}(k_{\theta^{(i)}}, k_{\theta_0})}_{\ell_0}.$$
 (4)

When considering multiple candidate EGFs, we parameterize k_{θ} as an array of independent deep networks. During the *i*-th M-step and for each EGF *e*, weights $\theta_e^{(i)}$ are optimized. We augment $\log p(\theta)$ with a prior $\mathcal{L}_{\text{multi}}$ that encourages the inferred EGFs to converge towards a single EGF $k_{\theta_{\text{best}}^{(i)}}$ that minimizes $\alpha_e^{(i)}$, the misfit between data and predictions:

$$\log p(\theta) = \mathcal{L}_{\phi} + \ell_2 + \mathcal{L}_{DTW} + \lambda_{\text{multi}} \underbrace{\sum_{e} \alpha_e^{(i)} \sum_{e} (k_{\theta_e^{(i)}} - k_{\theta_{\text{best}}^{(i)}})^2}_{e}, \tag{5}$$

with
$$k_{\theta_{\text{best}}^{(i)}} = \min_{e} \alpha_e^{(i)} \propto \sum (\mathbf{m}^{(i)} * k_{\theta_e^{(i)}} - \mathbf{d})^2$$
 (6)

Additionally, MSE loss for each EGF can be weighted with a user-defined parameter that can reflect the quality of each *a priori* selected EGF. Weights $\lambda_{\phi}, \lambda_{2}, \lambda_{DTW}$ and λ_{multi} were empirically chosen with a grid search on synthetic tests.

Sec- tion	Fig.	\mathbf{Tests}	Suppl. Mat.	STF	# of EGFs	True EGF	Assumed EGF	Main event
3.1	1	a to s	S2, Figs S1-S4	Stack of $N_{STF} = 10$ pulses with Gaussian shape, random weigths and widths	1	All arrivals, synthetic waveforms calculated with 1D crust, random source parameters, 10 Hz	All arrivals, source properties and crustal structure that randomly vary away from true EGF + white noise, PSNR 0 to 10 %	True EGF * STF
		a0 to $g4$	S2, Figs S5-S10	$N_{STF} = 3$	1	,,	"	"
3.2			S3, Fig. S18	$N_{STF}=3$	1	All arrivals of recorded waveforms, 20 or 25 Hz	True EGF + noise, PSNR 3 % different for each component	True EGF * STF + white noise, PSNR 3 $\%$
	2		S3, Figs S19- S20	$N_{STF} = 10$	1	P arrivals of recorded waveforms, 40 or 50 Hz	"	"
4.1	3		S4, Figs S21-33	$N_{STF} = 3$	3	Randomly weighted sum of selected recorded EGFs + white noise PSNR 3 %, all arrivals, 20 Hz	Selected recorded EGFs, all arrivals	"
			S4, Figs S34-39	$N_{STF} = 10$	4	Randomly weighted sum of selected recorded EGFs + white noise PSNR 3 %, P arrivals, 20 Hz	Selected recorded EGFs, P arrivals	"
	5		S4, Figs S38, S40	$N_{STF} = 10$	8	"	,,	"

 Table 1.
 Summary of design choices for synthetic tests. Details are in Supplementary Material.

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Figure 1. Inferred (colored) and target (black) Source-Time Functions (STFs) for toy models designed with fully synthetic waveforms. The mean of the STFs is colored by the peak cross-correlation value between prior and target Empirical Green's functions; the standard deviation is shaded (1σ darker, 3σ lighter).

3 Benchmark on toy models

We test the ability of DeepGEM-EGF to correctly infer STFs and EGFs on toy models,
based on fully synthetic waveforms or recorded waveforms. A summary of those tests is
presented in Table 3. For a few tests, we compare DeepGEM results with non-blind
deconvolution approaches; a multitaper deconvolution approach (e.g., Thomson, 1982;
Percival & Walden, 1993; Prieto et al., 2009) and an iterative Landweber approach
(Bertero et al., 1997). We choose both these algorithms as they are well-documented and
have been used extensively.

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3.1 Tests with synthetic waveforms

We first perform two series of tests with fully synthetic waveforms calculated at 10 Hz, using all of the arrivals (P, S and coda). The target STF consists of a stack of 10 (tests a to n, Fig. 1, Suppl. Mat. Table S3) or 3 (tests listed in Suppl. Mat. Tables S4 and S5) pulses parameterized as Gaussian functions with randomly varying width and heights. Source parameters for the main event (the largest) are random. Source properties (location, depth, moment tensor) for synthetic EGFs randomly vary away from the properties of the main event. Kagan angle (Kagan, 1991) between moment tensors of the source and assumed EGFs vary between 5 and 40°. We add white noise, with a peak signal to noise ratio ranging from 0-10%, to waveforms of both the main event and EGFs. Full results are detailed in Suppl. Mat. section S2.

The results are promising: target STFs are well recovered, including multiple peaks of 188 various frequencies (Fig. 1). The predicted waveforms fit the data well. The updated EGF 189 waveforms are usually close to the target ones, even when the prior is off, without 190 overfitting the additional noise. STF fits from a multitaper approach are outperformed, 191 even for very simple tests (e.g., Suppl. Mat. Fig. S4). The choice of the prior EGF has a 192 strong impact on the quality of the results: the larger the discrepancy between the 193 assumed and target EGF waveforms, the larger the misfit between inferred and target 194 STFs (see test b in Fig. 1). DeepGEM converges well even if the source of the prior EGF 195

sees a time shift (changes in depth, location, velocity structure) or variations of up to $\sim 20^{\circ}$ in Kagan angle (tests j to n, Suppl. Mat. Figs S1-S4).

Occasionally, our tests fail. These failures can be classified into two categories: either the 198 forward model does not deviate sufficiently from the incorrect prior, or it becomes 199 trapped in a local minimum. Good representatives of the first category are tests with 200 Kagan angles exceeding 30° (tests o to s). In that case, most peaks of the STFs are 201 correctly inferred, but the weight on the prior is too large for the EGF to fully update 202 towards the correct solution. The second category occurs when the true forward model is 203 too far from the prior, such as when the assumed velocity model significantly differs from the true one (e.g., tests 2b6 and 2g1 are accurate, but tests $2g^*$ are inconclusive) or when 205 initial EGFs are too noisy (added white noise PSNR reaches 10%, tests 2g7 and 2g8). In 206 practice, such mismatch between prior and true forward models are rare due to the careful 207 selection of EGFs. 208

3.2 Tests with recorded waveforms

We also perform a few synthetic tests using recorded waveforms as EGFs. We select, as 210 EGFs, P arrivals of recordings in Southern California for an ML 3.37 event that occurred 211 on 17 July 2014 near Borrego Springs, CA, USA (Figs. 2,3). To obtain waveforms for the 212 main toy model event, we convolve those P arrivals to synthetic multi-peaked STFs (10 213 pulses, similarly to the previous toy models). We then decimate all waveforms to 40 Hz 214 (for broadband) or 50 Hz (for accelerometers) to hasten the calculation process, and add 215 white noise. These tests differ from the fully synthetic toy models mostly because of the 216 duration of the data, which is close to the duration of the target STF, and the increased 217 frequency content: the problem is much more underdetermined. Full results are detailed 218 in Suppl. Mat., section S3. 219

The results of this experiment are overall consistent with those of the fully synthetic toy 220 models. We recover most features of the target STFs in all frequency bands, with 221 occasional misfit in amplitude, and a small tendency to overfit on noise at high frequencies. In general, target EGFs are well recovered. For stations BOR and TRO 223 (Fig. 2h,i) we cannot gain information on a few model parameters, as expressed by the 224 large posterior uncertainty on the boundaries of the inferred STF. This is probably 225 because the duration of the P-arrivals is less than a third of the assumed STF duration, 226 making the problem highly underdetermined. A larger weight on the boundary prior or a 227 decreased frequency band could solve the issue; but we choose here to use similar 228 hyperparameters for all the tests. 229

We also infer apparent STFs for each component with the Landweber approach (we visually select the best STF, orange in Fig. 2, more details in Suppl. Mat.). Baseline STFs are fair to poor, and only the first-order characteristics are retrieved: the frequency content is probably too high, and the problem too ill-posed for the method to perform correctly.

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4 Case study: the 2016-2019 Cahuilla swarm, CA, USA

The Cahuilla swarm outlined a complex but well-defined fault structure between the San 236 Jacinto and Elsinore fault zones (Fig. 5). Ross et al. (2020) produced a seismicity catalog 237 of more than 22,000 events with Mw ranging from 0.7 to 4.4, all of those event seemingly 238 affecting the same non-planar fault geometry. The main event (Mw 4.4) is ~ 5 km deep 239 and likely caused drastic changes in the evolution of the swarm by allowing fluids to 240 circulate at shallower depth. We take advantage of this detailed seismic catalog, with 241 many small events, to investigate how our forward model assumptions impacts our 242 knowledge of the source. Along with evaluating the robustness of our approach and how 243



Figure 2. Inferred (colormap) and target (black) Source-Time Functions (STFs) for toy models designed with waveforms (P arrivals only) recorded for a M_L 3.37 event that occured in 2014 near Borrego Springs (CA). The mean of the STFs is colored from the normalized misfit between inferred and target Empirical Green's functions; the standard deviation is shaded (1 σ darker, 3 σ lighter). Apparent best fitting STFs calculated with the approach by Bertero et al. (1997) are shown in orange. Stations locations are shown in Figure 3. In (h) and (i), the problem is highly underdetermined and there is no information gained on the boundaries.



Figure 3. (a-d) Inferred (colored) and target (black) Source-Time Functions (STFs) for toy models designed with waveforms (P arrivals only) recorded for four neighbor $M\sim2$ events that occurred during the Cahuilla swarm. The mean of the STFs is colored from the misfit between inferred and target Empirical Green's functions (EGFs); the standard deviation is shaded. (e) Best inferred (colored) and target (black) EGF for toy model (a). Stations locations are shown in the map on the right. CTW station is located to the north-west of TOR (see Fig. 5).

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characterize the source process of this puzzling event.

4.1 Toy models with multiple EGFs

the ad-hoc selection of EGFs can affect the inferred source parameters, we aim to better

We first design two sets of toy models that are different than the ones previously presented: we use several of the nearest Mw~2 events to the mainshock as prior EGFs. The "true" EGF, or the one convolved with synthetic STFs to obtain mainshock waveforms, is defined as a randomly weighted sum of the *a priori* selected EGFs. We use two sets of prior EGFs. For one set, we use all arrivals for 3 EGF events (Figs. S21-33). For the other set, we use P arrivals for 4 EGF events (results in Figs. 3, S34-38). Note that as EGFs are selected based on distance and not semblance, their waveforms can differ significantly, making the example more difficult than with carefully selected waveforms. Full results are detailed in Suppl. Mat. section S4.2.

The misfit between the inferred and target STFs and EGFs is usually below 20% of the maximum amplitude for simple tests (Fig. S21), and increases with distance between prior and target EGFs. For more complex STFs, fit remains very good even if a few high frequency artifacts appear (Fig. 3). Target EGF events are generally close to the mean inferred EGFs and within posterior uncertainties (Figs. 3 and S22-38).

For station TOR, we also perform tests using four more EGFs (Figs. 4 and S38-40). In 261 one case (Fig. 4c,d), the additional four waveforms are random linear combinations of the 262 initial four prior EGFs: in other words, the additional waveforms do not contain more 263 information than the initial four. In another case, the additional four waveforms are closer 264 to the target EGF: they are equal to the target EGF to which 10% of the random linear 265 combination is added (Fig. 4e,f). As expected, if additional EGFs do not add information 266 content, the inferred model is not improved. By contrast, adding informative EGFs (i.e. 267 more accurate EGFs) improves the quality of the inferred STF. Adding information only 268



Figure 4. Mean inferred (colored) and target (black) Source-Time Functions (STFs) and Empirical Green's Functions for toy models designed with waveforms (P arrivals only) recorded for four (a,b) or eight (c-f) neighboring M \sim 2 events that occurred during the Cahuilla swarm. In (e,f) the additional four waveforms are closer to the target than in (c,d). Standard deviation (1 σ) is shaded. EGFs are colored by the root mean squared error (RMSE) between data and average predictions normalized for each component.

very slightly improves the fit to the data. It is therefore more efficient to assume a few
 well selected prior EGFs than multiple poorly constrained priors. These tests further
 confirm the robustness of DeepGEM-EGF, and its capacity to use good information.

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4.2 Results

We compare the STFs inferred, at stations within 100 km of the mainshock, assuming 273 three different sets of prior EGFs decimated to 20 Hz: 1. minimizing distance (set A, 274 Fig 5a); 2. minimizing cross-correlation at one station (set B, Fig 5b); 3. minimizing 275 cross-correlation (set C, Fig 5c,d) between EGFs and the mainshock. In set B, selected 276 EGFs are less than 1 km away from the mainshock and maximize the cross-correlation 277 with the mainshock waveforms at station PLM. This choice of station is based on the 278 quality of the waveforms and the deconvolution with set A (Fig. S57). For set C, prior 279 EGFs are selected based on their cross-correlation with the mainshock waveforms at each 280 station, are less than 1 km away from the mainshock, and their number vary between 1 281 and 4 depending on their SNR. We use either P waves arrivals (Fig 5a-c) or S waves 282 arrivals (for set C only, Fig 5d), and assume similar priors and hyperparameters. We 283 expect the STFs inferred with set C to be our best estimates; and the ones estimated with 284 set B to be the worst, as EGFs are incorrectly constrained with a single station. Full 285 results are detailed in Suppl. Mat. sections S4.3 and S4.4. 286

For all sets of EGFs, inferred STFs share similar first-order characteristics: width, shape and azimuthal variations. As expected, STFs estimated with set B, while consistent between neighbor stations, show many high frequency peaks, which indicate their relatively poor quality compared to results with other sets (Figs S42, S46, S50, S54 for



Figure 5. Impact of the choice of candidate EGFs on the apparent STFs of the 2018 Mw 4.41 mainshock that occurred during the Cahuilla swarm. (e) Stations locations and (f) relocated catalog for the swarm and EGFs used. (a-d) STFs inferred assuming up to 4 M~2 event as prior EGFs. The posterior mean is in black, the posterior σ in gray. We assume as prior EGFs: (a) 4 events nearest to the mainshock, (b) 4 events maximizing EGF/mainshock waveform cross-correlation at station PLM, (c) and (d) up to 4 \underline{ege} ents maximizing EGF/mainshock waveform cross-correlation at each station. We use P wave arrivals in (a) to (c), and S wave arrivals in (d). STFs are aligned on their peak value: if the peak is off or shifted, the STF appears as partial.



Figure 6. STFs inferred for the 2018 Mw 4.41 mainshock of the Cahuilla swarm assuming the same prior EGFs as in Fig 5, using DeepGEM-EGF (mean in black) or with the approach of Bertero et al. (1997, mean for several components and EGFs in orange, standard deviation σ shaded). Note that σ estimates for the Landweber approach derive from the variability of multiples EGFs and their components, and not from posterior PDFs as for DeepGEM results. (left) Using P arrivals only (similar to Fig. 5e) or (right) using S arrivals only (same as Fig. 5f).

stations BOR, BLA2, PLS and LMH, respectively). Small posterior uncertainty on both 291 STFs and EGFs might suggest that over-constraining EGFs to an incorrect prior induced 292 the model to fall into a local minimum. In contrast, STFs estimated with sets A and C 293 are smoother and more similar, and furthermore have increased posterior uncertainty and 294 overall improved fit (particulary with set C, see Figs S41 to S60). For this latter set, STFs 295 shapes are more coherent across stations, for instance with two small pulses preceding the 296 largest peak around station BBS (Fig. 5e). In particular, using S arrivals, we infer similar 297 apparent STFs at two close stations (JEM and MTG or PSD and WWC), confirming the 298 robustness of our approach. 299

We finally compare our preferred STFs (set C) to apparent STFs estimated with a widely used frequency-domain deconvolution approach (or Landweber approach, from Bertero et al., 1997). With this latter method, we derive means and standard deviations from estimates of apparent STFs for each component of each prior EGF (Fig 6). The Landweber STFs are smooth, often largely over-estimate the STF duration (in particular for S arrivals), and show a large posterior uncertainty. P-wave pulse-like STFs are
 relatively well-imaged around station PSD but not to the north of the swarm, and
 apparent STFs for both methods are of similar shape for a few stations only (JEM, TOR,
 SAL). Even for those stations, STFs derived from the baseline approach are less
 informative: for instance, S-arrivals apparent STFs at station JEM (Fig. 6c,d) are
 symmetric, or by contrast clearly asymmetric, when estimated with the Landweber or
 DeepGEM approach, respectively.

Results from both methods indicate that apparent STFs are narrower for stations located 312 NE of the Cahuilla swarm. Our deconvolution approach allows us to reach a greater level 313 of detail: for stations to the north of the swarm (BBS, PSD), apparent STFs are complex 314 and long in duration, with two small peaks preceding an energetic pulse, especially around 315 stations PSD and MGE. To the South (stations SDD to TOR), apparent STFs are wider 316 and more Gaussian-like. The pulse-like shape of STFs around stations PSD and IDO 317 suggests a down-dip directivity of the Mw 4.4 event towards the northern portion of the 318 fault. Narrow STFs around station BBS might also suggest along-strike directivity 319 towards the north. 320

5 Discussion and conclusions

We introduce DeepGEM-EGF, a Bayesian joint inversion method for improving the source 322 deconvolution problem. We show through tests on toy models, synthetic tests with real 323 waveforms, and a case study in California, that this approach effectively discriminates 324 source parameters given approximations made in the forward model. Specifically, 325 DeepGEM-EGF robustly estimates complex apparent STFs with high-frequency content 326 (>40 Hz) and exhibits limited sensitivity to prior EGF selection. Additionally, 327 DeepGEM-EGF estimates posterior uncertainties for both STFs and EGFs (where 328 multiple priors are assumed), ensuring our ability to track the impact of epistemic 329 uncertainty on source estimates. 330

DeepGEM-EGF explicitly accounts for epistemic uncertainties and provides probabilistic estimates of source parameters in the time domain, advancing beyond deterministic and frequency-domain methods. Compared to these traditional deconvolution techniques, DeepGEM-EGF provides more robust, coherent, and informative results, which remain stable for various choices of prior EGFs. The proposed approach may underperform for highly underdetermined scenarios, or when prior EGFs deviate significantly from the true solution, i.e. when the EGF assumption does not hold anymore.

Understanding earthquake self-similarity across scales is crucial for advancing physical 338 models and improving hazard assessments. Self-similarity is often explored using two key 339 proxies: source complexity and stress drop. Both stress drop (Atkinson & Beresnev, 1997; 340 Abercrombie, 2021; Bindi et al., 2023; A. Baltay et al., 2024; Neely et al., 2024) and 341 metrics of source complexity (Vallée & Douet, 2016; Danré et al., 2019; Pennington et al., 342 2023; Neely et al., 2024) are uncertain and variable, and even more difficult to constrain 343 for smaller magnitude events due to increased noise and frequency content and the 344 scarcity of observations. DeepGEM-EGF can address these challenges by reliably imaging 345 complex, multi-peaked STFs under noisy conditions, while staying methodologically 346 consistent across magnitudes (e.g., Neely et al., 2024; A. Baltay et al., 2024). We aim for 347 this joint approach to become a valuable tool for disentangling the variability of the 348 seismic source, and stress drop, from artifacts introduced by modeling assumptions. 349 DeepGEM-EGF is available as an open-source tool.

		Code a
DeepGEM-EGF	is available	at https:

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Code and data availability

DeepGEM-EGF is available at https://github.com/thearagon/DeepGEM_EGF and stored at https://doi.org/10.5281/zenodo.14472786. A running example is available on the same repository. We used waveform data recorded by the regional (CI, California Institute of Technology and United States Geological Survey Pasadena, 1926), ANZA (AZ, Frank Vernon, 1982) and UCSB (SB, UC Santa Barbara, 1989) networks in Southern California.

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References

- Abercrombie, R. E. (2015). Investigating uncertainties in empirical Green's function
 analysis of earthquake source parameters. Journal of Geophysical Research:
 Solid Earth, 120(6), 4263–4277. doi: 10.1002/2015JB011984
- Abercrombie, R. E. (2021). Resolution and uncertainties in estimates of earthquake
 stress drop and energy release. *Philosophical Transactions of the Royal Society A: Mathematical, Physical and Engineering Sciences, 379*(2196), 20200131.
 doi: 10.1098/rsta.2020.0131
- Abercrombie, R. E., Poli, P., & Bannister, S. (2017). Earthquake Directivity,
 Orientation, and Stress Drop Within the Subducting Plate at the Hikurangi
 Margin, New Zealand. Journal of Geophysical Research: Solid Earth, 122(12),
 10,176–10,188. doi: 10.1002/2017JB014935
 - Abercrombie, R. E., & Rice, J. R. (2005). Can observations of earthquake scaling constrain slip weakening? *Geophysical Journal International*, 162(2), 406–424. doi: 10.1111/j.1365-246X.2005.02579.x
 - Aki, K. (1967). Scaling law of seismic spectrum. Journal of Geophysical Research (1896-1977), 72(4), 1217–1231. doi: 10.1029/JZ072i004p01217
 - Aki, K. (1996). Scale dependence in earthquake phenomena and its relevance to earthquake prediction. *Proceedings of the National Academy of Sciences*, 93(9), 3740–3747. doi: 10.1073/pnas.93.9.3740
- Anderson, J. G., & Brune, J. N. (1999). Probabilistic Seismic Hazard Analysis without the Ergodic Assumption. *Seismological Research Letters*, 70(1), 19–28. doi: 10.1785/gssrl.70.1.19
- Andrews, D. J. (1986). Objective Determination of Source Parameters and Similarity of Earthquakes of Different Size. In *Earthquake Source Mechanics* (pp. 259–267). American Geophysical Union (AGU). doi: 10.1029/GM037p0259
- Atkinson, G. M., & Beresnev, I. (1997). Don't Call it Stress Drop. Seismological Research Letters, 68(1), 3–4. doi: 10.1785/gssrl.68.1.3
- Baltay, A., Abercrombie, R., Chu, S., & Taira, T. (2024). The SCEC/USGS Community Stress Drop Validation Study Using the 2019 Ridgecrest Earthquake
 Sequence. Seismica, 3(1). doi: 10.26443/seismica.v3i1.1009
- Baltay, A. S., Hanks, T. C., & Beroza, G. C. (2013). Stable Stress-Drop Measurements and their Variability: Implications for Ground-Motion Prediction. Bulletin of the Seismological Society of America, 103(1), 211–222. doi: 10.1785/ 0120120161
- Bell-Kligler, S., Shocher, A., & Irani, M. (2019). Blind Super-Resolution Kernel Es timation using an Internal-GAN. arXiv e-prints, arXiv:1909.06581. doi: 10
 .48550/arXiv.1909.06581
- Bertero, M., Bindi, D., Boccacci, P., Cattaneo, M., Eva, C., & Lanza, V. (1997). Application of the projected Landweber method to the estimation of the source time function in seismology. *Inverse Problems*, 13(2), 465. doi: 10.1088/0266-5611/13/2/017
- Bindi, D., Pacor, F., Luzi, L., Massa, M., & Ameri, G. (2009). The Mw 6.3, 2009
 L'Aquila earthquake: Source, path and site effects from spectral analysis of

404	strong motion data. Geophysical Journal International, 179(3), 1573–1579.
405	Bindi D. Spallarogen D. Diagzzi M. Oth A. Maragan D. & Mayada K. (2022)
406	The Community Stress-Drop Validation Study—Part I: Source Propagation
407	and Site Decomposition of Fourier Spectra Seismological Research Letters
400	$g_{\ell}(4)$ 1980–1991 doi: 10.1785/0220230019
409	Bostwright I (1084) The effect of runture complexity on estimates of source
410	size Lowrnal of Coonhysical Research: Solid Earth 80(B2) 1132–1146 doi:
411	$101029/\mathrm{IB}089\mathrm{iB}02\mathrm{p}01132$
412	Boore D M (2003) Simulation of Ground Motion Using the Stochastic Method. In
415	Y Ben-Zion (Ed.) Seismic Motion Lithospheric Structures Earthquake and
414	Volcanic Sources: The Keijiti Aki Volume (pp. 635–676) Basel: Birkhäuser
415	doi: 10.1007/978-3-0348-8010-7_10
417	California Institute of Technology and United States Geological Survey Pasadena
418	(1926). Southern california seismic network. International Federation of Digi-
419	tal Seismograph Networks, doi: 10.7914/SN/CI
420	Cocco, M., Tinti, E., & Cirella, A. (2016). On the scale dependence of earthquake
421	stress drop. Journal of Seismology. 20(4), 1151–1170. doi: 10.1007/s10950-016
422	-9594-4
423	Cotton, F., Archuleta, R., & Causse, M. (2013). What is Sigma of the Stress Drop?
424	Seismological Research Letters, 84(1), 42–48. doi: 10.1785/0220120087
425	Courboulex, F., Vallée, M., Causse, M., & Chounet, A. (2016). Stress-Drop Vari-
426	ability of Shallow Earthquakes Extracted from a Global Database of Source
427	Time Functions. Seismological Research Letters, 87(4), 912–918. doi:
428	10.1785/0220150283
429	Courboulex, F., Virieux, J., Deschamps, A., Gibert, D., & Zollo, A. (1996). Source
430	investigation of a small event using empirical Green's functions and simu-
431	lated annealing. Geophysical Journal International, 125(3), 768–780. doi:
432	10.1111/j.1365-246X.1996.tb06022.x
433	Danré, P., Yin, J., Lipovsky, B. P., & Denolle, M. A. (2019). Earthquakes Within
434	Earthquakes: Patterns in Rupture Complexity. Geophysical Research Letters,
435	46(13), 7352-7360. doi: $10.1029/2019$ GL083093
436	Dinh, L., Sohl-Dickstein, J., & Bengio, S. (2016). Density estimation using Real
437	NVP. https://arxiv.org/abs/1605.08803v3.
438	Dreger, D. S. (1994). Empirical Green's function study of the January 17, 1994
439	Northridge, California earthquake. Geophysical Research Letters, 21(24),
440	2633-2636. doi: 10.1029/94GL02661
441	Frank Vernon. (1982). ANZA regional network. International Federation of Digital
442	Seismograph Networks. doi: 10.7914/SN/AZ
443	Frankel, A., Fletcher, J., Vernon, F., Haar, L., Berger, J., Hanks, T., & Brune, J.
444	(1986). Rupture characteristics and tomographic source imaging of ML ~ 3
445	earthquakes near Anza, southern California. Journal of Geophysical Research:
446	Source Eurin, 91(B12), 12055-12050. doi: 10.1029/JB0911B12D12055
447	drop for contheuplace in couthern Colifornia – Palletin of the Sciencelogical Soci
448	atu of America, 72(6A), 1527–1551, doi: 10.1785/BSSA07306A1527
449	Callagos A k Xia I (2020) A multichannel deconvolution method to retrieve
450	source-time functions: Application to the regional Lg wave <i>Geophysical Jour</i> -
452	nal International. 223(1), 323–347 doi: 10.1093/gii/ggaa303
452	Gao A F Castellanos J C Yue Y Ross Z E & Rouman K (2021) Deen-
454	GEM: Generalized Expectation-Maximization for Blind Inversion In Thirty-
455	Fifth Conference on Neural Information Processing Systems.
456	Garcia-Aristizabal, A., Caciagli, M., & Selva, J. (2016). Considering uncertainties in
457	the determination of earthquake source parameters from seismic spectra. Geo-
458	physical Journal International, 207(2), 691–701. doi: 10.1093/gji/ggw303

459	Gerstenberger, M. C., Marzocchi, W., Allen, T., Pagani, M., Adams, J., Danciu, L.,
460	Petersen, M. D. (2020). Probabilistic Seismic Hazard Analysis at Regional
461	and National Scales: State of the Art and Future Challenges. Reviews of Geo-
462	physics, 58(2), e2019RG000653. doi: 10.1029/2019RG000653
463	Godano, M., Bernard, P., & Dublanchet, P. (2015). Bayesian inversion of seismic
464	spectral ratio for source scaling: Application to a persistent multiplet in the
465	western Corinth rift. Journal of Geophysical Research: Solid Earth, 120(11),
466	7683–7712. doi: 10.1002/2015JB012217
467	Hardebeck, J. L. (2020). Are the Stress Drops of Small Earthquakes Good
468	Predictors of the Stress Drops of Moderate-to-Large Earthquakes? Jour-
469	nal of Geophysical Research: Solid Earth, 125(3), e2019JB018831. doi:
470	10.1029/2019JB018831
471	Hartzell, S. H. (1978). Earthquake aftershocks as Green's functions. Geophysical Re-
472	search Letters, 5(1), 1–4. doi: 10.1029/GL005i001p00001
473	Hough, S. E. (1997a). Empirical Green's function analysis: Taking the next step.
474	Journal of Geophysical Research: Solid Earth, 102(B3), 5369–5384.
475	10.1029/96JB03488
476	Hough, S. E. (1997b). Empirical Green's function analysis: Taking the next step.
477	Journal of Geophysical Research: Solid Earth, 102(B3), 5369–5384. doi:
478	10.1029/96JB03488
479	Hutchings, L., & Wu, F. (1990). Empirical Green's Functions from small earth-
480	quakes: A waveform study of locally recorded aftershocks of the 1971 San
481	Fernando Earthquake. Journal of Geophysical Research: Solid Earth, 95(B2).
482	1187–1214. doi: 10.1029/JB095iB02p01187
102	Ihmlé P. F. (1996) Monte Carlo slip inversion in the frequency domain Application
483	to the 1992 Nicaragua Slow Earthquake Geophysical Research Letters 23(9)
485	913–916. doi: 10.1029/96GL00872
486	Kagan V V (1991) 3-D rotation of double-couple earthquake sources <i>Geophysical</i>
400	<i>Journal International 106</i> (3) 709–716 doi: 10.1111/i.1365-246X 1991 tb06343
488	.X
480	Kanamori H & Brodsky E E (2004) The physics of earthquakes <i>Benorts on</i>
405	Progress in Physics 67(8) 1429 doi: 10.1088/0034-4885/67/8/R03
401	Kane D L Kilb D L & Vernon F L (2013) Selecting Empirical Green's Func-
491	tions in Regions of Fault Complexity: A Study of Data from the San Jacinto
493	Fault Zone, Southern California, Bulletin of the Seismological Society of Amer-
494	<i>ica</i> , 103(2A), 641–650, doi: 10.1785/0120120189
495	Kane, D. L., Prieto, G. A., Vernon, F. L., & Shearer, P. M. (2011). Quantifying
496	Seismic Source Parameter Uncertainties. Bulletin of the Seismological Society
497	of America, 101(2), 535–543, doi: 10.1785/0120100166
498	Kingma, D. P., & Ba, J. (2014). Adam: A Method for Stochastic Optimization. In
499	3rd International Conference for Learning Representations. San Diego.
500	Lambert V Lapusta N & Faulkner D (2021) Scale Dependence of Farthquake
500	Rupture Prestress in Models With Enhanced Weakening: Implications for
502	Event Statistics and Inferences of Fault Stress. Journal of Geophysical Re-
502	search: Solid Earth, 126(10), e2021JB021886, doi: 10.1029/2021JB021886
504	Lin Y-Y & Lapusta N (2018) Microseismicity Simulated on Asperity-Like Fault
505	Patches: On Scaling of Seismic Moment With Duration and Seismological Es-
506	timates of Stress Drops. Geophysical Research Letters, 45(16), 8145–8155. doi:
507	10.1029/2018GL078650
508	McGuire, J. J. (2004). Estimating Finite Source Properties of Small Earthquake
509	Ruptures. Bulletin of the Seismological Society of America 9/(2) 377–303
510	doi: 10.1785/0120030091
511	Mueller C S (1985) Source pulse enhancement by deconvolution of an empirical
512	Green's function. Geophysical Research Letters 12(1) 33–36 doi: 10.1029/
513	GL012i001p00033
-	T

Neely, J. S., Park, S., & Baltay, A. (2024).The Impact of Source Time Function 514 Complexity on Stress-Drop Estimates. Bulletin of the Seismological Society of 515 America. doi: 10.1785/0120240022 516 Oth, A., Miyake, H., & Bindi, D. (2017). On the relation of earthquake stress drop 517 and ground motion variability. Journal of Geophysical Research: Solid Earth, 518 122(7), 5474–5492. doi: 10.1002/2017JB014026 519 Pavic, R., Koller, M. G., Bard, P.-Y., & Lacave-Lachet, C. (2000). Ground motion 520 prediction with the empirical Green's function technique: An assessment of 521 uncertainties and confidence level. Journal of Seismology, 4(1), 59–77. doi: 522 10.1023/A:1009826529269 523 Pennington, C. N., Wu, Q., Chen, X., & Abercrombie, R. E. (2023).Quantify-524 ing rupture characteristics of microearthquakes in the Parkfield Area using a 525 high-resolution borehole network. Geophysical Journal International, 233(3), 526 1772–1785. doi: 10.1093/gji/ggad023 527 Percival, D. B., & Walden, A. T. Spectral Analysis for Physical Applica-(1993).528 tions. Spectral Analysis for Physical Applications. 529 Plourde, A. P., & Bostock, M. G. (2017). Multichannel Deconvolution for Earth-530 quake Apparent Source Time Functions. Bulletin of the Seismological Society 531 of America, 107(4), 1904–1913. doi: 10.1785/0120170015 532 Prieto, G. A., Parker, R. L., Vernon, F. L., Shearer, P. M., & Thomson, D. J. 533 (2006).Uncertainties in Earthquake Source Spectrum Estimation Us-534 ing Empirical Green Functions. In Earthquakes: Radiated Energy and the 535 Physics of Faulting (pp. 69–74). American Geophysical Union (AGU). doi: 536 10.1029/170GM08 537 Prieto, G. A., Parker, R. L., & Vernon III, F. L. (2009).A Fortran 90 library for 538 multitaper spectrum analysis. Computers & Geosciences, 35(8), 1701–1710. 539 doi: 10.1016/j.cageo.2008.06.007 540 Ross, Z. E., Cochran, E. S., Trugman, D. T., & Smith, J. D. (2020). 3D fault ar-541 chitecture controls the dynamism of earthquake swarms. Science, 368(6497), 542 1357–1361. doi: 10.1126/science.abb0779 543 Shearer, P. M., Abercrombie, R. E., Trugman, D. T., & Wang, W. (2019). Compar-544 ing EGF Methods for Estimating Corner Frequency and Stress Drop From 545 P Wave Spectra. Journal of Geophysical Research: Solid Earth, 124(4), 546 3966-3986. doi: 10.1029/2018JB016957 547 Shearer, P. M., Prieto, G. A., & Hauksson, E. (2006).Comprehensive analysis of 548 earthquake source spectra in southern California. Journal of Geophysical Re-549 search: Solid Earth, 111(B6). doi: 10.1029/2005JB003979 550 Sun, H., & Bouman, K. L. (2021). Deep Probabilistic Imaging: Uncertainty Quan-551 tification and Multi-modal Solution Characterization for Computational Imag-552 ing. In The Astrophysical Journal. doi: 10.48550/arXiv.2010.14462 553 Sun, H., Bouman, K. L., Tiede, P., Wang, J. J., Blunt, S., & Mawet, D. (2022). α -554 deep Probabilistic Inference (α -DPI): Efficient Uncertainty Quantification from 555 Exoplanet Astrometry to Black Hole Feature Extraction. The Astrophysical 556 Journal, 932(2), 99. doi: 10.3847/1538-4357/ac6be9 557 Supino, M., Festa, G., & Zollo, A. (2019). A probabilistic method for the estimation 558 of earthquake source parameters from spectral inversion: Application to the 559 2016–2017 Central Italy seismic sequence. Geophysical Journal International, 560 218(2), 988-1007. doi: 10.1093/gji/ggz206 561 Thomson, D. (1982). Spectrum estimation and harmonic analysis. Proceedings of the 562 IEEE, 70(9), 1055–1096. doi: 10.1109/PROC.1982.12433 563 Törnman, W., Martinsson, J., & Dineva, S. (2021).Robust Bayesian estimator 564 for S-wave spectra, using a combined empirical Green's function. Geophysical 565 Journal International, 227(1), 403-438. doi: 10.1093/gji/ggab184 566 Trugman, D. T. (2022).Resolving Differences in the Rupture Properties of M5 567 Earthquakes in California Using Bayesian Source Spectral Analysis. Journal 568

569	of Geophysical Research: Solid Earth, 127(4), e2021JB023526. doi: 10.1029/
570	2021JB023526
571	Trugman, D. T., & Shearer, P. M. (2017). Application of an improved spectral
572	decomposition method to examine earthquake source scaling in Southern Cali-
573	fornia. Journal of Geophysical Research: Solid Earth, 122(4), 2890–2910. doi:
574	10.1002/2017JB013971
575	UC Santa Barbara. (1989). UC santa barbara engineering seismology network. Inter-
576	national Federation of Digital Seismograph Networks. doi: 10.7914/SN/SB
577	Vallée, M. (2004). Stabilizing the Empirical Green Function Analysis: Develop-
578	ment of the Projected Landweber Method. Bulletin of the Seismological Soci-
579	ety of America, 94(2), 394–409. doi: 10.1785/0120030017
580	Vallée, M., & Douet, V. (2016). A new database of source time functions (STFs) ex-
581	tracted from the SCARDEC method. Physics of the Earth and Planetary Inte-
582	riors, 257, 149–157. doi: 10.1016/j.pepi.2016.05.012
583	Van Houtte, C., & Denolle, M. (2018). Improved Model Fitting for the Empirical
584	Green's Function Approach Using Hierarchical Models. Journal of Geophysical
585	Research: Solid Earth, 123(4), 2923–2942. doi: 10.1002/2017JB014943
586	Viegas, G., Abercrombie, R. E., & Kim, WY. (2010). The 2002 M5 Au Sable
587	Forks, NY, earthquake sequence: Source scaling relationships and energy
588	budget. Journal of Geophysical Research: Solid Earth, 115(B7). doi:
589	10.1029/2009JB006799
590	Viesca, R. C., & Garagash, D. I. (2015). Ubiquitous weakening of faults due to
591	thermal pressurization. Nature Geoscience, $8(11)$, $875-879$. doi: $10.1038/$
592	ngeo 2554
593	Zollo, A., Capuano, P., & Singh, S. K. (1995). Use of small earthquake records
594	to determine the source time functions of larger earthquakes: An alternative
595	method and an application. Bulletin of the Seismological Society of America,
596	85(4), 1249-1256. doi: 10.1785/BSSA0850041249